

Early summer hydroclimatic signals were well captured by tree-ring earlywood width in the eastern Qinling Mountains, central China

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10 **Abstract.** Tree-ring width (TRW) chronologies could only provide limited amount of moisture-related climatic information in the humid and semi-humid regions of China; thus, it is worth to explore the potentials of the intra-annul tree-ring width indices (i.e., the earlywood width (EWW) and latewood width (LWW)) to provide some additional climatic information. To fulfil this task, TRW, EWW and LWW were measured from the tree-ring samples of *Pinus tabulaeformis* in a semi-humid region, that is, the eastern Qinling Mountains, central China. Their standard (STD) and signal-free (SSF) chronologies were

15 created using different detrending methods including (1) negative exponential function together with linear regression with negative (or zero) slope (NELR), (2) cubic smoothed splines with a 50% frequency cutoff of 67% of the series length (SP67), and (3) age-dependent splines with an initial stiffness of 50 years (SPA50). The results showed that EWW chronologies were significantly negatively correlated with temperature, but positively correlated with precipitation and soil moisture conditions during the current early growing season. Comparatively, LWW and TRW chronologies had weaker relationships with these

20 climatic factors. EWW STD chronology with the detrending method of NELR contained the strongest climatic signal, explaining 50% variance of the May–July self-calibrated Palmer Drought Severity Index (MJJ scPDSI) during the instrumental period 1953–2005. Based on this relationship, the MJJ scPDSI was reconstructed back to 1868 using a linear regression function with strong statistical parameters, and the reconstruction was further validated by comparing with other hydroclimatic reconstructions and historical document records in adjacent regions. This highlighted the potentials of intra-annual tree-ring

25 indices for seasonal hydroclimatic reconstruction in humid and semi-humid regions of China. Furthermore, the reconstruction exhibited a strong in-phase relationship with a newly proposed East Asian summer monsoon index (EASMI) before the 1940s on the decadal and longer timescales, which may be due to the positive response of the local precipitation to EASMI. However, the cause for the weakened relationship after the 1940s is complicated, and cannot be solely attributed to the impacts of precipitation and temperature.

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1 Introduction

Most of the existing tree-ring width (TRW) based hydroclimatic reconstructions have been conducted in the regions between the 200-to 600-mm annual precipitation isolines in China (Liu et al., 2018b), close to the northern fringe of Asian Summer Monsoon. Comparatively, there are still a small amount of hydroclimatic reconstructions in the core monsoon region, for example, a few case studies in Southeast China (e.g. Cai et al., 2017; Chen et al., 2016a; Shi et al., 2015), North China (e.g. Chen et al., 2016b; Hughes et al., 1994; Lei et al., 2014; Liu et al., 2002), and the Hengduan mountains in Southwest China (Fan et al., 2008; Fang et al., 2010b; Gou et al., 2013; Li et al., 2016). Since precipitation is spatially variable (Ding et al., 2013), hydroclimatic variations in the monsoon fringe can not completely represent those in the core monsoon region (Liu et al., 2018b). Thus, more hydroclimatic reconstructions are needed in the monsoon region.

Some TRW chronologies within the monsoon region showed weak or unstable hydroclimatic signals (e.g. Li et al., 2016; Shi et al., 2012; Wang et al., 2018), unsuitable to be used for reliable reconstruction. Intra-annually resolved tree-ring width (i.e., earlywood width (EWW) and latewood width (LWW)), however, provided stronger hydroclimatic signals than TRW in some cases (Chen et al., 2012; Zhao et al., 2017a). This might be related to the seasonal movement of monsoon rainbelt which causes restrictions of water availability through the growing season (Liu et al., 2018a). Using EWW and LWW, researchers have successfully conducted reconstructions for seasonal rainfall (Hansen et al., 2017), standardized precipitation evapotranspiration index (SPEI; Zhao et al., 2017b), and streamflow (Guan et al., 2018) within the monsoon region.

The eastern Qinling Mountains are located within the core region of East Asian summer monsoon (EASM), and are characterized by a transitional climate from warm-temperate to subtropical. In this region, Shi et al. (2012) has built four TRW chronologies of *Pinus tabulaeformis* along an elevation gradient from 1200 to 1950 m above sea level (a.s.l.). The TRW chronologies from the two low-altitude sites, Baiyunshan and Longchiman, exhibited a positive response to precipitation and negative response to temperature during early summer, showing some kind of water stress. However, the dendroclimatic potentials of EWW and LWW were not explored.

Meanwhile, tree-growth at an adjacent site was found to be more restricted by the drought index, PDSI (Palmer Drought Severity Index), than precipitation and temperature (Peng et al., 2014). The PDSI monitors the cumulative departure of surface water balance in terms of the difference between the amount of precipitation required to retain a normal water-balance level and the amount of actual precipitation (Palmer, 1965; Wells et al., 2004). In addition, the SPEI, which represents a water balance as the form of difference between precipitation and potential evapotranspiration (Vicente-Serrano et al., 2010), has also been found to strongly limit latewood growth at a well-drained site in the core region of EASM (Zhao et al., 2017a, b). Therefore, the PDSI and SPEI should also be incorporated into the climate-tree-growth relationship assessment.

Recently, a new East Asian summer monsoon index (EASMI) was proposed by Zhao et al. (2015) based on the 200 hPa zonal wind anomalies, which shows a better performance in describing precipitation and temperature variations over East Asia than previous indices. When a strong EASM occurred, the new EASMI indicates abundant precipitation and relatively low temperature over the core region of EASM, particularly around the area of mei-yu/changma/baiu rainband (27.5°–32.5° N,

105°–120° E and 30°–37.5° N, 127.5°–150° E; Fig. 1a), which is the center of the leading mode of EASM precipitation (Wang et al., 2008; Zhao et al., 2015). It is worth to study the response of local hydroclimate to EASM from a long-term perspective in combination with these tree-ring materials and the new EASMI.

Since the TRW of *P. tabulaeformis* in Baiyunshan and Longchiman were mainly restricted by the early summer moisture condition, we hypothesize that the early summer hydroclimatic signals might be strengthened only using EWW, with the exclusion of LWW from TRW. Therefore, the objectives of this study are (1) to verify that EWW is more sensitive to early summer hydroclimatic factors than TRW and LWW for *P. tabulaeformis* at Baiyunshan and Longchiman, (2) to reconstruct early summer hydroclimate variations using EWW, and (3) to tentatively explore the relationship between the reconstructed hydroclimate variability and EASMI of Zhao et al. (2015).

2 Materials and Methods

2.1 Study sites

Dated tree-ring samples of *P. tabulaeformis* used in this study were previously presented in Shi et al. (2012). They were collected from two sampling sites in Mount Funiu in 2006 and 2008 separately: Baiyunshan (33.63° N, 111.85° E) and Longchiman (33.68° N, 112.05° E; Fig. 1b). The sampling sites are located on mountain tops, where soils are thin and well-drained. The elevations of Baiyunshan and Longchiman range from 1200–1300 m a.s.l., and 1340–1400 m a.s.l., respectively. The regional annual mean temperature and annual total precipitation are 14.1 °C and 822 mm, respectively. The majority part of the annual total precipitation drops during the warm season (Fig. 2). More detailed information of the study sites can be found in Shi et al. (2012).

2.2 Tree-ring data

The selected *P. tabulaeformis* is a widely distributed conifer species in North China with the extension from 31° 00' N to 43° 33' N, and 103° 20' E to 124° 45' E (Xu et al., 1981). Liang et al. (2009) studied the cambial dynamics of *P. tabulaeformis* in its northern distribution limit (43° 14.11' N, 116° 23.60' E, 1363 m a.s.l.), and found that the cell division in the cambial zone started within the third week of May and did not complete around mid-September. Zeng et al. (2018) found that the cambial cells of mature *P. tabulaeformis* in Northwest China (37° 02' N, 104° 28' E, 2456 m a.s.l.) started activity in late spring and ceased in late July to early August. Considering that our sampling sites are located at lower latitudes, the cambial activities of *P. tabulaeformis* in our study may start earlier and ends later than those found in above studies according to the temperature-controlled phenology theory (Chen and Xu, 2012).

P. tabulaeformis generally exhibits an abrupt transition from light-colored earlywood to dark-colored latewood (Liang and Eckstein, 2006; Fig. S1), and the transition can occur in mid-July in Beijing (39.9° N, 116.3° E; Zhang et al., 1982). Due to the characteristics of the ring anatomy, the earlywood and latewood segments of annual growth rings can be discriminated visually by the sudden change in cell size, lumen size, and color (Stahle et al., 2009). However, gradual transitions also occur

in a few samples, making the earlywood-latewood boundary difficult to discern. Therefore, only samples with distinct earlywood and latewood segments were used for subsequent measurements (Knapp et al., 2016). In total, 20 cores from 11 trees and 42 cores from 22 trees were selected from Baiyunshan and Longchiman, respectively. EWW and LWW were then measured using a LINTAB5 system at a resolution of 0.001 mm, and TRW was obtained by adding EWW and LWW together.

5 2.3 Development of tree-ring width chronologies

Non-climatic growth trends need to be fitted and removed from each “raw” (untreated) EWW, LWW and TRW series, which is known as detrending (Cook et al., 1990). In order to check the effects of detrending methods on the preservation of climatic signals, three detrending methods were selected for comparison. They were negative exponential function together with linear regression with negative (or zero) slope (NELR), cubic smoothed splines with a 50 % frequency cutoff of 67 % of the series length (SP67), and age-dependent splines with an initial stiffness of 50 years (SPA50). NELR is a deterministic method based on the assumption that tree radial growth declines monotonically (Cook et al., 1990). SP67 has a good ability in fitting the potential low-and middle-frequency perturbations contained in ring-width series (Cook et al., 1990). It allows no more than half of the amplitude of variations with wavelength of two-thirds of the length of series being preserved in resulting indices (Melvin et al., 2007). SPA50 specifies annually varying 50 % frequency cutoff parameter for each year by adding the initial stiffness with ring age. In comparison with SP67, it makes the resulting spline become more flexible in the early years and progressively stiffer in later years (Melvin et al., 2007). All raw ring-width series were divided by the estimated growth trends, and the resulting detrended ring-width series were averaged to generate the standard (STD) chronologies using the bi-weight robust mean method (Fig. S2). Since the traditional fitted curves may contain the climatic signals, which is termed as “trend distortion” problem (Melvin and Briffa, 2008), the signal-free (SSF) method is introduced to create the fitted growth curve free of climatic signals by dividing the raw ring-width series by the STD chronology by means of iterations (Melvin and Briffa, 2008). Therefore, the SSF chronologies were also developed for analysis (Fig. S3). The variance of each chronology was stabilized to minimize the effects of sampling depth according to the methods described in Osborn et al. (1997). The temporal extension for all width chronologies in Baiyunshan and Longchiman are 1841–2005 and 1850–2005, respectively. All above processes were performed with the program RCSsigFree Version 45_v2b (Melvin and Briffa, 2008; <http://www.ldeo.columbia.edu/tree-ring-laboratory/resources/software>). Since the number of cores per tree in our study are unequal, the signal of each chronology was estimated by the effective chronology signal ($R_{\text{bar}_{\text{eff}}}$) which incorporates the effective number of cores, within- and between-tree signals (Briffa and Jones, 1990). Besides, the expressed population signal (EPS), a function of $R_{\text{bar}_{\text{eff}}}$ and the number of trees, was used to estimate how well the sample chronology represents the theoretical chronology (Briffa and Jones, 1990; Wigley et al., 1984). The running $R_{\text{bar}_{\text{eff}}}$ and EPS for each chronology were calculated over a 51-year window with a 50-year overlap using the function “rwi.stats.running” in R package “dplR” version 1.6.9 (Bunn et al., 2018). The minimum number of common years in any pair of ring-width series required for their correlation was set to 30 (Briffa and Jones, 1990). The reliable period for each chronology was determined based on the generally accepted EPS threshold value of 0.85 (Wigley et al., 1984).

The width chronologies from the two sampling sites show high degree of coherence as evidenced by their significant positive correlations ($p < 0.001$) during their common period 1850–2005 (Table S1). Moreover, the positive correlations remain significant ($p < 0.001$) after removing the influence of autocorrelations and linear trends (Table S1). This indicates that the two sites share common climatic signals. Therefore, we pooled all raw ring-width series from the two sites, and developed composite STD and SSF chronologies for EWW, LWW and TRW using the three detrending methods as described above (Fig. S4). Statistics for each chronology including the starting year when $EPS \geq 0.85$, standard deviation, mean sensitivity and first-order correlation coefficient (AR1) are shown in Table S2. In addition, several statistics were calculated to assess the degree of similarity among detrended ring-width series over the common period 1915–2005 (Table S3). These statistics are variance explained by the first eigenvector (Var_{pc1}), $Rbar_{eff}$, signal-to-noise ratio (SNR), and EPS (Briffa and Jones, 1990; Trouet et al., 2006).

2.4 Climate data

Monthly mean maximum (Tmax), minimum (Tmin) and mean temperature (Tmean), and monthly total precipitation (Pre) were selected from four nearby meteorological stations (Table 1; Fig. 1b). These climate data were obtained from the China Meteorological Administration. Regional temperature values were calculated by directly averaging the temperature time series from the four stations over their common period 1957–2005. Regional precipitation was produced by firstly deriving the regional averages in terms of percentages, then multiplying the regional mean to transform the resulting series back to millimetre units (Jones and Hulme, 1996). The self-calibrated PDSI (scPDSI) and SPEI were also chosen as the hydroclimatic factors. Here we used the scPDSI instead of PDSI because it has solved the PDSI problems in spatial comparisons by calculating the duration factors (weighting coefficients for the current moisture anomaly and the previous drought severity) based on the characteristics of the climate at a given location (Wells et al., 2004). The regional scPDSI was calculated by averaging the CRU (Climate Research Unit) scPDSI grids (van der Schrier et al., 2013) over the area between 32° N to 34.5° N and 111° E to 112° E (Fig. 1b) where the meteorological stations utilized by CRU dataset were concentrated (Fig. S5; Table S4). The time span of scPDSI was selected as 1953–2005, because the stations used by CRU have a common period starting from July 1952 (Table S4). The SPEI has multi-timescales (Vicente-Serrano et al., 2010). Therefore, we calculated the regional SPEI at three timescales (1-month, 3-month and 12-month) using the R package “SPEI” version 1.7 (Beguería and Vicente-Serrano, 2012). The climatic factors used in SPEI calculation included regional Tmax, Tmin and Pre, which were derived from the four stations mentioned in Table 1. The time span of SPEI is 1957–2005.

In order to validate the reconstruction, we compared it with several hydroclimate time series and historical document records (Table 2), including (1) the June–August PDSI from the No. 370 grid point of the Monsoon Asia Drought Atlas (MADA) at 33.75° N, 111.25° E over the period 1868–2005 ($PDSI_{Cook}$; Cook et al., 2010), (2) the dryness/wetness index (DWI) from the grid point at 33.75° N, 111.25° E over the period 1868–2000 (DWI_{Yang} ; Yang et al., 2013), (3) reconstructed April–June precipitation based on TRW in Mount Hua over the period 1868–2005 (Pre_{Chen} ; Chen et al., 2016b), and (4) drought/wet events recorded in historical documents over the period 1868–2005 (He, 1980; Wen, 2006). The DWI dataset were reconstructed

from the historical documents and modern instrumental May–September precipitation in 120 sites over China (Chinese Academy of Meteorological Sciences, 1981). The dataset classified the degree of dryness and wetness into five grades: very wet (grade 1), wet (grade 2), normal (grade 3), dry (grade 4), and very dry (grade 5). Yang et al. (2013) has interpolated the DWI dataset into 2.5° latitude/longitude grid cells.

- 5 The EASM circulation was represented by a newly defined EASMI based on the 200 hPa zonal wind anomalies which was less affected by complex weather processes near the surface (Zhao et al., 2015). It was computed as Equation (1):
- $$\text{EASMI} = \text{Nor} [u(2.5^\circ - 10^\circ \text{ N}, 105^\circ - 140^\circ \text{ E}) - u(17.5^\circ - 22.5^\circ \text{ N}, 105^\circ - 140^\circ \text{ E}) + u(30^\circ - 37.5^\circ \text{ N}, 105^\circ - 140^\circ \text{ E})] \quad (1)$$
- where Nor and u represent standardization and mean 200 hPa zonal wind, respectively. To understand the possible impacts of local precipitation and temperature (32°–34.5° N and 111°–112° E) on the relationship between the scPDSI and EASMI, the
- 10 precipitation and temperature data were extracted from the gridded precipitation dataset Global Precipitation Climatology Centre Version 7 (GPCC v7; Schneider et al., 2015), and gridded temperature dataset Climatic Research Unit Time-Series Version 4.01 (CRU TS 4.01; Harris et al., 2014). The gridded dataset can represent the variations of precipitation and temperature over East China during the 20th century (Wang and Wang, 2017; Wen et al., 2006).

2.5 Statistical methods

- 15 To investigate the climate response of different tree-ring parameters (EWW, LWW, and TRW), we firstly calculated the Pearson correlation coefficients of the STD and SSF tree-ring width chronologies with monthly climate time series. The time window for the correlation analysis spanned from January of two years earlier to October of the current year. Secondly, correlations were calculated between the prewhitened and linearly detrended chronologies and climate time series to evaluate the possible effects of autocorrelations and secular trends. The prewhitening procedure was run with the “ar” function in R
- 20 package “stats” version 3.5.1 (R Core Team, 2018). The appropriate autoregressive order was automatically determined by the Akaike Information Criterion (Akaike, 1974). The linear detrending procedure was performed in Matlab R2016a with the “detrend” function (The MathWorks, 2016). Then, we analyzed the response of different tree-ring parameters to multi-month averaged scPDSI (which had the stronger impacts on tree-growth than other climatic factors; see the results for detail) to find the strongest climate-growth relationship. Finally, we used the wavelet coherence method (Grinsted et al., 2004) to test the
- 25 temporal stability and possible lags of the climate-growth relationship on different frequency domain.

- A simple linear regression model was applied to establish the transfer function using May–July (MJJ) scPDSI as the predictant, and the NELR based EWW STD chronology as the predictor (which had the strongest relationship; see the results for detail) over the period 1953–2005. Temporal stability of the model was tested by splitting the MJJ scPDSI into two sub-periods (1953–1979 and 1979–2005) for calibration and verification using the following statistics: correlation coefficient (r), explained
- 30 variance (R^2), reduction of error (RE), coefficient of efficiency (CE), and the sign-test (Meko and Graybill, 1995). Meanwhile, the possible autocorrelation and trend contained in the regression residuals were evaluated using the Durbin-Watson test (DW ; Durbin and Watson, 1950) with the “dwtest” function in R package “Lmtest” version 0.9-36 (Hothorn et al., 2018), and the two-sided Cox and Stuart trend test (CS ; Cox and Stuart, 1955) with the R package “snpar” version 1.0 (Qiu, 2014),

respectively. A *DW* value of 2 means no first order autocorrelation in the residuals, whereas values larger (less) than 2 are indicative for negative (positive) autocorrelation. The *DW* test has the null hypothesis that the autocorrelation of the residuals is 0. The two-sided *CS* trend test has the null hypothesis that there is no monotonic trend in the residuals. The variance of the MJJ scPDSI reconstruction was adjusted to match the variance of instrumental MJJ scPDSI during the calibration period using

Equation (2),

$$\text{Adj_Rec}_i = \frac{(\text{Rec}_i - \overline{\text{Rec}_{\text{cal}}})}{\sigma(\text{Rec}_{\text{cal}})} \times \sigma(\text{Ins}_{\text{cal}}) + \overline{\text{Ins}_{\text{cal}}} \quad (2)$$

where, the Rec_i and Adj_Rec_i indicate the reconstructed value and its variance adjusted value for a specific year i . The $\overline{\text{Rec}_{\text{cal}}}$ and $\overline{\text{Ins}_{\text{cal}}}$ indicate the arithmetic mean of the reconstructed and instrumental values during the calibration period (it is 1953–2005 in this study). The $\sigma(\text{Rec}_{\text{cal}})$ and $\sigma(\text{Ins}_{\text{cal}})$ are the corresponding standard deviations.

Spatial correlations were calculated between the reconstructed MJJ scPDSI and CRU scPDSI 3.25 dataset (van der Schrier et al., 2013) using the KNMI Climate Explorer (<http://climexp.knmi.nl/start.cgi>) to investigate the spatial representativeness of our reconstruction. All the hydroclimatic reconstructions were divided into interannual (< 10 years), and decadal and longer-term components (> 10 years) for comparison, respectively. The decadal and longer components were derived by lowpass filtering the original reconstructions using the the adaptive 10 point “Butterworth” low-pass filter at 0.1 cut-off frequency (Mann, 2008). Then, the interannual components were obtained by subtracting the decadal and longer-term components from the original reconstructions. The low-pass filtering technique has a good ability in preserving trends near time series boundaries (Mann, 2008).

We calculated the MJJ EASMI according to the definition of Zhao et al. (2015) using the 200 hPa zonal wind dataset which were obtained from the National Oceanic and Atmospheric Administration-Cooperative Institute for Research in Environmental Sciences Twentieth Century Reanalysis V2c (NOAA-20C; Compo et al., 2011) over the period 1868–2005. The relationship between EASMI and our reconstruction was firstly evaluated using the wavelet coherence method (Grinsted et al., 2004). In addition, 21-year moving window correlation analyses were calculated between the decadal-filtered MJJ EASMI, reconstructed scPDSI, and local precipitation to explore the connections of precipitation with scPDSI and EASMI. Moreover, empirical orthogonal function (EOF) analysis and spatial correlation analysis were performed to assess the impacts of the changed leading EASM mode on the relationship between decadal-filtered EASMI and local precipitation. The filtering procedure was conducted using the “Butterworth” low-pass filter (Mann, 2008) as mentioned above. The filtering, EOF and correlation analyses were performed in Matlab R2016a (The MathWorks, 2016) and the plots were drawn with Surfer 10 (Golden Software, 2011).

The significance tests for all observed correlation coefficients were conducted using Monte Carlo method (Efron and Tibshirani, 1986). In detail, modelled time series with the same structure as the original series were produced according to the frequency domain method of Ebisuzaki (1997). Then, correlation coefficients were computed between the modelled time series. The above processes were repeated 1000 times to obtain 1000 modelled correlation coefficients. The significance threshold was

estimated based on the probability distribution of the modelled 1000 correlation coefficients. The procedure was performed using the algorithms of Macias-Fauria et al. (2012).

3 Results and Discussion

3.1 Stronger hydroclimatic signals derived from EWW

- 5 As shown in Fig. 3, the EWW chronologies generated using different detrending and standardization methods were significantly negatively correlated with Tmax, Tmean during May–June, and significantly positively correlated with Pre in May. In terms of the drought indices, all EWW chronologies were significantly positively correlated with the 1-month SPEI in May, 3-month SPEI during May–June, 12-month SPEI during May–October, and scPDSI during April–October. It can be found that EWW showed a much longer-term response to the multi-month SPEI and scPDSI than to precipitation after May.
- 10 This may be because the summer temperatures still affected the soil water status as reflected by their negative correlations with EWW. Besides, soil has a memory effect on previous drought conditions, and this effect was considered by the multi-month SPEI and scPDSI (Vicente-Serrano et al., 2010; Dai, 2011). The scPDSI had higher correlations with tree-ring width than the SPEI. This indicates that the scPDSI has a better ability than the SPEI in monitoring the influence of soil moisture status on tree growth in our sampling sites, and the reasons remain unknown. However, the significant correlations found
- 15 between EWW and drought indices during autumn should not be regarded as a real drought impact, as the earlywood growth would terminate in the mid- and late-growing season (Larson, 1969). For LWW, during the current growing season, the highest correlation was found between the NELR based LWW STD and July scPDSI ($r = 0.37$, $p < 0.01$), which was much lower than that found between the NELR based EWW STD and the July scPDSI ($r = 0.62$, $p < 0.01$), indicating that LWW had less sensitivity to the scPDSI. TRW generally exhibited the similar climate response as EWW but with relatively lower correlations.
- 20 Taking the NELR based STD chronologies as an example, the correlation coefficients between TRW and the monthly scPDSI from May to July were 0.59 ($p < 0.01$), 0.58 ($p < 0.01$), 0.58 ($p < 0.01$), respectively. However, for EWW, the correlation coefficients were 0.66 ($p < 0.01$), 0.66 ($p < 0.01$) and 0.62 ($p < 0.01$). The above response patterns were also revealed by the correlation coefficients between the prewhitened and linearly detrended series (Fig. 4), indicating that autocorrelations and secular trends in the tree-ring width chronologies and climate time series have limited effects on the relationships.
- 25 Significant climate-growth relationships were also observed prior to the current growing season. For example, most of EWW and LWW chronologies exhibited negative response to Tmax and Tmean during the late summer and early autumn of last year (Figs. 3–4). This may be ascribed that high temperatures in the late growing season of last year could enhance soil water evaporation, thus inducing moisture stress and limiting the accumulation of photosynthetic products for the next year tree-growth (Peng et al., 2014). The influence of moisture status prior to the current growing season can also be reflected by the
- 30 significant positive correlations between LWW and drought indices from September of two years earlier to May of the previous year. However, EWW had lower correlations with these monthly drought indices. A possible explanation may be that the interannual variations of EWW were mainly contributed by the moisture status of the growth year.

Since the impacts of scPDSI on tree-growth can last for several months, we analysed the responses of various tree-ring width parameters to the multi-month averaged scPDSI. The strongest climate-growth relationship was found between the NELR based EWW STD chronology and the MJJ scPDSI ($r = 0.707$; $p < 0.01$; Fig. S6). Meanwhile, correlation coefficients derived from the methods SP67 and SPA50 were 0.67 ($p < 0.01$) and 0.68 ($p < 0.01$), respectively, which were lower than that based on NELR method (Fig. S6). This may be because the downward trend in MJJ scPDSI were better preserved using the NELR detrending method (Fig. S7). In addition, correlation coefficient between the NELR based EWW SSF chronology and the MJJ scPDSI was 0.705 ($p < 0.01$), which was quite close to that using the traditional STD method, indicating that the effects of so-called “trend distortion” in our tree-ring series were limited.

We further tested the temporal stability and possible lags (leads) in the relationships between NELR based STD chronologies and the MJJ scPDSI on different frequency domain (Fig. 5). It can be found that EWW generally showed high degree of coherence with the MJJ scPDSI on all timescales (2- to 18-year) except the periodicities between 3.5- and 6.5-year. Different from EWW, LWW only varied in-phase with the MJJ scPDSI but with some lags during the period from 1970s to 1990s on the timescales lower than 12-year. Moreover, LWW was inversely correlated with the MJJ scPDSI during the 1960s on the periodicities of 4- to 6-year. TRW showed an unstable relationship and certain lags to the MJJ scPDSI on the periodicities of 6- to 11-year. Therefore, it can be concluded that EWW has the most stable relationships with the MJJ scPDSI than LWW and TRW.

Previous studies based on TRW has evidenced that moisture status of the current growing season could strongly affect the radial growth of *P. tabulaeformis* (e.g. Cai and Liu, 2012; Cai et al., 2014; Cai et al., 2015; Chen et al., 2014; Fang et al., 2010a; Fang et al., 2012b; Li et al., 2007; Liang et al., 2007; Liu et al., 2017; Song and Liu, 2011; Sun et al., 2012). Fast radial growth of *P. tabulaeformis* usually happens in the early growing season (Liang et al., 2009; Shi et al., 2008; Zeng et al., 2018). Increased water deficiency due to the rising temperature and inadequate rainfall in the early growing season could induce water stress thus suppressing cell division and expansion (Fritts, 1976), and resulting in the formation of narrow earlywood bands. The less sensitivity of LWW to moisture status of the current growing season may be ascribed to that the moisture restrictions on tree growth was alleviated in the rainy season (July–August; Fig. 2). Meanwhile, the response sensitivity of TRW to moisture status of the current growing season was not as strong as EWW, although they shared a similar climatic response pattern because EWW represents the majority of TRW (on average, the portion of EWW of TRW accounts for 65.8%).

3.2 MJJ scPDSI reconstruction using NELR based EWW STD chronology

Based on the above analyses, we selected the MJJ scPDSI as the target for hydroclimate reconstruction, and the NELR based EWW STD chronology as the predictor (Fig. 6a). The transfer function was estimated using a simple linear regression model as expressed in Equation (3):

$$\text{MJJ scPDSI} = 4.74\text{EWW} - 4.32; (R^2 = 0.5, n = 53, p < 0.001), \quad (3)$$

The model explains 50 % of the actual MJJ scPDSI variance over the period of 1953–2005. The calibration-verification tests show that r , R^2 and the sign-test are significant at the 0.01 level, and that RE and CE values are positive (Table 3). In addition,

the p -value generated from DW and CS tests are larger than 0.05, indicating that there are no autocorrelations and long-term trends in the regression residuals (Table 3; Fig. 6b). All test results confirm that the model is valid (Cook et al., 1999; Fritts, 1976).

Based on the above model, the MJJ scPDSI of the study region was reconstructed back to 1868 (Fig. 6c). We adjusted the variance of the reconstruction to match the variance of instrumental MJJ scPDSI during the calibration period (1953–2005). Spatial correlation analysis indicates that the reconstruction most strongly represents Central China, including the western part of Henan, the northern part of Hubei, and the southern part of Shaanxi provinces (Fig. 1a).

3.3 Comparing the reconstructed MJJ scPDSI with other reconstructions and historical documents

On the interannual timescale (Figs. 7a–c), our reconstruction is significantly correlated with the $PDSI_{Cook}$ ($r = 0.37$; $p < 0.01$; 1868–2005), and Pre_{Chen} ($r = 0.52$; $p < 0.01$; 1868–2005). On the decadal and longer timescales, our reconstruction is significantly correlated with all other reconstructions (Figs. 7e–f). The common drought periods occurring in the 1870s and the 1920s were reflected in our reconstruction. These two drought periods were frequently observed in North and West China (Cai et al., 2014; Chen et al., 2014; Fang et al., 2012a; Kang et al., 2013; Liang et al., 2006; Liu et al., 2017; Zhang et al., 2017).

However, it should be noted that our reconstruction has some mismatches with others. On the interannual timescale, our reconstruction is not significantly correlated with the DWI_{Yang} over the whole period 1868–2000 ($r = -0.06$; $p = 0.64$; Fig. 7b). This probably due to the historical documents have limited ability in capturing the high frequency climatic variations (Zheng et al., 2014). On the decadal and longer timescales, our reconstruction varied out-of-phase with $PDSI_{Cook}$ during the period from the late 1940s to the early 1960s (Fig. 7d), weakly correlated with DWI_{Yang} after the 1940s (Fig. 7e), and leads Pre_{Chen} during the period of 1900s–1930s (Fig. 7f). The possible reasons might be that (1) the reconstructions aim to different target season (June–August for $PDSI_{Cook}$, May–September for DWI_{Yang} , and April–June for Pre_{Chen} which is before the rainy season); (2) the DWI_{Yang} after the 1940s was calculated using instrumental May–September precipitation and the chronology of Chen et al. (2016b) also reflects precipitation, while the scPDSI is influenced not only by precipitation but also temperature and previous drought conditions; and (3) in the MADA network, there are still limited tree-ring sites around our sampling sites, which may cause some difference on local scale.

We also compared the dry and wet events derived from our reconstruction with those recorded in historical document records. The moderately to severely dry (wet) events are defined based on the scPDSI values less than -2 (larger than 2) according to Palmer (1965). It can be found that all the dry events and 70 % of the wet events can be verified by corresponding descriptions in historical documents (Table 4). While, there are still some mismatches between our reconstruction and the historical records. For example, no relevant document record is found for the year 1983 when an extreme wet event is shown in our reconstruction. In addition, some historical events are not reflected in our reconstruction, such as the wet event in 1963 and the dry event in 1942 (Wen, 2006). These mismatches may reflect the uncertainties of historical documents records and tree-ring.

3.4 Connections with EASMI

The reconstructed MJJ scPDSI and EASMI in general exhibit an in-phase relationship before the 1940s on the decadal and longer timescales (Fig. 8). This in-phase relationship was further verified after conducting a 21-year moving window correlation analysis on the decadal-filtered scPDSI and EASMI (Fig. 9a, b and e). As EASM directly drives precipitation rather than the scPDSI, we compared the EASMI with the local precipitation (32° N to 34.5° N and 111° E to 112° E). It was found that the local precipitation also exhibits the similar variation as the scPDSI and EASMI before the 1940s (Fig. 9a, b and c). Therefore, the in-phase relationship between the decadal-filtered EASMI and scPDSI before the 1940s may be ascribed to the fact that a stronger EASM could induce enhanced local precipitation, thus increasing the soil moisture content. However, interestingly, the correlation between the local precipitation and EASMI weakened after the 1940s, and even became negative since the 1970s (Fig. 9e). We attributed this unstable relationship to the nature of EASMI. In fact, the EASMI was designed to capture the leading mode of EASM precipitation variability, whose largest loading is in general located at the mei-yu/changma/baiu rainband (Zhao et al., 2015). In other words, the mei-yu/changma/baiu rainfall has the most robust relationship with the EASMI. Moreover, although the leading pattern of EASM precipitation can change on interdecadal timescale and may even much longer timescales (which is still elusive due to the limited paleo-precipitation record), the unique importance of mei-yu/changma/baiu in EASM most likely remains (Wang et al., 2008). Therefore, due to the change of EASM precipitation pattern, the precipitation outside of the mei-yu/changma/baiu rainband could be in-phase, out-of-phase and uncorrelated with mei-yu/changma/baiu rainfall (Wang et al., 2008), thus manifesting an unstable relationship with EASMI. The EASM experienced an abrupt shift in the late 1970s, which caused a change of the leading mode of EASM precipitation (Wang, 2001; Ding et al., 2008). We demonstrated how this mode change affects the relationship between EASMI and precipitation in the eastern Qinling Mountains. As shown in Fig. 10a, the anomalies of the decadal-filtered MJJ precipitation exhibited similar variations over the Yangtze River basin and Yellow-Huaihe River basins during 1901–1978, but they were divided by the Yangtze River, showing a dipole pattern during 1979–2005 (Fig. 10b). The south of the Yangtze River basin (27° – 30° N) were the loading centres during both periods, and the decadal-filtered MJJ precipitation over this area were well captured by the designed EASMI as manifested by their significant positive correlations ($p < 0.1$; Figs. 10c–d). On the contrary, the decadal-filtered MJJ precipitation over the north of the Yangtze River, including our sampling sites, varied out-of-phase with those over the south of the Yangtze River basin after the late 1970s, thus being negatively correlated with the EASMI. The weakened scPDSI-EASMI relationship after the 1940s cannot be solely attributed to the change of EASM precipitation mode, because the scPDSI also showed weakened relationship with the local precipitation simultaneously (Fig. 10f). It cannot be ascribed simply to that the variations of scPDSI became dominated by temperature either, as no stably enhanced scPDSI-temperature relationship was found (Fig. 10f). This may be because the scPDSI is not a simple formula based on precipitation and temperature, but a complex function incorporating previous drought conditions and current moisture departure (Wells et al., 2004). In addition to the precipitation and temperature, the available energy, humidity and wind speed can also affect the

scPDSI via controlling evapotranspiration (Sheffield et al., 2012). Therefore, a variety of climate data is required to identify the specific cause for the weakened relationship. While, this would be difficult due to the limited climate records.

4 Conclusions

Besides TRW, climatic responses of EWW and LWW were also explored for the tree-ring samples of *P. tabulaeformis* in the eastern Qinling Mountains, Central China. Regardless of the detrending and standardisation methods used, the resulting EWW chronologies are more sensitive to early summer soil moisture conditions (scPDSI) than LWW and TRW during the instrumental period 1953–2005. The MJJ scPDSI (1868–2005) reconstructed from the NELR based EWW STD chronology captures the past early summer hydroclimatic fluctuations, further validated by other proxy-based reconstructions and historical document records in adjacent regions. This indicates that EWW has great potentials to reconstruct early summer hydroclimatic conditions in this area. Moreover, on the decadal and longer timescales, the reconstruction showed a strong in-phase relationship with the EASMI before the 1940s, which may be related to the positive response of the local precipitation to EASM. Our finding in this study is different from that found at a well-drained site in South China, where strongest moisture signals were contained in LWW with a different tree species (Zhao et al., 2017a, b). Therefore, more EWW and LWW related studies should be conducted in terms of tree species differences, different environmental conditions, in humid and semi-humid regions of China, that provides a possibility to understand EASM variations at longer time periods beyond the meteorological records.

Data availability

The tree-ring data used in this study are available on request (shijf@nju.edu.cn). DWI, precipitation reconstruction, and dry/wet events recorded in historical documents are available from corresponding authors or publications. MADA is available from <https://www.ncdc.noaa.gov/paleo-search/study/10435>. The 200 hPa zonal wind dataset of NOAA-20c is available from https://www.esrl.noaa.gov/psd/data/gridded/data.20thC_ReanV2c.html and <https://rda.ucar.edu/datasets/ds628.1/>. The gridded dataset CRU scPDSI 3.25 is available from <https://crudata.uea.ac.uk/cru/data/drought/>. The gridded precipitation dataset GPCC v7 is available from https://opendata.dwd.de/climate_environment/GPCC/html/fulldata_v7_doi_download.html. The gridded temperature dataset CRU TS 4.01 is available from https://crudata.uea.ac.uk/cru/data/hrg/cru_ts_4.01/cruts.1709081022.v4.01/tmp/.

Author contributions

YZ and JS designed the study. JS provided the tree-ring samples. YZ performed tree-ring width measurement, data analyses and interpretation. JS, SS, XS and HL assisted in data interpretation. YZ wrote the first draft of the paper. All authors revised the paper.

5 Competing interests

The authors declare that they have no conflict of interest.

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Table 1. Characteristics of climate data.

Climate data	Source	Longitude (° E)	Latitude (° N)	Elevation (m a.s.l.)	Temporal cover
Tmax, Tmean,	Luanchuan (LC) meteorological station	111.6	33.8	751	1957–2005
Tmin, Pre	Xixia (XX) meteorological station	111.5	33.3	250	1957–2005
	Ruyang (RY) meteorological station	112.5	34.2	311	1957–2005
	Nanzhao (NZ) meteorological station	112.6	33.6	198	1956–2005
scPDSI	CRU scPDSI 3.25 (van der Schrier et al., 2013)	111–112	32–34.5	—	1953–2005
SPEI	Calculated in R using the SPEI package with the Tmax, Tmin and Pre data (Beguería and Vicente-Serrano, 2012)	—	—	—	1957–2005

Table 2. Long-term hydroclimatic reconstructions and East Asian summer monsoon index (EASMI) selected for comparison with the reconstructed MJJ scPDSI.

Time series	Source	Longitude (° E)	Latitude (° N)	Temporal cover
June–August PDSI	Monsoon Asia Drought Atlas (MADA, Cook et al., 2010)	111.25	33.75	1868–2005
April–June precipitation in Mount Huashan (HS)	Chen et al., 2016	110.08	34.48	1868–2005
Dryness/wetness index (DWI)	Yang et al., 2013	111.25	33.75	1868–2000
EAMSI	Calculated using the 200 hPa zonal wind anomalies (NOAA-20c; Compo et al., 2011) according to the definition of Zhao et al. (2015)	—	—	1868–2005

Table 3. Statistics for split calibration-verification of the regression model.

Calibration period	r	R^2	DW value (p -value)	CS p -value	Verification period	RE	CE	Sign-test
Full period (1953–2005)	0.71**	0.50**	2.03 (0.53)	0.85	—	—	—	—
Early half (1953–1979)	0.68**	0.46**	2.02 (0.49)	1	Late half (1979–2005)	0.53	0.53	22+/5—**
Late half (1979–2005)	0.73**	0.54**	2.02 (0.50)	1	Early half (1953–1979)	0.46	0.45	21+/6—**

** $p < 0.01$; r , Pearson correlation coefficient; R^2 , explained variance; DW , Durbin-Watson test; CS , Cox and Stuart trend test; RE , reduction of error; and CE , coefficient of efficiency.

Table 4. Moderately to severely dry ($\text{scPDSI} \leq -2$) and wet ($\text{scPDSI} \geq 2$) events derived from the MJJ scPDSI reconstruction and corresponding descriptions from historical documents.

Event type	Year	scPDSI	Description
Dry	1879	-3.61	A mega-drought occurred has caused a great famine over Henan, Shaanxi and other provinces in North China in the early Guangxu reign (1876–1879) ^a
	1900	-2.24	Severe drought from spring to Autumn over Henan and Shaanxi
	1923	-2.28	Drought over Henan and Shaanxi
	1926	-2.33	No harvest at Ruyang (West Henan) due to severe drought
	1929	-2.53	Summer drought over Henan and Shaanxi
	1994	-2.12	Severe drought occurred in April, May and July over west Henan
	1995	-2.10	Intensified drought severity since April 22 over Henan
	2000	-2.94	The drought from Feburary to May is the worst one since 1950 over Henan
Wet	1869	2.31	Flood in summer and autumn over Henan
	1883	2.62	Persistent rainfall in summer at Shanxian and Mianchi (Northwest Henan)
	1885	3.07	Flood in summer at Lingbao and Shanxian (Northwest Henan)
	1894	3.06	Not available
	1895	2.11	Flood over the Qinhe River (Northwest Henan) in summer
	1898	3.77	Severe flood in summer at Lushi (Northwest Henan), Shangnan (Southeast Shaanxi) and Danjiang (Northwest Hubei)
	1905	2.26	Persistent rainfall in spring and summer over Henan

1906	3.65	Heavy rainfall in summer over Henan
1910	2.05	Flood in summer and autumn over Henan
1911	3.84	Heavy rainfall in summer over Henan
1912	2.01	Heavy rainfall and flood in summer over Nanyang (Southwest Henan)
1933	2.51	Heavy rainfall in summer over Henan and Shaanxi
1934	3.11	Summer rainfall over Henan, South Shaanxi and Northwest Hubei
1936	3.72	Not available
1944	2.38	Flood over Henan; Rainstorm in Zhenan (Southeast Shaanxi) on July 8; The Tianhui Channel (Southeast Shaanxi) was destroyed by flood on May 13
1948	2.81	Wheat loss caused by summer rainfall
1949	2.95	Not available
1973	2.97	Not available
1980	2.07	Rainfall in June is higher than usual for most regions over Henan
1983	4.15	Not available
1984	2.33	From June to September, there are 5 large-scale rainstorms over Henan
1990	2.32	Not available

^a Historical description of the 1879 drought event is cited from He (1980), and others from Wen (2006)

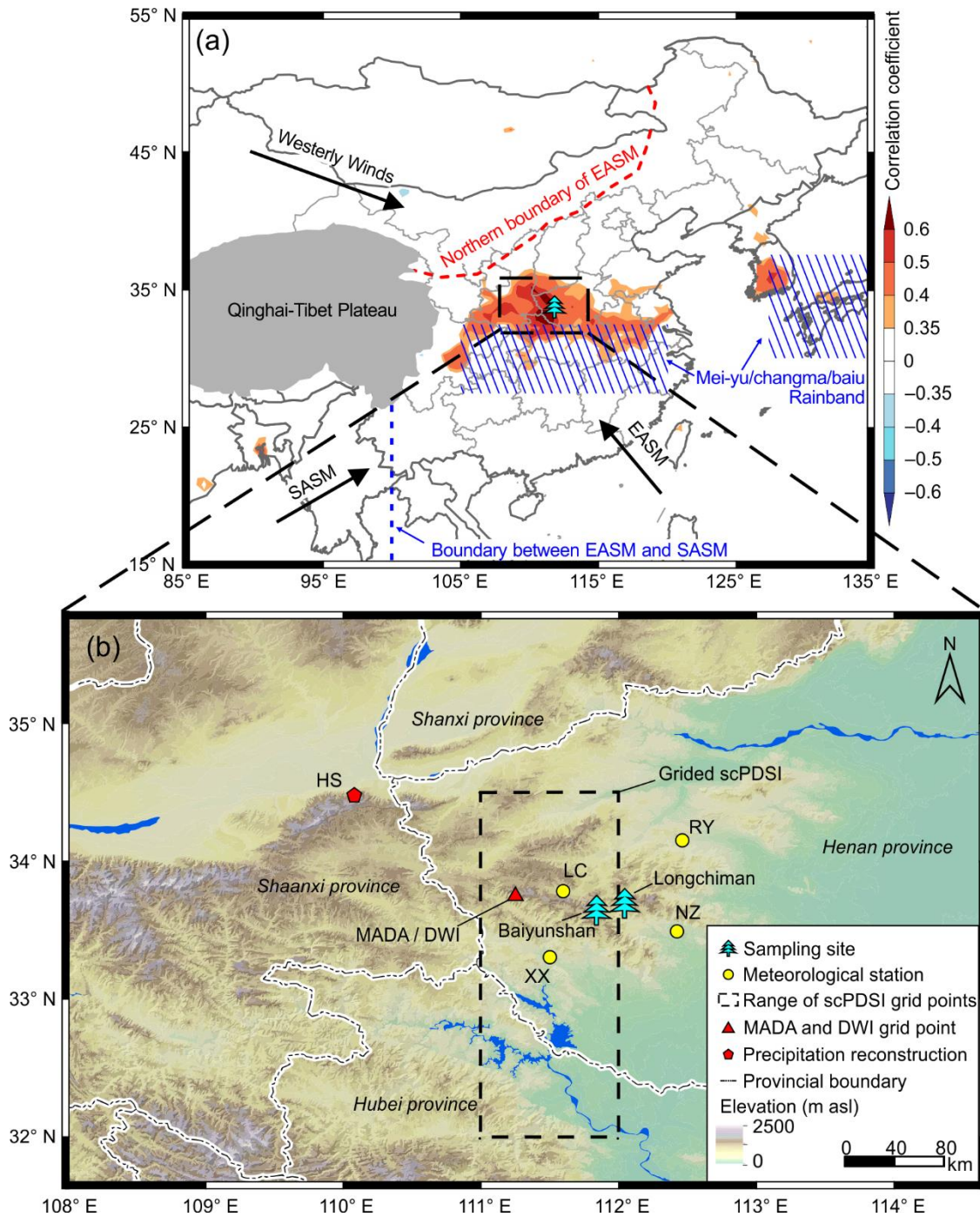


Figure 1. Map of the study region. (a) Location of the sampling site (tree symbol), and the spatial correlations between the May–July (MJJ) scPDSI reconstruction and the gridded scPDSI dataset (van der Schrier et al., 2013) during the period 1953–2005. The color bar indicates the correlation coefficient. The blue dashed line (100° E) indicates the boundary between East

Asian summer monsoon (EASM) and South Asian summer monsoon (SASM; Tang et al., 2010). The red dashed line indicates the northern boundary of EASM (Chen et al., 2018). The blue shaded area represents the mei-yu/changma/baiu rainband (Zhao et al., 2015). **(b)** Partial enlargement of the study region. The circles indicate the locations of the four meteorological stations (LC: Luanchuan, XX: Xixia, RY: Ruyang, and NZ: Nanzhao). The triangle indicates the location of selected grid data from the datasets Monsoon Asia Drought Atlas (MADA; Cook et al., 2010) and Dryness/Wetness Index (DWI; Yang et al., 2013). The pentagon indicates a tree-ring width based precipitation reconstruction in Huashan Mount (HS, Chen et al., 2016b). The dashed rectangle indicates the gridded scPDSI obtained from the gridded scPDSI dataset (van der Schrier et al., 2013).

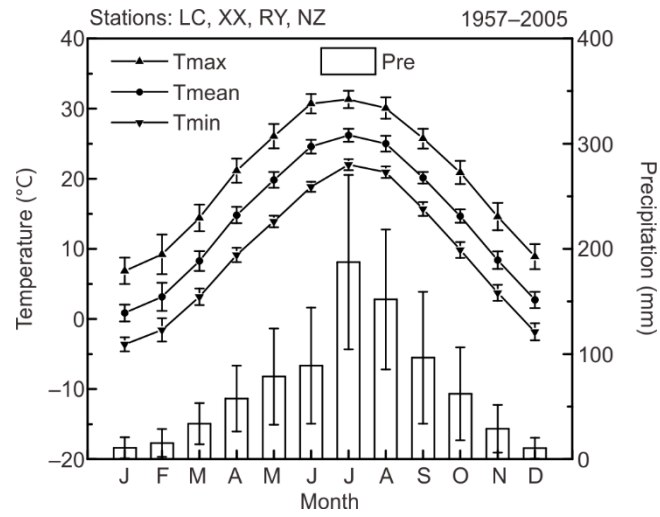


Figure 2. Monthly maximum, mean, minimum temperature (Tmax, Tmean, Tmin), and total precipitation (Pre) averaged from the four selected meteorological stations (LC, XX, RY, NZ) during the period 1957–2005. Error bar denotes \pm one standard deviation.

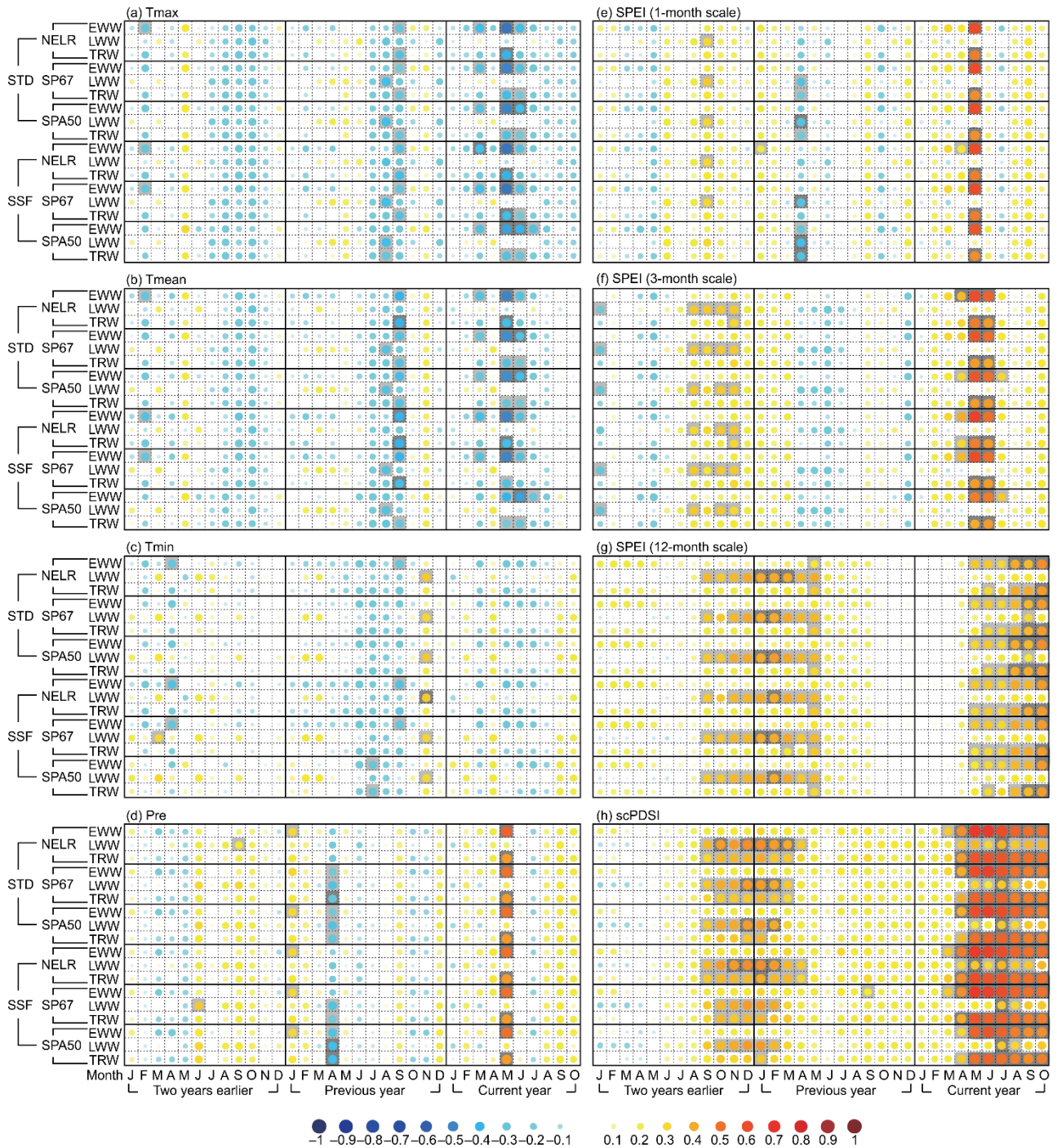


Figure 3. Matrix plots for the correlation coefficients between tree-ring width chronologies and monthly climate time series from January of two years earlier to October of the current year. The climatic factors are monthly (a) maximum temperature

(Tmax), **(b)** mean temperature (Tmean), **(c)** minimum temperature (Tmin), **(d)** total precipitation (Pre), **(e)** SPEI of 1-month scale, **(f)** SPEI of 3-month scale, **(g)** SPEI of 12-month scale, and **(h)** scPDSI. EWW, LWW, and TRW indicate the earlywood width, latewood width, and total tree-ring width, respectively. The correlation analyses were conducted during the period 1953–2005 for the scPDSI, and 1957–2005 for the other climatic factors. NELR, SP67, and SPA50 indicate the three
5 detrending methods: (1) negative exponential function together with linear regression with negative (or zero) slope (NELR), (2) cubic smoothed splines with a 50 % frequency cutoff of 67 % of the series length (SP67), and (3) age-dependent splines with an initial stiffness of 50 years (SPA50). STD and SSF indicate the two standardization methods “standard” and “signal-free”, respectively. The correlation coefficients are reflected by the colorful and different-size circles, which can be referred to the color bar as shown at the bottom of the figure. The squares filled with light and dark gray color indicate that the
10 correlation coefficients are statistically significant at the 0.05 and 0.01 level, which are tested using the Monte Carlo method (Efron and Tibshirani, 1986; Macias-Fauria et al., 2012).

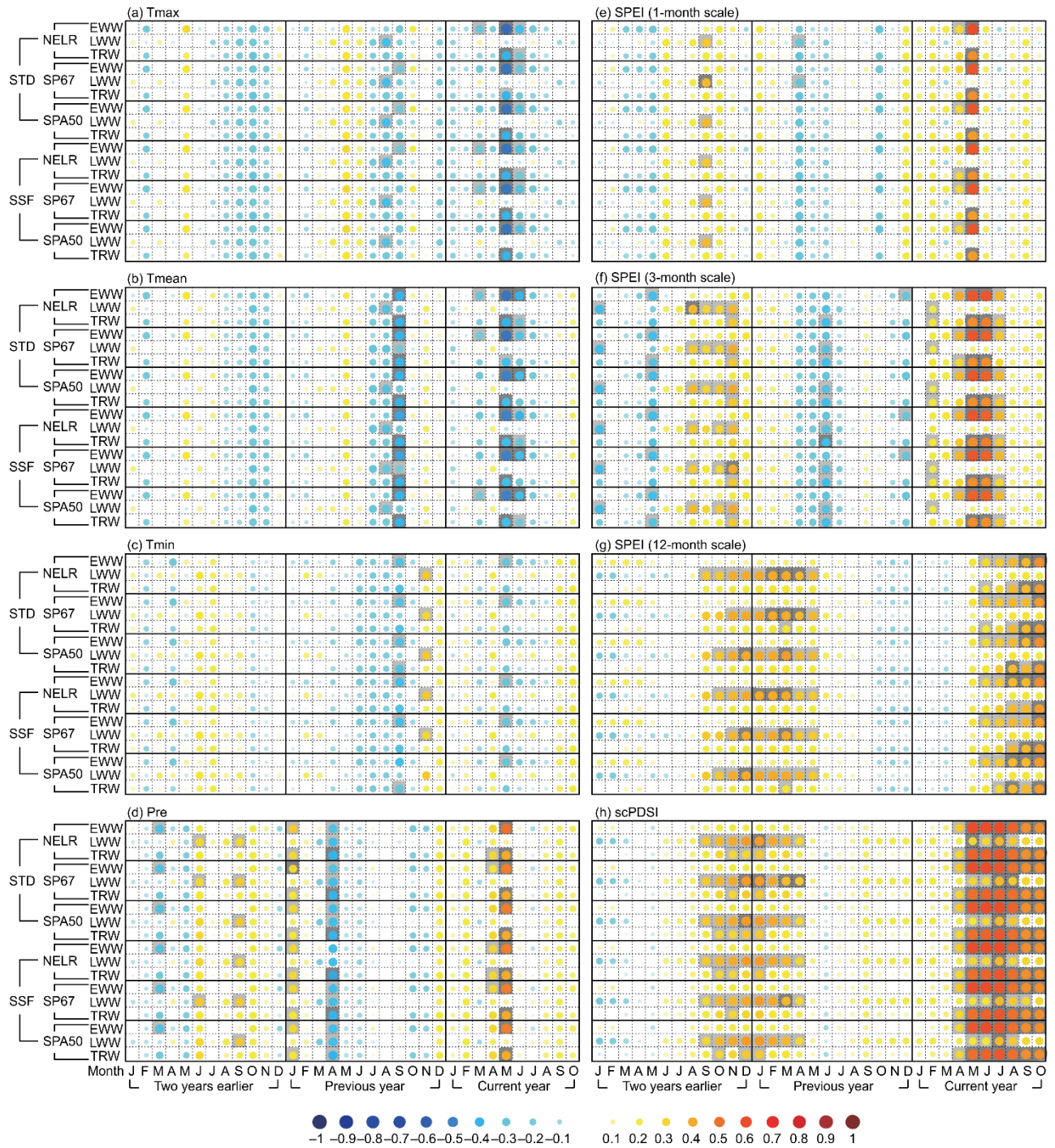


Figure 4. Correlation coefficients between the prewhitened and linearly detrended tree-ring width chronologies and climate time series. The explanations and legends are the same as Figure 3.

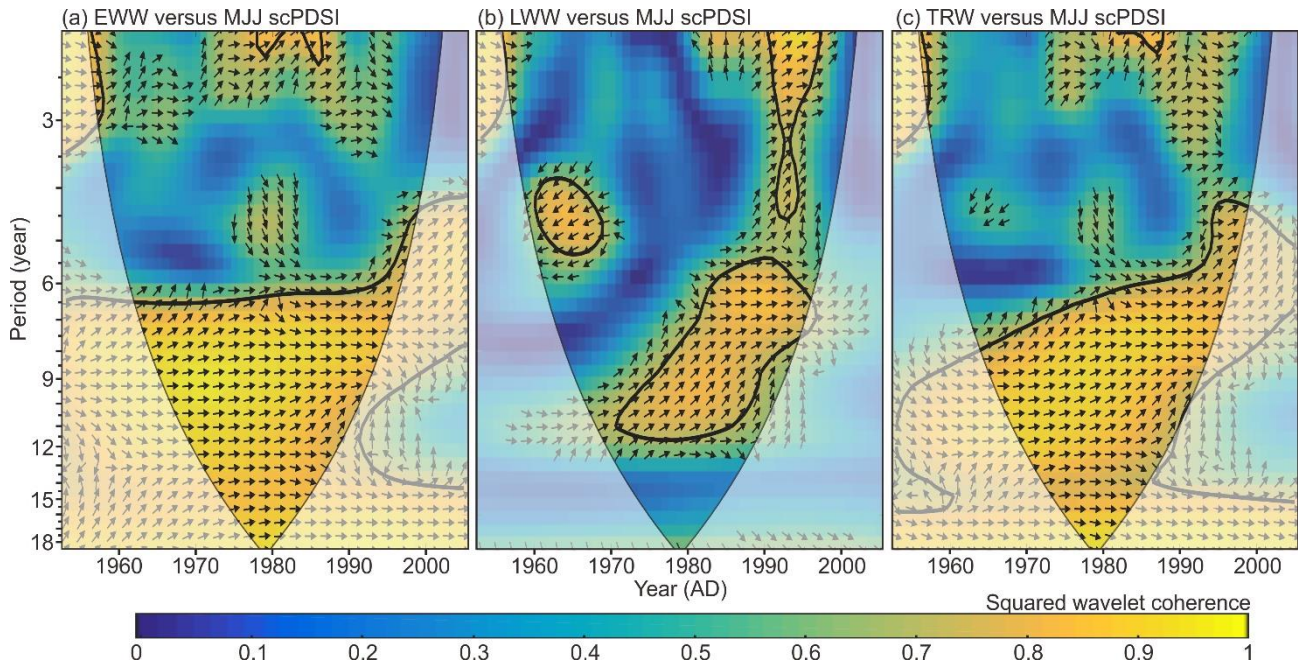


Figure 5. Squared wavelet coherence and phase relationship between the NELR based tree ring-width STD chronologies and MJJ scPDSI. (a–c) represent the results for EWW, LWW, and TRW, respectively. The color bar indicates the squared wavelet coherence. The arrows indicate the phase relationship with in-phase (anti-phase) pointing right (left), and MJJ scPDSI leading (lagging) tree-ring width with 90° pointing straight up (down). The thick contour indicates the 5% significance level against red noise. The cone of influence (COI) where edge effects might distort the picture is shown as a lighter shade.

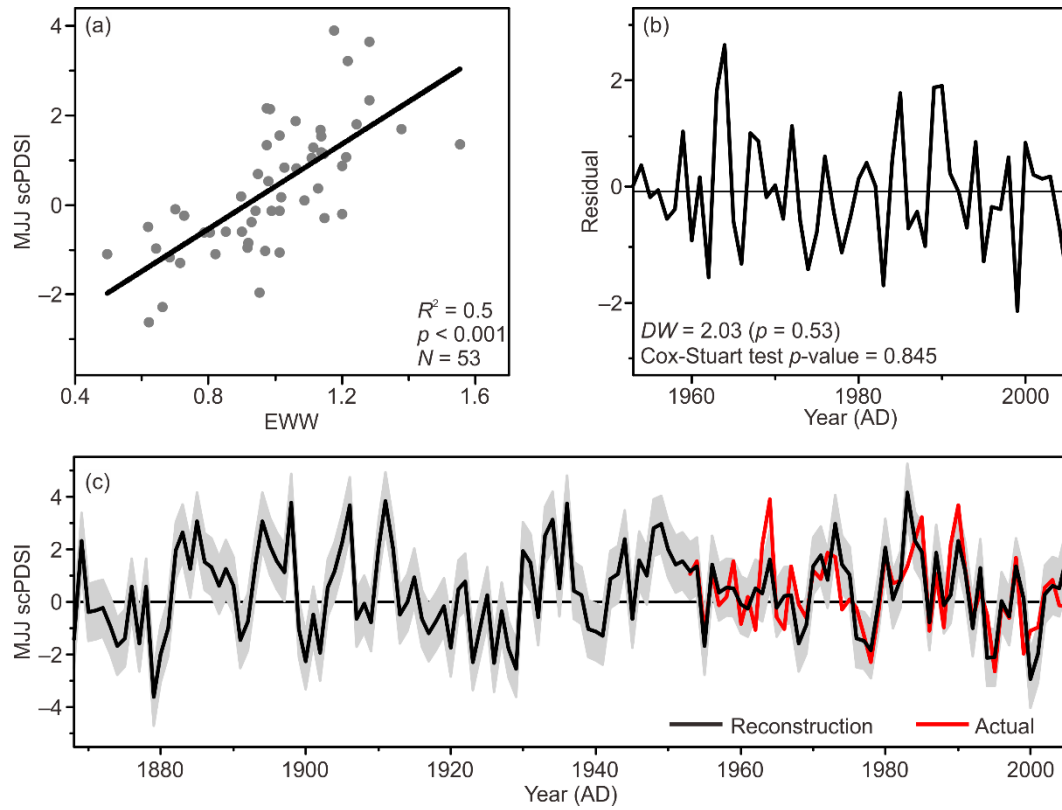


Figure 6. MJJ scPDSI reconstruction using NELR based EWW STD chronology. **(a)** Scatter diagram during the period 1953–2005, and **(b)** the resulting residuals. **(c)** MJJ scPDSI reconstruction (black line, after variance adjusted) and instrumental MJJ scPDSI (red line). Shaded area denotes the uncertainties of reconstruction in the form of ± 1 root mean square error.

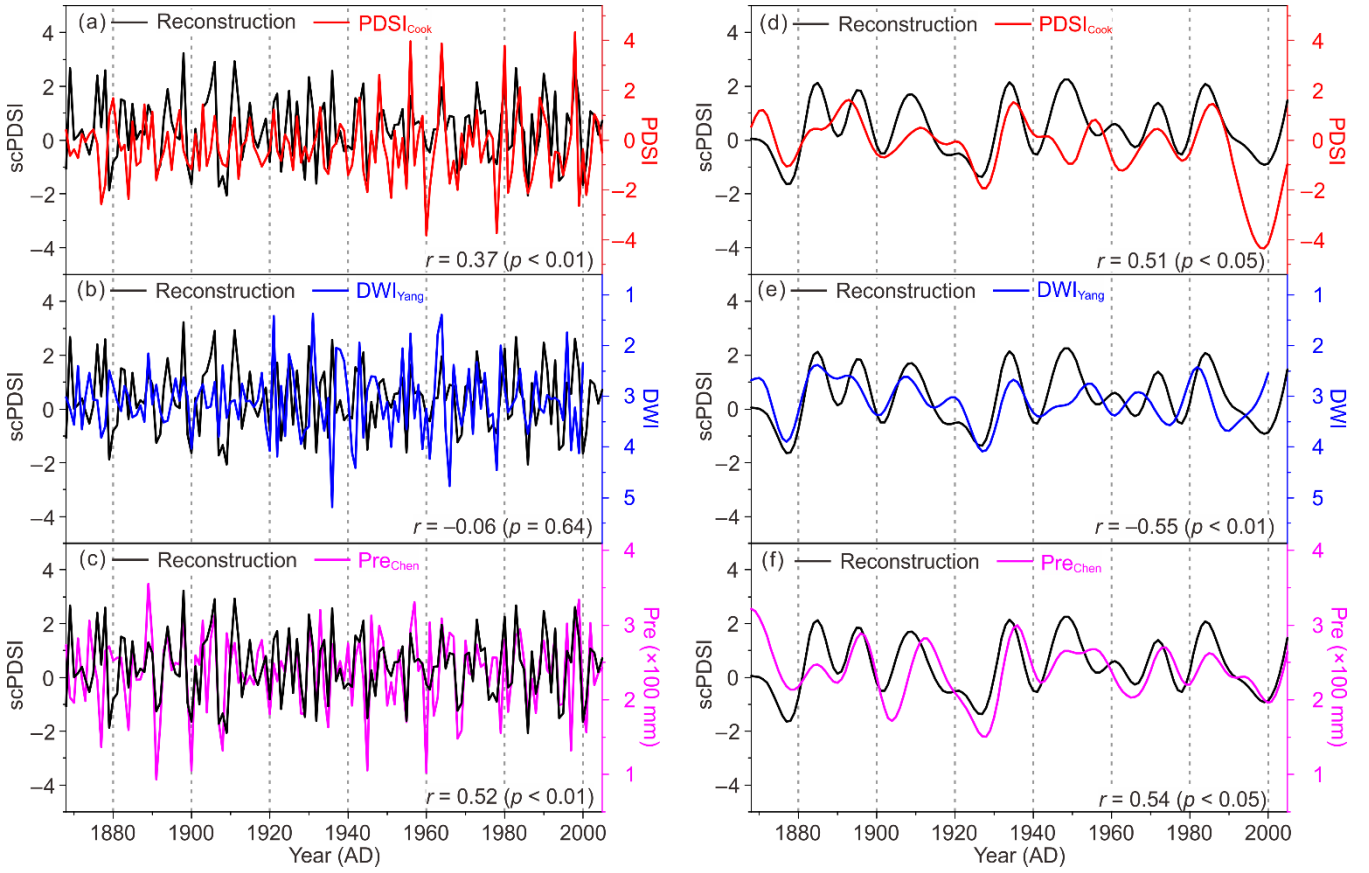


Figure 7. Comparison of the reconstructed MJJ scPDSI (black line) with other hydroclimatic reconstructions in adjacent regions on the interannual (left panels), and decadal and longer timescales (right panels). The referenced reconstructions are (a, d) June–August PDSI of MADA NO. 370 point (Cook et al., 2010), (b, e) reversed DWI (Yang et al., 2013), and (c, f) TRW based April–June precipitation (Pre) reconstruction (Chen et al., 2016b). The interannual and decadal and longer fluctuations were separated using the adaptive 10 point “Butterworth” low-pass filter with 0.1 cutoff frequency (Mann, 2008). r represents the Pearson correlation coefficient between the reconstructed MJJ scPDSI and other hydroclimatic reconstruction over their common period. The significance level for all correlation coefficients were tested using the Monte Carlo method

(Efron and Tibshirani, 1986; Macias-Fauria et al., 2012).

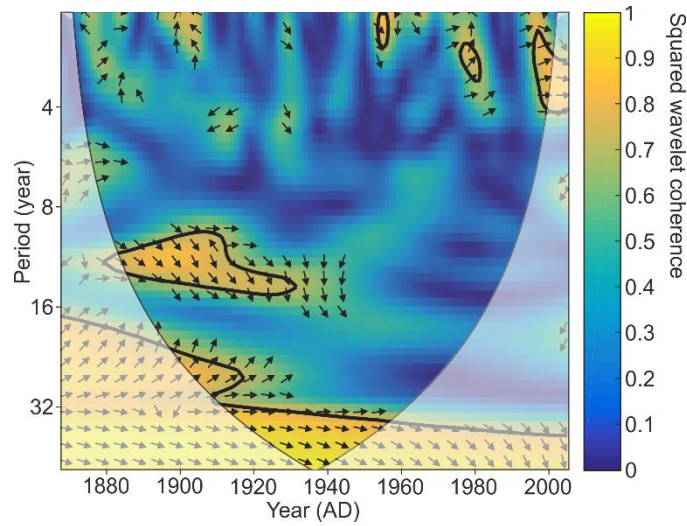


Figure 8. Squared wavelet coherence and phase relationship between the reconstructed MJJ scPDSI and EASMI (Zhao et al., 2015). The color bar indicates the squared wavelet coherence. The arrows indicate the phase relationship with in-phase (anti-phase) pointing right (left), and EASM leading (lagging) scPDSI with 90° pointing straight up (down). The thick contour indicates the 5 % significance level against red noise. The cone of influence (COI) where edge effects might distort the picture is shown as a lighter shade.

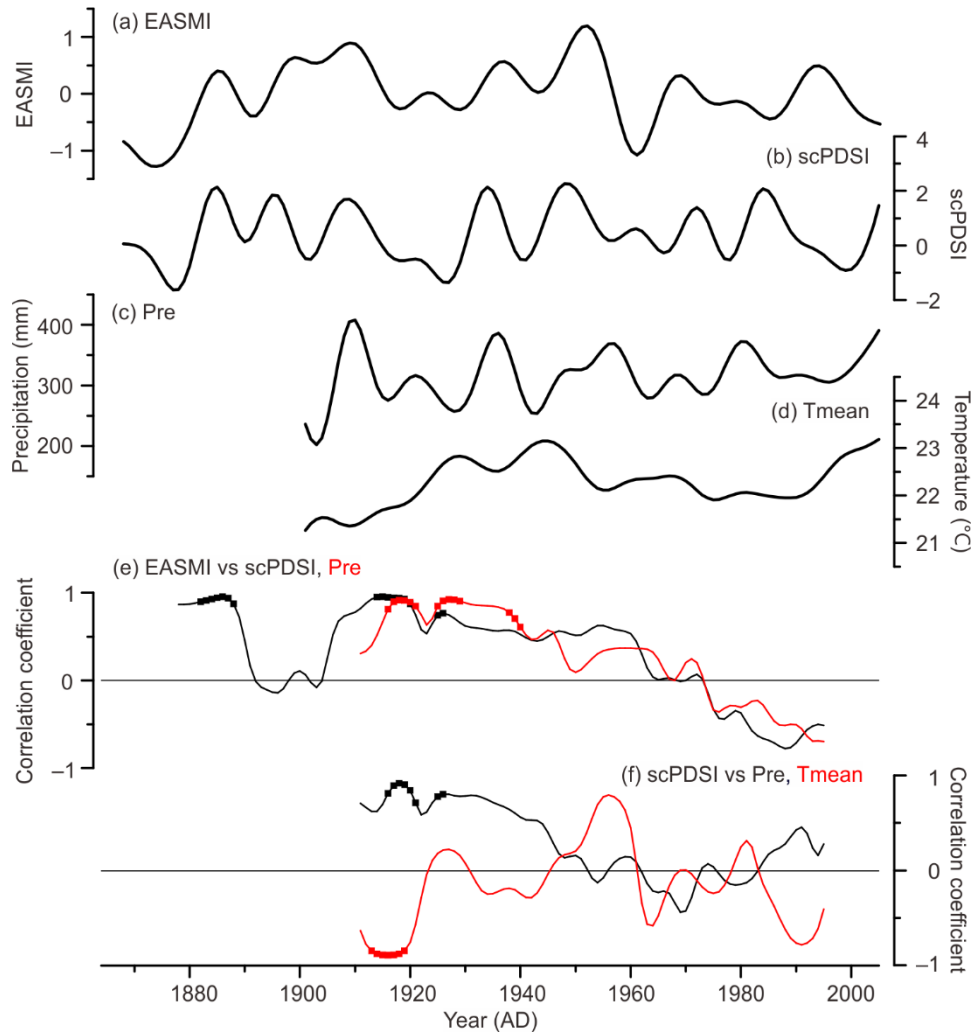


Figure 9. Comparison between the decadal and longer fluctuations of May–July **(a)** EASMI (Zhao et al., 2017), **(b)** reconstructed scPDSI, **(c)** precipitation (Pre; GPCC v7; Schneider et al., 2015), and **(d)** mean temperature (Tmean; CRU TS 4.01; Harris et al., 2014) over the reconstructed area. **(e)** 21-year moving Pearson correlation coefficients of the decadal-filtered EASMI with scPDSI (black), and Pre (red). **(f)** 21-year moving Pearson correlations of the decadal filtered scPDSI with Pre (black), and Tmean (red). The decadal and longer fluctuations were derived using the adaptive 10 point “Butterworth” low-pass filter with 0.1 cutoff frequency (Mann, 2008). Statistically significant ($p < 0.05$) correlations are denoted as squares, which were tested using the Monte Carlo method (Efron and Tibshirani, 1986; Macias-Fauria et al., 2012).

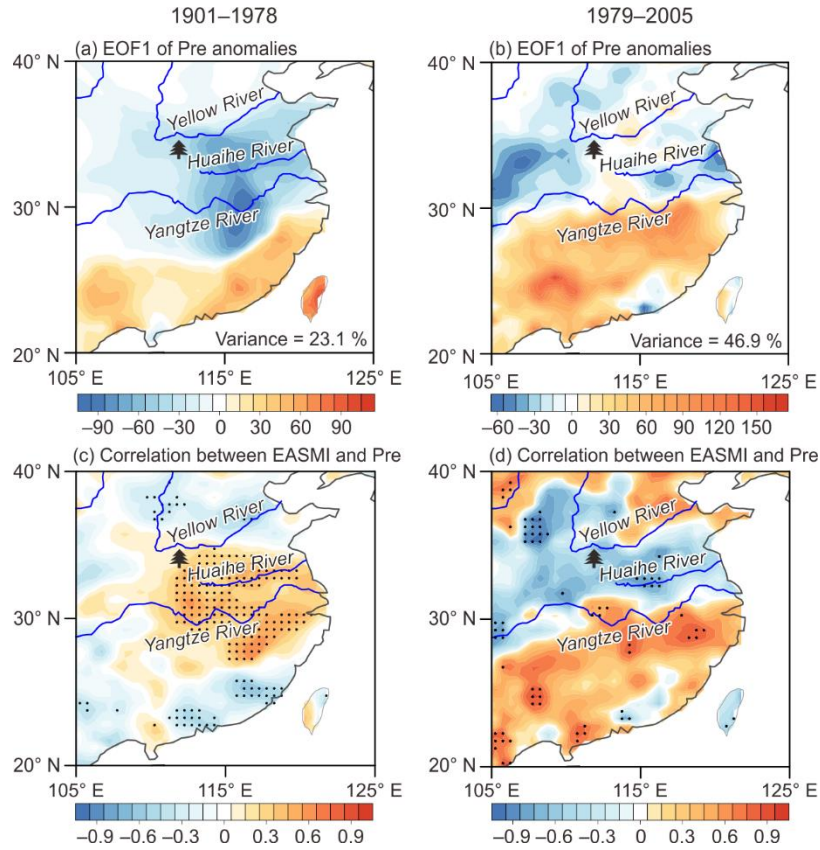


Figure 10. (a–b) The leading empirical orthogonal function (EOF) modes of decadal filtered May–July GPCC Precipitation anomalies for the periods 1901–1978 (left panel) and 1979–2005 (right panel). The color bar indicated the EOF values. **(c–d)** Spatial correlations between the decadal filtered May–July EASMI defined by Zhao et al. (2015) and Precipitation (Pre) for the periods 1901–1978 (left panel) and 1979–2005 (right panel). The color bar indicates the correlation coefficient. The dot indicates that the correlation is statistically significant ($p < 0.1$) which was tested using the Monte Carlo method (Efron and Tibshirani, 1986; Macias-Fauria et al., 2012). The decadal and longer fluctuations of precipitation were derived using the adaptive 10 point “Butterworth” low-pass filter with 0.1 cutoff frequency (Mann, 2008). The tree symbol denotes the study region.