

Submission of revised version

cp-2018-112: What climate signal is contained in decadal to centennial scale isotope variations from Antarctic ice cores? by Thomas Münch and Thomas Laepple

Dear Lukas Jonkers,

along with this document we hand in the revision of our manuscript entitled above.

In the revision, we have addressed all issues raised by both reviewers and the notes on the code and data availability mentioned by yourself. We have revised the respective manuscript parts as suggested in our reviewer responses; some changes slightly differ from the ones suggested in the responses and for those we have added respective statements to the original response marked up in **red**. Emphasis in our revision was put on the Results section to add some more explanatory text in order to facilitate understanding by a broader audience of CP readers, as suggested by Reviewer 2. The final url and doi's for the code availability are in the process of being registered and will be added in the later process.

Please find attached the one-to-one response to the reviewer comments with all detailed changes we made, as well as a marked-up manuscript version created with `latexdiff`.

Thank you again for considering our manuscript.

On behalf of Thomas Lapple and with kind regards,
Thomas Münch

Author Reply to the Review Comments by Dmitry Divine (Referee #1)

on the manuscript

cp-2018-112: What climate signal is contained in decadal to centennial scale isotope variations from Antarctic ice cores?

by Thomas Münch and Thomas Laepple

Thank you very much, Dmitry Divine, for the time you spent on reading and reviewing our manuscript; we are happy about the positive outcome. Below we include a point-by point response to both the general and to all specific comments. The original referee comments are set in normal font, our answers (author comment, AC) are typeset in *green italic font*.

Overall:

In this manuscript the authors present a method for calculating a timescale dependent SNR for an array of climate proxy records with a common climatic signal and a physical mechanism(s) behind.

For a particular case presented, the study uses ice cores based isotopic records and accounts for the effects of stratigraphic noise, diffusion in firn and timescale uncertainties elaborating the respective power spectral densities for the background climate signal and the aforementioned contributing noise factors. The proposed technique is then applied to ice core arrays from DML and WAIS. Opposite timescale behavior of SNR for the two core networks is linked to the homogeneity/heterogeneity of distillation trajectories between the two regions associated, for example, with the effects of sea ice on isotopes in precipitation.

In general the paper is clearly written and results are well presented. Moreover, my general impression over this study, is that this was one of the rare cases I had so far as a reviewer that can be published almost “as is”. When reading the manuscript, I left a number of remarks/suggestions/question marks that I planned to list later when writing this review, yet it turned out in the end that almost all of them the authors have already addressed in Discussion and Conclusions.

This study is certainly recommended for those who deals with multiproxy archives – this is an explicit demonstration of a value of a single proxy (ice core) record and a clear warning against overinterpreting single spikes/events on the shorter timescales. Therefore, I consider the manuscript deserves to be published after some very minor modifications to the content if the authors/editor finds them relevant.

AC:

Thank you very much for this positive evaluation of our work. We are happy that you find our study relevant, well presented and deserving publication.

Minor Comments:

Page 2, line 5: “...to a first approximation, changes in isotopic composition are only recorded in the ice if there is snowfall.” Recent studies suggest the effects of air (and hence water vapor) exchange across the firn–air interface in between the precipitation events may have a larger impact on the final d18O in snow than previously thought, see for example Stenni et al., 2016 , 10.5194/tc-10-2415-2016. It actually increases the role of SAT variability throughout the accumulation season even given the intermittency of precipitation itself.

AC:

We agree that there are indications that vapour exchange processes possibly further affect the isotopic composition of the surface snow and, if connected to SAT variations, could actually increase (again) the role of SAT variability for the isotope variability. Nevertheless, we could show for the Kohnen Station region in DML that, even after averaging away the local stratigraphic noise, the near-surface isotope variability is still very discrepant from the local interannual SAT variability (Münch et al. 2017). This suggests either precipitation intermittency to be a main driver of the remaining (after reducing the

stratigraphic noise) isotope variability, or isotope modifications occurring directly at the surface which are not controlled by SAT variations.

In order to briefly reflect this discussion in the manuscript, we changed the respective sentence to: "To a first approximation, changes in isotopic composition are only recorded in the ice if there is snowfall, while the role of water vapour exchange processes in between precipitation events is still debated (Steen-Larsen et al., 2011; Stenni et al., 2016; Casado et al., 2018; Ritter et al, 2016; Münch et al, 2017)."

Page 6, line 4: Please provide a ref to eq. (7).

AC:

The expression directly follows from the definition of the correlation coefficient as shown in Fisher et al. (1985). We added this reference to the paper here.

Page 6, line 9: "... for display purposes... smoothed using a Gaussian kernel". Still the motivation is not clear, would it be possible to see an unsmoothed signal (in the letter of response for example). What is the kernel bandwidth used?

AC:

We agree that the formulation in the respective sentence was unclear. We do not apply the smoothing only for display purposes. In general, a strong smoothing is necessary in order to improve the quality of the spectral estimates. If one assumes that the spectrum of the climate signal (which one aims to reconstruct with proxy data) can be described by piecewise power laws, smoothing in logarithmic space is reasonable, which we have adopted here. This logarithmic smoothing is applied by taking weighted averages of spectral power over a Gaussian smoothing window with a scale factor in log units (Kirchner et al., 2005) which we choose to be 0.1 for the WAIS data and 0.15 for the DML data. The scaling in log units is proportional to the frequency at which the smoothing is applied, thus, at the higher frequencies more data points are averaged resulting in a stronger smoothing. We added a more thorough motivation for the logarithmic smoothing at this point of the paper.

Below you see an unsmoothed version of paper figure 1 where you can observe the strong spectral uncertainty when no smoothing is applied.

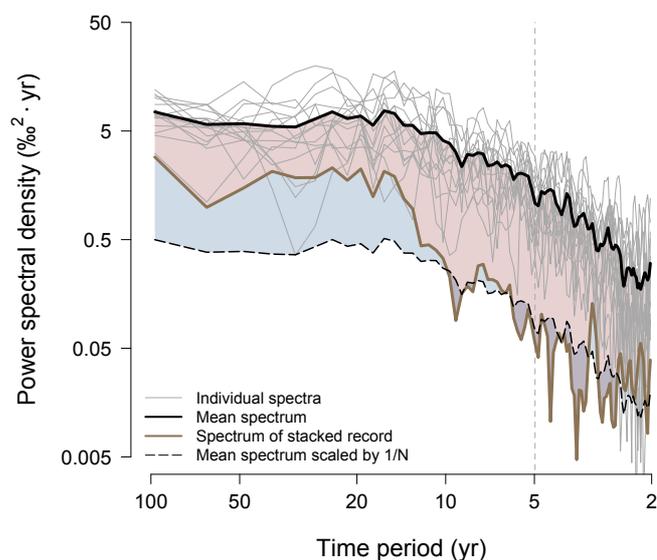


Figure 1: As Fig. 1 from the paper but without logarithmic smoothing applied to the spectral estimates.

Page 11, lines 10-20. Quality of ERA precipitation needs to be briefly discussed. How reliable are the estimates based on this variable?

AC:

We agree with the reviewer that the quality of the used ERA reanalysis data should be discussed but we think that the appropriate place for this would be the beginning of Appendix C. However, since our results do not critically depend on the accuracy of the estimated decorrelation scales, we included only a short and general discussion of the ERA-Interim quality of temperature and precipitation in Antarctica at the beginning of Appendix C.

Page 14, lines 30-33. The effects of sea ice on the modelled isotopic composition of precipitation in Antarctica can be found in the studies by Noone, e.g. Noone, D., and I. Simmonds (2004), Sea ice control of water isotope transport to Antarctica and implications for ice core interpretation, *J. Geophys. Res.*, 109, D07105, doi:10.1029/2003JD004228. The authors are recommended to see if these results can be used to elaborate more on the potential controls of the different patterns in SNR found between the two study regions.

AC:

Thank you very much for pointing us at this work. We included a reference to Noone and Simmonds and a short discussion of their findings at this point of the manuscript in order to improve our discussion of the role of sea ice changes for the isotopic composition of precipitation.

Page 19, lines 3-5. The authors present the winter and summer precipitation results. It is highly recommended to do the same analysis for the fall and spring seasons. The semi-annual oscillation (SAO) tends to modulate the seasonal distribution of precipitation depending on the strength of the semiannual harmonic. In addition, for West Antarctica (though shown for Faraday only in Broeke et al., 2000, part 4) the sea ice extent in the Amundsen and Bellingshausen seas (also linked with SAO strength) was shown to modulate the seasonal precipitation too. One can speculate that a long term variability in the strength/position of the low in contraction phase of the SAO (March and September) can actually be one of the mechanisms responsible for disruption of the coherence between the isotopic records on the longer timescales.

See in the series of earlier publications by Van den Broeke.

AC:

Thank you very much for pointing us at these interesting results. We extended the decorrelation analysis of the ERA-Interim precipitation data for the fall and spring seasons. This, however, does not change the results: also for the distribution of spring or fall precipitation (e.g., spring minus fall or spring/total precipitation, etc.) the decorrelation scales lie roughly between 300 and 500 km for both regions. Thus, whatever season drives the precipitation intermittency, the typical spatial scale of the intermittency should be of the order of 300-500 km, and thus a factor of 3-4 below that of the temperature fields.

Thomas Münch and Thomas Laepple

References to the author comments:

Casado, M., et al., *The Cryosphere*, 12, 1745–1766, <https://doi.org/10.5194/tc-12-1745-2018>, 2018.

Fisher, D. A., et al., *Ann. Glaciol.*, 7, 76–83, <https://doi.org/10.1017/S0260305500005942>, 1985.

Kirchner, J. W., *Phys. Rev. E*, 71, 066110, <https://doi.org/10.1103/PhysRevE.71.066110>, 2005.

Münch, T., et al., *The Cryosphere*, 11, 2175–2188, <https://doi.org/10.5194/tc-11-2175-2017>, 2017.

Ritter, F., et al., *The Cryosphere*, 10, 1647–1663, <https://doi.org/10.5194/tc-10-1647-2016>, 2016.
Steen-Larsen, H. C., et al., *J. Geophys. Res.*, 116, D06108, <https://doi.org/10.1029/2010JD014311>, 2011.

Stenni, B., et al., *The Cryosphere*, 10, 2415–2428, <https://doi.org/10.5194/tc-10-2415-2016>, 2016.

Author Reply to the Review Comments by an anonymous reviewer (Referee #2)

on the manuscript

cp-2018-112: What climate signal is contained in decadal to centennial scale isotope variations from Antarctic ice cores?
by Thomas Münch and Thomas Laepple

We appreciate a lot the very thorough and detailed reading and reviewing of our manuscript by the anonymous reviewer. The many comments and suggestions will be of great help to improve our first version of the paper. Below we include a point-by-point response to both the general and to all specific as well as technical comments raised by the referee. The original referee comments are set in normal font, our answers (author comment, AC) are typeset in *green italic font*.

General comments:

The article submitted deals with quantification of climate signal versus noise in ice cores from Antarctica. It is therefore well within the scope of *Climate of the Past*, and addresses an important issue for climatologists. Its aim is not to present new ice core data, but to present a methodology to evaluate (quantify) the climate signal contained in a series of records.

The methodology is based on a spectral analysis of the dataset, where the spectrum of the stacked record is compared to the mean spectrum and to white noise. The method also includes a correction for diffusion and for time uncertainty. The Methodology section is concise, because details are described in Appendixes. The paragraph 3.1 in Results is a useful complement to the methodology section, as it applies the method to an example, and provides a figure where the various steps are represented. It is well suited to an article that aims a large audience, not necessarily with statistical background, and who might overlook the equations in the methodology section and Appendixes.

AC:

We are happy that the structure of the manuscript and the steps at which we present the results are positively evaluated. Since not every reader of CP might be familiar with the details of the applied spectral analyses, it is indeed our intention to still present our results in a way that is understandable to a broader audience, as we envision our approach to be applicable in many fields of paleoclimatology/proxy research.

In the continuation of the Results section, the figures are described in less details. Some more precision is needed here, so that the main message is not obscured by unanswered questions on the parts of the figures that are not described. The results are different for the two studied regions. At EDML, the signal to noise ratio is found to increase for longer time scales (0.2 to 1), whereas at WAIS, it is relatively stable, and even seems to decrease at long (centennial) time scales. For the first region, the authors therefore recommend to use single cores only for multidecadal or longer timescales. For the second region, oppositely, they conclude that single cores yield good regional information at interannual and decadal scales, but give a more local information at longer time scales.

AC:

Thank you for these comments. Also in the light of what was said above, we edited the second part of the results section to ensure that the results are described in adequate precision in order to be comprehensible for a broad audience; see also our answers to the detailed comments.

In the Discussion, the authors consider the possible contribution of four processes to the climate signal, by looking at their spatial scale of coherence. They note that precipitation intermittency acts as noise or contributes to signal, depending on the

scale considered.

They also discuss the unexpected decrease of signal power at WAIS for longer timescales. This coastal region is particularly sensitive to the variability in atmospheric circulation. They suggest that slow processes modifying the topography of the region may reduce the spatial coherence of the signal over long timescales.

The conclusions of the article are important regarding the confidence that we can attribute to one or several ice core records. The results for WAIS are unexpected, and therefore may trigger more research in the area, or allow to consider differently the results from previous drillings. Lastly, as noted by the authors, the new methodology can be applied to other records, and therefore be of interest to a large audience.

The manuscript is well written and well structured, with a good use of appendixes to procure relevant information. Given that the conclusions of the article are interesting, and well supported by the Method and Results sections, I recommend that this article should be accepted by *Climate of the Past* with minor corrections.

AC:

We appreciate this positive evaluation.

Specific comments:

P2, 112: 'the diffusion of vapor through the firn column smoothes the isotope variations, reducing the power spectral density at high frequencies'

The first part of the sentence is very clear but the second part is less straightforward. Is it possible to add an intermediary step? '...smoothes the isotope variations, with stronger impact at short distances and therefore on the high frequencies.'

AC:

We will add the additional information, as suggested, and split the sentence into two parts for the sake of clarity: "Finally, once the snow is deposited, the diffusion of vapour through the firn column smoothes the isotope variations. This has a larger impact on short distances in the firn and therefore reduces the power spectral density of the variations strongest at the high frequencies (ref), which substantially shapes the isotope variability (ref)."

We changed the sentence to: "Finally, once the snow is deposited, the diffusion of vapour through the firn column smoothes the isotope variations (Johnsen, 1977; Whillans and Grootes, 1985). This has the strongest effect on short distances in the firn, significantly reduces the high-frequency power spectral density of the variations (Johnsen et al., 2000; van der Wel et al., 2015) and thereby substantially shapes the isotope variability (Laepfle et al., 2018)."

P2, l. 24: 'Furthermore, knowing the timescale dependence, and thus the spectral shape of the noise, would allow the correction of the isotope-inferred variability estimates for the noise contribution.'

This sentence is complicated, it takes some time to find the subject. Is it possible to split it or reverse it?

'Furthermore, in order to correct the noise contribution to the isotope inferred variability estimates, it is necessary to know the noise distribution over various timescales, i.e. its spectral shape.'

AC:

We changed the respective sentence to: "Furthermore, in order to correct the isotope inferred variability estimates for the noise contribution, it is necessary to know the variance of the noise across timescales, i.e., its spectral shape."

P3, 114: 'The time uncertainty of the record chronologies, ..., has been reported with 2 % of the time interval to the nearest tie point.'

This sentence is unclear, is there something missing? Is it 'has been reported to be 2 % of the time interval'?

AC:

The sentence indeed was misleading and it should read "...reported to be 2% of the time interval". For the sake of clarity, we corrected this and also rearranged the sentence to: "The record chronologies were established from seasonal layer counting of chemical impurity records constrained by tie points from the dating of volcanic ash layers (Graf et al., 2002a). The resulting uncertainty of the chronologies has been reported to be 2 % of the time interval to the nearest tie point (Graf et al., 2002a), which ..."

P5, 18: Equations 4

Although it was only a few lines above, please repeat here what M and S are (respectively the mean of spectra, and the spectra of the stacked record along time). Also, is it possible to get another symbol for the noise (or at least a very specific font) so that it is easier to distinguish between N and N?

AC:

We repeated the definition of M and S here, as suggested, and changed the symbol for the number of records from upper case 'N' to lower case 'n' throughout the manuscript. We also note here that in Eq. (2), the lower limit of the sum was erroneously given by "i=0" instead of the correct "i=1". This was corrected.

P5, l. 19: 'we restrict our analyses to the frequency region where $G < 2$. This avoids large uncertainties and yields cutoff frequencies of...'

Based on Figure B1 in annex, G is always below 2. Its maximum value is one. This seems logical since a filter will eliminate some frequencies, not add more frequencies to the signal. Should it be 'restrict to the region where $G > 0.5$ '? (so that the short frequencies will not triple in power after correction)?

Please clarify this point in the text or in the Appendix B.

Please define the term 'cut-off frequency' (and maybe draw the limits on Figure B1).

AC:

We apologize for the confusion. At this point of the manuscript we mixed up \bar{G} and \bar{G}^{-1} (the inverse of \bar{G} , which is actually applied as the correction function in Eqs. (4)). We will correct the sentence to "restrict to the region where $\bar{G} \geq 0.5$ " and clarify in the appendix that this means a correction factor less than 2 in Eqs. (4). Additionally, we will define the phrase "cutoff frequencies" by paraphrasing the sentence as "This avoids large uncertainties and translates to a maximum frequency that is used for the analyses (hereafter: cutoff frequency) of... (Fig. B1)", and we will add the respective frequency limits as vertical lines in Fig. (B1a).

We revised the entire paragraph to clearly explain the approach.

P6, 12: 'the value of F is related to the correlation between...'

Is the correlation r providing more information than F (on the quality of the data) or the same information? Is the r value only for intercomparing with other studies?

AC:

The correlation is mathematically equivalent to the value of \bar{F} ; however, we think the correlation in addition provides a more direct and intuitive way of expressing the amount of variability contained in a stacked isotope record that is related to the common (climate) signal in relation to its total variability. In order to stress the equivalence of both quantities, we rephrased the sentence to "The value of \bar{F} can then be used to obtain the correlation between the time series of the common signal c and a "stacked" record \bar{x} built from the spatial average of n individual records:".

P6, 120-28:

1. Is it possible to insert here the symbols used previously in the methodology section (M, S) to facilitate the identification of the various terms?
2. The mean over all individual spectra, M, (figure1, black line)...
3. The mean spectrum divided by the number of records (M/15, figure 1, dashed line)...

4. ...averaging across record that contain noise and additionally... (S..., Figure 1, brown line)

AC:

We inserted the respective symbols here, as suggested.

Furthermore, please precise 'averaging across records along the time dimension' or 'in time space' to refer to the 'time stack' S, by opposition to the spectrum mean M.

AC:

We clarified the sentence as follows: "In comparison, averaging in the time domain across records that contain noise and additionally a common (i.e. spatially coherent) signal, ...".

P6, 129: 'For short timescales (2 to seven years), ..., is consistent with the null hypothesis of independent noise.'

The curves are closest between 3 and 5 years; they diverge again around 2. What causes this divergence from white noise? Is it an artefact?

AC:

This divergence from the white noise level close to the 2-yr Nyquist period is probably indeed an artefact caused by noise added in the measurement process, i.e. from the measurement uncertainty and more likely from the "jitter error", thus uncertainty in the definition of annual depth increments upon dating which translates into uncertainty of the annual averages. We added a respective remark to the manuscript here.

P7, 16-7: 'Unlike the average spectrum across all individual isotope variations (fig1), the corrected DML1 signal spectrum shows an increase of power spectral density with increasing time scale (fig. 2a).'

1. Please precise which line is described here (color, symbol).

AC:

We will add the respective descriptions.

2. Both the black and brown line show an increase of power spectral density towards higher timescales on Fig1 (strongest between 2 and 20, then flat area at longer timescales). This proposition has to be nuanced.

AC:

Please note that both these spectra in Fig1 have not been corrected for the loss of spectral power by diffusion, so most of the increase in spectral power between the 2 and 20 year period is attributable to the decreasing diffusional smoothing towards longer timescales. What we wanted to stress at this part of the manuscript is the difference in spectral shape between the mean spectrum (M, black line in Fig1) and the estimated signal (blue for DML1 in Fig2) for longer timescales, which stems from correcting the isotope variability for the noise contribution. We will rewrite the respective sentence to clarify this.

3. For the brown line, there is again an increase at very long timescales (>50) still on Figure 1.

AC:

This is indeed expected given that the estimated signal (blue line in Fig2), which is contained in the spectrum of the stack (see Eq. 3), increases in power towards the longer timescales.

4. On Figure 2, the 3 DML1 curves are very similar for $t > 20$. So, the correction does not seem to have a huge effect. Moreover, the increase of power between 50 and 200 years exist with/ without correction...

AC:

It is indeed expected for the diffusion and time uncertainty corrections to have no effect for periods

>20 years, given their estimated transfer functions (Fig. B1). What is still relevant to correct on the longer timescales is the residual amount of noise still contained in the spectrum of the stack (since we only average a finite number of records). This residual noise is subtracted when calculating (solving for) the signal spectrum in Eq. (4a).

5. Lastly, regarding the increase between 3 and 20 years on Figure 2, the correction actually reduces the slope...

AC:

This is true and already noted in the manuscript (p.7 ll.11-13).

Possible correction to this sentence:

‘Unlike the average spectrum across all individual variations (M, black line on Fig 1), the corrected DML1 signal spectrum (C, solid blue line on Figure 2a) shows an increase of power spectral densities with increasing time scales at time scales larger than 50 years.’

AC:

We added the respective descriptions, as suggested, and rewrote the paragraph to clarify that the corrections include both the correction for residual noise as well as diffusion+time uncertainty, and that they lead to a more steady increase in signal power as compared to the mean spectrum. For further clarification, we in addition changed the legend of the dashed line in Fig. 1a from "Raw" to "Uncorrected signal".

P7, l. 7-8: ‘This is confirmed by the three 1000-yr long records from the DML2 data set whose signal exhibits a similar power spectrum in the range of timescales that overlap.’ The two curves are not really a perfect match.

There is common behavior between 60 and 100 years (increasing) and between 5 and 15 years (increasing) but the middle part (15 to 60) is different between the two curves. DML1 is slowly decreasing (or flat), while DML2 is strongly decreasing until 30, then strongly increasing. Maybe add: ‘...for time scales longer than 50 years.’ Or split the paragraph in two, one dealing with long time scales, and the other with short time scales.

AC:

We acknowledge that the two curves are indeed not a perfect match, but only show a similar slope. We weakened the statement accordingly.

P9, l.6: ‘...decrease in signal power on centennial timescales’

Is it possible that this decrease is an artefact? How certain are we of the power of the 100-year frequency on a 200-year record?

AC:

It is indeed possible that this decrease is an artefact from the spectral estimates, since log-smoothing uses less data points for lower frequencies and thus leads to higher spectral uncertainties there. We added this remark to the manuscript here.

P11, l.11: ‘In fact, the average correlation of 0.87 (0.94) over distances up to the maximal intercore distance...’

Please add here a reference to Appendix C, Figure C1. Otherwise the origin of these numbers (and their meaning) is unknown.

Based on the figure caption, this correlation is an average (for various intercore distances). Thus, maybe the formulation ‘up to the maximum intercore distance’ is misleading here. Why not say the correlation for average distance between cores instead?

AC:

We apologize for the missing figure reference here. We added the reference to Fig. C1, and included the average correlation values as symbols in this figure. In fact, the values are the average across all correlations for distances smaller than the maximum intercore distances, i.e. the average of all grey

dots within the shaded regions in Fig. C1 (so not the correlation at the average core distance). We rephrased the sentence to clarify this.

P12, l. 11-13: 'The raw noise spectra derived from the two DML data sets exhibit a clear imprint from the diffusional smoothing in the firn, as suggested by the smaller PSD for periods <20 years of the raw DML2 spectrum, i.e. prior to correction, in comparison to DML1 (fig. 2b).'

This is too fast. Please split in two steps:

1. first, raw to corrected for both records (DML1 and DML2): the effect of diffusion is to decrease power at high frequencies;
2. second, comparison of the raw data for DML1 and DML2: DML2 is more affected by diffusion, since frequencies between 12 and 30 are affected (low power) only for DML2.

AC:

We agree that the sentence condensed many pieces of information and might be hard to follow. We thus split the sentence into several steps: "The raw noise spectra, i.e. prior to correction, derived from the two DML data sets exhibit a clear imprint from the diffusional smoothing in the firn. This is suggested by their common decrease in PSD towards shorter periods (Fig. 2b), since diffusion acts stronger on higher frequencies. It is corroborated by comparing the loss in PSD between the two data sets, which for DML2 is stronger towards the high-frequency end and also extends further towards longer periods. This is due to the stronger diffusional smoothing in the older sections of the cores that are only contained in the DML2 records, since the diffusion process had more time to act there since deposition of the snow."

P12, l30: 'yields a slope of the DML2 signal spectrum of roughly $\beta=0.6$ '

Is it possible to add this fit to Figure 2 or 3?

AC:

This would of course be possible, but in our opinion it impairs the visual appearance of Fig. 2a which already contains many different line plots. Furthermore, since the slope is here used only as a diagnostic means, we do not want to place special focus on its value. Therefore, we would rather refrain from adding the fit to Fig. 2a (neither to Fig. 3 since it shows a different quantity than discussed at this point of the manuscript).

P13, l. 7: 'minimum averaging period constrained by the diffusion correction (2.5 years)'

Is it correct to compare r values at different averaging times (1 year and 2.5 years)?

Could this also contribute to the higher correlation compared to previous study of the same data?

AC:

Of course you are right that we compare correlation values at slightly different averaging times, which also could contribute to the slight difference between the two values. We added this as a remark here.

P14, l. 1-3: 'For WAIS, the higher SNR at interannual timescales as compared to DML is consistent with the general notion that higher accumulation rates result in higher SNR.'

1. Please insert the SNR values for both.

AC:

We added the average SNR value for periods from 5-10 years for both DML and WAIS (as obtained from the data in Fig. 3) for better comparison here.

2. For $N=1$ and $\Delta t=5$ years, $r=0.45$ at DML and $r=0.51$ at WAIS. The difference looks very small (especially considering the 3-times accumulation at WAIS). Why is the increase in SNR not scaled to the large increase in accumulation?

AC:

Please note that we compare at this point the SNR values (Fig. 3) between the two regions and not the correlations in Fig. 4, since the latter cannot be compared one-to-one as the correlation values, which are currently shown in the figure, are based on different lower integration limits in Eq. (6) (DML lower limit of 500 yr period, WAIS lower limit of ~100 yr period), which is due to the different lengths of the data sets.

However, this review comment made us aware of the fact that it is more appropriate to use the same integration limits for both regions for Fig. 4. Thus, we provide an updated version of the figure where we will use the same lower integration limits (i.e. 100 yr period), which we also mention now upon describing Eq. (6).

Additionally, we added the statement that the interannual correlations on Fig.4 are only slightly different between the datasets since the correlation is based on the integrated SNR values.

P14, l26-28: 'Together with the observed spectral shapes on the longer timescales (30-100y) our results therefore might indicate a true increase in local variability at the WAIS sites towards longer time scales, and a close-to-constant or even decreasing coherent signal variability.'

This sentence is strange. It looks unbalanced with a first part dealing with long time scales and a second part dealing with... probably shorter time scales?

Is it possible to correct this sentence?

To facilitate reading, maybe this paragraph could be limited to the issue of high frequencies at WAIS, since the discussion for the long timescales already existed in the previous one.

Is it possible to separate the issue of high frequencies for noise and for signal?

AC:

The sentence was intended to summarize both the findings on short and high frequencies: we argue before that deficiencies in the correction approach likely cannot explain the remaining decrease in noise level towards high frequencies. Taking into account additionally the results for the longer timescales, could then therefore suggest that the WAIS noise level tends to increase across the timescales studied and that the signal tends to stay constant or even decrease.

We suggest to clarify the sentence as follows: "By including the found spectral shapes of the signal and noise on the longer timescales (periods from 30-100 yr), together our results for the WAIS sites might therefore indicate that, across the timescales studied, there is a true increase in local variability and a close-to-constant or even decreasing coherent signal variability."

We changed the sentence to: "Taking additionally into account the spectral shapes we find for the signal and noise on the longer timescales (periods from ~30 –100 yr), our results might therefore overall indicate a true increase in local variability at the WAIS sites across the timescales studied and a close-to-constant or even decreasing coherent signal variability."

What are the hypotheses explaining this decrease in noise at high frequencies? Could it be something like wind blowing that would homogenize only the surface over large areas? Or large-scale evaporation of surface snow homogenizing its composition?

AC:

We see no obvious explanation for the decreasing noise at high frequencies and, given the uncertainty of our results, we would refrain here from suggesting such explanations which would be purely speculative.

2.1 Appendix A

P16, l 23: 'a Gaussian convolution kernel whose standard deviation is the diffusion length σ_i that is a function of depth (time) and depends on site i.'

This is very synthetic, but not very clear. Please add a reference to Appendix B where more details are provided. Please give some values of σ_i (and possibly the associated frequencies) at various depths. Is there a formula that relates σ_i to time (for each site)? If not, the use of the term 'function' might be misleading here. Indeed, σ_i seems to answer on density, temperature and pressure, and therefore to be more similar to an

adjustable parameter in the densification model.

AC:

We added the reference to Appendix B here and clarified the sentence as: "a Gaussian convolution kernel whose standard deviation is the diffusion length σ_i that is a function of depth (or, equivalently, time since deposition) and depends on site i due to its dependence on local temperature, atmospheric pressure and accumulation rate (Appendix B)."

Please note that to our understanding it is correct to state that σ_i is a function of depth: For constant temperature and pressure, the diffusion length is solely a function of firn density (Gkinis et al., 2014), and since density is a function of depth z , also $\sigma_i = \sigma_i(z)$. Assuming a Herron-Langway model for the firn density, there is an analytical solution which relates σ_i to the firn density (van der Wel, 2012).

We added some typical diffusion length values (both in cm and years) to the manuscript, which we think is most suitable in Appendix B (p.18 l.16).

In addition, we added some more details here on how we exactly calculate the diffusion lengths.

2.2 Appendix C

P19, l8: This sentence is complicated. It would be simpler to describe the procedure at EDML, and then say that the same method is applied to WAIS.

AC:

We rephrased the sentence as suggested.

Technical corrections:

1. P5, l. 12: "individual" typo

AC:

Thank you for spotting this typo which was corrected.

2. P6, l. 14: Please remove 'exemplarily' which is redundant with 'to illustrate our method'.

AC:

We agree and rephrased the sentence to: "In order to illustrate our method (Sect. 2.2), we first use the DML1 data set to demonstrate the individual steps involved in the analysis."

3. P11, l. 8: 'estimated signal term' Please add a reference to the equation (5 or 6).

AC:

We do not see a motivation for referencing Eq. 5 or 6 here, since these provide different quantities (i.e. the signal-to-noise ratio (SNR) and its integrated value). Instead we assume that the reviewer would welcome a reference to the relevant spectral quantities in the methods section 2.2 again for clarification. Since the term "estimated signal" has already been mentioned earlier in this paragraph, we provided this reference not here but directly in the first paragraph of the current section 4.1 (p.11, l.6-7): "We presented a method and the results of separating the variability recorded by Antarctic isotope records into two contributions: local variations ("noise", Eq. 4b) and spatially coherent variations ("signal", Eq. 4a)."

4. P16, l. 12: 'core sites' typo

AC:

Thank you for spotting this typo was corrected.

5. P17, l.8: It is $F(\epsilon_i)F(\epsilon_j) \neq 0$ only for $i=j$, and hence... Maybe 'we have' instead of 'it is'?

AC:

We changed the sentence as suggested.

Thomas Münch and Thomas Laepple

References to the author comments:

Gkinis, V., et al., Earth Planet. Sci. Lett., 405, 132–141, [https://doi.org/ 10.1016/j.epsl.2014.08.022](https://doi.org/10.1016/j.epsl.2014.08.022), 2014.

van der Wel, L. G., Doctoral Thesis, University of Groningen, 146 pp., <http://hdl.handle.net/11370/72cf3b0b-d258-44a1-8830-b0c355ddb90>, 2012.

Author Reply to the Editor Comment by Lukas Jonkers
on the manuscript

cp-2018-112: What climate signal is contained in decadal to centennial scale isotope variations from Antarctic ice cores?

by Thomas Münch and Thomas Laepple

Dear Lukas Jonkers,

below you find our answer to your comment on the above mentioned manuscript; your comment is set in normal font, our answers (author comment, AC) in *green italic font*.

Even though the manuscript is based on data that have already been published elsewhere, the following points will add to the reproducibility of your results:

- Please include a data availability statement, including data citations, for the data used (DML and WAIS ice cores and ERA Interim reanalysis data).

AC:

We included a data availability statement to the revised manuscript including the data citations for the used DML, WAIS and ERA-Interim data and add the respective references to the reference list.

- Please consider to make code used in your study publicly available in online repository.

AC:

We added code availability statements to the revised manuscript. Final url and doi's of the respective code repositories are in the process of being registered and will be added in the later process.

Thomas Münch and Thomas Laepple

What climate signal is contained in decadal to centennial scale isotope variations from Antarctic ice cores?

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Abstract. Ice-core-based records of isotopic composition are a proxy for past temperatures and can thus provide information on polar climate variability over a large range of timescales. However, individual isotope records are affected by a multitude of processes that may mask the true temperature variability. The relative magnitude of climate and non-climate contributions is expected to vary as a function of timescale, and thus it is crucial to determine those temporal scales at which the actual signal dominates the noise. At present, there are no reliable estimates of this timescale dependence of the signal-to-noise ratio (SNR). Here, we present a simple method that applies spectral analyses to stable-isotope data from multiple cores to estimate the SNR, and the signal and noise variability, as a function of timescale. The method builds on separating the contributions from a common signal and from local variations and includes a correction for the effects of diffusion and time uncertainty. We apply our approach to firn-core arrays from Dronning Maud Land (DML) in East Antarctica and from the West Antarctic Ice Sheet (WAIS). For DML and decadal to multicentennial timescales, we find an increase of the SNR by nearly one order of magnitude (~ 0.2 at decadal and ~ 1.0 at multicentennial scales). The estimated spectrum of climate variability also shows increasing variability towards longer timescales, contrary to what is traditionally inferred from single records in this region. In contrast, the inferred variability spectrum for WAIS stays close to constant over decadal to centennial timescales, and the results even suggest a decrease in SNR over this range of timescales. We speculate that these differences between DML and WAIS are related to differences in the spatial and temporal scales of the isotope signal, highlighting the potentially more homogeneous atmospheric conditions on the Antarctic Plateau in contrast to the marine-influenced conditions on WAIS. In general, our approach provides a methodological basis for separating local proxy variability from coherent climate variations which is applicable to a large set of palaeoclimate records.

1 Introduction

Ice cores represent key archives for studying polar climate variability on timescales beyond instrumental observations. The isotopic composition of water stored in the ice serves as a proxy for reconstructing past temperature variations (Dansgaard, 1964; Jouzel and Merlivat, 1984; Jouzel et al., 2003) over a wide range of timescales ranging from subannual to glacial–interglacial variations. This can provide valuable insights into the timescale dependence of climate variability which is thought

to be an inherent property of the climate system (Hasselmann, 1976; North et al., 2011; Lovejoy et al., 2013; Rypdal et al., 2015).

However, the interpretation of isotope records in terms of local atmospheric temperatures is complicated by a multitude of processes that distort the original relationship present in precipitation (e.g., Fujita and Abe, 2006; Sjolte et al., 2011; Steen-Larsen et al., 2011; Touzeau et al., 2016; Casado et al., 2018). To a first approximation, changes in isotopic composition are only recorded in the ice if there is snowfall, while the role of water vapour exchange processes in between precipitation events is still debated (Steen-Larsen et al., 2011; Stenni et al., 2016; Casado et al., 2018; Ritter et al., 2016; Münch et al., 2017a). Both the seasonality and interannual variability in the seasonality of precipitation events is known to strongly affect the isotopic composition of snow layers, introducing a bias to the mean values, e.g., the annual average (Steig et al., 1994; Sime et al., 2009; Laepple et al., 2011), and adding variability to the signal (Persson et al., 2011; Laepple et al., 2018). Furthermore, uneven deposition of snow in combination with steady and strong surface winds lead to constant erosion, drift and vertical mixing of the surface snow, creating a strong spatial variability (Fisher et al., 1985; Münch et al., 2016, 2017a). Finally, once the snow is deposited, the diffusion of vapour through the firn column smoothes the isotope variations, ~~reducing the~~ (Johnsen, 1977; Whillans and Grootes, 1985). This has the strongest effect on short distances in the firn, significantly reduces ~~the high-frequency~~ power spectral density at high frequencies (Johnsen, 1977; Whillans and Grootes, 1985; Johnsen et al., 2000; van der Wel et al., 2015) and thereby substantially shapes the isotope variability (Laepple et al., 2018). Overall, these processes, being to a first approximation not directly linked to variations in temperature, add a significant amount of noise to the isotope records, especially in low-accumulation regions on the Antarctic Plateau. This has been demonstrated by the low representativity of individual ice-core measurements (Karlöf et al., 2006; Münch et al., 2016) and questions the usability of a direct interpretation of high-resolution isotope variations in terms of temperature variability (Laepple et al., 2018), particularly when the climate signal itself is only relatively weak. While deep ice cores show a consistent picture for the strong glacial–interglacial variations (e.g., Jouzel et al., 2007), it may be questionable whether the Holocene variability recorded in individual cores depicts the true temperature variability (Kobashi et al., 2011). Spatial or temporal averaging of isotope records thus provide important tools to reduce overall noise (Fisher et al., 1996; Münch et al., 2016; Stenni et al., 2017).

While previous studies provided first insights into the relationship between climate signal and noise for short spatial and temporal scales (Fisher et al., 1985; Münch et al., 2016), an extension to longer scales, which is important for the interpretation of Holocene temperature reconstructions, is still missing. Furthermore, ~~knowing the timescale dependence, and thus the spectral shape of the noise, would allow the correction of the isotope-inferred variability estimates for the noise contribution.~~ in order to correct the isotope inferred variability estimates for the noise contribution, it is necessary to know the variance of the noise across timescales, i.e. its spectral shape.

Here, we present a simple spectral method to separate the local noise from spatially coherent signals using information from several isotope records, including a correction for diffusion and time uncertainty. We apply our model to two spatial arrays of firn cores: (1) from Dronning Maud Land in East Antarctica, and (2) from the West Antarctic Ice Sheet. Our objective is to derive an improved estimate of the temperature variability in both regions and to learn about the timescale dependence of

the signal-to-noise ratio in ice-core isotope data. For Dronning Maud Land, our results confirm the noise levels inferred in previous studies on short temporal scales (Münch et al., 2016, 2017a) and also suggest white noise on longer timescales, which results in an increase of the isotopic signal-to-noise ratio (SNR). Unexpectedly, the SNR ~~as~~ inferred from the West Antarctic data is found to show the opposite timescale dependence. These results may point towards marked differences in the spatial and temporal scales of the isotope signals and reveal gaps in our current understanding of the underlying processes.

2 Data and methods

2.1 Isotope records from Dronning Maud Land and West Antarctica

We analyse published oxygen isotope records of annually dated firn cores from two contrasting Antarctic regions (Table 1): (1) Dronning Maud Land (DML) on the Antarctic Plateau, and (2) the West Antarctic Ice Sheet (WAIS). While the DML core sites are located in a rather flat region relatively isolated from the coast (average elevation 2900 m), the WAIS core sites are lower in elevation (1600 m on average) and therefore potentially more influenced by marine conditions.

For DML, we use a total of 15 records which were collected during the EPICA (European Project for Ice Coring in Antarctica) pre-site survey (Oerter et al., 2000) and published in Graf et al. (2002a). All records cover at least the last ~ 200 yr and form our data set DML1. Three of these records (B31–B33) span the last ~ 1000 yr and are therefore included in a second separate data set for this time span (DML2). Core B32 is in close proximity (~ 1 km) to the deep EPICA DML (EDML) ice-core site at Kohnen Station (EPICA community members, 2006). The ~~time uncertainty of the record chronologies, which record chronologies~~ were established from seasonal layer counting of chemical impurity records constrained by tie points from the dating of volcanic ash layers (Graf et al., 2002a). ~~The resulting uncertainty of the chronologies~~ has been reported ~~with to be~~ $\sim 2\%$ of the time interval to the nearest tie point (Graf et al., 2002a). ~~This, which~~ translates to a maximum time uncertainty of ~ 1.2 yr for the short and of ~ 3.5 yr for the long records. Since our method (Sect. 2.2) relies on all records having equal lengths, we restrict the time span for the DML1 data set to the period from 1801–1994 CE and to the period from 1000–1994 CE for the DML2 data set.

The WAIS data set selected for this study consists of five isotope records (Steig et al., 2013) collected during the US ITASE (International Trans-Antarctic Scientific Expedition) project (Mayewski et al., 2005), including the core WDC2005A from the WAIS Divide ice core (WDC; WAIS Divide Project Members, 2013), and cover the time interval from 1800–2000 CE. The oxygen isotope data of cores ITASE-2000-4 and ITASE-2000-5, previously published at subannual resolution (Steig et al., 2013), has been block-averaged in this study to obtain annual resolution data. The core selection has been based on the constraint to cover a similarly large area and sufficiently long time period to allow a meaningful comparison with the DML1 records. The relative time uncertainty between the WAIS records, based on dating through annual layer counting of chemical trace species validated by identification of volcanic marker horizons, was reported to be no more than 1 yr (Steig et al., 2005).

Table 1. Overview of the firn cores (sorted into three data sets) used in this study. Listed are the covered time span of each core array (in yr CE), the number of records in each array (N_n), the region of origin (latitude/longitude range), the range of site elevations, local accumulation rates (\dot{b}) and 10 m firn temperatures (T_{firn}) as reported in the original publications, and the range of intercore distances (d). The range of WAIS firn temperatures is based on ERA-Interim annual mean anomalies (Dee et al., 2011) with respect to the WDC2005A site.

Core array (Time span)	N_n	Region (Lat./Lon.)	Elevation m a.s.l.	\dot{b} $\text{kg m}^{-2} \text{yr}^{-1}$	T_{firn} °C		d km	Data reference
					min	max		
DML1 ^a (1801–1994)	15	74.5–75.6° S 6.5° W–6.5° E	2630–3160 ^d	40–90 ^d	–46 ^d	–40 ^d	1–370	Graf et al. (2002b)
DML2 ^b (1000–1994)	3	75.0–75.6° S 3.4° W–6.5° E	2670–3160 ^d	50–60 ^d	–46 ^d	–44 ^d	120–280	Graf et al. (2002b)
WAIS ^c (1800–2000)	5	77.7–80.6° S 111.2–124.0° W	1350–1830 ^{e,f}	140–220 ^{e,f}	–30 ^{f,g}	–28 ^{f,g}	20–340	Steig (2013)

^aFirn cores FB9804, FB9805, FB9807–FB9811, FB9813–FB9817, B31–B33 ^bFirn cores B31–B33 ^cFirn cores WDC2005A, ITASE-1999-1, ITASE-2000-1, ITASE-2000-4, ITASE-2000-5 ^dOerter et al. (2000) ^eKaspari et al. (2004) ^fWAIS Divide Project Members (2013) ^gDee et al. (2011)

2.2 Model for the separation of signal and noise in the spectral domain

Our approach in general assumes that individual isotope records from a certain region contain two contributions: (i) a signal common to all cores from that region, and (ii) independent noise components which are, for example, related to spatial variability from redistribution of snow (stratigraphic noise) or to precipitation intermittency. By utilising several records we can disentangle both contributions and estimate the underlying common and noise signals. The approach is similar to the analysis of variance (e.g., Fisher et al., 1985) but uses the power spectra of the time series to derive timescale-dependent estimates.

More formally, given a core array of N_n isotope records, the power spectral density (PSD) of an individual record from site i , $\mathcal{X}_i(f)$, where f denotes frequency, is the sum $\mathcal{X}_i(f) = \mathcal{C}(f) + \mathcal{N}_i(f)$, where $\mathcal{C}(f)$ and $\mathcal{N}_i(f)$ are the original spectra of the common signal and the noise component, respectively, prior to the smoothing by molecular diffusion of water vapour within the porous firn (Johnsen et al., 2000; van der Wel et al., 2015). To relate this with the actually measured signal $\hat{\mathcal{X}}_i(f)$, we additionally have to account for the measurement process which adds additional noise to the diffused record. $\hat{\mathcal{X}}_i(f)$ is thus given by

$$\hat{\mathcal{X}}_i(f) = \mathcal{G}_i(f) [\mathcal{C}(f) + \mathcal{N}_i(f)] + \Sigma. \quad (1)$$

Here, $\mathcal{G}_i(f)$ is a linear transfer function that acts as a low-pass filter and accounts for the diffusion process (Appendix A), and Σ is the measurement error. At annual resolution, Σ is typically at least one order of magnitude smaller than the stratigraphic noise level (Münch et al., 2016) and is therefore neglected in the following. We now calculate the mean spectrum, $\mathcal{M}(f)$, of all individual spectra $\hat{\mathcal{X}}_i(f)$. Assuming that the statistical properties of the individual noise terms are identical for all N_n records,

we obtain

$$\mathcal{M}(f) = \frac{1}{N} \frac{1}{n} \sum_{i=0}^N \hat{\mathcal{X}}_i(f) \hat{\mathcal{X}}_i(f) = \bar{\mathcal{G}}(f) [\mathcal{C}(f) + \mathcal{N}(f)], \quad (2)$$

where $\bar{\mathcal{G}}(f) = 1/N \sum_{i=1}^N \mathcal{G}_i(f)$ $\bar{\mathcal{G}}(f) = 1/n \sum_{i=1}^n \mathcal{G}_i(f)$ is the average diffusion transfer function. In contrast to forming the spectral mean, we can also average the $N-n$ isotope records in the time domain and then calculate the PSD of this “stacked” record. Here, the noise component will be reduced by a factor of $N-n$ compared to the mean spectrum in (2) since we assume independent noise between the sites. Additionally, we have to take into account the time uncertainty of the dated records. This does not affect the overall shape of broadband spectra (Rhines and Huybers, 2011) but diminishes the correlation between the records (Haam and Huybers, 2010) and thus their spectral coherence, which reduces the variance of the common signal in the stacked record. The PSD of the stacked record is thus

$$\mathcal{S}(f) \approx \bar{\mathcal{G}}(f) \left[\Phi(f) \mathcal{C}(f) + \frac{1}{N} \frac{1}{n} \mathcal{N}(f) \right], \quad (3)$$

where $\Phi(f)$ is a linear transfer function accounting for the effect of time uncertainty. Applying the average diffusion transfer function, $\bar{\mathcal{G}}(f)$, also to the spectrum of the stacked record is a valid approximation in the case of similar $\mathcal{G}_i(f)$ which we confirmed for our records (Appendix A). From [equations and the expressions for the mean spectrum \(\$\mathcal{M}\$, Eq. 2\) and the spectrum of the stacked record \(\$\mathcal{S}\$, Eq. 3\)](#) we can now derive expressions for the spectra of the common signal \mathcal{C} and the noise \mathcal{N} (omitting the explicit frequency dependence in the notation here),

$$\mathcal{C} = \frac{N}{N - \Phi^{-1}} \frac{n}{n - \Phi^{-1}} \Phi^{-1} \bar{\mathcal{G}}^{-1} \left[\mathcal{S} - \frac{N-n}{N-n} \mathcal{M} \right], \quad (4a)$$

$$\mathcal{N} = \frac{N}{N - \Phi^{-1}} \frac{n}{n - \Phi^{-1}} \bar{\mathcal{G}}^{-1} \left[\mathcal{M} - \Phi^{-1} \mathcal{S} \right]. \quad (4b)$$

2.3 Transfer functions for vapour diffusion and time uncertainty

For estimating the transfer functions for diffusion and time uncertainty ([Eq.s, whose inverses are used to correct the signal and noise spectra \(Eqs. 4\)](#)), we use numerical simulations since no closed form expressions are available. For this, we create surrogate records mimicking the [individual individual](#) core arrays and simulate the effects of diffusion and time uncertainty on the surrogate spectra of the stacked records. We model the firm diffusion as the convolution of the original time series with a Gaussian kernel (Johnsen et al., 2000). The width of the kernel is set by the diffusion length σ , which is site-specific and a function of depth. For modelling the time uncertainty we use the approach of Comboul et al. (2014). Model parameters are the rates of missing and [doubly counted](#) annual layers which we set to match the reported time uncertainties of the isotope records. Appendix B gives a detailed description of the simulations including the estimated transfer functions (Fig. B1). Because the diffusion correction ($\bar{\mathcal{G}}^{-1}$ in Eqs. 4) strongly amplifies both the raw spectra as well as their uncertainties on the fast frequencies, we restrict our analyses to the frequency region where $\bar{\mathcal{G}} \leq 2$. [the estimated transfer function \$\bar{\mathcal{G}} > 0.5\$, equivalent to a correction factor \$\leq 2\$ in Eqs. 4.](#) This avoids large uncertainties and [yields cutoff frequencies translates to a maximum frequency that is](#)

used for the analyses (hereafter: cutoff frequency) of $1/5 \text{ yr}^{-1}$ for DML1, $1/8.5 \text{ yr}^{-1}$ for DML2 and $1/2.8 \text{ yr}^{-1}$ for WAIS (Fig. B1).

2.4 Timescale-dependent estimate of the signal-to-noise ratio

The frequency-dependent SNR, $F(f)$, is defined as the ratio of the signal and noise spectra,

$$5 \quad F(f) = \frac{\mathcal{C}(f)}{\mathcal{N}(f)}. \quad (5)$$

As we explicitly neglect measurement noise, this quantity is unaffected by the effect of diffusion or its correction (Eqs. 4) but directly influenced by time uncertainty which biases the uncorrected SNR towards zero. Typically, firn or ice-core isotope records are averaged onto a fixed temporal resolution Δt (the ‘‘averaging period’’) during preprocessing. ~~Therefore, we additionally provide the SNR~~ The signal-to-noise variance ratio after this averaging step ~~, using is given by~~

$$10 \quad \bar{F}(f_{\text{Nyq}}) = \frac{\int_{f_0}^{f_{\text{Nyq}}} \mathcal{C}(f) \, df}{\int_{f_0}^{f_{\text{Nyq}}} \mathcal{N}(f) \, df}, \quad (6)$$

where f_0 is the lowest frequency of the spectral estimates and f_{Nyq} the Nyquist frequency for the chosen averaging period, i.e. $1/(2\Delta t)$. Since our records have different lengths, we choose a common minimum value for f_0 of $1/100 \text{ yr}^{-1}$ for the integrations in (6). The value of \bar{F} ~~is related to~~ can then be used to obtain the correlation between the time series of the common signal c and a ~~‘‘stacked’’ stacked~~ record \bar{x} built from the spatial average of ~~N individual records, with the correlation~~ given by n individual records (Fisher et al., 1985):

$$15 \quad r(c, \bar{x})(f_{\text{Nyq}}) = \frac{1}{\sqrt{1 + (N\bar{F}(f_{\text{Nyq}}))^{-1}}} \frac{1}{\sqrt{1 + (n\bar{F}(f_{\text{Nyq}}))^{-1}}}. \quad (7)$$

2.5 Spectral analysis

Missing years in the published records ($\sim 1.6\%$ of all data points) are linearly interpolated. ~~Spectra~~ Power spectra are then estimated using Thomson’s multitaper method with three windows (Percival and Walden, 1993) with linear detrending before
 20 analysis. This approach is known to introduce a small bias at the lowest frequencies, and we omit the lowest frequency from all plots ~~. For display purposes, the spectra are shown on logarithmic axes and smoothed using a Gaussian kernel of constant width and calculations.~~ In general, spectral smoothing is necessary to improve the quality of the estimates from short time series. Here, we use a Gaussian smoothing kernel with varying width proportional to the applied frequency and thus constant in logarithmic frequency space (Kirchner, 2005). The smoothing width in logarithmic units is chosen to be 0.1 for the WAIS
 25 data and 0.15 for the DML data. To avoid biased estimates at the low- and high-frequency boundaries, the kernel is truncated on both sides to maintain its symmetry. Logarithmic smoothing preserves power-law scaling of spectra and is thus useful for climate spectra.

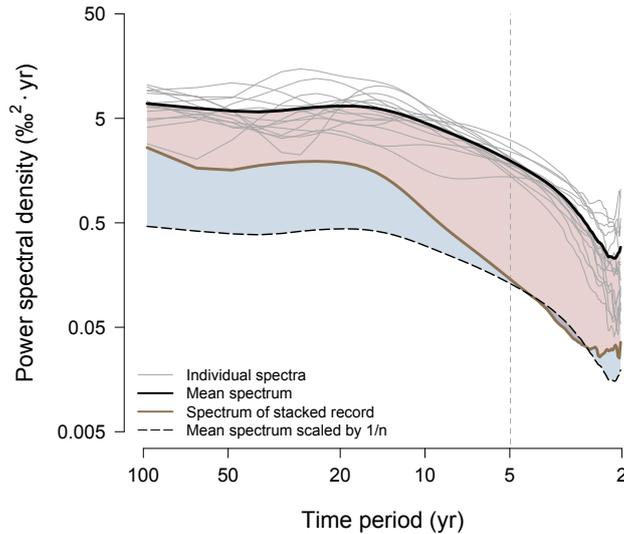


Figure 1. Detailed results of estimated PSD for the DML1 data set. Thin grey lines show the individual power spectra for each record with the mean spectrum indicated by the thick black line. The dashed black line shows the null hypothesis according to which all isotope variations are noise; the brown line depicts the spectrum from averaging all records in the temporal-time domain (the “stacked” record). The extent of the blue (red) shadings is proportional to the uncorrected PSD of the signal (noise) (Eq. 4). The vertical dashed line denotes the cutoff period for the diffusion correction (see Methods).

3 Results

3.1 Illustrating the methodological approach

In order to illustrate our method (Sect. 2.2), we first present exemplarily for use the DML1 data set to demonstrate the individual steps involved in the analysis.

- 5 Each power spectrum derived from an individual record of the DML1 firm-core array provides a timescale-dependent representation of the isotope variations in the study region (Fig. 1, thin grey lines). The differences between the individual spectra are not only due to differences in the isotope variability between the records, but also caused by the spectral uncertainty from the finite length of each record. The mean over-spectrum, \mathcal{M} , of all individual spectra (Fig. 1, black line) reduces this spectral uncertainty and thus provides a more robust estimate of the single records’ spectrum. This mean spectrum shows a (Fig. 1, black line). It shows a two-part structure with a near constant PSD above decadal timescales (i.e. is “white”) and a strong decrease in spectral power towards shorter timescales which is expected from vapour diffusion in the firm (Eq. 2 and Appendix B).

- The mean spectrum divided by the number of records (here, $N=15$ \mathcal{M}/n ; here, $n=15$; Fig. 1, dashed line) is the expected spectrum if all isotope variations present in the firm-core records were independent noise. In comparison, averaging in the time domain across records that contain noise and additionally a common (i.e. spatially coherent) signal, will result in a spectrum \mathcal{S} , where the noise component is also reduced by $1/N$ $1/n$ but with the common signal left unaltered (Eq. 3, neglecting time

uncertainty), and which is thus located between the mean spectrum and the mean spectrum divided by N . For short timescales (periods from 2 to ~ 7 yr), the spectrum of n . This spectrum \mathcal{S} for the DML1 stack (Fig. 1, brown line) is for short timescales (periods from ~ 3 to 5 yr) nearly identical to the mean spectrum divided by Nn , which suggests that here the variability of the individual records is consistent with the null hypothesis of independent noise. The divergence from the white-noise level close to the 2-yr Nyquist period is likely an artefact from the jitter of the annual averages as a result of the uncertainty in defining annual depth increments upon dating. In contrast to these short periods the short periods below 10 yr, the individual records clearly contain a common isotope signal on longer timescales that “survives” the averaging process and when building the stacked record and which increases in power with increasing timescale (Fig. 1, brown line). The differences between the spectrum of the stack and the mean spectrum \mathcal{S} and \mathcal{M} , or, respectively, the mean spectrum divided by $N\mathcal{M}/n$, hence inform us about the average signal and noise content of the individual isotope records but need to be corrected for the residual amount of noise contained in \mathcal{S} and for the loss of variance from diffusion and time uncertainty (EqEqs. 4).

3.2 Timescale dependence of DML and WAIS signal and noise variability

After this detailed description for DML1, we now turn towards the results of the estimated signal and noise spectra for all three data sets. Unlike the average spectrum across all In general, the shape of the signal spectra is, as a result of the corrections, clearly distinct from the mean spectrum of the individual isotope variations (Fig. 1). This is seen, for example, in the corrected DML1 signal spectrum shows an increase of power spectral density with increasing timescale which indicates a much more steady increase in PSD from short to long timescales (Fig. 2a), solid blue line) as compared to the mean spectrum (Fig. 1, black line). This is confirmed by the three 1000 yr long records from the DML2 data set records whose signal exhibits a roughly similar power spectrum in the range of timescales that overlap. This is partly expected We partly expect this since the longer cores are also part of the DML1 data set, but the increase in signal power in this frequency band seems to be a general feature over the entire last 1000 yr. In addition, the DML2 signal spectrum shows a further and similar increase for timescales beyond the 100 yr period. In the frequency band where the corrections for The change in spectral shape from the mean to the signal spectrum results from two contributions in the correction process: Firstly, the average isotope variability is corrected for the noise contribution (“uncorrected signal”, dotted lines in Fig. 2a). This correction is mostly apparent on the longer timescales leading to a higher slope in PSD of the signal spectrum compared to the mean spectrum (for DML1, the signal increases from the 10 to 100 yr period by a factor of ~ 2.6 , the mean spectrum by a factor of ~ 1.4). Secondly, the corrections for the loss in spectral power by the effects of diffusion and time uncertainty are important (here important for time periods from ~ 20 to 5 yr periods), the corrections cause lead to a smaller increase in PSD of the signal power spectra with increasing timescale as compared to the raw estimates (compare dotted, dashed and solid lines in Fig. 2a). The corrected signal spectrum of the WAIS records In sharp contrast to DML, the corrections for the WAIS records yield a signal spectrum (Fig. 2c) is in stark contrast to the DML spectra, exhibiting, solid line) that exhibits a rather flat appearance throughout and indicating indicates decreasing PSD towards centennial timescales. Much of this the flat character is caused by the diffusion and time uncertainty correction which causes a strong amplification of corrections which strongly amplify the raw signal power for subdecadal timescales. on the subdecadal timescales; the decrease

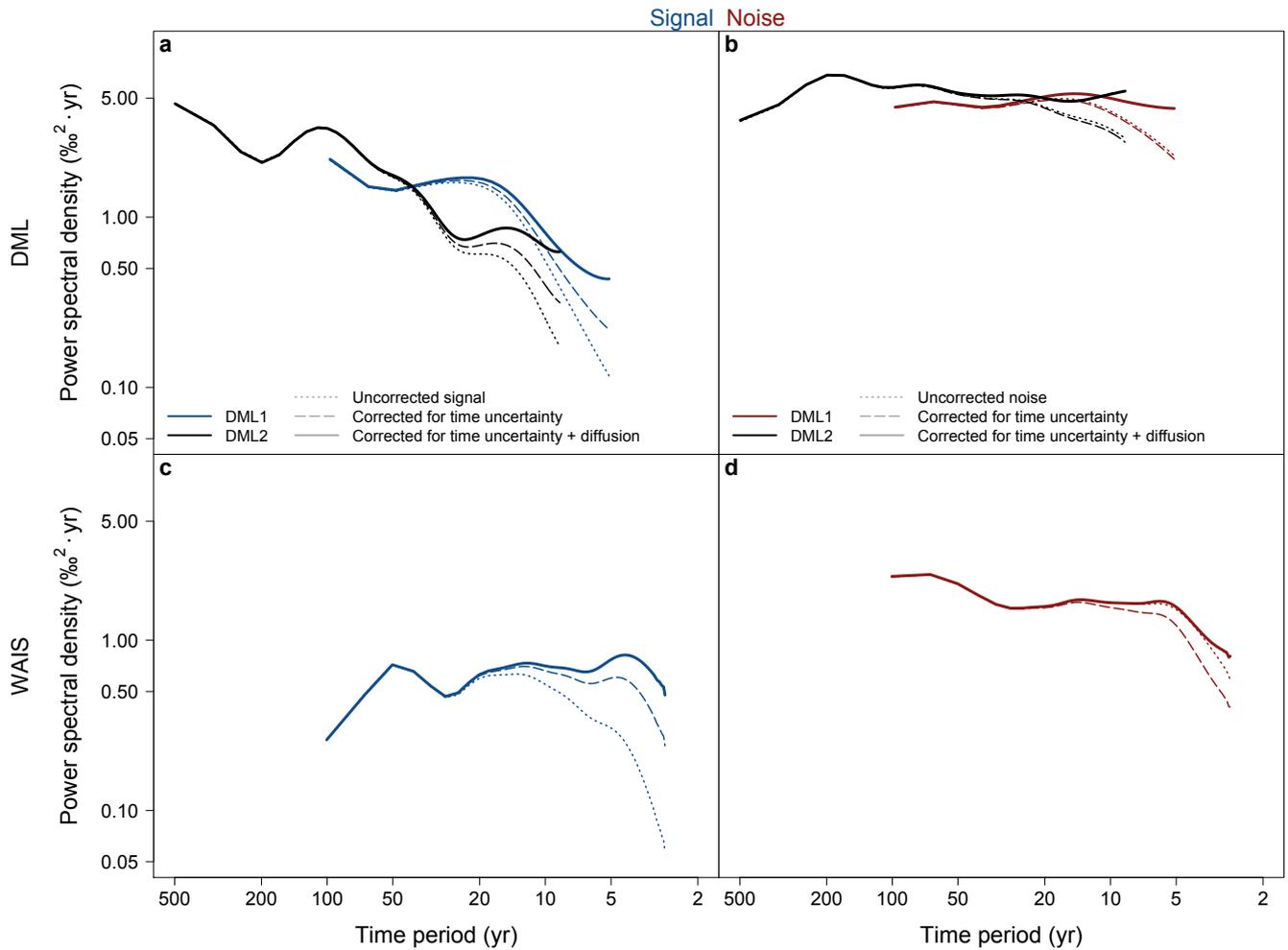


Figure 2. Estimated signal (left) and noise (right) spectra of Antarctic isotope records. Results are shown based on the spectral correction model for the records from Dronning Maud Land (DML, top row) and from the West Antarctic Ice Sheet (WAIS, bottom row). The raw estimates (dotted lines) show the spectra prior to any correction, while corrected estimates differentiate between the correction for time uncertainty only (dashed lines) and the full correction including time uncertainty and diffusion (solid lines). For DML, results include both the short, 200 yr long records (DML1; blue lines in **a**, red lines in **b**) and the longer records covering the last 1000 yr (DML2; black lines).

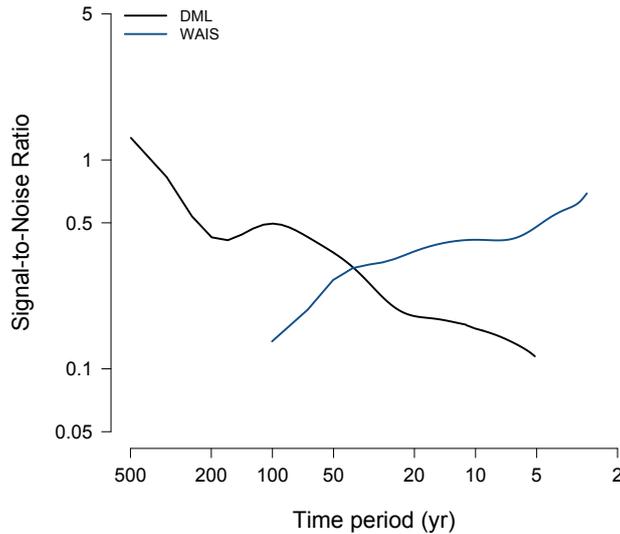


Figure 3. Estimated timescale dependence of ice-core isotope signal-to-noise ratios. Results are shown for the DML (black) and WAIS (blue) isotope records. The results for DML are based on combining the spectra from DML1 and DML2 (see text).

in PSD on the long timescales is a result of the correction for the noise contribution. We note, however, that the spectral uncertainty at the lowest frequency is high since less data points contribute to the estimated PSD for lower frequencies when using logarithmic smoothing.

A ~~similar~~ difference between both regions can also be seen in the noise spectra (Fig. 2b+d). Prior to any correction, the DML1 and DML2 noise spectra are different on shorter timescales ($\lesssim 20$ yr period), but become consistent to each other after the correction showing essentially white PSD (Fig. 2b). In comparison, the corrected WAIS noise spectrum shows an increase of spatially incoherent isotope variations towards longer timescales, which correlates with the decrease in signal power on centennial timescales (Fig. 2c).

The corrected signal and noise spectra directly provide an estimate of the timescale dependence of the SNR (Eq. 5). To derive a single estimate for DML, we combine the DML2 spectra for timescales above the decadal period with the respective DML1 spectra from the subdecadal frequency band. Again, the results are very different between DML and WAIS (Fig. 3). For DML, the SNR shows a continuous increase with timescale with values ranging from ~ 0.1 at the 5 yr period to $\lesssim 0.2$ at the decadal timescale and increasing further to ~ 1 at multicentennial timescales. The estimate for WAIS exhibits the opposite timescale dependence with SNR values of ~ 0.5 – 0.7 for interannual timescales that continuously decrease to ~ 0.1 at centennial timescales.

A complementary picture is obtained from the expected correlation between the time series of a stacked record and the underlying common signal as a function of both the number of averaged records and the temporal averaging period (Eq. 7, Fig. 4). For DML, the correlation for an individual record at interannual resolution is rather low with roughly 0.4–0.5 (Fig. 4a); for WAIS, this correlation is slightly higher (~ 0.5 – 0.6 , Fig. 4b). For longer averaging periods, the correlation for DML shows

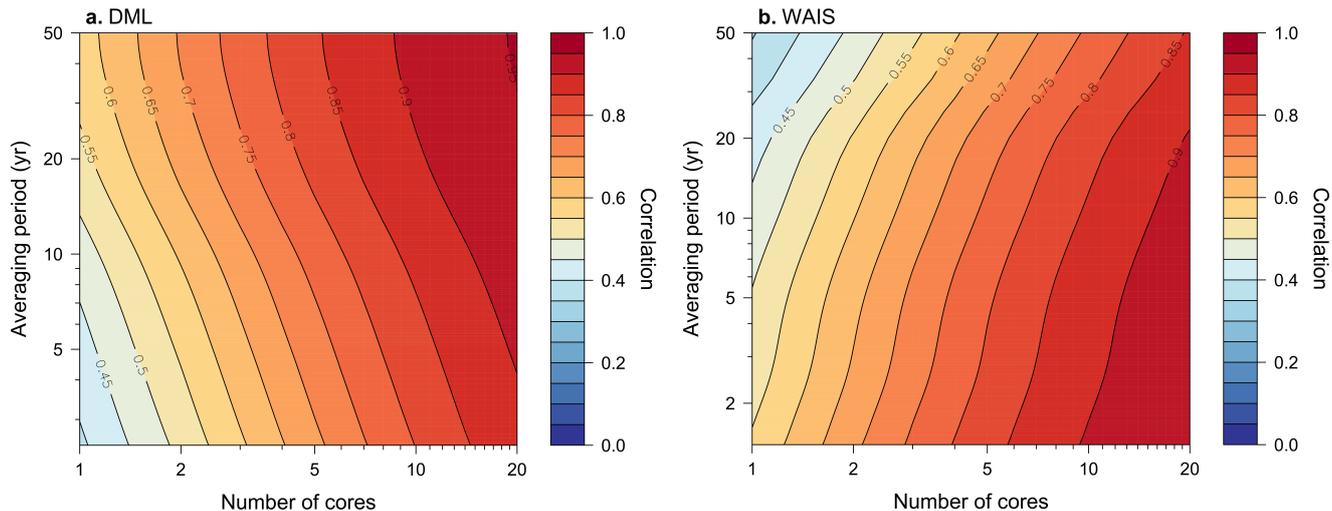


Figure 4. Estimated correlation of stacked isotope records from (a) DML and (b) WAIS with their common signal as a function of the averaging period and the number of firm cores included in the average. The correlations are based on Eq. (7) with the same signal and noise spectra that were used for Fig. 3.

a steady increase with the averaging period, in line with the increase in SNR, reaching values of about 0.7 around 0.6 for single records and centennial-multidecadal averaging periods (Fig. 4a). Naturally, the correlation further increases with the number of averaged cores. For WAIS, the correlation with the underlying signal decreases with the averaging period and only increases by averaging more records (Fig. 4b).

5 4 Discussion

4.1 Physical interpretation of the analysis

We presented a method and the results of separating the variability recorded by Antarctic isotope records into two contributions: local variations (“noise”, Eq. 4b) and spatially coherent variations (“signal”, Eq. 4a). We now assess the physical meaning of these terms by discussing the relevant spatial scales of the major processes that influence the isotopic composition of the records: (i) temperature variations, (ii) atmospheric circulation, (iii) precipitation intermittency, and (iv) stratigraphic noise.

Classically, the isotopic composition of firm and ice cores is interpreted as being related to variations in local air temperature (Dansgaard, 1964; Jouzel and Merlivat, 1984; Jouzel et al., 2003). The extent to which spatially distributed isotope records are influenced by a common temperature signal then depends on the decorrelation scale of the temperature anomalies, which is, for annual mean temperatures, typically of the order of hundreds to thousands of kilometres and increases with timescale (Jones et al., 1997). For our study regions, annual mean temperature variations from the ERA-Interim reanalysis (Dee et al., 2011) exhibit decorrelation scales of ~ 1200 km (Appendix C) which are much larger than the individual intercore distances

(< 400 km, Table 1). This implies that, if the temperature variations were fully recorded by the array of firn-core records, they would to a large extent contribute to the estimated signal spectrum. In fact, [for the temperature reanalysis fields in the study regions, the average of all correlations between sites separated by less than the average correlation of 0.87 \(0.94\) over distances up to the maximum intercore distances for DML \(WAIS \(Fig. C1\)\)](#) suggests that the estimated signal spectra [for DML \(WAIS\)](#) would contain 87% (94%) of the total temperature variability while the remaining fraction would be interpreted as noise.

However, changes in large-scale atmospheric circulation can lead to variations in the source and the pathways of the moisture, which can affect the isotopic composition of the precipitation that is formed, independent of local temperature changes (e.g. [Schlosser et al., 2004](#)) (e.g., [Schlosser et al., 2004](#)). In general, the spatial scales of such variations are unclear. For the studied DML core array, which is located on the rather flat and remote Antarctic Plateau, one could expect that the effect is small and possibly spatially coherent, thus contributing to the estimated signal term. For WAIS, isotope data have been interpreted to also reflect changes in atmospheric circulation and sea-ice cover of the adjacent oceans (Küttel et al., 2012; Steig et al., 2013). These processes might exhibit, through the ice-sheet topography, a regional expression at the individual firn-core sites, which are generally lower in elevation and located near the ice divide, as suggested by the significant spatial differences in observed intercore correlations (Küttel et al., 2012). Such effects would affect the estimated noise component.

Major additional contributions to the overall variability of isotope data arise from precipitation intermittency, i.e. interannual variations in the seasonality of precipitation events (Persson et al., 2011; Laepple et al., 2018), and from stratigraphic noise (Fisher et al., 1985; Karlöf et al., 2006; Münch et al., 2016) created during deposition of the snow. Both processes exhibit significantly different decorrelation scales. The decorrelation scale of precipitation intermittency is related to the decorrelation scale of the precipitation itself, which is expected to be coherent on a local scale (i.e. ~ 100 m) and to decorrelate on scales generally smaller than the temperature decorrelation scales. In both regions, the similar analyses of ERA-Interim data as for the temperature variations suggests decorrelation scales of 300–500 km (Appendix C). This is supported by an analysis of measurements from autonomic weather stations (Reijmer and van den Broeke, 2003) that indicate that accumulation in DML arises from many small (1–2 cm) and few larger events ($\gtrsim 5$ cm) of snowfall, where the major events occur only a few times per year without any clear seasonality but often over large areas. In contrast, stratigraphic noise is a short-scale phenomenon. Its generation is related to the uneven deposition (Fisher et al., 1985) and the constant erosion and redeposition of the surface snow by wind (Münch et al., 2016). Both processes are connected to the local surface undulations as these directly interact with the snow deposition but also strongly shape the near-surface wind field. This is suggested by the observed decorrelation scale of the stratigraphic noise at EDML (< 5 m, Münch et al., 2016) which is similar to the typical scale of the surface undulations.

These differences in decorrelation scales also become apparent when analysing the relative contributions of both processes to the total isotope variability. At EDML, stratigraphic noise provides $\sim 50\%$ of the total variance at the seasonal timescale, as suggested from the average correlation of individual shallow isotope profiles separated above (> 5 m) the decorrelation scale of the stratigraphic noise but below 1 km (Münch et al., 2016, 2017a). A much higher noise level of at least 80% of the total variance (i.e. higher by a factor of ≥ 1.6) has been independently inferred from comparing the observed seasonal isotope variability to the expectation from a profile that is a mixture of a deterministic seasonal cycle and diffused noise (Laepple et al., 2018). The apparent mismatch with the noise level from stratigraphic noise can be reconciled by asserting that a significant

part of the common isotope signal at EDML on the local scale is coherent noise from precipitation intermittency, leading to an underestimation of the total noise level when analysing only the interprofile correlations. On the larger spatial scales of the here analysed firn-core arrays (“array scale”), the effect of precipitation intermittency should then appear, at least partly, as a noise term since the spatial precipitation patterns are expected to decorrelate on these scales. Therefore, in general, the variability contribution by stratigraphic noise will fully appear in the noise spectra, while precipitation intermittency is expected to partly contribute to the noise spectra but also partly appear in the signal spectra.

In summary, given the large decorrelation scales of atmospheric temperature variations and the generally smaller scales of the other terms, one could interpret the estimated isotope signal spectra to a first approximation as temperature signals. However, this clearly will be an upper bound of the true temperature signal since also other processes can lead to spatially coherent isotope signals. Furthermore, we have neglected the transfer function from isotopic ratios to temperatures, and other less constrained effects that affect the isotope signal from the atmosphere to the snow (e.g., Casado et al., 2018).

4.2 Interpretation of the estimated signal and noise spectra

The raw noise spectra, i.e. prior to correction, derived from the two DML data sets exhibit a clear imprint from the diffusional smoothing in the firn, ~~as suggested from the smaller PSD for periods < 20 yr of the raw DML2 noise spectrum, i.e. prior to correction, in comparison to DML1.~~ This is suggested by their common decrease in PSD towards shorter time periods (Fig. 2b): This can be explained by the stronger, since diffusion acts stronger on higher frequencies. It is corroborated by comparing the loss in PSD between the two data sets, which for DML2 is stronger towards the high-frequency end and also extends further towards longer time periods. This is due to the stronger diffusional smoothing in the older sections of the cores that are only contained in the DML2 records, since the diffusion process had more time to act there since deposition of the snow. The applied correction reconciles both noise estimates, as it takes these differences into account, and reveals a close-to-constant noise level (“white noise”) across the range from subdecadal to multicentennial timescales.

This near constancy of the noise level suggests that the noise-creating processes are independent of the timescale. This seems plausible for stratigraphic noise as it is also indicated by the observed small memory in the interannual variations of the surface topography at EDML (Laepfle et al., 2016; Münch et al., 2016). Furthermore, we would not expect strong changes in surface wind regimes or depositional characteristics for the rather stable climatic conditions over the studied time periods. To test whether the effect of precipitation intermittency additionally contributes to the noise spectrum, as suggested from the decorrelation scales of the ERA-Interim precipitation fields, we apply our spectral analysis to the available isotope profiles on the local scale (~ 100 m) at EDML (Münch et al., 2017a) (Münch et al., 2017a, b) and compare the estimated noise spectra between the local scale and the array scale (Fig. 5). Although the spectral estimates do not overlap, the results indicate an offset between the noise levels at both spatial scales. At the 5 yr period, the PSD of the noise on the array scale is about 1.7 times higher than at the local scale. Some increase in the noise level is expected from the small contribution of uncorrelated temperature variability, but the ~ 1.7 -fold increase suggests that most of the additional noise arises from spatially incoherent precipitation intermittency.

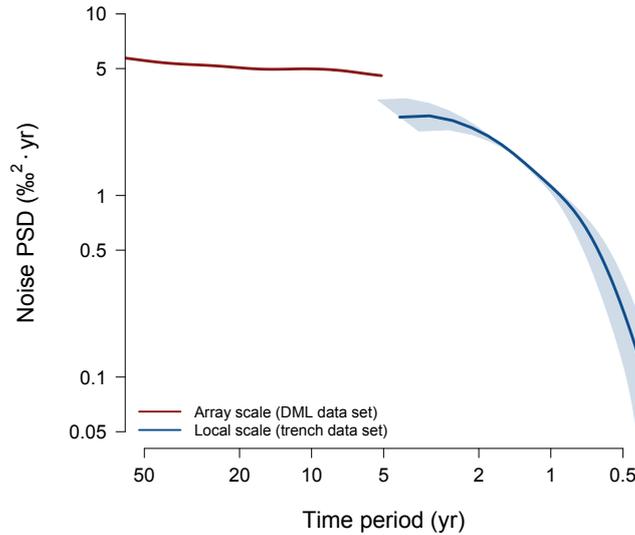


Figure 5. Comparison of DML noise spectra as a function of intersite distance. The red curve (array scale, ~ 100 km) shows the section for periods ≤ 50 yr of the composite noise spectrum from the DML1 and DML2 firn-core records (Fig. 2b and text). The blue curve (local scale, ~ 100 m) depicts the noise spectrum as estimated from 22 shallow (~ 3.5 m depth) trench profiles from EDML (Münch et al., 2017a) (Münch et al., 2017a, b). The depth scale of the trench data has been converted into time units using a constant accumulation rate of 25 cm of snow per year, neglecting the small compression by densification (Münch et al., 2017a). Blue shadings denote the spectral range from an assumed 20% uncertainty of the accumulation rate. Note that the trench noise spectrum is not corrected for the diffusional smoothing whose effect is negligible for the trench records at the 5 yr period and thus does not affect the comparison of the noise levels.

This supports the interpretation of the estimated DML signal spectra (Fig. 2a) as temperature variability, if the influence by atmospheric circulation changes is small. Fitting a power law of the form $f^{-\beta}$, where f denotes frequency, yields a slope of the DML2 signal spectrum of roughly $\beta \sim 0.6$ for variations above the 20 yr period. This is higher than the slopes of the mean spectra of the single records ($\beta \sim 0.2$ for DML2, $\beta \sim 0$ for DML1) which shows underlines again that the high noise level in individual records strongly masks the true spectral shape of the signal. Such a power-law increase in temperature variability is not unexpected since it was also observed in spectra from marine-proxy-inferred sea surface temperature variations (Laepple and Huybers, 2014) and from other proxy and instrumental temperature records (e.g., Pelletier, 1998; Huybers and Curry, 2006).

The resulting SNR (Fig. 3) of the DML data illustrates at which timescales the signal dominates the isotope variability recorded in individual records, and how this translates into the representativity of isotope variations when averaging records in space and time (Fig. 4). We can test the validity of our spectral approach by comparing the estimated SNR to published estimates. Graf et al. (2002a) analysed the same data set as the present study and found a SNR at annual resolution of $\overline{F} = 0.14$ based on the average intercore correlations ($r = 0.35$). Similar values have been obtained using a larger set of cores from the same region that cover a shorter time period (Altnau et al., 2015). From integrating the estimated signal and noise spectra up

to the minimum averaging period constrained by the diffusion correction (~ 2.5 yr) we find a SNR of $\bar{F} \sim 0.2$ (correlation $r \sim 0.41$, $r \sim 0.4$, Fig. 4a), which is slightly higher than the published estimates, as expected, since we corrected for the effect of time uncertainty, besides the effect from the slightly different underlying averaging periods. Our results further explain the strong agreement between individual ice cores on glacial–interglacial timescales. Averaging to centennial–multidecadal resolution results in a correlation between a single core and the signal of around 0.7 – 0.6 (Fig. 4a), which rises to ~ 0.7 for the available centennial averaging periods, and the steady increase in SNR on the analysed timescales (Fig. 3) might indicate a further increase in SNR towards even longer timescales. However, clearly our results also underline again that in order to assess for assessing the temperature variability on shorter timescales within the Holocene, the averaging of records is essential to increase their representativity (Fig. 4a).

For WAIS, the higher SNR at interannual timescales as compared to DML (average SNR between 5 and 10 yr periods of 0.43 for WAIS compared to 0.13 for DML; Fig. 3) is consistent with the general notion that higher accumulation rates (on average, ~ 3 times as high as at DML; Table 1) result in higher SNR (Fisher et al., 1985; Steen-Larsen et al., 2011). However, one would expect, to a first approximation, a similar increase in SNR with timescale in both regions. In contrast to this expectation, the WAIS results show strongly different timescale dependencies with the tendency of a decrease in signal power and an increase in noise level on longer timescales (Fig. 2), resulting in an overall decrease in SNR (Fig. 3). If these findings This explains the only slightly higher correlations at interannual timescales of a stacked isotope record with the common signal at WAIS as compared to DML (Fig. 4), since the correlation is based on the integrated value of the SNR (Eqs. 6 and 7). These findings suggest, if they are correct, they suggest no simple scaling of the SNR with accumulation rate but that additional effects need to be involved.

The shape of the signal and noise spectra at subdecadal timescales are sensitive to the diffusion and time uncertainty corrections (Fig. 2), which could thus contribute to the discrepant SNR estimates. While the diffusion correction for DML led to a consistent white-noise level of both noise estimates (Fig. 2b) at these timescales, lending support to our approach, the WAIS noise spectrum keeps decreasing towards higher frequencies even after applying the diffusion correction (Fig. 2d). This result could be caused either by a too weak diffusion correction, or by an overestimation of the time uncertainty leading to an excessive reduction in high-frequency noise levels that cannot be compensated by the low diffusion correction. To test both hypotheses, we varied the strengths of the diffusion and time uncertainty corrections. This has different impacts on the estimated spectra: while the diffusion correction equally applies to both raw signal and noise spectra, the time uncertainty correction has a proportional influence on the signal spectrum but only an indirect influence on the noise spectrum (Eqs. 4). Indeed, halving the time uncertainty would increase the noise spectrum only by a maximum factor of ~ 1.4 close to the cutoff frequency, and this therefore cannot reconcile the remaining decrease in the corrected noise spectrum. By contrast, an overall doubling of the diffusion length for all WAIS records would amplify the noise spectrum much stronger with a maximum factor of ~ 4 at the cutoff frequency, leading to interannual noise levels similar to those observed for centennial periods. However, such a strong diffusional smoothing at WAIS compared to the expectation seems implausible given that the same physical mechanisms of the diffusion process should be valid for East as well as West Antarctica, and no anomalously high diffusional smoothing has been observed for the upper part of the WAIS Divide ice core (Jones et al., 2017). Additionally, a much stronger

diffusion correction would also imply a much steeper decrease of the signal spectrum towards longer timescales (Fig. 2c), contrary to the expectation. We conclude that there is no obvious reason for the applied correction approach to strongly over- or underestimate the WAIS signal and noise spectra on the interannual timescales. ~~Together with the observed spectral shapes~~ Taking additionally into account the spectral shapes we find for the signal and noise on the longer timescales (periods from 5 ~ 30–100 yr), our results ~~therefore might might therefore overall~~ indicate a true increase in local variability at the WAIS sites ~~towards longer timescales, across the timescales studied~~ and a close-to-constant or even decreasing coherent signal variability.

These results suggest firstly that an additional noise process, apart from stratigraphic noise and precipitation intermittency, must contribute to the noise spectrum towards longer timescales. Secondly, the shape of the signal spectrum either implies a nearly white-noise temperature signal, or some process that destroys the coherence of the large-scale temperature field on 10 longer timescales upon recording by the firn-core isotope records. Since there is no obvious reason for a fundamentally different Holocene climate variability in West compared to East Antarctica, the second possibility seems more likely.

WAIS isotope data have been reported to covary with local temperatures, but also with the large-scale atmospheric circulation and the sea-ice cover of the adjacent oceans (~~Küttel et al., 2012; Steig et al., 2013~~)(Noone and Simmonds, 2004; Küttel et al., 2012; Steig et al., 2013). A pronounced regional imprint on the recorded isotope variability is also suggested from the spatial differences in intercore 15 correlations for an extended set of the available WAIS firn cores (Küttel et al., 2012). Especially ~~the eastern cores~~ cores east of the WAIS Divide display strong differences in the recorded isotope variability despite their rather small spatial separation (Gregory and Noone, 2008), which is suggested to be related to slow variations in the dominant circulation patterns. While such differences have not explicitly been reported ~~between our western core sites for the core sites west of the divide studied here,~~ we hypothesise that signatures of sea-ice variations or meridional inflow could affect the isotopic composition at the in- 20 dividual firn-core sites to different extents, ~~possibly due to topographic differences.~~ This is motivated by model-based results (Noone and Simmonds, 2004) linking an increase in sea-ice cover to an earlier ascent of long-range transported air masses to the continent, reducing the influence of local air from the surface, while less sea ice inhibits this early ascent and allows more mixing of surface air. Variations in sea ice could thus influence the isotopic composition of air masses by controlling the influence of local moisture, which might be recorded only by certain WAIS sites depending on the elevation of the air mass transport and the topography of the core sites such as distance to the coast or elevation. If such variations exhibit decadal trends or slower ~~variations~~ changes, it could destroy the recording of the large-scale coherent temperature field by the firn cores and thus cause a loss in spectral ~~power~~ signal power in the isotope record towards longer timescales and an increase in the noise level (Fig. 2c, d). ~~This might~~ In contrast, the East Antarctic Plateau including the DML firn-core sites is higher in elevation and might be more shielded from marine influences by the steep topographic slopes leading to a more coherent signal. Together, 30 this might also explain, besides differences in core quality, the rather low agreement among deep West Antarctic cores on millennial timescales compared to East Antarctic cores (WAIS Divide Project Members, 2013). However, since ~~these our~~ in-ferences are speculative, further studies are needed to help disentangle the role of variations in West Antarctic temperature ~~and atmospheric circulation,~~ atmospheric circulation and sea ice on the recorded isotope variability.

5 Conclusions

We presented a simple spectral method to separate the variations recorded by isotope records into a local (“noise”) and a common (i.e. spatially coherent) “signal” component. We applied this method to firn cores from the East Antarctic Dronning Maud Land (DML) and the West Antarctic Ice Sheet (WAIS) to estimate, for the first time, the isotopic signal-to-noise ratio (SNR) as a function of timescale. This is of fundamental interest for interpreting isotope records obtained from individual ice cores, since it provides an upper limit on the SNR of the temperature signal recorded by the cores. For DML, the SNR at the interannual timescale is very low, but it steadily increases with timescale reaching values above 0.5 for multidecadal and slower variations. Therefore, only on these timescales isotope changes from individual cores should be interpreted in terms of regional climate variations. For WAIS, the results are counterintuitive. On interannual to decadal timescales, the estimated SNR is higher than 0.5, which would support the regional climate interpretation of the isotope records. For longer timescales, however, the estimated SNR decreases, suggesting that local variations start to dominate the recorded isotope variability.

Our method further allows the estimation of the power spectra of the coherent isotope signal. For DML, the spectra of single cores largely resemble white noise. In contrast, the derived signal spectrum shows increasing variability towards longer timescales. Such an increase is also observed in instrumental temperature records and other climate proxies. The marked difference between the raw interpretation of single cores – as it is usually done – and the signal spectra derived from the core array demonstrates the relevance of the signal and noise separation. The interpretation of the WAIS isotope signal is more challenging, since the signal shows a close-to-constant spectral power even after applying our method. We speculate that this might be due to atmospheric circulation variations that create a local imprint at the different firn-core sites. This might prevent a coherent recording of the large-scale atmospheric temperature field. To test this hypothesis, we suggest to analyse firn-core arrays as a function of the average separation distance between the individual core sites within each array. This could allow us to investigate whether the stable-isotope data record a stronger coherent signal on a regional scale (e.g., ~ 1–10 km) than on the larger scales analysed here. A similar approach in DML would also help to better constrain the spatial correlation scales of the precipitation intermittency.

We conclude that the pronounced differences seen between East and West Antarctica could thus be related to the differences in the topographic settings and the different marine influence (WAIS Divide Project Members, 2013)(Noone and Simmonds, 2004; WAIS D). Attempts to reconstruct the natural temperature variability in these regions based on firn and ice cores should thus not only aim at averaging as many records as possible, but also consider the spatial coherence of the circulation and precipitation patterns in order to establish an optimal strategy for averaging, or obtaining new, firn-core isotope records. Additionally, extending our analyses to data derived from non-isotope-based temperature proxies could give further insights into the true spectral shape of temperature variability in Antarctica.

Finally, our approach of separating signal and noise from a set of spatially distributed records is also applicable beyond Antarctic ice cores. The challenge of low and timescale-dependent SNR is common to many high-resolution climate archives, and the number of nearby core sites is continuously increasing. Therefore, we envision our approach to better constrain

the reliability of proxy records as a function of timescale in general, and to allow a significant improvement of our knowledge of past climate variability.

Code and data availability. Software for the spectral analyses, the implementation of the method and the plotting of the results is available as the R package *proxysnr*; source code for the package is hosted in the public git repository at url-xxx, a snapshot of the code used for this publication (version 0.1.0) is archived under <https://doi.org/xxx> (Münch, 2018a). Software to model the time uncertainty of the isotope records is based on the MATLAB code by Comboul et al. (2014), which has been adapted for this publication and implemented in R as the package *simproxysnr*, its source code is available in the public git repository at url-xxx, a snapshot of the code used for this publication (version 0.1.0) is archived under <https://doi.org/xxx> (Münch, 2018b). The diffusion length calculations have been performed with the R package *FirmR* which is available on request from the authors.

10 The original DML firn-core and trench oxygen isotope data are archived at the PANGAEA database (<https://www.pangaea.de>) under <https://doi.org/10.1594/PANGAEA.728240> (Graf et al., 2002b) and <https://doi.org/10.1594/PANGAEA.876639> (Münch et al., 2017b), respectively. PANGAEA is hosted by the Alfred Wegener Institute Helmholtz Centre for Polar and Marine Research (AWI), Bremerhaven and the Center for Marine Environmental Sciences (MARUM), Bremen, Germany. The original WAIS firn-core oxygen isotope data are archived at the U.S. Antarctic Program Data Center (USAP-DC; <http://www.usap-dc.org/>) under <https://doi.org/10.7265/N5QJ7F8B> (Steig, 2013).
 15 USAP-DC is hosted by Lamont-Doherty Earth Observatory of Columbia University, Palisades, USA. The processed isotope data as used in this study, together with generating code, are provided with the package *proxysnr*. The ERA-Interim reanalysis data is available from the European Centre for Medium-Range Weather Forecasts (ECMWF) under <http://apps.ecmwf.int/datasets/data/interim-full-mode/levtype=sfc/> (European Centre for Medium-Range Weather Forecasts, 2018). All relevant site parameter data for the diffusion length calculations are available from the publications referenced in Table 1 and, together with the simulated diffusion length estimates, provided with the package
 20 *proxysnr*.

Appendix A: Theoretical spectra for site-specific diffusion

We derive the effect of site-specific diffusion on the estimates of the mean spectrum and the spectrum of the stack given a core array consisting of n isotope records.

The spectrum of a time series $x(t)$ is given by the absolute square of its Fourier transformation, $\mathcal{X} = \mathcal{F}(x)\mathcal{F}^*(x)$, where
 25 $\mathcal{F}^*(x)$ denotes the complex conjugate of $\mathcal{F}(x)$. Given an isotope record at site i , $x_i(t) = c(t) + \varepsilon_i(t)$, where $c(t)$ denotes the common signal and $\varepsilon_i(t)$ is noise, the time series after diffusion is $x'_i(t) = g_i \star (c(t) + \varepsilon_i(t))$. Here, \star denotes convolution and g_i is a Gaussian convolution kernel whose standard deviation is the diffusion length σ_i that is a function of depth ([time or, equivalently, time since deposition](#)) and depends on site i [due to its dependence on local temperature, atmospheric pressure and accumulation rate \(Appendix B\)](#). The spectrum of x' is then

$$\begin{aligned}
 \mathcal{X}_i &= \mathcal{F}(g_i)\mathcal{F}(c + \varepsilon_i) \cdot \mathcal{F}^*(g_i)\mathcal{F}^*(c + \varepsilon_i) \\
 30 \quad &= \mathcal{G}_i(\mathcal{C} + \mathcal{N}_i),
 \end{aligned}
 \tag{A1}$$

with $\mathcal{C} = \mathcal{F}(c)\mathcal{F}^*(c)$, $\mathcal{N} = \mathcal{F}(\varepsilon_i)\mathcal{F}^*(\varepsilon_i)$ and where the linear site-specific diffusion transfer function was defined as $\mathcal{G}_i := \mathcal{F}(g_i)\mathcal{F}^*(g_i)$ ¹. Assuming that the statistical properties of the individual noise terms are the same for all sites, $\mathcal{N}_i \equiv \mathcal{N}$, the mean across N/n such spectra is

$$\mathcal{M} = \frac{1}{N} \frac{1}{n} \sum_{i=1}^N \mathcal{X}_i^n \mathcal{X}_i = \bar{\mathcal{G}}(\mathcal{C} + \mathcal{N}) \quad (\text{A2})$$

5 with the average diffusion transfer function $\bar{\mathcal{G}} := \frac{1}{N} \sum_{i=1}^N \mathcal{G}_i$, $\bar{\mathcal{G}} := \frac{1}{n} \sum_{i=1}^n \mathcal{G}_i$.

We now calculate the spectrum of the average of N/n independent noise “signals” subject to site-specific diffusion, $y_{\text{noise}}(t) = \frac{1}{N} \sum_{i=1}^N g_i \star \varepsilon_i(t)$
 $y_{\text{noise}}(t) = \frac{1}{n} \sum_{i=1}^n g_i \star \varepsilon_i(t)$

$$\begin{aligned} \mathcal{S}_{\text{noise}} &= \mathcal{F}(y_{\text{noise}})\mathcal{F}^*(y_{\text{noise}}) \\ &= \frac{1}{n^2} \sum_i \sum_j \mathcal{F}(g_i)\mathcal{F}^*(g_j)\mathcal{F}(\varepsilon_i)\mathcal{F}^*(\varepsilon_j). \end{aligned} \quad (\text{A3})$$

It is We have $\mathcal{F}(\varepsilon_i)\mathcal{F}^*(\varepsilon_j) \neq 0$ only for $i = j$, and hence

$$10 \quad \mathcal{S}_{\text{noise}} = \frac{1}{N} \frac{1}{n} \bar{\mathcal{G}}\mathcal{N}, \quad (\text{A4})$$

which shows that averaging N/n independent noise records reduces the spectral power by $1/N \cdot 1/n$, as expected.

The average of N/n coherent signals subject to site-specific diffusion, $y_{\text{coh}}(t) = \frac{1}{N} \sum_{i=1}^N g_i \star c(t)$, $y_{\text{coh}}(t) = \frac{1}{n} \sum_{i=1}^n g_i \star c(t)$ has a spectrum of

$$\begin{aligned} \mathcal{S}_{\text{coh}} &= \mathcal{F}(y_{\text{coh}})\mathcal{F}^*(y_{\text{coh}}) \\ &= \frac{1}{n^2} \mathcal{F}(c)\mathcal{F}^*(c) \sum_i \sum_j \mathcal{F}(g_i)\mathcal{F}^*(g_j) \\ &= \frac{1}{n} \mathcal{C} \left(\bar{\mathcal{G}} + \frac{1}{n} \sum_{i \neq j} \mathcal{F}(g_i)\mathcal{F}^*(g_j) \right). \end{aligned} \quad (\text{A5})$$

15 For small differences in the diffusion lengths between the sites, the transfer functions are approximately the same and we can simplify the second term in brackets in Eq. (A5),

$$\frac{1}{N} \frac{1}{n} \sum_{i \neq j} \mathcal{F}(g_i)\mathcal{F}^*(g_j) \approx \frac{1}{N} \frac{1}{n} \sum_{i \neq j} \mathcal{F}(g_i)\mathcal{F}^*(g_i). \quad (\text{A6})$$

Rearranging the summation terms, this is identical to $(N-1)\bar{\mathcal{G}}(n-1)\bar{\mathcal{G}}$, and Eq. (A5) becomes

$$\mathcal{S}_{\text{coh}} \approx \bar{\mathcal{G}}\mathcal{C}. \quad (\text{A7})$$

20 In this approximation, the average over diffused coherent signals does not reduce the spectral power, as one would expect. We tested the approximation for our core arrays by comparing simulation results with surrogate data between the cases of

¹ We note that there is no closed form expression for \mathcal{G}_i ; however, in case of a constant diffusion length the transfer function is a Gaussian (van der Wel et al., 2015).

independent and coherent noise. We find that (A7) is a reasonable approximation for the full expression (A5). However, we note that (A5) has slightly less power on the high-frequency end as compared to (A7) since the site-specific diffusion destroys part of the coherence between the sites.

Appendix B: Estimates of the transfer functions for diffusion and time uncertainty

5 To estimate the transfer functions for diffusion (\bar{G}) and time uncertainty (Φ), we simulate for each core array N surrogate records of the same time duration as the original isotope records. For each record, we simulate the effects of diffusion and time uncertainty, respectively, and then calculate in each case and for each core array the spectrum of the average record ($\mathcal{S}_{\text{diffusion}}$ or $\mathcal{S}_{\text{time}}$). At the same time, the spectrum of the average record of the initial surrogate data (\mathcal{S}_0) is calculated without applying diffusion or time uncertainty. The entire process is repeated n times and the resulting spectra are averaged to reduce the spectral uncertainty. We report $\bar{G} = \langle \mathcal{S}_{\text{diffusion}} \rangle / \langle \mathcal{S}_0 \rangle$ and $\Phi = \langle \mathcal{S}_{\text{time}} \rangle / \langle \mathcal{S}_0 \rangle$ as the transfer functions, where $\langle \cdot \rangle$ denotes the average over the n simulations where we use $n = 10^5$.

For the diffusion simulations, each surrogate record is smoothed by convolution with a Gaussian kernel of width σ , where σ is the diffusion length, measured in yr, sensitive to ambient temperature, pressure and the density of the firn (Whillans and Grootes, 1985) and therefore site-specific and a function of depth. We treat the dependency on density according to Gkinis et al. (2014) with diffusivity after Johnsen et al. (2000); firn density is modelled according to the Herron–Langway model (Herron and Langway, 1980). Surface air pressure is calculated from the barometric height formula using the local elevation and firn temperature. For the surface firn density we assume a constant value of 340 kg m^{-3} for all sites. Range of input parameters and their references are given in Table 1. Missing temperature information for the cores FB9807, FB9811 and FB9815 is filled with the temperatures from the nearby cores B32 (1 km), FB9812 (19 km) and FB9805 (28 km), respectively. We simulate diffusion lengths measured in cm for core lengths of 500 m at a resolution of 1 cm, convert the values into yr, and interpolate them onto a regular time axis at a resolution of 0.5 yr. This higher temporal resolution compared to the annual target resolution of the isotope records is necessary for numerical stability of the diffusion results at the high-frequency end. The diffusion transfer functions are then interpolated in frequency space onto the frequency axis of the isotope records. Simulated diffusion lengths at a depth corresponding to 200 yr after deposition of the snow range, across the firn-core sites, from 0.6 to 1.3 yr (8.3–11.4 cm) for DML and from 0.3 to 0.6 yr (8.9–11.4 cm) for WAIS; for DML and 1000 yr after deposition of the snow diffusion lengths range from 1.1 to 1.6 yr (8.2–9.1 cm).

To simulate the effect of time uncertainty, we apply the modelling approach by Comboul et al. (2014). This model represents an unconstrained process, meaning that the time uncertainty monotonically increases with modelled age, thus deeper in the core. To account for volcanic tie points where the chronologies of the isotope records are fixed (neglecting time uncertainties of the volcanic chronologies themselves), we modify the Comboul approach by forcing the model back to zero time uncertainty at the reported volcanic tie points (constrained process) in form of a Brownian Bridge process. In contrast to diffusional smoothing, it is not a priori clear whether time uncertainty acts as a linear transfer function on the spectrum of the average record. We test this by using different spectral characteristics of the surrogate data, adopting the cases of white noise as well as power

laws with spectral slopes of $\beta = 0.5, 1, 1.5$ and 2 . ~~We find, and found~~ that the modelled effect of time uncertainty is insensitive to the spectral shape of the input data and indeed acts as a linear transfer function. The Comboul model includes as parameters the rates of missing and doubly counted annual layers. We assume equal rates for these processes which we set such so that the yielded maximum time uncertainties (maximum standard deviation over simulated chronologies) match the reported time

5 uncertainties of the isotope records. The reported value of the DML records ($\sim 2\%$ of the time interval to the nearest tie point, Graf et al., 2002a) erroneously implies that the uncertainty increases linearly with time (Comboul et al., 2014). Here, as a best guess, we choose a process rate that yields a maximum standard deviation of 2% for a constrained process of duration equal to the mean interval length between the tie points. In summary, we use process rates of $\gamma = 0.013$ for the DML and of $\gamma = 0.027$ for the WAIS cores, respectively.

10 The resulting transfer functions as used for the main results are shown in Fig. (B1). Diffusion shows a stronger smoothing for the longer records of DML2 than for DML1 as it had more time to act. WAIS shows less smoothing as compared to DML1 due to the higher accumulation rates. The influence of diffusion is negligible for frequencies below the decadal period for all core arrays. The effect of time uncertainty is negligible for frequencies below the decadal period for DML1 and WAIS but only for frequencies below the 20 yr period for DML2 due to the longer distances between the volcanic tie points.

15 Appendix C: ~~ERA-interim~~ ERA-Interim temperature and precipitation field decorrelation scales

To estimate the present-day spatial decorrelation scales of the atmospheric temperature and the precipitation fields in our study regions, we resort to ERA-Interim reanalysis data (~~Dee et al., 2011~~) (Dee et al., 2011; European Centre for Medium-Range Weather Foreca
, since direct observations are sparse. ~~We~~ ERA-Interim is in general regarded the most reliable reanalysis product both for
temperature and precipitation in Antarctica; the reanalysis temperatures are, despite significant biases in mean values, well
20 correlated with station observations at monthly to interannual timescales (Bracegirdle and Marshall, 2012; Jones and Lister, 2015)
, and the reanalysis produces the closest match to Antarctic snowfall regarding both mean values and seasonal to interannual
variability (Bromwich et al., 2011; Palerme et al., 2017).

To obtain deccorelation scales, we calculate the correlations between ~~the respective~~ reanalysis time series from the ~~nearest~~
~~gridbox of either gridbox closest to~~ the location of the EDML site ($75^\circ\text{S}, 0^\circ\text{W}$) ~~, or the location of the WDC site ($79.5^\circ\text{S},$~~
25 ~~112°W), with, respectively, the~~ with the respective time series of all other ~~continental Antarctic gridbox~~ gridboxes of the data
set below 60°S . Using an exponential fit to the correlations of the form $\exp(-d/\Delta)$, where d is the distance between EDML
~~(WDC)~~ and a particular gridbox, yields the decorrelation scale through the fit parameter Δ . The same approach is applied to
the WDC site ($79.5^\circ\text{S}, 112^\circ\text{W}$).

For annual mean temperatures, the analysis yields decorrelation scales of around 1200 km for both regions (Fig. C1). The
30 decorrelation scales of precipitation intermittency are expected to depend on the decorrelation scales of the seasonal precipitation
distribution rather than on the total annual precipitation amount (Persson et al., 2011). However, our analysis suggests that the
precipitation decorrelation scales are actually more ore less insensitive on the chosen variable with estimates between roughly
300 and 500 km for both regions for a variety of chosen fields (total precipitation amount, difference between summer (months

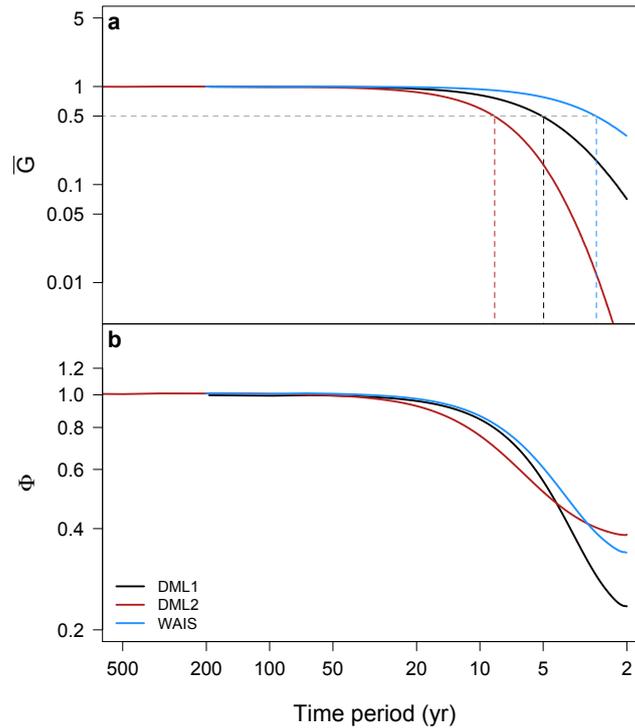


Figure B1. Estimates of the spectral transfer functions for the effects of site-specific diffusion **(a)** and time uncertainty **(b)** for the three studied core arrays DML1 (black), DML2 (red) and WAIS (blue). The estimates are based on simulations with surrogate data as explained in the text. Both effects are largely negligible beyond the decadal period. [The horizontal dashed line in \(a\) at 0.5 denotes the chosen transfer function value for constraining the diffusion correction \(see main text\), corresponding to minimum time periods until which we analyse the respective spectral data as indicated by the vertical dashed lines.](#)

DJF) and winter (months JJA) precipitation amount, fraction of summer over winter (winter over summer) precipitation amount, and fraction of summer (winter) over total precipitation amount).

Author contributions. TM and TL designed the research, developed the methodology and interpreted the results. TM reviewed relevant literature, established the database and performed all analyses. TM wrote a first version of the manuscript which was revised by both authors.

5 *Competing interests.* The authors declare that they have no conflict of interest.

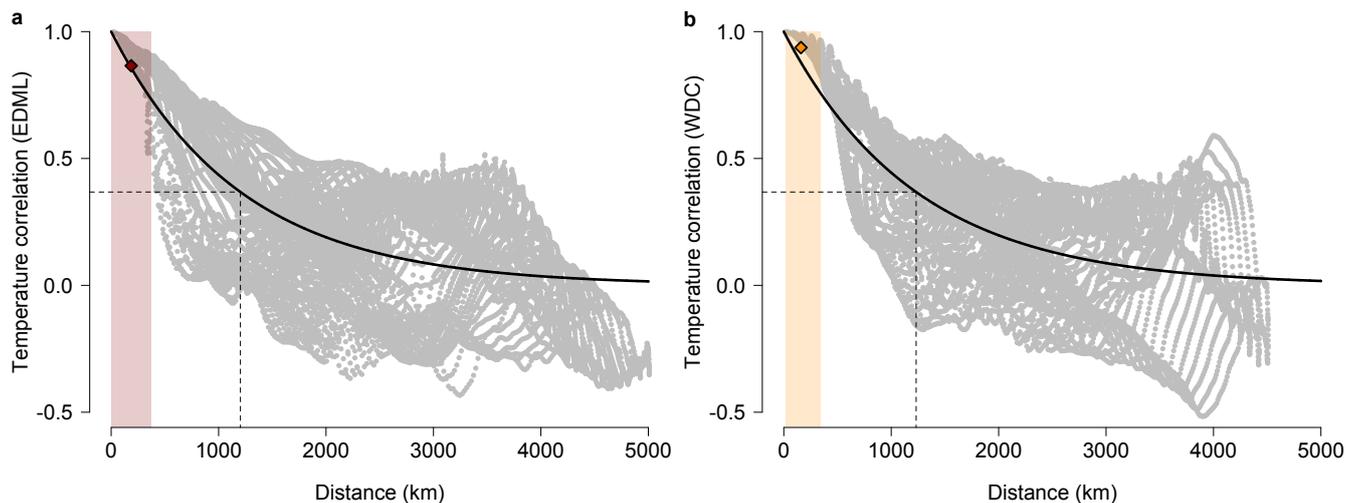


Figure C1. Present-day temperature decorrelation across DML and WAIS. Shown are the correlations (grey dots) of ERA-Interim annual-mean temperatures (Dee et al., 2011) at (a) EDML and (b) WDC with all other [gridbox on the Antarctic continent](#) [gridboxes below 60° S](#). Black lines show an exponential fit to the data; dashed lines indicate the point at which the correlations in the model have dropped to $1/e$ (at 1202 (1232) km for EDML (WDC)). Coloured [shadings indicate shading indicates the range of site distances between the sites](#) of the studied firn-core arrays (Table 1); the average correlation [over these distances across each range is marked by the filled diamond \(average correlation of 0.87 \(0.94\) for EDML \(WDC\)\).](#)

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