1	Optimizing sampling strategies in high-resolution paleoclimate records
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10 Abstract

The aim of paleoclimate studies to-resolvinge climate variability from noisy proxy records can in essence 11 12 be reduced to a statistical problem. The challenge is to extract meaningful information about climate 13 variability from these records by reducing measurement uncertainty through a combination combining 14 measurements of for proxiesy data while retaining the temporal resolution needed to assess the timing and 15 duration of variations in climate parameters. In this study, we explore the limits of this compromise by testing 16 different methods for combining proxy data (smoothing, binning and sample size optimization) on a 17 particularly challenging paleoclimate problem: resolving seasonal variability in stable isotope records. We test and evaluate the effects of changes in the seasonal temperature and the hydrological cycle as well as 18 19 changes in accretion rate of the archive and parameters such as sampling resolution and age model 20 uncertainty on the reliability of seasonality reconstructions based on clumped and oxygen isotope analyses 21 in 33 real and virtual datasets. Our results show that strategic combinations of clumped isotope analyses 22 can significantly improve the accuracy of seasonality reconstructions compared to conventional stable 23 oxygen isotope analyses, especially in settings where the isotopic composition of the water is poorly 24 constrained. Smoothing data using a moving average often leads to an apparent dampening of the seasonal 25 cycle, significantly reducing the accuracy of reconstructions. A statistical sample size optimization protocol 26 yields more precise results than smoothing. However, the most accurate results are obtained through 27 monthly binning of proxy data, especially in cases where growth rate or water composition cycles obscure 28 the seasonal temperature cycle. Our analysis of a wide range of natural situations reveals that the effect of 29 temperature seasonality on oxygen isotope records almost invariably exceeds that of changes in water 30 composition. Thus, in most cases, oxygen isotope records allow reliable identification of growth seasonality 31 as a basis for age modelling in the absence of independent chronological markers in the record. These 32 specific findings allow us to formulate general recommendations for sampling and combining data in 33 paleoclimate research and have implications beyond the reconstruction of seasonality. We briefly discuss the implications of our results for solving common problems in paleoclimatology and stratigraphy. 34

36 1. Introduction

37 Improving the resolution of climate reconstructions is a key objective in paleoclimate studies because it 38 allows climate variability to be studied on different timescales and sheds light on the continuum of climate 39 variability (Huybers and Curry, 2006). However, the temporal resolution of climate records is limited by the 40 accretion rate (growth or sedimentation rate) of the archive and the spatial resolution of sampling for climate 41 reconstructions, which is a function of the sample size required for a given climate proxy. This tradeoff 42 between sample size and sampling resolution is especially prevalent when using state-of-the-art climate 43 proxies which require large sample sizes, such as the carbonate clumped isotope paleothermometer (Δ_{47} ; see applications in Rodríguez-Sanz et al., 2017; Briard et al., 2020; Caldarescu et al., 2021) or stable 44 45 isotope ratios in specific compounds or of rare isotopes (e.g. phosphate-oxygen isotopes in tooth apatite, 46 triple oxygen isotopes in speleothems or carbon isotopes of CO2 in ice cores; Jones et al., 1999; Schmitt 47 et al., 2012; Sha et al., 2020). The challenge of sampling resolution persists on a wide range of timescales: 48 from attempts to resolve geologically short-lived (kyr-scale) climate events from deep sea cores with low 49 sedimentation rates (e.g. Stap et al., 2010; Rodríguez-Sanz et al., 2017) to efforts to characterize tidal or daily variability in accretionary carbonate archives (e.g. Warter and Müller, 2017; de Winter et al., 2020a). 50 51 What constitutes "high-resolution" is therefore largely dependent on the specifics of the climate archive.

52 Sample size limitations are especially important in paleoseasonality reconstructions. Reliable archives for 53 seasonality (e.g. corals, mollusks and speleothem records) are in high demand in the paleoclimate 54 community, because the seasonal cycle is one of the most important cycles in Earth's climate and 55 seasonality reconstructions complement more common long-term (kyr to -Myr) records of past climate 56 variability (e.g. Morgan and van Ommen, 1997; Tudhope et al., 2001; Steuber et al., 2005; Steffensen et 57 al., 2008; Denton et al., 2005; Huyghe et al., 2015; Vansteenberge et al., 2019). A more detailed 58 understanding of climate dynamics at the human timescale is increasingly relevant for improving climate projections (IPCC, 2013). Unfortunately, the growth and mineralization rates of archives that capture high-59 60 resolution variability (rarely-only exceeding 10 mm/yr in rare exceptions, e.g. Johnson et al., 2019) limit the 61 number and size of samples that can be obtained at high temporal resolutions (e.g. Mosley-Thompson et al., 1993; Passey and Cerling, 2002; Treble et al., 2003; Goodwin et al., 2003). In addition, accurate dating 62

63 of climate archivespositioning of samples within the seasonal cycle is challenging. In absence of fine-scale 64 growth markings (e.g. daily laminae in mollusk shells; e.g. Schöne et al., 2005; de Winter et al., 2020a), 65 this dating problem relies on modelling or interpolation of the growth of the archive, which introduces 66 uncertainty on the age of samples (e.g. Goodwin et al., 2009; Judd et al., 2018). Theise problems is are 67 exacerbated by the fact that accurate methods for climate reconstructions often may require comparatively 68 large sample sizes, while methods relying on smaller sample sizes or rely on uncertain assumptions. A case 69 in point is the popular carbonate stable oxygen isotope temperature proxy ($\delta^{18}O_c$) which relies on 70 assumptions of the water composition ($\delta^{18}O_w$) that become progressively more uncertain further back in geological history (e.g. Veizer and Prokoph, 2015). ContrarilyIn contrast, the clumped isotope proxy (Δ47) 71 72 does not rely on this assumption but requires larger amounts of sample (e.g. Müller et al., 2017)

73 A promising technique for circumventing sample size limitations is to analyze larger numbers of small 74 aliquots from the same sample or from similar parts of the climate archive. These smaller aliquots typically 75 have poor precision but averaging multiple aliquots into one estimate while propagating the measurement 76 uncertainty leads to a more reliable estimate of the climate variable (Dattalo, 2008; Meckler et al., 2014; 77 Müller et al., 2017; Fernandez et al., 2017). This approach yields improved sampling flexibility since aliquots 78 can be combined in various ways after measurement. It also allows outlier detection at the level of individual 79 aliquots, thereby spreading the risk of instrumental failure and providing improved control on changes in 80 measurement conditions that may bias results.

81 Previous studies have applied several different methods for combining data from paleoclimate records to reduce analytical noise or higher order variability, and extract variability with a specific frequency (e.g. a 82 83 specific orbital cycle or seasonality; e.g. Lisiecki and Raymo, 20054; Cramer et al., 2009). These data 84 reduction approaches can in general be categorized into smoothing techniques, in which a sliding window 85 or range of neighboring datapoints is used to smooth high resolution records (see e.g. Cramer et al., 2009) or binning techniques, in which the record is divided into equal bins in the sampling direction (e.g. time, 86 87 depth or length in growth direction; e.g. Lisiecki and Raymo, 2004; Rodríguez-Sanz et al., 2017). In addition, 88 a third approach is proposed here based on optimization of sample size for dynamic binning of data along 89 the climate cycle using a moving window in the domain of the climate variable (as opposed to the sampling

domain) combined with a T-test routine (see section 2.1). All three approaches have advantages and
caveats.

In this study, we explore the (dis)advantages of these three data reduction approaches by testing their reliability in resolving seasonal variability in sea surface temperature (SST) and water stable oxygen isotope composition ($\delta^{18}O_w$), both highly sought-after variables in paleoclimate research. We compare reconstructions of SST and $\delta^{18}O_w$ in real and virtual datasets from accretionary carbonate archives (e.g. shells, corals and speleothems) using the clumped isotope thermometer (Δ_{47}) combined with stable oxygen isotope ratios of the carbonate ($\delta^{18}O_c$).

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99 2. Methods

100 2.1 Reconstruction approaches

101 Throughout the remainder of this work, the three approaches for combining data for reconstructions are 102 defined as follows (see also **Fig. 1**):

103 Smoothing refers to the reconstruction of SST and $\delta^{18}O_w$ based on moving averages of Δ_{47} and $\delta^{18}O_w$ 104 records (Fig. 1B). For every dataset, the full possible range of moving window sizes (from 1 sample to the full length of the record) for SST and $\delta^{18}O_w$ reconstructions was explored. The window size that resulted in 105 106 the most significant difference between maximum and minimum Δ_{47} values (based on a student's T-test) 107 was applied to reconstruct SST and $\delta^{18}O_w$ from Δ_{47} and $\delta^{18}O_c$ records. SST and $\delta^{18}O_w$ were calculated for 108 all case studies using a combination of empirical temperature relationships by Kim and O'Neil (1997; $\delta^{18}O_c$ -109 δ^{18} Ow-temperature relationship) and Bernasconi et al. (2018; Δ₄₇-temperature relationship). <u>To obtain δ^{18} Ow</u> 110 values, the $\delta^{18}O_{e^-} \delta^{18}O_{w^-}$ temperature relationship (Kim and O'Neil, 1997) was solved for $\delta^{18}O_{w}$ using the 111 temperature reconstruction obtained from Δ_{47} measurements. Here and in other approaches, a typical 112 analytical uncertainty on measurements of Δ_{47} (one standard deviation of 0.04‰) and $\delta^{18}O_c$ (one standard 113 deviation of 0.05‰) was used to include uncertainty due to measurement precision. These analytical 114 uncertainties were chosen based on typical uncertainties reported for these measurements in the literature 115 (e.g. Schöne et al., 2005; Huyghe et al., 2015; Vansteenberge et al., 2016) and long-term precision 116 uncertainties obtained by measuring in-house standards using the MAT253+ with Kiel IV setup in the 117 clumped isotope laboratory at Utrecht University (e.g. Kocken et al., 2019). The measurement uncertainty 118 was propagated through all calculations using a Monte Carlo simulation (N = 1000) in which Δ_{47} and $\delta^{18}O_c$ 119 records were randomly sampled from a normal distribution with the virtual Δ_{47} and $\delta^{18}O_c$ values as means 120 and analytical uncertainties as standard deviations. <u>Resulting SST and $\delta^{18}O_w$ values were grouped into</u> 121 monthly time bins using the age model of the archive.

Binning refers to reconstructions of SST and $\delta^{18}O_w$ based on binning of Δ_{47} and $\delta^{18}O_c$ records into monthly 122 123 time bins (**Fig. 1C**). The Δ_{47} and $\delta^{18}O_c$ data from each case study were grouped into monthly time bins and 124 converted to SST and $\delta^{18}O_w$ using the Kim and O'Neil (1997) and Bernasconi et al. (2018) formulae. Here 125 too, Monte Carlo simulation (N = 1000) was applied to propagate measurement uncertainties onto monthly 126 SST and $\delta^{18}O_w$ reconstructions. Note that the prerequisite for this method is that the data is aligned using 127 a (floating) age model accurate enough to allow samples to be placed in the right bin. The age of virtual 128 samples in this study is known so this prerequisite poses no problems in this case. However, in the fossil 129 record this alignment might be less certain in the absence of accurate chronologies within the archive (e.g. 130 through daily growth increments in mollusk shells; e.g. Schöne et al., 2008; Huyghe et al., 2019; see 4.1.3).

131 **Optimization** refers to reconstructions of SST and $\delta^{18}O_w$ based on sample size optimization in Δ_{47} records 132 (Fig. 1D). In this approach aliquots of each virtual dataset are ordered from warm (low $\delta^{18}O_c$) to cold (high 133 $\delta^{18}O_c$ data) samples, regardless of their position relative to the seasonal cycle. From this ordered dataset, 134 increasingly large samples of multiple aliguots (from 2 aliguots to half the length of the record) are taken from both the warm ("summer") and the cold ("winter") side of the distribution. Summer and winter samples 135 were kept equal (symmetrical grouping) to reduce the number of possible sample size combinations and 136 137 allow for more efficient computation. However, asymmetrical grouping with differing sample sizes on the 138 summer and winter ends of the $\delta^{18}O_c\text{-spectrum}$ are possible (see 4.1.3 and 4.2.2). Sample sizes with significant difference in Δ_{47} value between summer and winter groups (p \leq 0.05 based on a student's T-139 140 test) were selected as optimal sample sizes. The moving window T-test in the proxy domain ensures that 141 an optimal compromise is reached between high precision and resolving differences between seasonal 142 extremes. For each successful sample size, SST and $\delta^{18}O_w$ values were calculated from Δ_{47} and $\delta^{18}O_c$ data according to Kim and O'Neil (1997) and Bernasconi et al. (2018) formulae. The relationship between SST and $\delta^{18}O_w$ obtained from these reconstructions was used to convert all Δ_{47} and $\delta^{18}O_c$ data to SST and $\delta^{18}O_w$, which are then grouped into monthly SST and $\delta^{18}O_w$ reconstructions along the archive's age model. Measurement uncertainties were propagated through the entire approach by Monte Carlo simulation (N = 1000).

For comparison, we also include reconstructions based solely on $\delta^{18}O_c$ measurements with an (often inaccurate) assumption of a constant $\delta^{18}O_w$ (equal to the modern ocean value of 0‰ VSMOW), which form the most common method for carbonate-based temperature reconstructions in paleoclimate research (see e.g. Schöne et al., 2005; Westerhold et al., 2020; **Fig. 1A**; hereafter: $\delta^{18}O$). For these reconstructions, $\delta^{18}O_c$ records were grouped into monthly time bins with analytical uncertainties propagated using the Monte Carlo approach (N = 1000) and were directly converted to SST using the Kim and O'Neil (1997) temperature relationship.

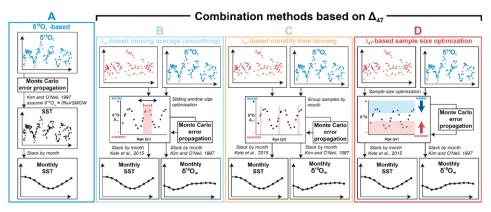
For each reconstruction, SST and $\delta^{18}O_w$ results were aggregated into monthly averages, medians, standard deviations, and standard errors. Step by step documentation of calculations made for the three Δ_{47} -based reconstruction approaches and the $\delta^{18}O_c$ reconstructions are given in **Supplmentary Data S7** and in the complementary R package (de Winter, 2021a).

159 2.2 Benchmarks for accuracy and precision

Accuracy and precision of reconstructions were evaluated against official USGS definitions of climate
 parameters (O'Donnell et al.and Ignizio, 2012):

- 162 1. mean annual SST (MAT), defined as the average of all 12 monthly temperature reconstructions.
- 163 2. seasonal range in SST, defined as the temperature difference between warmest and coldest164 month.
- 165 3. mean annual δ^{18} Ow, defined as the average of all 12 monthly δ^{18} Ow reconstructions.
- seasonal range in δ¹⁸O_w, defined as the δ¹⁸O_w difference between most enriched (highest δ¹⁸O_w)
 and most depleted (lowest δ¹⁸O_w) monthly reconstruction.

Accuracy was defined as the absolute offset of the reconstructed climate parameter from the "true" value. Precision was defined as the (relative) standard deviation of the reconstruction, as calculated from the variability within monthly time bins resulting from Monte Carlo error propagation (see **2.1**). An overview of monthly SST and $\delta^{18}O_w$ reconstructions using the four approaches in all cases is given in **S4**. Raw data and figures of reconstructions of all cases using all sampling resolutions are compiled in **S8**.



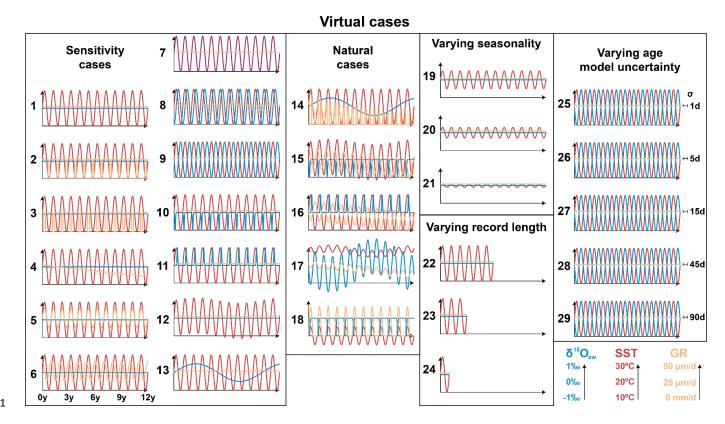
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Figure 1: Schematic overview of the four approaches for seasonality reconstructions: (A) δ¹⁸O-based
reconstructions, assuming constant δ¹⁸O_w. (B) Reconstructions based on smoothing δ¹⁸O_c and Δ₄₇ data
using a moving average. (C) Reconstructions based on binning δ¹⁸O_c and Δ₄₇ data in monthly time bins.
(D) Reconstructions based on optimization of the sample size for combining δ¹⁸O_c and Δ₄₇ data (see
description in 2.1). Colored curves-points represent virtual δ¹⁸O_c (blue) and Δ₄₇ (red) series in sampling
domain. Black curves represent reconstructed monthly SST and δ¹⁸O_w averages.

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181 2.3 SST and $\delta^{18}O_w$ datasets

182	The three reconstruction approaches were tested and compared based on three types of data $\underline{::}$ Firstly, $\underline{:a}$
183	set of datasets based on fully artificial environmental SST and $\delta^{18}O_w$ data (case 1-29; see Fig. 2) converted
184	to virtual Δ_{47} and $\delta^{18}O_c$ records. data from a real specimen of a Pacific syster (Crassostrea gigas, syn.
185	Magallana gigas) reported in Ullmann et al. (2010). Secondly, data based on actual measurements of
186	natural variability in SST and sea surface salinity (SSS; case 30-33) converted to virtual Δ_{47} and $\delta^{18}O_c$
187	records. Thirdly, measured proxy data from a real specimen of a Pacific oyster (Crassostrea gigas, syn.
188	<u>Magallana gigas</u>) compared to measured environmental (SST and $\delta^{18}O_w$) data reported in Ullmann et al.
189	(2010). a set of datasets based on fully artificial SST and $\delta^{19}O_w$ data (case 1-29; see Fig. 2) converted to
190	virtual Δ_{47} and $\delta^{48}\Theta_{e}$ records.



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Figure 2: Overview of time series of all virtual test cases. Colored curves represent time series of SST (red), $\delta^{18}O_w$ (blue) and growth rate (orange, abbreviated as "GR"). Horizontal axes in all plots are 12 years long (see legend below case 6). Vertical axis of all plots has the same scale (SST: 10 to 30°C; $\delta^{18}O_w$: -1 to +1‰; Growth rate: 0—50 µm/day; see legend in bottom right corner). Horizontal error bars and labels on the right side of cases 25-29 represent standard errors introduced on the age model (bars not to scale). The $\delta^{18}O_c$ and Δ_{47} records resulting from these virtual datasets are provided in **S6** (see also **Fig. 3** for natural examples).

Sensitivity cases		Sensitivity cases Natural cases		Varying age model uncertainty
	7. δ ¹⁸ O _w seasonality in phase with SST		19. Control case with reduced SST amplitude (~5°C)	
1. Control	8. $\delta^{18}O_w$ seasonality in antiphase with SST	14. Full marine case with ontogenetic GR trend	20. Control case with reduced SST amplitude (~3°C)	25. Case 9 with ±1 day age model uncertainty
2. Growth stops <12°C	9. δ ¹⁸ O _w seasonality lags SST by ¼ year	15. Coastal case with spring $\delta^{18}O_w$ decrease and decreasing GR trend	21. Control case with reduced SST amplitude (~1°C)	26. Case 9 with ±5 days age model uncertainty
3. Growth stops >28°C	10. Negative $\delta^{18}O_w$ in spring	16. Lagoonal case with summer $\delta^{18}O_w$ increase	Varying record length	27. Case 9 with ±15 days age model uncertainty
4. Linear decrease in GR	11. Positive $\delta^{18}O_w$ in summer	17. Tropical monsoon case with confined SST seasonality and strong multi-annual SST cycle	22. Control case shortened to 6 yr	28. Case 9 with ±45 days age model uncertainty
5. GR seasonality in phase with SST	12. Multi- annual (5 yr) SST cycle	18. Worst-case scenario with growth limited to summer half of the year	23. Control case shortened to 3 yr	29. Case 9 with ±90 days age model uncertainty
6. GR seasonality lags SST by ¼ year	13. Multi- annual (5 yr) δ ¹⁸ O _w cycle		24. Control case shortened to 1 yr	

Table 1: Overview of virtual cases 1-29 used to test the reconstruction methods. Case descriptions are abbreviated. Details on the SST, growth rate and δ¹⁸O_w included in each case are described in detail in S1.
SST, growth rate and δ¹⁸O_w records of all cases are shown in Fig. 2. "GR" = growth rate.

201	2.3.1	Virtual cases	Cases	1-29:	Virtual	environmental	data,	virtual	proxy	/ data

202	Virtual SST and $\delta^{18}O_w$ time series were artificially constructed to test the effect of various SST and $\delta^{18}O_w$
203	scenarios on the effectivity of the reconstruction methods. The default test case (case 1) contained an ideal,
204	12-year sinusoidal SST curve with a period of 1 year (seasonality), a mean value of 20°C and a seasonal
205	amplitude of 10°C, a constant $\delta^{18}O_w$ value of 0‰ and a constant growth rate of 10 mm/yr. Other cases
206	contain various deviations from this ideal case (see also Fig. 2, Table 1 and S1):
207	Linear and/or seasonal changes in growth rate, including growth stops (cases 2-6, 14-18)

208	 Seasonal and/or multi-annual changes in δ¹⁸O_w (cases 7-11, 13-18)
209	 Multi-annual trends in SST superimposed on the seasonality (cases 12, 15 and 17)
210	 Variations in the seasonal SST amplitude (cases 19-21)
211	 Change in the total length of the time series (cases 22-24).
212	 Variation in uncertainty on the age of each virtual datapoint (cases 25-29)
213	Comparison of the virtual time series (case 1-29; Fig. 2) with the natural variability (case 30-33; Fig. 3)
214	shows that the virtual cases are not realistic approximations of natural variability in SST and $\delta^{18}O_{\underline{w}}$. Natural
215	SST and $\delta^{18}O_w$ variability are not limited to the seasonal or multi-annual scale but contain a fair amount of
216	higher order (daily to weekly scale) variability. To simulate this natural variability, we extracted the seasonal
217	component of SST and $\delta^{18}O_w$ variability from our highest resolution record of measured natural SST and
218	SSS data (case 30: data from Texel, the Netherlands, see 2.3.2 and Fig. 3). The standard deviation of
219	residual variability of this data after subtraction of the seasonal cycle was used to add random high-
220	frequency noise to the SST and $\delta^{18}O_w$ variability in virtual cases. Note that while sub-annual environmental
221	variability can be approximated by Gaussian noise (Wilkinson and Ivany, 2002), this representation is an
222	oversimplification of reality. In the case of our Texel data, the SST and SSS residuals are not normally
223	distributed (Kolmogorov-Smirnov test: $D = 0.010$; $p = 7.2^{*}10^{-14}$ and $D = 0.039$; $p < 2.2^{*}10^{-16}$ for SST and
224	SSS residuals respectively; see S2-4). SST and $\delta^{18}O_w$ data from cases 1-29 was converted to the sampling
225	domain and subsampled at a range of sampling resolutions following the same procedure applied to cases
226	<u>30-33 (see 2.3.2).</u>
227	
228	Modorn oystor data
229	Environmental SST and $\delta^{19}O_w$ data from the List Basin in Denmark (54°59.25N, 8°23.51E), where the
230	modern oyster specimen lived, were obtained from local in situ measurements of SST and SSS described
231	in Ullmann et al. (2010). Since direct, <i>in situ</i> measurements of δ ¹⁸ O _w variability at a high temporal resolution

232 were not available, 5⁴⁶O_{**} was estimated from more widely available SSS data using a mass balance

233 (equation 1 and 2; following e.g. Ullmann et al., 2010):

$\delta^{18} \mathcal{O}_{sw} = \delta^{18} \mathcal{O}_{w, freshwater} * f + \delta^{18} \mathcal{O}_{w, ocean} * (1 - f) - \underbrace{(1)}{(1)}$ 234

 $f = \frac{SSS_{sample} - SSS_{ocean}}{SSS_{freshwater} - SSS_{ocean}} \tag{2}$ 235

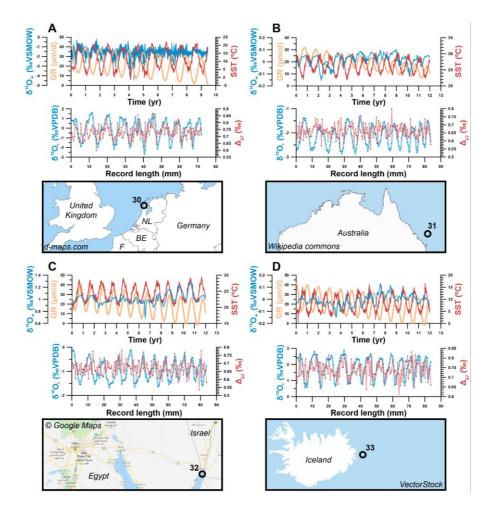
236	Here, we assume salinity (SSS _{sample}) results from a mixture of a fraction (f) isotopically light and low-salinity
237	$(\delta^{48}\Theta_{w,\text{free/water}} = -8.5\%; SSS_{\text{free/water}} = 0)$ freshwater and a fraction (1-f) ocean water ($\delta^{48}\Theta_{w,\text{eccarr}} = -0\%;$
238	$SSS_{ocearr} = 35$), with negative amounts of freshwater contribution (f < 0) representing net evaporation
239	$(SSS_{example} > SSS_{eccar})$. The value for $\delta^{18}O_{w,trestwater}$ was based on the discharge weighted average $\delta^{18}O_w$ of
240	water in the nearby Elbe and Weser rivers (see Ullmann et al., 2010). All $\delta^{48}O_{*}$ values throughout the text
241	are with reference to the VSMOW scale. Contrary to the virtual datasets (cases 1-33; see 2.3.2 and 2.3.3),
242	the Ullmann et al. (2010) data was already available in the sampling domain, hence no subsampling was
243	required.
244	2.3.2 Cases <u>30-33: based on real climate dataMeasured environmental data, virtual proxy data</u>
245	Four test cases were based on time series of real measured SST and SSS data from four different locations,
246	selected to capture a variety of environments with different SST and SSS variability (see Fig. 3):
247	1. Tidal flats of the Wadden Sea near Texel, the Netherlands (case 30)
248	2. Great Barrier Reef in Australia (case 31)
249	3. Gulf of Aqaba between Egypt and Saudi Arabia (case 32)
250	4. Northern Atlantic Ocean east of Iceland (case 33).
251	Daily measurements of SST and SSS for case 31-33 were obtained from worldwide open-access datasets
252	of the National Oceanic and Atmospheric Administration (NOAA, 2020) and European Space Agency (ESA,
253	2020) respectively. Hourly SST and SSS measured in situ in the Wadden Sea (case 30) were obtained
254	from the Dutch Institute for Sea Research (NIOZ, Texel, the Netherlands). Since direct, in situ
255	measurements of δ^{18} O _w variability at a high temporal resolution are scarce, δ^{18} O _w was estimated from (more
256	widely available) SSS data using the same mass balance described in 2.3.1. Since direct, in situ
257	measurements of $\delta^{18}O_w$ variability at a high temporal resolution were not available, $\delta^{18}O_w$ was estimated
258	from more widely available SSS data using a mass balance (equation 1 and 2; following e.g. Ullmann et
259	<u>al., 2010):</u>
1	12

260 $\delta^{18}O_{sw} = \delta^{18}O_{w,freshwater} * f + \delta^{18}O_{w,ocean} * (1-f)$ (1)

 $261 f = \frac{SSS_{sample} - SSS_{ocean}}{SSS_{freshwater} - SSS_{ocean}} (2)$

Here, we assume salinity (SSS_{sample}) results from a mixture of a fraction (f) isotopically light and low-salinity 262 263 $(\delta^{18}O_{w,freshwater} = -8\%; SSS_{freshwater} = 0)$ freshwater and a fraction (1-f) ocean water ($\delta^{18}O_{w,ocean} = 0\%;$ 264 <u>SSS_{ocean} = 35</u>), with negative amounts of freshwater contribution (f < 0) representing net evaporation 265 (SSS_{sample} > SSS_{ocean}). The value for δ^{18} Ow, freshwater was based on the δ^{18} Ow of rain in the Netherlands (-8%; Mook, 1970; Bowen, 2020). Applying this mass balance on the SSS record of the Wadden Sea tidal flats 266 267 (case 30) results in $\delta^{18}O_w$ values and a SSS- $\delta^{18}O_w$ relationship in agreement with measurements in this 268 region (Harwood et al., 2008). SST and $\delta^{18}O_w$ time series for all cases are given in Supplementary Data 269 S4 and natural cases are plotted in Fig. 3.

270 For all virtual proxy datasets (cases 1-33), records of SST and $\delta^{18}O_w$ were converted to the sampling 271 domain (along the length of the record) by defining a virtual growth rate in the sampling direction. Adding 272 this growth rate as a variable allowed us to test the sensitivity of approaches to changes in the extension 273 rate of the archive, including hiatuses (growth rate = 0). This is important, because fluctuations in linear 274 extension rate and periods in which no mineralization occurs (hiatuses or growth cessations) are common 275 in all climate archives (e.g. Treble et al., 2003; Ivany, 2012). After conversion to the sampling domain, virtual 276 aliquots were subsampled at equal distance from the SST and $\delta^{18}O_w$ series of all cases using six sampling 277 intervals: 0.1 mm, 0.2 mm, 0.45 mm, 0.75 mm, 1.55 mm and 3.25 mm. The four largest sampling intervals 278 were chosen such that the standard growth rate (10 mm/yr) was not an integer multiple of the sampling 279 interval (e.g. 0.45 mm instead of 0.5 mm, and 3.25 mm instead of 3 mm). This decision prevents sampling 280 the same parts of the seasonal cycle (e.g. same months) every year, which biases both the mean value and the precision of monthly SST and $\delta^{18}O_w$ reconstructions. This bias towards certain parts of the seasonal 281 282 cycle is much stronger at low sample sizes (large sampling intervals) and is illustrated in the 283 Supplementary InformationFigure S2.



284

Figure 3: Overview of the four cases of virtual data based on natural SST and SSS measurements explored in this study. (**A**) Case 30: Tidal flats on the Wadden Sea, Texel, the Netherlands. (**B**) Case 31 Great Barrier Reef, Australia). (**C**) Case 32: Gulf of Aqaba between Egypt and Saudi Arabia. (**D**) Case 33: Atlantic Ocean east of Iceland. For all cases, graphs on top show environmental data, with SST plotted in red, $\delta^{18}O_w$ in blue and growth rate (abbreviated as "GR") in orange (as in **Fig. 2**). The graph below shows virtual $\delta^{18}O_c$ (blue) and Δ_{47} (red) records created from these data series using a sampling interval of 0.45 mm and including analytical noise (se **3.3**). Note that the scale of vertical axes varies between plots.

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293 2.3.3 <u>Modern oyster-data</u>: Measured environmental data, measured proxy data

294	Environmental SST and $\delta^{18}O_w$ data from the List Basin in Denmark (54°59.25N, 8°23.51E), where the	
295	modern oyster specimen lived, were obtained from local in situ measurements of SST and SSS described	
296	in Ullmann et al. (2010). Since direct, in situ measurements of $\delta^{18}O_w$ variability at a high temporal resolution	
297	were not available, $\delta^{18}O_w$ was estimated from more widely available SSS data using athe mass balance	
298	described in 2.3.2. (equation 1 and 2; following e.g. Ullmann et al., 2010):	
299	$\delta^{\underline{18}} \theta_{\underline{sw}} = \delta^{\underline{18}} \theta_{\underline{w, freshwater}} * f + \delta^{\underline{18}} \theta_{\underline{w, ocean}} * (1 - f) \underline{(1)}$	
300	$f = \frac{SSS_{sumple} - SSS_{acean}}{SSS_{freshwater} - SSS_{acean}} $ (2)	
301	Here, we assume salinity (SSSsample) results from a mixture of a fraction (f) isotopically light and low-salinity	
302	$(\underline{\delta^{18}O_{w_{\underline{freshwater}}} = -8.5\%; SSS_{\underline{freshwater}} = 0) \text{ freshwater and a fraction (1-f) ocean water } (\underline{\delta^{18}O_{w_{\underline{ocean}}} = 0\%;}$	
303	<u>SSS_{ecean} = 35), with negative amounts of freshwater contribution ($f < 0$) representing net evaporation</u>	
304	$\frac{(SSS_{sampler} > SSS_{secont})}{\delta}$. The value for $\delta^{18}O_{w, freshwater}$ was based on the discharge weighted average $\delta^{18}O_{w}$ of	
305	water in the nearby Elbe and Weser rivers (see Ullmann et al., 2010). All $\delta^{18}O_w$ values throughout the text	
306	are with reference to the VSMOW scale. Contrary to the virtual datasets (cases 1-33; see 2.3.12 and	
307	2.3.32), the Ullmann et al. (2010) data was already available in the sampling domain, hence no subsampling	
308	was required.	
309	Virtual-cases	
310	Virtual SST and δ^{48} Ow time series were artificially constructed to test the effect of various SST and δ^{48} Ow	
311	scenarios on the effectivity of the reconstruction methods. The default test case (case 1) contained an ideal,	
312	12-year sinusoidal SST curve with a period of 1 year (seasonality), a mean value of 20°C and a seasonal	
313	amplitude of 10°C, a constant δ^{18} O _w -value of 0‰ and a constant growth rate of 10 mm/yr. Other cases	
314	contain various deviations from this ideal case (see also Fig. 2, Table 1 and S1):	
315	 Linear and/or seasonal changes in growth rate, including growth stops (cases 2-6, 14-18) 	Forma
316	 Seasonal and/or multi-annual changes in δ⁴⁹O_w (cases 7-11, 13-18) 	
317	 Multi-annual trends in SST superimposed on the seasonality (cases 12, 15 and 17) 	

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318	 Variations in the seasonal SST amplitude (cases 19-21)
319	Change in the total length of the time series (cases 22-24).
320	 Variation in uncertainty on the age of each virtual datapoint (cases 25-29)
321	Comparison of the virtual time series (case 1-20; Fig. 2) with the natural variability (case 30-33; Fig. 3)
322	shows that the virtual cases are not realistic approximations of natural variability in SST and δ^{48} O _w . Natural
323	SST and $\delta^{48}O_w$ -variability are not limited to the seasonal or multi-annual scale but contain a fair amount of
324	higher order (daily to weekly scale) variability. To simulate this natural variability, we extracted the seasonal
325	component of SST and $\delta^{18}O_w$ variability from our highest resolution record of measured natural SST and
326	SSS data (case 30: data from Texel, the Netherlands, see 2.3.2 and Fig. 3). The standard deviation of
327	residual variability of this data after subtraction of the seasonal cycle was used to add random high-
328	frequency noise to the SST and $\delta^{40}O_w$ variability in virtual cases. Note that while sub-annual environmental
329	variability can be approximated by Gaussian noise (Wilkinson and Ivany, 2002), this representation is an
330	oversimplification of reality. In the case of our Texel data, the SST and SSS residuals are not normally
331	distributed (Kolmogorov-Smirnov test: D = 0.010; p = $7.2^{*10^{-44}}$ and D = 0.039; p < $2.2^{*10^{-46}}$ for SST and
332	SSS residuals respectively; see S2-4). SST and $\delta^{49}O_w$ data from cases 1-29 was converted to the sampling
333	domain and subsampled at a range of sampling resolutions following the same procedure applied to cases
334	30-33 (see 2.3.2).

336 2.4 Conversion to $\delta^{18}O_c$ and Δ_{47} data

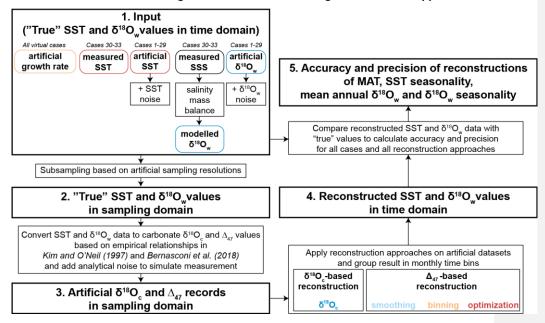
After subsampling, SST and $\delta^{18}O_w$ series (cases 1-33) were converted to $\delta^{18}O_c$ and Δ_{47} using a carbonate model based on empirical relationships between Δ_{47} and $\delta^{18}O_c$ with and SST and $\delta^{18}O_w$ (equation 3 and 4; Kim and O'Neil, 1997; Kele et al., 2015; Bernasconi et al., 2018) and the conversion of $\delta^{18}O$ values from VSMOW to VPDB scale (equation 5; Brand et al., 2014).

341
$$\Delta_{47} = \frac{0.0449 * 10^6}{(SST + 273.15)^2} + 0.167$$
 (3)

342
$$1000 * \ln \frac{\binom{^{18}o}{_{/16}o}_{caCO_3}}{\binom{^{^{18}o}_{_{/16}o}}_{_{H_2O}}} = 18.03 * \left(\frac{^{10^3}}{^{(SST+273.15)}}\right) - 32.42$$
 (4)

343
$$\delta^{18}O_{VPDB} = 0.97002 * \delta^{18}O_{VSMOW} - 29.98$$
 (5)

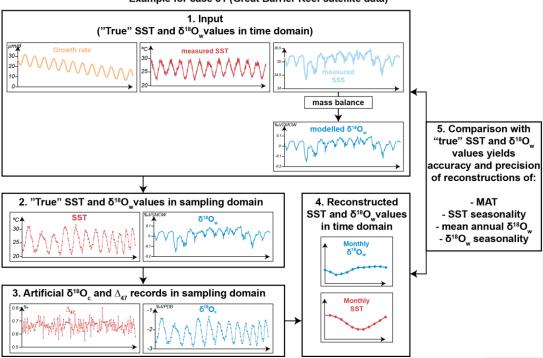
344 For the real-modern oyster data (Ullmann et al., 2010; see 2.3. 3^{1}), only the Δ_{47} data needed to be created 345 because $\delta^{18}O_c$ was directly measured. As a result, each case study yielded records of Δ_{47} and $\delta^{18}O_c$ in the 346 sampling domain and corresponding "true" SST and δ18Ow records in the time domain, allowing assessment 347 of the reliability of the reconstruction approaches in different scenarios- (Figure Fig. 4). The result of 348 applying these steps is illustrated on case 31 (Great Barrier reef data, Fig. 5). All calculations for creating 349 Δ_{47} and $\delta^{18}O_c$ series in sampling domain were carried out using the open-source computational software R 350 (R core team, 2013), and scripts for these calculations are given in Supplementary Data S7 and compiled 351 in the documented R package "seasonalclumped" (de Winter, 2021a). All Δ_{47} and $\delta^{18}O_c$ datasets are 352 provided in Supplementary Data SS6.



Workflow for creating virtual datasets and testing reconstruction approaches

354

Figure 4: Flow diagram showing the steps taken to create virtual data (Δ_{47} and $\delta^{18}O_c$; cases 1-33) and 355 compare results of SST and $\delta^{18}O_w$ reconstructions with the actual SST and $\delta^{18}O_w$ data the record was 356 357 based on (counterclockwise direction). Steps 1-3 outline the procedure for creating virtual Δ_{47} and $\delta^{18}O_c$ 358 datasets (see sections 2.3 and 2.4), step 4 shows the application of the different reconstruction methods 359 on this virtual data (see Fig. 2 for details) and step 5 illustrates how the reconstructions are compared with 360 the original ("true") SST and $\delta^{18}O_w$ data to calculate accuracy and precision of the reconstruction approaches. Note that step 1 is different for cases 1-29 (based on fully artificial SST and $\delta^{18}O_w$ records; 361 362 **2.3.31**) than for cases 30-33 (SST and $\delta^{18}O_w$ records based on real SST and SSS data; see **2.3.2**).



Workflow for creating virtual datasets and testing reconstruction approaches: Example for case 31 (Great Barrier Reef satellite data)

363

Figure 5: An example of the steps highlighted in **Fig. 4** using case 31 (Great Barrier Reef data) to illustrate the data processing steps. Virtual data plots include normally distributed measurement uncertainty on Δ_{47} and $\delta^{18}O_c$

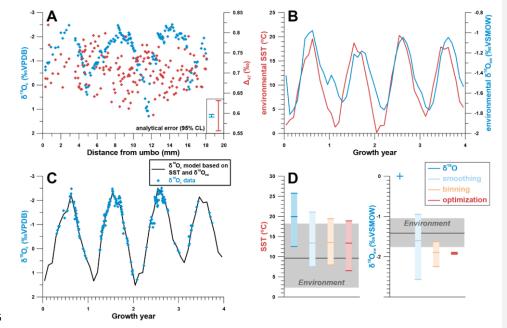
368 3. Results

369 3.1 Real example

370 Measured ($\delta^{18}O_c$) and simulated (Δ_{47}) data from the Pacific oyster from the Danish List Basin yielded 371 estimates of SST and δ¹⁸O_w seasonality using all reconstruction approaches (Fig. 6). While a model of shell 372 $\delta^{18}O_c$ based on SST and SSS data closely approximates the measured $\delta^{18}O_c$ record (Fig. 6C), basing SST 373 reconstructions solely on $\delta^{18}O_c$ data without any *a priori* knowledge of $\delta^{18}O_w$ variability (assuming constant 374 $\delta^{18}O_w$ equal to the global marine value) leads to high inaccuracy in SST seasonality and mean annual SST 375 (Fig. 6D). Note that, in absence of significant $\delta^{18}O_w$ seasonality (as in this case study), seasonal 376 temperature range reconstructions from $\delta^{18}O_c$ measurements can be very accurate. However, assuming 377 constant δ¹⁸Ow year-round may introduce considerable bias (see Fig. 7 and 8). The in-phase relationship 378 between SST and SSS (Fig. 6B) slightly dampens the seasonal $\delta^{18}O_c$ cycle, causing underestimation of 379 temperature seasonality, while a negative mean annual $\delta^{18}O_w$ value in the List Basin biases SST 380 reconstructions towards higher temperatures. In terms of SST reconstructions, the smoothing, binning 381 and **optimization** approaches based on Δ_{47} and $\delta^{18}O_c$ data yield more accurate reconstructions, albeit with 382 a reduced seasonality and a bias towards the summer season. The latter is a result of severely reduced 383 growth rates in the winter season, which was therefore undersampled (see Fig. 6A and 6C). Approaches 384 including Δ_{47} data also yield far more accurate $\delta^{18}O_w$ estimates than the $\delta^{18}O$ approach. However, the accuracy of $\delta^{18}O_w$ seasonality and mean annual $\delta^{18}O_w$ estimates is low in these approaches too, largely 385 386 because of the limited sampling resolution, especially in winter. The optimization approach suffers from the strong in-phase relationship between SST and SSS, which obscures the difference between the $\delta^{18}O_w$ 387 388 effect and the temperature effect on shell carbonate. Yet, disentangling SST from $\delta^{18}O_w$ seasonality is 389 central to the success of the approach (see 3.4). Fig. 6D does not show the reproducibility errorprecision 390 on SST and $\delta^{18}O_w$ estimates, which is much larger lower for the **smoothing** approach than for the **binning** 391 an optimization approaches due to the limited data in the winter seasons (see Supplementary Data S56). 392 These results show that several properties of carbonate archives, such as growth rate variability, phase 393 relationships between SST and $\delta^{18}O_w$ seasonality and sampling resolution, can impact the reliability of

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394 paleoseasonality reconstructions. The virtual and real data cases in this study were tailored to test the



395 effects of these archive properties more thoroughly.



405

Figure 6: (**A**) Plot of $\delta^{18}O_c$ and (virtual) Δ_{47} data from a modern Pacific oyster (Crassostrea gigas; see Ullmann et al., 2010). (**B**) shows SST and $\delta^{18}O_w$ data from the List Basin (Denmark) in which the oyster grew. (**C**) shows the fit between $\delta^{18}O_c$ data and modelled $\delta^{18}O_c$ calculated from SST and $\delta^{18}O_w$ on which the shell age model was based. (**D**) Shows a summary of the results of different approaches for reconstructing SST and $\delta^{18}O_w$ from the $\delta^{18}O_c$ and Δ_{47} data. The vertical colored bars show the reconstructed seasonal variability using all methods with ticks indicating warmest month, coldest month, and annual mean. The grey horizontal bars show the actual seasonal variability in the environment. Precision errors standard deviation on monthly reconstructions are not shown but are given in **S4**.

406 3.2 Case-specific results

407 A case-by-case breakdown of the precision (Fig. 7) and accuracy (Fig. 8) of reconstructions using the four

408 approaches shows that reliability of reconstructions varies significantly between approaches and is highly

409 case-specific. In general, precision is highest in δ^{18} O reconstructions, followed by optimization and

410 **binning**, with **smoothing** generally yielding the worst precision. Average precision standard deviations of

411 the underperforming methods (**binning** and **smoothing**) are up to 2-3 times larger than those of δ^{18} O (e.g.

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412 respectively 3.9°C and 3.5°C vs. 1.3°C for $\delta^{18}O$ MAT reconstructions; see also Supplementary 413 **Information**). It is worth noting that precision on δ^{18} O-based estimates is mainly driven by measurement 414 precision (which is better for $\delta^{18}O_c$ than for Δ_{47} measurements, see section 4.1.1). Δ_{47} -based reconstructions 415 lose precision due to the higher measurement error on Δ_{47} measurements and the method used for 416 combining measurements for seasonality reconstructions. On a case-by-case basis, the hierarchy of approaches can vary, especially if strong variability in growth rate is introduced, such as in case 14, where 417 418 the size of hiatuses in the record increases progressively, or in case 18, in which half of the year is missing 419 due to growth hiatuses (see Table 1, Supplementary Data S1 and S4). Of the Δ_{47} -based methods 420 (smoothing, binning and optimization), optimization is rarely outcompeted in terms of precision in both 421 SST and $\delta^{18}O_w$ reconstructions.

422 The comparison based on precision alone is misleading, as the most precise approach ($\delta^{18}O$) runs the risk 423 of being highly inaccurate (offsets exceeding 4°C on some MAT reconstructions; see Fig. 7C8A), especially 424 in cases based on natural SST and SSS measurements (case 30-33). The smoothing approach also often 425 yields highly inaccurate results, especially in cases with substantial variability in δ^{18} Ow (e.g. case 9-11; Fig. 426 8). Accuracy of optimization and binning outcompete the other methods in most circumstances. Binning 427 outperforms **optimization** in reconstructions of $\delta^{18}O_w$ seasonality, making it overall the most accurate 428 approach. Interestingly, optimization is less accurate specifically in cases with sharp changes in growth 429 rate in summer (e.g. cases 11, 14, 16 and 17), while binning performs better in these cases. Reconstructions of mean annual SST and $\delta^{18}O_w$ in case 18 are especially inaccurate regardless of which 430 431 method is applied. This extreme case with growth only during one half of the year combined with seasonal 432 fluctuations in both SST and $\delta^{18}O_w$ presents a worst-case scenario for seasonality reconstructions leading 433 to strong biases in mean annual temperature reconstructions. In situations like case 18, the optimization 434 approach is most accurate in MAT and SST seasonality reconstructions, but $\delta^{18}O_w$ is more accurately 435 reconstructed using the binning approach. Finally, it is worth noting that in natural situations (Fig. 3), variability in SST almost invariably has a larger influence on $\delta^{18}O_c$ and Δ_{47} records than $\delta^{18}O_w$, such that 436 fluctuations in $\delta^{18}O_c$ records closely follow the SST seasonality even in cases with relatively large $\delta^{18}O_w$ 437 438 variability (e.g. case 30). Chronologies based on these $\delta^{18}O_c$ fluctuations are therefore generally accurate.

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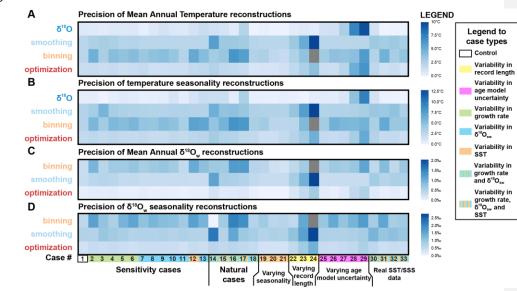
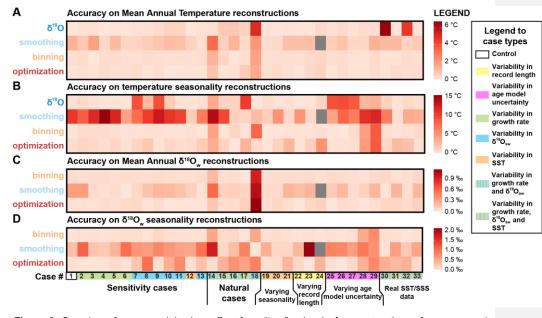


Figure 7: Overview of precision (propagated standard deviation of variability within reconstructions, see 2.2) of reconstructions of mean annual temperature (**A**), seasonal temperature range (**B**), mean annual $\delta^{18}O_w$ (**C**) and seasonal range in $\delta^{18}O_w$ (**D**), with higher values (darker colors) indicating lower precision (more variability between reconstructions) based on average sampling resolution (sampling interval of 0.45 mm). The different cases on the horizontal axis are color coded by their difference from the control case (case 1; see legend on the right-hand side). Grey boxes indicate cases for which reconstructions were not successful. All data on precision (standard deviation values) is provided in <u>Supplementary Data</u> S4.

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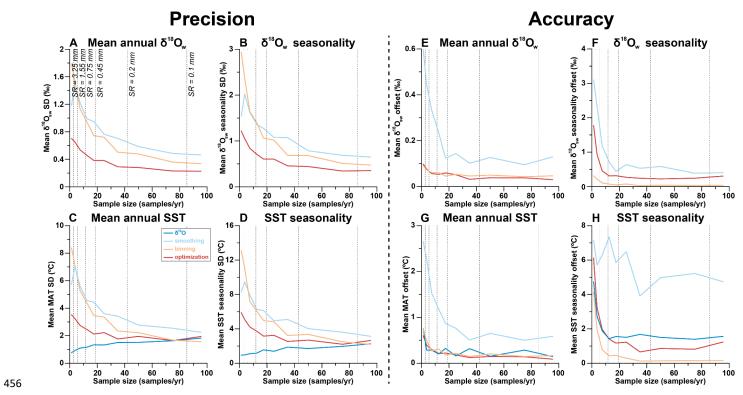


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Figure 8: Overview of accuracy (absolute offset from "true" values) of reconstructions of mean annual temperature (**A**), seasonal temperature range (**B**), mean annual $\delta^{18}O_w$ (**C**) and seasonal range in $\delta^{18}O_w$ (**D**), with higher values (darker colors) indicating lower accuracy (higher offsets) based on average sampling resolution (sampling interval of 0.45 mm). The different cases on the horizontal axis are color coded by their difference from the control case (case 1; see legend on the right-hand side). Grey boxes indicate cases for which reconstructions were not successful. All data on accuracy (difference between

454 reconstructed and "true" values) is provided in <u>Supplementary Data S4</u>.

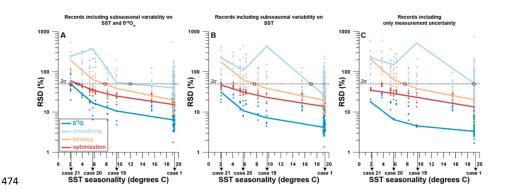
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457 **Figure 9**: Effect of sampling resolution (in samples per year, see **S5**) on the precision (one standard deviation) of results of reconstructions of mean 458 annual $\delta^{18}O_w$ (**A**), seasonal range in $\delta^{18}O_w$ (**B**), mean annual SST (**C**) and seasonal range in SST (**D**). Effect on the accuracy (absolute offset from 459 actual value) of results of reconstructions of mean annual $\delta^{18}O_w$ (**E**) and seasonal range in $\delta^{18}O_w$ (**F**), mean annual SST (**G**) and seasonal range in 450 SST (**H**). Color coding follows the scheme in **Fig. 1** and **Fig. 4**.

461 3.3 Effect of sampling resolution

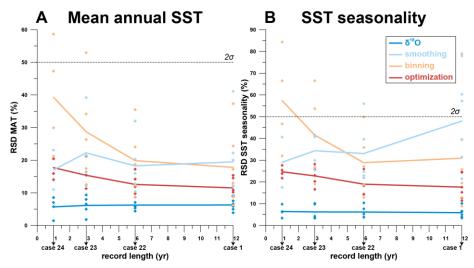
462 As expected, increasing the temporal sampling resolution (i.e. number of samples per year) almost 463 invariably increases the precision and accuracy (Fig. 9) of reconstructions using all methods. An exception 464 to this rule is the precision of $\delta^{18}O$ reconstructions, which decreases with increasing sampling resolution 465 (see Fig. 9C-D). Precision errors-standard deviations of all Δ_{47} -based approaches eventually converge with 466 the initially much lower higher precision error of $\delta^{18}O$ reconstructions when sampling resolution increases. 467 However, the sampling resolution required for Δ_{47} -based reconstructions to rival or outcompete the $\delta^{18}O$ reconstructions differs, with optimization requiring lower sampling resolutions than the other methods (e.g. 468 20-40 samples/year compared to 40-80 samples/year for smoothing and binning; Fig. 9A-D). Accuracy 469 470 also improves with sampling resolution (Fig. 9E-H). When grouping all cases together, it becomes clear 471 that δ^{18} O reconstructions can only approach the accuracy of Δ_{47} -based approaches for reconstructions of 472 MAT. Seasonality in both SST and $\delta^{18}O_w$ is most accurately reconstructed using binning, and the 473 smoothing approach once again performs worst.

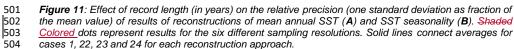


475 Figure 10: Effect of SST seasonality range (difference between warmest and coldest month) in the record 476 on the relative precision of SST seasonality reconstructions ("RSD", defined as one standard deviation 477 divided by the mean value). Panel A shows precision results if random variability ("weather patterns") in 478 both SST and $\delta^{18}O_w$ as well as measurement uncertainty is added to the records (see 2.3.3 and S1). Panel 479 B shows precision of records with random variability in SST and measurement uncertainty only. Panel C 480 shows precision if only measurement uncertainty is considered. Color coding follows the scheme in Fig. 1 481 and Fig. 4. Shaded dots represent results at various sampling resolutions, while bold lines are averages 482 for all reconstruction approaches. Black circles highlight the places where curves cross the threshold of two 483 standard deviations, which indicates the minimum SST seasonality that can be resolved within 2 standard 484 deviations (~95% confidence level) using the reconstruction approach.

486 3.4 Resolving SST seasonality

487 Comparison of cases 19, 20 and 21 (SST seasonality of 9.7°C, 5.7°C and 2.1°C respectively) with control 488 case 1 (SST seasonality of 19.3°C) shows how changes in the seasonal SST range affect the precision of 489 measurements (Fig. 10; see also Table 1 and Supplementary Data S1). The data reconfirms that δ^{18} O 490 reconstructions are most precise; a deceptive statistic given the risk of highly inaccurate results this 491 approach yields (see Fig. 8). Taking into consideration only analytical uncertainty, all approaches except 492 for smoothing can confidently resolve at least the highest SST seasonality within a significance level of 493 two standard deviations (~95%) using a moderate sampling resolution (mean of all resolutions shown in 494 Fig. 10). Increasing sampling resolution improves the precision of Δ_{47} -based reconstructions (see Fig. 9D), 495 so high sampling resolutions (0.1 or 0.2 mm) allow smaller seasonal differences to be resolved. When random sub-annual variability is added to the SST and $\delta^{18}O_w$ records (see 2.3.3), the minimum seasonal 496 497 SST extent that can be resolved decreases for all approaches (Fig. 10B and 10C). Nevertheless, $\delta^{18}O$ and 498 optimization reconstructions remain able to resolve a relatively small SST seasonality of 2-4°C.





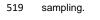
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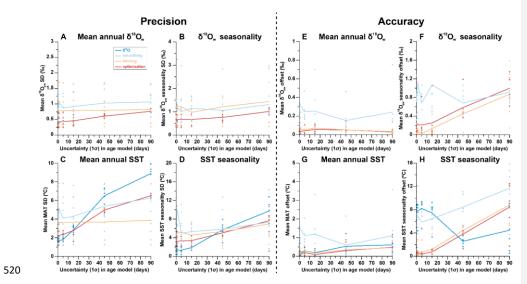
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506 3.5 Effect of record length

507 The effect of variation in the length of the record was investigated by comparing cases 22, 23 and 24 (record 508 lengths of 6 years, 3 years and 1 year, respectively) with the control case (record length of 12 years; see Fig. 11 and Table 1). Precision of MAT and SST seasonality reconstructions slightly increase in larger 509 datasets (longer records) for **optimization** and **binning**, but not for **smoothing** and δ^{18} O reconstructions. 510 511 Differences between reconstruction approaches remain relatively constant regardless of the length of the record, with general precision hierarchy generally remaining intact ($\delta^{18}O$ > optimization > binning > 512 513 smoothing). However, in very short records (1-2 years) smoothing generally gains an advantage over 514 other Δ_{47} -based methods due to its lack of sensitivity to changes in the record length, and binning 515 reconstructions are not precise enough to resolve MAT and SST seasonality within two standard deviations (~95% confidence level). Variation in precision is largely driven by very high low precision errors of 516 517 reconstructions in records with low sampling resolutions (sampling intervals of 1.55 mm or 3.25 mm; see

s18 also Fig. 9A-D). As a result, most of the reduction in precision in shorter records can be mitigated by denser





521Figure 12: Effect of uncertainty in age model on the reproducibility precision (standard deviation on522estimate) of results of reconstructions of mean annual $\delta^{18}O_w$ (**A**) and seasonal range in $\delta^{18}O_w$ (**B**), mean523annual SST (**C**) and seasonal range in SST (**D**). Effect of uncertainty in age model on the accuracy (offset524from true value) of results of reconstructions of mean annual $\delta^{18}O_w$ (**E**) and seasonal range in $\delta^{18}O_w$ (**F**),525mean annual SST (**G**) and seasonal range in SST (**H**). Color coding follows the scheme in **Fig. 1** and **Fig.**526**4**.

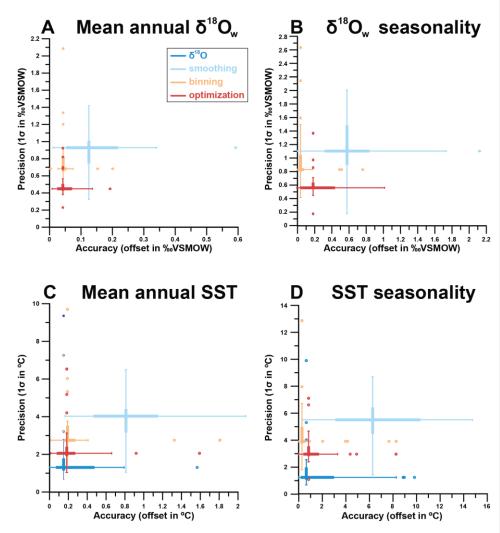
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528 3.6 Effect of age model uncertainty

529	Uncertainty on-in the age model has a significant effect on both the precision and the accuracy (Fig. 12) of
530	reconstructions using all approaches. The $\delta^{18}O$ reconstructions are most strongly affected by uncertainties
531	in the age model and suffer from a large decrease in precision with increasing age model uncertainty (Fig.
532	12C-D). The high reproducibility precision of the $\delta^{18}O$ approach in comparison with the Δ_{47} approaches
533	quickly disappears when age model uncertainty increases beyond 20-30 days. Accuracy of $\delta^{18}\text{Oc}\text{-based}$
534	SST seasonality reconstructions initially improves with age model uncertainty (Fig. 12H). However, this
535	observation is likely caused by the fact that age model uncertainty was compared based on conditions in
536	case 9, which features a phase offset between SST and $\delta^{18}O_w$ seasonality causing the $\delta^{18}O$ method to be

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537	highly inaccurate even without age model uncertainty. The precision of $\ensuremath{\textit{smoothing}}$ and $\ensuremath{\textit{optimization}}$
538	approaches also decreases with increasing age model uncertainty (Fig 12A-D), and the $\ensuremath{\textit{optimization}}$
539	approach loses its precision advantage over the $\ensuremath{\textbf{binning}}$ and $\ensuremath{\textbf{smoothing}}$ approaches when age model
540	uncertainty increases beyond 30 days. The monthly binning approach is most resilient against increasing
541	age model uncertainty. Seasonality reconstructions through both the binning and optimization approach
542	quickly lose accuracy when age model uncertainty increases but the accuracy of the smoothing approach
543	remains the worst of all Δ_{47} -based approaches in regardless of age model uncertainty except in the case
544	of $\delta^{18}O_w$ seasonality at exceptionally high (>60 days) age uncertainty (Fig. 12E-H).



545

Figure 13: Overview of averages and ranges of accuracy (absolute offset from real value) and precision (one standard deviation from the mean) on mean annual δ^{18} O_w (**A**) and seasonal range in δ^{18} O_w (**B**), mean annual SST (**C**) and seasonal range in SST (**D**) within all cases using the four different reconstruction approaches. Color coding follows the scheme in **Fig. 1** and **Fig. 4**. Box-whisker plots for precision and accuracy cross at their median values and outliers (colored symbols) are identified based on 2x the interquartile difference <u>(thick lines)</u>.

553 4. Discussion

554 4.1 Performance of reconstruction approaches

555 4.1.1 $\delta^{18}O_c$ vs Δ_{47} -based reconstructions

556 Figure 13 summarizes the general reliability of the four approaches. $\delta^{18}O$ reconstructions are generally 557 less accurate than Δ_{47} -based reconstructions (especially binning and optimization; see also <u>Supplementary Data</u> S9). This is a consequence of the assumption that $\delta^{18}O_w$ remains constant year-558 559 round, and that one knows its true value. Both these assumptions are problematic in the absence of 560 independent evidence of the value of $\delta^{18}O_w$, especially in deep time settings (see e.g. Veizer and Prokoph, 561 2015; Henkes et al., 2018). The risk of this assumption is made clear when comparing cases in which $\delta^{18}O_w$ 562 is indeed constant year-round at the assumed value (0%; e.g. cases 1-6 and 19-24) with cases in which 563 shifts in $\delta^{18}O_w$ occur, especially when these shifts are out of phase with respect to the SST seasonality (e.g. cases 9-11, 18 and 25-33; Fig. 8C-D). Cases mimicking or based on natural SST and SSS variability (cases 564 565 14-18 and 30-33) as well as the modern oyster data (Fig. 6) yield stronger inaccuracies in MAT and seasonality reconstructions, showing that even in many modern natural circumstances the assumption of 566 567 constant $\delta^{18}O_w$ is problematic.

It is important to consider that the value of mean annual $\delta^{18}O_w$ remained very close to the assumed value 568 569 of 0‰ (within 0.15‰) in all cases except for natural data cases 30 (-1.55‰), 32 (1.01‰; see 570 Supplementary Data S5) and the real oyster data (-1.42%; Fig. 5). The SST values of these cases 571 reconstructed using δ¹⁸O_c data show large offsets from their actual values (+6.7°C, -4.7°C and +10.3°C for 572 case 30, case 32 and the real oyster data respectively; see Fig. 6 and 8 and Supplementary Data S5). 573 These offsets are equivalent to the temperature offset one might expect from inaccurately estimating δ^{18} Ow 574 (~-4.6 °C/‰; Kim and O'Neil, 1997) and are only rivaled by the offset in MAT reconstructions of case 18 575 (+5.0°C), which has growth hiatuses obscuring the coldest half of the seasonal cycle. The fact that such 576 differences in $\delta^{18}O_w$ exist even in modern environments should not come as a surprise, given the available 577 data on worldwide variability of $\delta^{18}O_w$ (at least -3% to +2%; e.g. LeGrande and Schmidt, 2006) and SSS 578 (30 to 40; ESA, 2020) in modern ocean basins. However, it should warrant caution in using $\delta^{18}O_c$ data for 579 SST reconstructions even in modern settings. Implications for deep time reconstructions are even greater,

580 given the uncertainty on and variability in global average (let alone local) $\delta^{18}O_w$ values (Jaffrés et al., 2007; 581 Veizer and Prokoph, 2015). The complications of using $\delta^{18}O_c$ as a proxy for marine temperatures in deep 582 time are discussed in detail in O'Brien et al. (2017), and Tagliavento et al. (2019). Complications arising 583 from variability in $\delta^{18}O_w$ are more serious in climate records from euryhaline carbonate producers (e.g. 584 oysters) than from stenohaline organisms (e.g. corals), as they are mainly driven by salinity fluctuations. 585 For example, seasonal salinity variability in the North Sea in offshore sites away from freshwater sources 586 can be as low as 0.25 (Harwood et al., 2008), compared to 3-4 in the coastal Texel site simulated in case 587 30. Given this variability, studies using the $\delta^{18}O_c$ proxy for SST reconstructions are recommended to either 588 reconstruct $\delta^{18}O_w$ through additional measurements (e.g. including clumped isotope analysis) or constrain 589 δ¹⁸O_w variability through isotope-enabled modelling (e.g. Williams et al., 2009)

590 The analytical uncertainty of individual δ¹⁸O_c aliquots (typically 1 S.D. of 0.05‰; e.g. de Winter et al., 2018) 591 represents only ~1.1% of the variability in $\delta^{18}O_c$ over the seasonal cycle (~4.3‰ for the default 20°C 592 seasonality in case 1, following Kim and O'Neil, 1997). This is much smaller than the analytical uncertainty 593 of Δ_{47} (typically 1 S.D. of 0.02-0.04%; e.g. Fernandez et al., 20187; de Winter et al., 2020b), which equates to 25-50% of the seasonal variability in Δ_{47} (~0.08‰ for 20°C seasonality, following Bernasconi et al., 2018; 594 595 see Supplementary Data S7). This roughly 20-fold difference in relative precision causes $\delta^{18}O_c$ based SST 596 reconstructions to be much more precise (see Figs 7, 9-12) than those based on Δ_{47} , and forces the 597 necessity for grouping Δ_{47} data in reconstructions. However, as discussed above, the high precision of δ^{18} O 598 reconstructions is a misleading statistic if they are highly inaccurate.

599 Our results show that paleoseasonality reconstructions based on $\delta^{18}O_c$ can only be relied upon if there is 600 strong independent evidence of the value of δ^{18} O_w and if significant sub-annual variability in δ^{18} O_w (>0.3%, 601 equivalent to a 2-3°C SST variability; see Fig. 9-10; Kim and O'Neil, 1997) can be excluded with confidence. 602 Examples of such cases include fully marine environments unaffected by influxes of (isotopically light) freshwater or evaporation (increasing 5¹⁸Ow; Rohling, 2013). Carbonate records from environments with 603 604 more stable $\delta^{18}O_w$ conditions include, for example, the *A. islandica* bivalves from considerable depth (30-605 50m) in the open marine Northern Atlantic (e.g. Schöne et al., 2005, on which case 33 is based). However, 606 even here variability in $\delta^{18}O_{sw}$ due to, for example, shifting influence of different bottom water masses

cannot be fully excluded. Previous reconstruction studies show that $\delta^{18}O_w$ in smaller basins are heavily influenced by the processes affecting $\delta^{18}O_w$ on smaller scales, such as local evaporation and freshwater influx from nearby rivers (e.g. Surge et al., 2001; Petersen et al., 2016). Consequently, accurate quantitative reconstructions of seasonal range in shallow marine environments with extreme seasonality may not be feasible using the $\delta^{18}O$ approach, because these environments are invariably characterized by significant fluctuations in $\delta^{18}O_w$ and growth rate.

613 While variability in $\delta^{18}O_w$ compromises accurate $\delta^{18}O$ -based seasonality reconstructions, the compilation 614 in **Fig. 3** shows that its influence on the $\delta^{18}O$ records is too small to affect the shape of the record to such 615 a degree that seasonality is fully obscured. While natural situations with $\delta^{18}O_w$ fluctuations large enough to 616 totally counterbalance the effect of temperature seasonality on $\delta^{18}O$ records are imaginable, these cases 617 are likely rare. This means that chronologies based on $\delta^{18}O$ seasonality, which are a useful tool to anchor 618 seasonal variability in absence of independent growth markers (e.g. Judd et al., 2018; de Winter, 2021b), 619 are reliable in most natural cases.

620 4.1.2 Seasonality reconstructions using moving averages (smoothing)

621 Of the three methods for combining Δ_{47} data, the **smoothing** approach clearly performs worst in all four 622 reconstructed parameters (MAT, SST seasonality, mean annual $\delta^{18}O_w$ and $\delta^{18}O_w$ seasonality), both in 623 terms of accuracy and precision (Fig. 13). While applying a moving average may be a good strategy for 624 lowering the uncertainty of Δ_{47} -based temperature reconstructions in a long time series (e.g. Rodríguez-625 Sanz et al., 2017), the method underperforms in cases where baseline and amplitude of a periodic 626 component (e.g. MAT and SST seasonality) are extracted from a record. This is likely due to the smoothing 627 effect of the moving average, which reduces the seasonal cycle and causes highly inaccurate seasonality 628 reconstructions (offsets mounting to >6°C; Fig. 13). This bias is especially detrimental in cases where the 629 seasonal cycle is obscured by seasonal growth halts (e.g. case 18), multi-annual trends in growth (e.g. 630 case 4, 14 and 17) and multi-annual trends in SST (e.g. case 15 and 17; see Fig. 7 and 8). The poor 631 performance of the smoothing approach can be slightly mitigated by increasing sampling resolution (Fig 632 9), but even at high sampling resolutions (every 0.1 or 0.2 mm) the method still fails to reliably resolve 633 seasonal SST ranges below 5°C even in idealized cases (case 19-21; Fig. 10). Increasing the number of samples by analyzing longer records does not improve the result, because smoothing of the seasonal cycle
by a moving average window introduces the same dampening bias if the temporal sampling resolution
(number of samples per year) remains equal (Fig. 11).

637 More critically, employing the smoothing method may give the illusion that seasonality is more reduced, 638 and severely bias reconstructions. This bias highlights the importance of using the official meteorological 639 definition of seasonality as the difference between the averages of warmest and coldest month in 640 paleoseasonality work (O'Donnell et al.and Ignizio, 2012). This definition is much more robust than the 641 "annual range" often cited based on maxima and minima in $\delta^{18}O_c$ records. This "annual range" strongly 642 depends on sampling resolution, which is typically <12 samples/yr (Goodwin et al., 2003), equivalent to the 643 third lowest sampling interval (0.75 mm) simulated in this study. Therefore, we strongly recommend future 644 studies to adhere to the monthly definition of seasonality to foster comparison between studies. While inter-645 annual variability is lost by combining data from multiple years into monthly averages, this approach 646 increases precision, accuracy and comparability of paleoseasonality results. Inter-annual variability can still 647 be discussed from plots of raw data plotted in time or sampling domain.

648 4.1.3 Monthly binning, sample size optimization and age model uncertainty

649 Overall, the most reliable paleoseasonality reconstructions can be obtained from either binning or 650 optimization (Fig. 13). In general, optimization is slightly more precise, while binning yields more 651 accurate estimates of seasonal range in SST and $\delta^{18}O_w$ (Fig. 13B and D). The more flexible combination 652 of aliquots in the optimization routine yields improved precision (especially on mean annual averages) in 653 cases where parts of the record are undersampled or affected by hiatuses and simultaneous fluctuations 654 in both SST and $\delta^{18}O_w$ (e.g. case 3-6, 14-18, 30-33). The downside of this flexibility is that in the case of 655 larger sample sizes, the seasonal variability may be dampened, like in the smoothing approach (see 4.1.2). 656 This apparent dampening effect may be reduced by allowing the sample size of summer and winter samples 657 to vary independently in the optimization routine, at the cost of higher computational intensity due to the 658 larger number of sample size combinations (see 2.1 and 4.2.2). The rigid grouping of data in monthly bins 659 in binning prevents this dampening and therefore yields slightly more accurate estimates of seasonal 660 ranges in SST and $\delta^{18}O_w$. A caveat of **binning** is that it requires a very reliable age model of the record, at least on a monthly scale. If the age model has a large uncertainty, there is a risk that samples are grouped in the wrong month, which compromises the accuracy of **binning** reconstructions, especially for reconstructions of seasonal range (**Fig 12H**). This problem is exacerbated by potential phase shifts between seasonality in paleoclimate variables (SST and δ^{18} Ow) and calendar dates, which may occur in the presence of a reliable age model.

666 Previous authors attempted to circumvent the dating problem by analyzing high-resolution $\delta^{18}O_c$ transects 667 and subsequently sampling the seasonal extremes for clumped isotope analyses (Keating-Bitonti et al., 668 2011; Briard et al., 2020). While this approach does not require sub-annual age models, it has several 669 disadvantages compared with the binning and optimization approaches: Firstly, it requires separate 670 sampling for $\delta^{18}O_c$ and Δ_{47} , which may not be possible in high-resolution carbonate archives due to sample 671 size limitations. Analyzing small aliquots for combined $\delta^{18}O_c$ and Δ_{47} analyses consumes less material. 672 Secondly, individual summer and winter temperature reconstructions require large (> 1.5 mg; e.g. 673 Fernandez et al., 2017) Δ₄₇ samples from seasonal extremes, which causes more time-averaging than the 674 approaches combining small aliquots. Finally, the position of seasonal extremes estimated from the $\delta^{18}O_c$ 675 record may not reflect the true seasonal extent if seasonal SST and $\delta^{18}O_w$ cycles are not in phase (as in 676 case 9), causing the seasonal Δ_{47} -based SST reconstructions to underestimate the temperature 677 seasonality. In such cases, $\delta^{18}O_c$ and Δ_{47} analyses on small aliquots allow the seasonality in SST and $\delta^{18}O_w$ 678 to be disentangled, yielding more accurate seasonality reconstructions.

679 Techniques for establishing independent age models for climate archives range from counting of growth layers or increments (Schöne et al., 2008; Huyghe et al., 2019), modelling and extracting of rhythmic 680 variability in climate proxies through statistical approaches (e.g. De Ridder et al., 2007; Goodwin et al., 681 682 2009; Judd et al., 2018; de Winter, 2021b) and interpolation of uncertainty on absolute dates (e.g. Scholz 683 and Hoffman, 2011; Meyers, 2019; Sinnesael et al., 2019). While propagating uncertainty in the data on which age models are based onto the age model is relatively straightforward, errors on underlying a priori 684 685 assumptions such as linear growth rate between dated intervals, (quasi-)sinusoidal forcing of climate cycles 686 and the uncertainty on human-generated data such as layer counting are very difficult to quantify (e.g. 687 Comboul et al., 2014) and may not be normally distributed. Results of cases 25-29 show that uncertainties

688 in the age domain can significantly compromise reconstructions (Fig. 12). Within the scope of this study, 689 only the effect of symmetrical, normally distributed uncertainties on an artificial case with phase decoupled SST and $\delta^{18}O_w$ seasonality (case 9) was tested. The effects of other types of uncertainties on the 690 691 reconstructions remain unknown, highlighting an unknown uncertainty in paleoseasonality and other high-692 resolution paleoclimate studies that may introduce bias or lead to over-optimistic uncertainties on 693 reconstructions. Future research could quantify this unknown uncertainty by propagating estimates of 694 various types of uncertainty on depth values of samples and on the conversion from sampling to time 695 domain in age models.

696 4.2 Conditions influencing success of reconstructions

697The reliability (accuracy and precision) of SST and $\delta^{18}O_w$ reconstructions depend on case-specific698conditions. The range of case studies tested in this study allowed us to evaluate the effect of variability in699SST, growth rate, $\delta^{18}O_w$, sampling resolution and record length relative to the control case (case 1; see700Supplementary Data S1). A summary of the effects of these changes is given in Table 2.

701

Variable	cases	Metric	Effect on reconstructions			
variable			δ ¹⁸ Ο	smoothing	binning	optimization
SST	12 15	Precision	0	+++	+	0
	17 19-21 30-33	Accuracy	+	+	0	+
Growth rate	2-6	Precision	+	++	++	+
	14-18 30-33	Accuracy	+	++	0	+
-18-	7-11	Precision	+	++	0	0
δ ¹⁸ O _w	13-18 30-33	Accuracy	+++	+++	+	++
Sampling resolution	1-33	Precision	0	+++	++	++
Sampling resolution		Accuracy	+	+	+++	+
Record length	22-24	Precision	0	0	+++	++
Record length		Accuracy	+	0	++	++
Age model	I 25-29	Precision	+++	++	0	++
uncertainty?		Accuracy	+	+	++	++

702 703 704

Table 2: Qualitative summary of the effects of changes in variables relative from the ideal case on reconstructions using the four approaches. The "cases" column lists cases in which the changes in the respective variable relative to the control case (case 1) were represented (see Table 1 and S1). "0" = 705 negligible effect, "+" = weak increase in uncertainty, "++" = moderate increase in uncertainty, "+++" = strong 706 increase in uncertainty. Precision and accuracy of all tests is given in S9.

707

708 4.2.1 SST variability

709 Variability in water temperature most directly affects the proxies under study. By default (case 1), SST 710 varies sinusoidally around a MAT of 20°C with an amplitude of 10°C (see 2.3.3, Fig. 2 and Supplementary 711 Data S1). In cases in which multi-annual variability in SST is simulated (e.g. case 15 and 17), the accuracy 712 of SST reconstructions using δ^{18} O and optimization are reduced, while the binning approach is less 713 strongly affected. Examples of such multi-annual cyclicity are El-Niño Southern Oscillation (ENSO; 714 Philander, 1983) or North Atlantic Oscillation (NOA; Hurrell, 1995). The effect is especially large in case 17, which simulates a tropical environment with reduced SST seasonality and a strong multi-annual cyclicity. 715 716 This type of environment is analogous to the environment of tropical shallow water corals, which are often 717 used as archives for ENSO variability (e.g. Charles et al., 1997; Fairbanks et al., 1997) and is similar to 718 tropical cases from the Australian Great Barrier Reef (case 31) and Red Sea (case 32; see Fig. 3). We 719 therefore recommend using the binning approach on carbonate records where multi-annual cyclicity is

720 prevalent and if a reliable age model can be established for these records (as in e.g. Sato, 1999; Scourse

721 et al., 2006; Miyaji et al., 2010).

722 4.2.2 Growth rate variability and hiatuses

723 Figures 7 and 8 show that variations in the growth rate of records, including the occurrence of hiatuses, 724 have a strong effect on reconstructions, especially using the smoothing approach. In general, hiatuses 725 and slower growth reduce precision of monthly SST and $\delta^{18}O_w$ reconstructions by reducing mean temporal 726 sampling resolution (samples/yr; see Fig. 9), and because parts of the record are undersampled. The effect 727 on accuracy depends strongly on the timing of changes in growth rate or the occurrence of hiatuses. Cases 728 2-6 simulate specific growth rate effects and can be used to test these differences. The smoothing method 729 is especially sensitive to changes in growth rate that take place in specific seasons, such as hiatuses in 730 winter (case 2) or summer (case 3) and growth peaks in summer (case 5) or spring (case 6). The other 731 reconstruction approaches are less affected by this bias, because they generally do not mix samples from 732 different seasons. The $\delta^{18}O$ method is especially well suited to deal with changes in growth rate because it does not require combining different aliquots for accurate SST reconstructions. The binning and 733 734 optimization approaches are slightly less reliable in cases where growth rate decreases linearly or 735 seasonally along the entire record (cases 4-6; Fig. 2). Because these two methods consider all samples in 736 the records at once, they are more sensitive to changes in temporal sampling resolution along the record. 737 It is worth noting that optimization is especially sensitive to sharp changes in growth rate in summer (e.g. 738 cases 11, 14, 16 and 17) because those conditions force the optimization routine to use larger sample 739 sizes or include samples outside the warmest month for summer temperature estimates. A potential solution 740 to this problem could be to allow sample sizes of summer and winter groups to vary independently in the 741 optimization routine (see 2.1). This would allow sample size in the undersampled season (in this case: 742 summer) to become larger than that at the other end of the $\delta^{18}O_c$ spectrum, reducing uncertainty on the 743 more densely sampled season and therefore improving the entire seasonality reconstruction.

A worst-case scenario is represented by case 18, where the cold half of the year is not recorded. Such cases result in strong biases in reconstructions of mean annual and seasonal ranges in SST and $\delta^{18}O_w$, regardless of which method is used. In such extreme cases the record simply contains insufficient information to reconstruct variability in growth rate, SST and $\delta^{18}O_w$, and it seems that no statistical method would enable this missing information to be recovered. The solution for these reconstructions would be to establish reliable age models, independent of $\delta^{18}O$ or Δ_{47} data, which show that a large part of the seasonal cycle is missing. All methods used in this study rely on a conversion of SST and $\delta^{18}O_w$ reconstructions to the time domain to define monthly time bins. This conversion breaks down in fossil examples when the seasonal cycle cannot be extracted from the archive, which happens when half of the seasonal cycle or more is obscured by growth hiatuses, as exemplified in case 18.

754 While hiatuses encompassing half of the seasonal cycle are uncommon, changes in growth rate are 755 common in accretionary carbonate archives because conditions for (biotic or abiotic) carbonate 756 mineralization often vary over time. This variability is either driven by biological constraints, such as 757 senescence (e.g. Schöne, 2008; Hendriks et al., 2012), the reproductive cycle (Gaspar et al., 1999) or 758 stress (Surge et al., 2001; Compton et al., 2007) or by variations in the environment that promote or inhibit 759 carbonate production, such as seasonal variations in temperature (Crossland, 1984; Bahr et al., 2017) or 760 precipitation (Dayem et al., 2010; Van Rampelbergh et al., 2014). In general, such conditions occur more 761 frequently in mid- to high-latitude environments than in low-latitudes, and in more coastal environments 762 rather than in open marine settings, because these environments contain stronger variations in the factors 763 that influence growth rates (e.g. temperature, precipitation or freshwater influx; e.g. Surge et al., 2001; 764 Ullmann et al., 2010). This difference was simulated in the cases representing natural variability (case 14-765 18 and 30-33). Accuracy in the coastal high-latitude settings (cases 16, 18 and 29) are indeed more strongly 766 affected by changes in growth rate. Because in such highly variable environments growth rate variability 767 often co-occurs with variability in $\delta^{18}O_w$, using $\delta^{18}O_c$ -based reconstructions is not advised, unless $\delta^{18}O_w$ 768 variability can be constrained or neglected (which is rare in these environments).

Additional complications include that the lack of constraint on growth rate variability because of uncertainties in the record's age model (see **4.1.3**) and the effect of growth rate variability on the sampling resolution. The effect of growth rate on time-averaging within samples was not specifically tested in this study but introduces uncertainty in practice when archives with variable growth rate are sampled at a constant sampling resolution in the depth domain. In this case, parts of the archive with a lower growth rate yield more time-averaged samples, potentially dampening one extreme of the seasonal cycle (e.g. Goodwin et al., 2003). In highly dynamic environments it is challenging to isolate all variables that introduce bias, and irregular variability in growth rate and $\delta^{18}O_w$ will invariably introduce uncertainty in SST reconstructions, even when applying the best Δ_{47} -based approaches (e.g. **binning** and **optimization**). In such examples, the results of natural variability cases (14-18 and 30-33) and of the real oyster data (**Fig. 6**) serve as benchmarks for the degree of uncertainty that may remain unexplained in these records.

780 4.2.3 Variability in δ¹⁸O_w

As discussed in 4.1.1, these variations in $\delta^{18}O_w$ have a large effect on the accuracy of $\delta^{18}O_c\text{-based}$ 781 782 reconstructions, and their occurrence constitutes the main advantage of applying the Δ_{47} thermometer 783 (Eiler, 2011). However, results of cases 7-11 in Fig. 8 and Table 2 show that $\delta^{18}O_w$ variations can also bias 784 Δ_{47} -based reconstructions, especially those of seasonal ranges and those using the **smoothing** approach. 785 Smoothing reconstructions are biased by these $\delta^{18}O_w$ shifts in much the same way as they are affected 786 by shifts in growth rate (see **4.2.2**). The **optimization** approach is sensitive to seasonal changes in $\delta^{18}O_w$ 787 in antiphase with SST seasonality and by increases in $\delta^{18}O_w$ in summer (e.g. due to excess evaporation; 788 e.g. case 11), especially when used for reconstructions of $\delta^{18}O_w$ seasonality. This effect arises because 789 the **optimization** approach orders data based on $\delta^{18}O_c$ and Δ_{47} seasonality to isolate the $\delta^{18}O_w$ -SST 790 relationship. Both antiphase $\delta^{18}O_w$ seasonality and summer evaporation dampen the seasonal $\delta^{18}O_c$ cycle and therefore influences the reconstruction of the $\delta^{18}O_w$ -SST relationship. A good example of this is seen 791 792 in the real oyster data (Fig. 6), where $\delta^{18}O_w$ and SST vary in phase and $\delta^{18}O_w$ dampens the SST 793 seasonality. The binning approach is more robust against 5¹⁸Ow variability that dampens the seasonal 794 cycle and is therefore a better choice for absolute SST reconstructions in environments where summer 795 evaporation or other $\delta^{18}O_w$ variability in phase with SST seasonality is expected to occur, if the age model 796 is reliable enough to allow monthly binning of raw data (see 4.1.3). Indeed, reconstructions from the 797 lagoonal environment (case 16) and Red Sea case (case 32 which is characterized by strong summer 798 evaporation; e.g. Titschack et al., 2010) show that binning is the most reliable choice in these 799 environments.

800 4.2.4 Variability in sampling resolution and record length

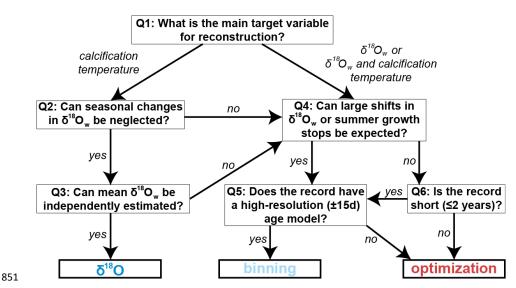
801 Other factors influencing the effectivity effectiveness of reconstructions are the sampling resolution and the 802 length of the record. Many of the cases discussed in this study represent idealized cases with comparatively 803 high sampling resolutions over comparatively long (12 yr) paleoseasonality records, which yield large 804 sample sizes. By comparison, the typical age of mollusks, which are often used as paleoseasonality 805 archives, is 2-5 years (Ivany, 2012). Records with the highest sampling resolutions (0.1 and 0.2 mm) contain up to 1200 samples. Generating such records is not impossible, but it is highly unlikely to be applied in 806 807 paleoclimate studies given the limitation of resources (e.g. instrument time) and the desire to analyze 808 multiple records from different specimens, species, localities or ages to gain a better understanding of the 809 variability in paleoseasonality (e.g. Goodwin et al., 2003; Schöne et al., 2006; Petersen et al., 2016). In 810 some cases large datasets are meticulously collected from single carbonate records (e.g. Schöne et al., 811 2005; Vansteenberge et al., 2016; de Winter et al., 2020a; Shao et al., 2020). However, in such studies, 812 the aim is often to investigate variability at a higher (e.g. daily; de Winter et al., 2020a) resolution or longer 813 timescales (e.g. decadal to millennial; Schöne et al., 2005; Vansteenberge et al., 2016; Shao et al., 2020) 814 in addition to the seasonal cycle, rather than to improve the reliability of reconstructing one type of variability 815 (e.g. seasonality) alone.

816 Fig. 9 shows that increasing temporal sampling resolution (samples/yr) improves both the accuracy and 817 precision of all Δ_{47} -based reconstructions. This occurs because Δ_{47} samples have a large analytical 818 uncertainty (see 4.1.2) and grouping of data therefore improves reconstructions. The decrease in precision 819 of $\delta^{18}O_c$ -based reconstructions (Fig. 9C-D) is explained by the fact that the analytical uncertainty of $\delta^{18}O_c$ 820 measurements is much smaller than the variability introduced by natural sub-annual variability in SST and 821 $\delta^{18}O_w$ unrelated to the seasonal cycle (see **<u>Supplementary Data</u>S4**). Therefore, higher sampling 822 resolutions allow $\delta^{18}O_c$ records to better capture this sub-seasonal variability, which introduces more noise 823 on to the seasonal cycle (reducing precision) but causes monthly mean SST and $\delta^{18}O_w$ to be more 824 accurately reconstructed. Towards higher sampling resolutions, the gap in precision between $\delta^{18}O_{c}$ - and 825 Δ47-based reconstructions closes, eventually (in an ideal case) diminishing the advantage of high analytical 826 precision in $\delta^{18}O_c$ measurements (Fig. 9C-D).

827 An optimum sample resolution can be defined for each method after which improving sampling resolution 828 does not significantly improve the reliability of the reconstruction (as in de Winter et al., 2017). Figure 9 shows that this optimum varies depending on which variable (MAT, SST seasonality, mean annual $\delta^{18}O_w$ 829 830 or $\delta^{18}O_w$ seasonality) is reconstructed. Therefore, Fig. 9 will allow future researchers to determine the 831 sampling resolution that is tailored to their purpose. In general, the improvement after a sample size of 20-30 samples per year is negligible for the binning and optimization methods if the total number of samples 832 833 (depending on both sampling resolution and record length) is sufficient for monthly temperature 834 reconstructions. Our data show that 200-250 paired $\delta^{18}O_c$ and Δ_{47} measurements are in general sufficient 835 for a standard deviation of 2-3°C on monthly SST reconstructions using the binning or optimization 836 approach, preferably when spread over multiple growth years to eliminate the effect of short-term weather 837 events or years with exceptional seasonality (Fig. 10; Supplementary Data S5).

838 Record length only has a minimal influence on the optimization method but for very short records (≤2 839 years) binning becomes very imprecise, especially at low sampling resolutions (Fig. 11). The reason is 840 that the sample size within monthly time bins becomes too small in these cases, while the more flexible 841 sample size window of the optimization routine circumvents this problem. The choice between these two 842 approaches should therefore be based on a tradeoff between the length of the record (in time) and the 843 number of samples that can be retrieved from it. As a result, shorter-lived, fast-growing climate archives, 844 such as large or fast-growing (e.g. juvenile) mollusk shells, are best sampled using a high temporal 845 resolution (>30 samples/yr) sampling strategy with the optimization approach. Longer lived archives with 846 a lower mineralization rate, such as annually laminated speleothems, corals and gerontic mollusks, are 847 best sampled using long time series at monthly resolution using the **binning** approach.

A simplified decision tree that could guide sampling strategies for future paleoseasonality studies is shown in **Figure 14**. Note that choices and tradeoffs for these reconstructions may differ depending on the archive and environment in which it formed (see discussion above).



Schematic guide to reconstructing SST and $\delta^{18}O_w$ from accretionary carbonate archives

852 Figure 15: Schematic guide to choosing the right approach for reconstructing annual mean or seasonality 853 in SST and $\delta^{18}O_w$ from accretionary carbonate archives. Recommendations are based on the results of testing all four approaches on the entire range of cases. Researchers can follow the six steps (questions 854 855 Q1-6) to decide on the right approach for reconstructing the target variable. Guidelines are based on minimizing maximizing both accuracy and precision (see details in Supplementary Data S9). Note that the 856 857 smoothing approach is never the best choice. The choice between the two remaining Δ_{47} -based 858 approaches (binning and optimization) relies heavily on the situation and may be driven by a preference for more accurate or more precise results. 859

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861 **4.3 Implications for clumped isotope sample size**

The **optimization** technique for grouping Δ_{47} aliquots for accurate SST and $\delta^{18}O_w$ reconstructions allows us to assess the limitations of the clumped isotope thermometer for temperature reconstructions from highresolution carbonate archives. The optimal sample size given by the approach is different for different cases and depends on the temporal sampling resolution and the characteristics of the record (see **S4**). As expected, in cases more like the ideal case (case 1), optimal sample sizes are low (~14-24), while sample sizes increase in more complicated cases based on simulated natural environments (case 14-18) or cases based on actual SST and SSS data (cases 30-33). More confined SST seasonality (cases 19-21) also Formatted: Font: Italic

869 requires larger samples to reconstruct (up to 100 samples in some cases). This is not surprising, because 870 variability within samples will increase in records in which the seasonality is smaller or more obscured by 871 other environmental variability. The optimal sample size between cases and sampling resolutions is not 872 normally distributed but tails towards high sample sizes with some extreme outliers (Shapiro Wilk test p << 873 0.05; Supplementary Data S10). The median sample size of all our simulations is 17 aliquots. This number 874 lies between the minimum number of 14 ~100 μg replicates of standards calculated by Fernandez et al. 875 (2017) and the minimum of 20-40 ~100 µg aliquots required for optimal paleoseasonality reconstruction 876 from fossil bivalves by de Winter et al. (2020b). This is to be expected since many of the cases explored in 877 this study represent ideal cases compared with the natural situation. However, in these virtual cases a 878 measure of random sub-annual variability in SST and $\delta^{18}O_w$ was added (see Fig. 4 and Supplementary 879 Data S2), simulating a more realistic environment and resulting in poorer precision than replicates of a 880 carbonate standard (as in Fernandez et al., 2017). Our simulations show that the optimum number of 881 samples to be combined in seasonality studies depends on both the analytical uncertainty of Δ_{47} 882 measurements (as represented by the estimate in Fernandez et al., 2017) and the variability between 883 aliquots pooled within a sample that is attributed to actual variability within the record (as represented by 884 our simulations and the estimate in de Winter et al., 2020b). The optimal sample size is therefore a good 885 measure for the limitations of temperature variability that can be resolved in a record and can help 886 researchers decide which strategy to apply for combining measurements to obtain the most reliable 887 paleoseasonality estimates, or to decide whether extra sampling is required, even if the chosen approach 888 is not to use the optimization routine itself. Note that the optimum sample size is kept equal for summer and winter samples in this study, and that the optimization approach can likely achieve better performance 889 890 by considering unequal sample sizes in opposite seasons (see 4.1.3 and 4.2.2). While this added flexibility 891 comes at a higher computational cost due to the increased number of possible sample size combinations 892 to be considered, future studies should investigate whether this updated optimization approach could yield 893 more reliable seasonality reconstructions.

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894 4.4 Implications for other sample size problems

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895	While the discussion above focuses on optimizing approaches for combining samples for clumped
896	isotope analyses in paleoseasonality reconstructions, the problem of combining samples to reduce
897	uncertainty and isolate variation in datasets is very common (e.g. Zhang et al., 2004; Merz and Thieken,
898	2005; Tsukakoshi, 2011). Therefore, the approaches outlined and tested in this study have applications
899	beyond paleoseasonality reconstructions. Examples of other problems that could benefit from applying
900	similar approaches for reducing the uncertainty of estimates of target variables while minimizing the
901	number of analyses required to meet analytical requirements include: (1) reconstructing
902	paleoenvironmental variability in the terrestrial realm from tooth bioapatite (e.g. Passey and Cerling,
903	2002; Kohn, 2004; Van Dam and Reichart, 2009; de Winter et al., 2016), (2) quantitative time series
904	analysis of orbital cycles in stratigraphic records (e.g. Lourens et al., 2010; de Vleeschouwer et al., 2017;
905	Noorbergen et al., 20172018; Westerhold et al., 2020), (3) strontium isotope dating (e.g McArthur et al.,
906	2012; de Winter et al., 2020c), (4) reconstructing sub-seasonal variability from ultra-high-resolution
907	records (e.g. from fast-growing mollusks and gastropods; e.g. Sano et al., 2012; Warter and Müller, 2017,
908	de Winter et al., 2020d; Yan et al., 2020), and (5) reconstructing sea surface and deep-sea temperatures
909	across short-lived (10–100 kyr) episodes of climate change or climate shifts from deep marine archives
910	characterized by low sedimentation rates (e.g. Lear et al., 2008; Jenkyns, 2010; Stap et al., 2010;
911	Lauretano et al., 2018). A more detailed discussion of the implications for other sample size problems is

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provided in the Supplementary Information.

914 5. Conclusions and recommendations

915 The performance of three Δ₄₇-based approaches to reconstruct seasonality from accretionary carbonate 916 archives was evaluated in comparison with conventional 518Oc-based reconstructions in a wide range of 917 case studies. From the results, we conclude that while $\delta^{18}O_c$ -based reconstructions ($\delta^{18}O$) yield superior 918 precision for SST reconstructions, this method runs a high risk of yielding inaccurate results due to innate 919 assumptions about the value of $\delta^{18}O_w$, which must be estimated and assumed constant year-round. Unless 920 $\delta^{18}O_w$ can be independently constrained or variability in $\delta^{18}O_w$ can be neglected, Δ_{47} -based reconstructions 921 should be the method of choice for absolute mean annual temperature and SST seasonality 922 reconstructions. Various techniques for combining Δ_{47} data were evaluated. Our findings suggest that 923 smoothing Δ_{47} data using a moving average almost allways cases results in a dampening of the seasonal 924 cycle which severely hampers recovery of seasonality. Applying the smoothing approach results in 925 inaccuracies in reconstructions of MAT as well, especially in cases where part of the seasonal cycle is 926 obscured by variability in growth rate or multi-annual trends. More reliable seasonality reconstructions are 927 achieved with two approaches for combining Δ_{47} data using time binning (**binning**) or applying a flexible 928 sample size optimization (optimization) approach. Of these two approaches, optimization achieves better 929 precision and can resolve smaller seasonal temperature differences with confidence. However, binning is 930 often more accurate, and outperforms optimization as the most reliable approach. This is especially true 931 in cases with growth stops or $\delta^{18}O_w$ changes in phase with temperature seasonality (e.g. strong seasonal 932 evaporation or freshwater influx) and in longer multi-annual time series with a reliable age model. 933 Optimization is the better choice for shorter (<3 years) records, especially if the sampling resolution can 934 be increased, such as in short, fast growing climate archives.

Despite the focus on the problem of resolving seasonality in carbonate archives, the findings in this study have applications for other problems in earth science where sample size and sampling resolution put limits on the ability to resolve specific trends, events, and cycles from time series. While the above-mentioned recommendations of the **optimization** and **binning** methods are likely valid for most studies aiming to quantify the mean and amplitude of a specific cycle or event (equivalent to MAT and SST seasonality),

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940 (dynamic) moving averages (smoothing) are expected to yield the best results in studies quantifying941 aperiodic trends from longer data series.

942

943 Code availability

All scripts used to make the calculations described in this study are compiled in the documented R package "seasonalclumped", which is freely available on the open-source online R-database CRAN (de Winter, 2021a; <u>https://cran.r-project.org/web/packages/seasonalclumped</u>). Annotated R scripts used to make calculations for this study are available in the digital supplement uploaded to the open-source online repository Zenodo (<u>www.doi.org/10.5281/zenodo.3899926</u>).

949

950 Data availability

951 Supplementary data, figures and tables as well as all scripts used to do the calculations and create the 952 virtual datasets used in this study are deposited in the open-source online repository Zenodo 953 (www.doi.org/10.5281/zenodo.3899926). Virtual datasets generated within the context of this study are also 954 made available as datafiles within the R package that contains the scripts used for this study 955 ("seasonalclumped"; de Winter, 2021a; see https://cran.r-project.org/web/packages/seasonalclumped).

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957 Author contributions

958 NJW designed the study, wrote the scripts for all calculations, and created a first draft of the manuscript 959 text and figures. MZ, TA and NJW worked together from the first draft towards the final manuscript. All 960 authors contributed to the representation of the data and methods in figures and to the discussion of the 961 implications of the data in the discussion.

962

963 Competing Interests

964 The authors have no potential conflicts of interest to declare with regards to this study.

965

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