Interactive discussion on 'Temperature variability of the Iberian Range since 1602 inferred from tree-ring records'. Cp-2016-9

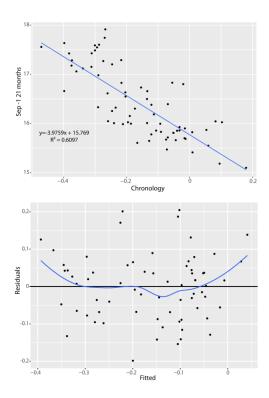
Anonymous Referee #1. 12 February, 2016.

In this study, a set of tree-ring chronologies from the Iberian Range is used to develop a maximum temperature reconstruction spanning the period 1602-2012. This topic is potentially very interesting since the temperature reconstructions in this region are rare. However, I see relevant issues that raise a number of (serious) concerns related to the composite chronology used for the reconstruction, the climate variable reconstructed and, particularly to the statistics of the calibration-verification. Considering these concerns, I unfortunately cannot recommend this manuscript for publication, and I think that addressing these concerns would entail the preparation of a whole new manuscript. I will just focus on those main issues starting from the statistics of the calibration verification. In addition, the manuscript would also require a careful editing since there are spelling errors, repetitions and inaccuracies, particularly related to the definition of correlations (r) and coefficient of determination (r2). I loathe to be so critical but Figure 8 and Table 2 give the impression that the numbers provided do not match with the series shown in the Figure and likely something went wrong when calculating and interpreting some statistics.

Thank you for the comments. About the calibration-verification statistics, you are right, some of the values included in the submitted version of the manuscript were wrong (sorry about that). We perhaps should have delved deeper into the development of the chronology and the climate variable reconstructed. Through this comment we aim to answer all the questions that the manuscript has generated.

- 1. As an example, Figure 8 shows that the r2 of the later period is 0.54 (or a correlation of 0.73). This value does not seem to match the poor interannual synchrony between the series that can visually be seen in the figure. It seems to me that either the correlation is spurious and largely inflated by the similar trend; or a correlation of 0.54 was mistakenly labelled as r2. Please note that correlation (r) and coefficient of determination (r2) are used in the manuscript and figures in both upper and low case letters and sometimes mixed (i.e., in Figure 11. R2 is defined as adjusted correlation; and text between lines 27-31 in page 8 mention correlations but show values labelled as r2) and I wonder if this could has been a potential source of confusion.
- 1. Regarding Figure 8, the correlation was not mistakenly labelled as r2. However, as suggested by the referee, we have correctly labelled the Pearson correlation (r) and the coefficient of determination (r2). The inter-annual synchrony that can be seen between the series denotes that the reconstruction is better at mid to low frequencies than at high frequency.
- 2. The validation statistics seem also too high. The reduction of error (RE) value of 0.99 is just hard to believe. A RE value of 0.99 (considering that the theoretical maximum value is 1) would basically mean that trees are recording temperature with the precision of a thermometer and, unfortunately, this is not realistic. It is very likely that something went wrong in the calculation, and this would need to be re-checked and re-interpreted.
- 2. Regarding the calibration/verification statistics, we apologize for the error; the included in the ms were calculated using unstandardized series. The revised are now shown in Table2: RE for the period 1945-2012 is 0.56, so substantially lower than reported, but still indicating reconstruction skill.

- 3. The reconstructed climate variable is the mean temperature over 21 months. This variable will presumably have a strong autocorrelation. It is not clear to me whether and how the authors statistically addressed the calculation of the significant levels considering the reduction in the degrees of freedom associated to a high autocorrelation. On the other hand, the authors stated on the manuscript that the chronology used for the reconstruction (BasPois) displays a first-order autocorrelation of 0 which implies that the proxy record does not mimic the autocorrelation of the temperature series used for calibration. Hence, there is a clear mismatch in the statistical properties of the predictor and the predictand. At this point, I am missing an analysis of the residuals from the regression modelling (trend, autocorrelation, etc) that will provide critical information on the adequacy of the predictand included in the model. I would expect that the residuals derived from such a regression will show a strong autocorrelation which will question the estimations of the uncertainties and statistical significance.
- 3. Regarding autocorrelation, the correct value is 0.83, which is crucial for the development of a reconstruction retaining information of the past 21 months. To further assess the accuracy of the model we included a new figure (Model_Residuals) detailing the transfer model and regression residuals.

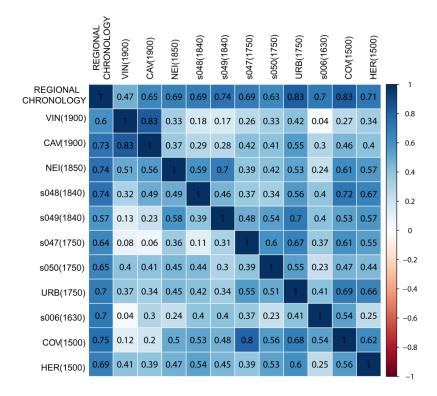


4. In view of the clear visual mismatch in high frequency between the tree-ring chronology and the instrumental temperature record, I would recommend to do a comparison of both series at different time scales to make sure that the correlation observed in the calibrations are due to synchrony in both, low and high-frequency domains, and it is not a spurious correlation due to similar long-term trends. This would definitely help to know if the climate variable chosen for the calibration is the correct one. In the background there are a couple of other issues that are not as relevant as the one with the statistics of the calibration-verification, but also critical in the general context of the paper.

- 4. Regarding visual mismatch in the high frequency domain. Similar long-term trends between temperature and tree-rings are not necessarily indicative of a spurious relation, but might simply suggest that trees are, responding to long-term temperature trends. Sure this is difficult to assess due to limited degrees of freedom. However, preserving such trends remains a key challenge in tree-ring based climate reconstructions (Briffa et al. 1992, Esper et al. 2003b). We employed a running correlation analysis (Fig. 7) not only to test temporal stability, but also to support the selection of climate variable.
- 5. The authors combined data from different chronologies into a single sort of regional chronology using different methods, which is always an interesting exercise. However, having a look to Figure 3, I wonder why all chronologies have been included into the final regional composite instead of discarding the chronologies that clearly showed poor correlations (i.e., s047). According to the information currently available on the paper, a reader cannot be sure whether chronologies encoding different climate signals

have been merged into a final composite. To answer this question and reinforce the methodological decision adopted, I would suggest to check whether all chronologies encode the same climate signal before building regional chronologies, particularly if some chronologies clearly show a limited agreement with the rest. In this way, potential doubts on the quality and regional representativeness of the composite regional chronology will be minimized.

5. Regarding the regional chronology, we develop this timeseries by combining 11 sites including two pine species within an area of 90 square kilometers ranging from 1,500 to 1,900 masl. A new column in figure 3 showing the correlation between the single sites and the regional chronology provides perhaps useful information. Despite local differences among the sites, the group of chronologies shares common variance, and the mean chronology contains a clear climate signal. Data integration from this treering network enabled the development of a regional rather than local reconstruction. Figure 11 shows the spatial extent of the reconstruction indicating r2 > 0.4 for the central and Mediterranean regions of the Iberian Peninsula.



6. The link between the climate variable reconstructed in the paper and the proxy record lacks a consistent physiological explanation. The explanation given between lines 13-16 of page 11, though correct in general terms, is too general and seems insufficient to explain the selection of a climate season that is quite unusual in the context of tree-ring based climate reconstructions. In fact, the explanation given could be applied to any lagged climate season. However, and independently of the physiological explanation, calibrating with a 21 cumulative monthly mean of temperature when the chronology shows a first order autocorrelation of 0, seems totally contradictory to me. I do not doubt that the authors have a consistent reason for all the decisions adopted in the paper. However, the present version of the manuscript gives the impression that the selection of the composite chronology and the climate season used for the reconstruction were purely based on the highest correlation obtained, and all other considerations and potential implications were somehow overlooked.

6. Regarding physiological explanation. Extended between lines 25 of page 11 to to 21 page of 12.in the new version of the ms.

Anonymous Referee #2. 25 February, 2016.

The manuscript presents climate reconstruction for the 1602-2012 period for previously less explored part of the Iberian Peninsula. The authors used a new standardization method based on the basal area. The applied detrending improved the quality of the reconstruction since the size-based standardization maximizes the common signal. The presented reconstruction fills the gap in climatic reconstructions in the area vastly affected by climate change where long term climatic records (and old trees) are very rare.

The first version of the ms contained numerous small errors. Some of them were already commented by the referee #1. The authors have responded to them. Their response is to my opinion adequate. The authors presented their response and submitted an improved version of the manuscript. The errors are now corrected, although they generally did not affect the message (result and conclusions) of the manuscript.

The chronology is to my opinion well-constructed and I agree with the response of the authors. The fact that they used different species has been adequately justified already in the first version of the ms. This is a common praxis if the chronologies match well - this is the case in this ms. Both versions of the ms contain small errors (mainly language and style) which should be corrected at the end of the interactive review process.

S. Klesse; stefan.klesse@wsl.ch; 01 March, 2016

While the authors have addressed the remarks of Reviewer #1 in a reasonable and adequate way, I see some methodological problems, mainly with the RCS (baspois) application:

1. Did you use pith offset (or for your case of Basal-Area-RC distance to pith) estimates? I cannot find it in the text. If not, why? Omitting pith offset estimates will lower your RC and ultimately introduce a fake negative trend in the early years of your chronology. In your case of inversion a positive biased trend, which could be amplified when using Basal-Area. See Briffa & Melvin 2011 _"A closer look on RCS..." and Klesse & Frank 2013 (attached).

Dear Stefan:

1. Regarding pith offset. On the one hand, we use a set of sites (VIN, CAV, NEI, URB, COV, HER) of which the samples are available, and therefore pith offsets can be estimated. On the other hand, we use a set of sites from the ITRDB (s047, s048, s049, s050, s006) from which the samples cannot be accessed. Here, for each tree, we have assumed PO = 1 in the oldest sample, and adjusted the shorter series accordingly. Sure this procedure introduces uncertainties, but this is true for all studies using data from the ITRDB. We believe, however, that these uncertainties do not generate a systematic bias, but are minimized using the new BasPois detrended method based on basal area instead of age.

- 2. You include old ITRDB datasets from the 1980s. Do you have the samples or pith offset estimates, or at least correct for the a- and b-sample difference of starting year? For example: a- and b-samples of the ITRDB series SPAI047 have quite large differences between their starting years (mean: 33 years). For their RCS curve that would mean, that those samples are overestimated on average already by 50mm (should be probably 0.5mm).
- 2. Perhaps already addressed in the previous paragraph. Since we do not have the ITRDB samples, pith offsets need to be estimated. Assuming an age of 1 in the oldest sample of every tree (of the ITRDB data) is indeed a compromise, and perhaps there are other possible solutions, however, we believe that for the development of a chronology representative of the regional climate of the western Iberian Range, all available samples should be included.
- 2b. The y axis in figure 4b is presumably off by a factor of 100 and should range from zero to 2.5mm instead of 250mm.

2b. We apologize for the mistake. The axis is now correctly labelled (Fig.4).

- 3. Do the trees have the same growth rates at all 11 sites? If not then a use of a single RC might introduce false trends, when sample and site replication changes. From originally 11 sites, 5 drop out in the 1980s, 4 in the 90s and you are left with only two sites. Do these sites have the same growth level as the ones that drop out before? See also Figure 6 in Klesse & Frank for an example of falsely introduced trends. I have attached a figure showing this potential problem including the 5 Iberian Range (IR) ITRDB chronologies and 2 chronologies from Büntgen near Madrid. Although 250km to SW they grew at similar elevation and correlate with the mean chronology of the other 5 series quite good (r=0.52, 1701-1985, 30-year spline detrended). So, well in the range of your observed site to site correlations and only a little bit weaker than your weakest site to regional chronology corrlelation, but completely independent (one could actually argue to include them to increase the regional representation, but that's beside the point here). I applied a single RC and no pith offsets, split the IR and Büntgen series and averaged them separately with an arithmetic mean. It is obvious that the mean of Büntgen have permanently lower values over the IR series. So if the IR series drop out, the overall RCS chronology gets heavily drawn towards lower values, while the Büntgen series remain constant. This effect could probably also be enlarged using the size/basal area detrending. Can you show that this does not cause a problem in your data?
- 3. About growth rates, the variability among sites is lower than the variability within a site. Besides, we are not joining chronologies, but develop the regional chronology from all 316 individual TRW series. It is generally unavoidable to add some noise when integrating TRW series from different locations and species in a regional chronology. However, the high correlation between each site and the regional chronology suggests a general climatic signal. Similar approaches have been detailed in Briffa et al. 1998 and numerous climatic reconstructions have been developed using networks of different sites and species i.e. Wilson et al. 2003, Battipaglia et al. 2010, Büntgen et al. 2011 or, Esper et al. 2012. However, to prove that the variable end dates are without effect in the trend we develop a regional chronology with the sites ending in 1993 (using the BasPois detrended method) (Fig1 on this comment).

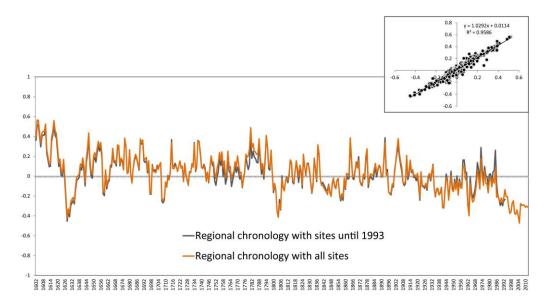


Fig.1

- 4. Figure 5a) Why do you compare your residual AC-free chronology with raw temperature data? That does not make sense if you want to highlight common signal in the high-frequencies. The simplest method would have been to detrend both series with a flexible spline (e.g. 30 years). That actually comes back to Remark 4 from Reviewer #1. If the TRW signal is truly representing pSep21 temperature, than it still should at least have reasonable negative correlations on the high-frequencies. A running correlation with raw temperature and BasPois does not answer Remark 4 and still contains trend-in-signal and not necessarily causal effect. I believe the authors might have kept things too easy during the RCS application, which might have led to erroneous conclusions. I would be really happy if my concerns don't have an impact on the conclusions, but without showing that I remain cautious.
- 4. As suggested, we have detrended the climate data using a flexible spline (30 years) and correlated with ArstanRES and ArstanSTD chronologies emphasizing high frequency variance. The results show an increase in correlation with pSep21 temperature; for ArstanRES the correlation is r=-0.39, while for ArstanSTD the correlation increases to r=-0.56. These correlations indicate that the reconstruction contains some skill in the high frequency domain. Nevertheless, in order to assess long-term climate variability, we prefer using the BasPois chronology, in which the climate signal is enhanced, and both high and low frequency pSep21 variance and forcing (volcanic and solar) is retained. The changes in methodology and results have been included in the manuscript as well as the new Figure 5 (Fig2 on this comment).

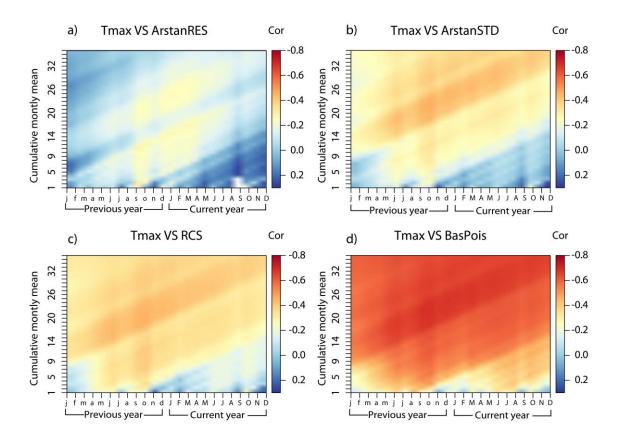


Fig.2.

Anonymous Referee #1 interactive comment. 08 March, 2016

The effort of the authors commenting and answering my questions and assuming mistakes in the calculation of some statistics is laudable. However, relevant issues described in detail in my initial review have not been answered and the reliability of the calibration is still questionable. The key component of this manuscript is the development of a climate reconstruction based on tree rings. If the calibration raises serious doubts, then the whole manuscript is dubious. I am aware of the effort that is needed to develop a proxy-based climate reconstructions, however, I still believe that the reconstruction presented in this manuscript is not fully reliable mainly due to three main issues:

Referee #1: Thank you for your interest and comments, which we aim to answer here.

1. How good or bad is the agreement between the tree-ring record and the climate record on the high-frequency domain remains unanswered. A poor visual agreement on the high frequency is obvious having a look to Figure 8, and the author's reply ("the reconstruction is better at mid and low frequencies than at high frequencies") is insufficient and does not provide any additional or new information to clarify this point. The way to solve the doubt is, as suggested in point 4 of my initial review and also suggested by S.Klesse, to do a comparison (correlation) of both series at different time scales and make sure that the correlation observed in the calibration are due also to synchrony in the high frequency and not only to a similar long-term trend. If the series correlate on the low frequency but do not show a reasonable agreement on the high frequency, then the correlation shown for the calibration period would be spurious and the reconstruction simply incorrect. The running correlation analysis will only answer this question if the series would have been high-pass filtered, which is

not the case. In addition, the residual analysis now included on the paper would also require to include some test on the trend and autocorrelation of the residuals (i.e., Durbin-Watson test). If the calibration fulfils all above (agreement on the high-frequency domain and test of residuals), then we could talk about a statistical reliable calibration.

- 1. In order to test the agreement between the tree-ring chronology and climate record in the high-frequency domain, and in line with Stefan Klesse's suggestions, we correlated the ArstanRES and ArstanSTD timeseries with the detrended (30-year spline) climatic data. The results show an increase in correlation with pSep21 temperature. For ArstanRES the correlation is r= -0.39, whereas for ArstanSTD the correlation increases to r=-0.56. These results validate correlation in the high-frequency domain and indicate that the reconstruction signal is not spurious. However, we intend to reconstruct both high and low frequency climate variations and prefer using the BasPois chronology as it enhances the climatic signal (r=-0.78) and reproduces the full variance spectrum of the pSep21 variable very well. We performed a Durbin-Watson test, as suggested, to assess residual autocorrelation. The results were added to Table2. The Durbin Watson value for the period 1945-2012 is 1.45 (p<0.001) indicating no substantial autocorrelation in the residuals. We believe that both the correlations after removal of low frequency variance as well as the insignificant autocorrelation in the residuals support the pSep21 reconstruction.
- 2. Whether chronologies encoding different climate signals have been merged into a final composite remains also unanswered. The new column added to Figure 3 containing the values of the correlations between the single and regional chronologies does not answer my initial question. Checking whether all chronologies encode the same climate signal means to correlate each individual chronology with climate. This is the way to know if tree growth is limited by the same climatic factor at all sites or different climate signals are being mixed in the regional chronology. Considering that the chronologies used are from different tree species, derived from different elevations and some chronologies do show poor correlation with the others, testing potential different climate signals is advisable, particularly because solving such a question is extremely easy.
- 2. About chronology development, we did not merge site chronologies, but applied the standardization methods to all 316 individual TRW series to produce a regional chronology. Nonetheless, we added climate calibrations for each site to validate that the climate signal is regionally consistent. We developed a chronology for each of the 11 sites (detrended with the BasPois method) and correlated with the climatic variables. Highest correlations in the 11 sites appear for pSep20, pSep21 and pOct21. Since we chose to reconstruct pSep21 we also performed running correlations using a 30-year window to assess correlation stability within the calibration period. Results are shown in the Fig.1 (of this comment) and chronologies are sorted by elevation, VIN and CAV are *Pinus uncinata*, while the rest are *Pinus sylvestris*. The correlation never drops below r= -0.2. There are also periods surpassing r=-0.80. However, we would like to reemphasize that the aim of this study it is not to develop a local climate reconstruction, but to reconstruct the regional climate of the western lberian Range.

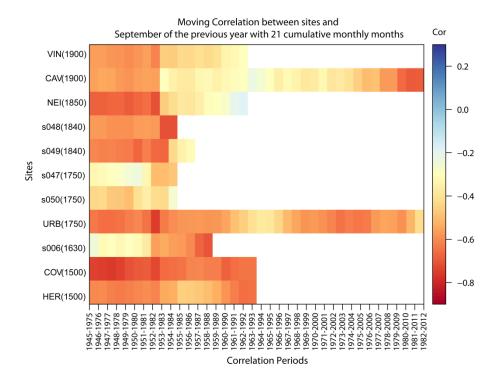


Fig.1.

- 3. The physiological explanation is still too general (in fact, has not substantially changed) and not very convincing. It is hard for me to picture how tree growth can be negatively influenced by the cumulative mean of temperature from the current and previous year of growth: how trees manage to grow then? How did they survive for centuries and did not die by carbon starvation if cumulative temperature of the previous 21 months have no positive effect on growth? Physiologically seems quite unlikely to me but still, I was hoping for a good explanation or answer that could challenge my thoughts on this regard.
- 3. We would like to remark that tree-ring growth it is not negatively influenced by temperature. It is, however, negatively correlated with temperature of the previous year using a cumulative monthly mean of 21 months. That would mean that within the environment in which trees are growing and with respect to the mean, they will grow more in cold years than in hot years.

The negative temperature correlation is already shown for the previous September (r=-0.56) without any cumulative monthly mean. This negative temperature correlation has been reported in numerous dendroclimatic studies (i.e. Büntgen et al. 2006 or van der Werf et al. 2007) including the most recently developed climatic reconstruction for the Iberian Peninsula by Dorado-Liñán et al. 2014 showing a negative correlation with previous summer temperatures. One of the strengths of this paper is precisely adding the cumulative monthly mean to the climate variables which maximizes the correlation to r=-0.78.

The ecophysiological explanation of previous year's influence on current's year tree-ring growth was already related with the storage of starch and sugar in parenchyma ray tissue and the remobilization of carbohydrates from root structures. Memory effects on TRW data have also been studied regarding the delayed response in TRW to post volcanic eruptions (1â'Lij5 years) associated with a decrease in current's year temperature (D'Arrigo et al., 2013, Esper et al., 2014).

We agree on the need to conduct further studies to better understand the full range of ecophysiological processes of pine and other species. To this extend, we are aware of an experiment conducted by a colleague (Dr. Eustaquio Gil Pelegrin; ttps://www.researchgate.net/profile/Eustaquio_Pelegrin) in which they try to demonstrate

that the generation of pinecones and needles in pine trees is very slow and itgenerally takes two years.

C.Salt; christ.salt77@gmail.com: 09 May, 2016

After reading this discussion paper, I was left with the doubt on whether the authors have reconstructed temperature or precipitation. Considering that all or most of these sites are probably sensitive to variations in soil moisture given their location in Mediterranean mountains, at least a mixed precipitation-temperature signal could be expected and should be analyzed and discussed. One must be extremely careful when analyzing negative effects of temperature on tree growth, particularly at sites >1500 m asl where temperatures are most likely not warm enough to cause direct damage to plant cells.

What would be the biological mechanism of a 21-month long cumulative negative effect of temperature, if it were not for an indirect effect through hydric stress of the trees? It makes sense that these relationships are driven by temperature increasing the evaporative demand or vapor pressure deficit. Thus, precipitation or a drought index should be considered in the analysis. I don't think this issue has been addressed in the original file or in the author's response to referee #1. Results for SEA for volcanic eruptions would show lower temperature 3 years decrease after eruptions. That would mean wider tree rings. But those could also be caused by increased precipitation as shown for other parts of the Mediterranean Basin (Köse, N. et al., 2013. An improved reconstruction of May-June precipitation using treering data from western Turkey and its links to volcanic eruptions. International Journal of Biometeorology, 57(5): 691-701.)

I would suggest the use of partial correlations for temperature (secondary variable), controlling for the effect of precipitation (primary variable). Using something like the seascorr function [Meko, D.M., Touchan, R. and Anchukaitis, K.J., 2011. Seascorr: A MATLAB program for identifying the seasonal climate signal in an annual tree-ring time series. Computers & Geosciences, 37(9): 1234-1241] should be straightforward. I would recommend the same time periods and lags be analyzed for precipitation or a drought index (similar to figure 5), before performing a temperature reconstruction from negative correlations with tree rings. It may be that the correlations with temperature are ok, but I think this deserves better explanations and justifications.

Dear Chris, thank you for your interest and comments, which we aim to answer here.

Speaking about the Iberian Peninsula can sometimes generate misconceptions. The Iberian Peninsula is a very large territory with a broad set of climates ranging from a dry Mediterranean climate with 200 mm/year and a dry season during summer to an Atlantic climate with more than 2,500 mm/year and no dry season. The study area, as described in lines 23 to 27 (page 3) belongs to a Continental bioclimatic belt which is characterized by moderate mean temperatures (9.5C) and a mean annual precipitation which exceeds 1,000 mm/years very frequently (Fig 2A,Fig.2AC in the manuscript).

Therefore, there is no dry season within the study area. As well as in other mountain forests in Spain (see Büntgen et al., 2008, Dorado-Liñán et al., 2014), trees in the study area are limited by temperature. In Dorado-Liñán et al. 2014 they reconstruct the previous year's summer temperature for the past 800 years in the southeast of Spain using tree-ring width. During the conduct of the first analysis, we also took into account precipitation and drought indices such as SPI (McKee et al., 1993) and SPEI (Vicente-Serrano et al., 2010). However, due to the poor correlation values (see Fig.1 in the comment) we decided to focus on the maximum temperature signal. In Fig.1 of this comment SPEI (1 to 24) and SPI (1 to 24) values are correlated with the BasPois Chronology. The maximum correlation (r=0.35) is shown for

the SPEI19 of August and it is very much related with the temperature, since the SPEI drought index integrates temperature, in terms of evapotranspiration, to the equation.

There are, however, as suggested, some mountain areas in the Iberian Peninsula with Mediterranean climate conditions including a dry season with its trees limited by precipitation. For instance, in Tejedor et al., 2015 we developed a drought reconstruction using the SPI index.

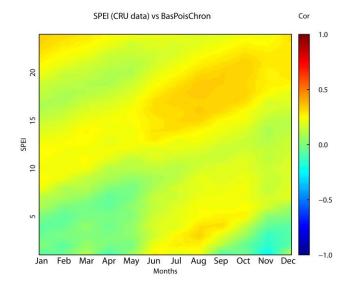
Büntgen, U., Frank, D., Grudd, H., Esper, J.: Long-term summer temperature variations in the Pyrenees. Climate Dynamics, 31 (6), pp. 615-631, 2008.

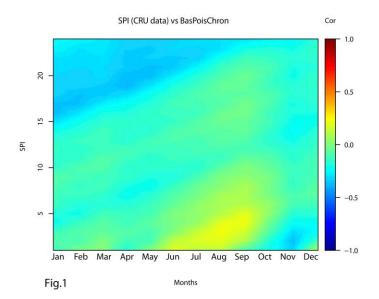
Dorado Liñán, I., Zorita, E., González-Rouco, J.F., Heinrich, I., Campello, F., Muntán, E., Andreu-Hayles, L., Gutiérrez, E.: Eight-hundred years of summer temperature variations in the southeast of the Iberian Peninsula reconstructed from tree rings. Climate Dynamics, 44 (1-2), pp. 75-93, 2014.

McKee TB, Doesken NJ, Kliest J (1993) The relationship of drought frequency and duration to time scales. In: Proceedings of the 8th Conference on Applied Climatology, Anaheim, CA, USA, 17–22. American Meteorological Society, Boston, MA, USA, pp 179–184

Vicente-Serrano SM, Beguería S, López-Moreno JI (2010) A multiscalar drought index sensitive to global warming: the standardized precipitation evapotranspiration index. J Clim 23(7):1696–1718.

Tejedor, E., de Luis, M., Cuadrat, J.M., Esper, J. & Saz, M.A. 2015. Tree-ring-based drought reconstruction in the Iberian Range (east of Spain) since 1694. International Journal of Biometeorology, DOI:10.1007/s00484-015-1033-7.





Anonymous Referee #3: 06 July, 2016.

I find this paper intends to show and interesting matter, it is well written, and with a good quality of figures. Only some typos need to be corrected. However, there are some aspects that need clarification before the paper can be ready for publication, and some of them make me doubt about the validity of the results.

First of all, I wonder what the real objective of the paper is. On the one hand, different standardization methods are tested, and on the other hand the authors perform a reconstruction of a climatic variable. Though I can understand that this is a necessary step to provide a reliable reconstruction, I do not see that has this been brought in detail into the discussion, especially as regards the first two methods. But my main concern regards the variable selected for reconstruction. Though statistics seem to me optimal, I am not able to figure out what the causes for the existent relationship could be (21 month temperature). Explanations in the Discussion are too weak to be convincing, and I think this aspect needs to be much better clarified or hypothesized.

Referee #3: Thank you for your interest and comments. The issues raised in your comments have been deeply treated throughout the open discussion process. We believe that in the latest version of the manuscript questions related with the standardization method and the 21 months climate variable have been clarified. In any case, we will be glad to address any particular aspect if there is a further suggestion. Attached is the latest version of the manuscript.

Temperature variability of the Iberian Range since 1602

2 inferred from tree-ring records

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11

Abstract

- 12 Tree-rings are an important proxy to understand the natural drivers of climate variability in
- 13 the Mediterranean basin and hence to improve future climate scenarios in a vulnerable region.
- 14 Here, we compile 316 tree-ring width series from 11 conifer sites in the western Iberian
- 15 Range. We apply a new standardization method based on the trunk basal area instead of the
- 16 tree cambial age to develop a regional chronology which preserves high to low frequency
- 17 variability. A new reconstruction for the 1602-2012 period correlates at -0.78 with
- 18 observational September temperatures with a cumulative mean of the 21 previous months
- 19 over the 1945-2012 calibration period. The new IR2T_{max} reconstruction is spatially
- 20 representative for the Iberian Peninsula and captures the full range of past Iberian Range
- 21 temperature variability. Reconstructed long-term temperature variations match reasonably
- 22 well with solar irradiance changes since warm and cold phases correspond with high and low
- 23 solar activity, respectively. In addition, some annual temperatures downturns coincide with
- volcanic eruptions with a three year lag.

25

26 1 Introduction

- 27 The IPCC report (IPCC, 2013) highlighted a likely increase of average global temperatures in
- 28 upcoming decades, and pointed particularly to the Mediterranean basin, and therefore in the

1 Iberian Peninsula (IP), as a region of substantial modelled temperature changes. The 2 Mediterranean area is located in the transitional zone between tropical and extra-tropical 3 climate systems, characterized by a complex topography and high climatic variability (Hertig 4 and Jacobeit 2008). Taking into account these features, even relatively minor modifications of 5 the general circulation, i.e. a shift in the location of sub-tropical high pressure cells, can lead 6 to substantial changes in Mediterranean climate (Giorgi and Lionello 2008), making the study 7 area a potentially vulnerable region to anthropogenic climatic changes by anthropogenic 8 forces, i.e. increasing concentrations of greenhouse gases (Lionello et al., 2006a; Ulbrich et 9 al., 2006) 10 Major recent efforts have been made in understanding trends in temperatures throughout the IP over the instrumental period (Kenaway et al., 2012; Pena-Angulo et al., 2015; Gonzalez-

11 12 Hidalgo et al., 2015) and future climate change scenarios (Sánchez et al., 2004; López-13 Moreno et al., 2014). However, the fact that most of the observational records do not begin until the 1950s (Gonzalez-Hidalgo et al., 2011) is limiting the possibility of investigating the 14 15 inter-annual to multi-centennial long-term temperature variability. Therefore, it is crucial to 16 explore climate proxy data and develop long-term reconstructions of regional temperature 17 variability to evaluate spatial patterns of climatic change and the role of natural and 18 anthropogenic forcings on climate variations (Büntgen et al., 2005). In the IP, much progress 19 has been made to reconstruct past centuries climate variability, including analysis of documentary evidences for temperature (i.e. Camuffo et al., 2010) and droughts 20 reconstruction (i.e. Barriendos et al. 1997; Cuadrat and Vicente, 2007; Domínguez-Castro et 21 22 al., 2010). Additionally, progress has been made to further understanding of long-term climate 23 variability of the IP through dendroclimatological studies focussing on drought (Esper et al., 2014; Tejedor et al., 2015) and temperature (Büntgen et al., 2008; Dorado-Liñán et al., 2012, 24 25 2014; Esper et al. 2015a). Nevertheless, a high-resolution temperature reconstruction for 26 central Spain is still missing.

Several studies have been made to develop a temperature reconstruction for the Iberian Range (IR) using *Pinus uncinata* tree-ring data (Creus and Puigdefabreas, 1982; Ruiz, 1989). The results, in fact, showed a pronounced inter-annual to century scale chronology variability. However, their main result was a complex growth response function due to a mixed climate signal instead of a temperature reconstruction. Furthermore, Saz (2003) developed a 500-year temperature reconstruction for the Ebro Depression (North of Spain), but this chronology is

- 1 based on a reduced number of cores and a standardized methodology that did not retain the
- 2 medium and low frequency variance.
- 3 Here we present the first tree-ring dataset combining samples from three different sources
- 4 from the eastern IR extending back from the Little Ice Age (1465) to present (2012). The aim
- 5 of this study is to develop a temperature reconstruction representing the IR, and thereby fill
- 6 the gap between records located in the northern and southern IP. A new methodology, based
- 7 on basal area instead of the cambial-age, was applied to preserve high-to-low frequency
- 8 variance in the resulting chronologies. Furthermore, the relationship between the tree-ring and
- 9 climate data is reanalysed by adding memory to the climate parameters, since memory effects
- 10 on tree-ring data are much less acknowledged (Anchukaitis et al., 2012). This analysis is
- challenging because of the mix of tree species and their unidentified responses to climate. The
- 12 resulting reconstruction of September maximum temperatures over the past four centuries is
- 13 compared with latest findings from the Pyrenees and Cazorla, and the relationship with solar
- and volcanic forcings at inter-annual to multi-decadal timescales.

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2.1 Site description

Material and methods

- 18 We compiled a tree ring network from 11 different sites in the western IR (Table 1) in the
- 19 province of Soria. Urbión is the most extensive forest of the IP including 120,000 ha between
- 20 the Burgos and Soria provinces. It has a long forest management tradition. Therefore, all sites
- 21 are situated at high elevation locations where forests are least exploited and maximum tree
- age is reached (Fig.1). The altitude of the sampling sites ranges from 1,500 to 1,900 meters
- above sea level (masl) with a mean of 1,758 masl. These forests belong to the Continental
- 24 Bioclimatic Belt (Guijarro, 2013) characterized by moderate mean temperatures (9.5°C,
- 25 Fig.2B) and a large seasonal range including more than 90 frost days and summer heat
- 26 exceeding 30°C. Mean annual precipitation for the period 1944-2014 is 927 mm (CRU TS.3
- 27 v.23 dataset by Harris et al., 2014) and reaches its maximum during December (Fig. 2AC).
- 28 Although scotts pine (Pinus sylvestris) is the dominant tree species of the region, other
- 29 pinaceaes are found such as Pinus pinaster, Pinus nigra or Pinus uncinata. Especially
- 30 remarkable is occurrence of Pinus uncinata growing above 1,900 masl and reaching its

- 1 European southern distribution limits in the IR. The lithology of the study area consists of
- 2 sandstones, conglomerates and lutites.

3 2.2 Tree ring chronology development

- 4 The new dataset is composed by 316 tree-ring width (TRW) series of *Pinus uncinata* (56) and
- 5 Pinus sylvestris (260) located in the western IR (Tab. 1, Fig. 1). The most recent samples
- 6 were collected during the field campaign in 2013 including old dominant and co-dominant
- 7 trees with healthy trunks and no sign of human interference. We extracted two core samples
- 8 from each tree at breast height (1.3 m) when possible, otherwise, we try to avoid compression
- 9 wood due to steep slopes, compiling a set of 96 new samples from two sites, i.e. the outermost
- 10 ring is 2012. Core samples were air-dried and glued onto wooden holders and subsequently
- 11 sanded to ease growth ring identification (Stokes and Smiley 1968). The samples were then
- 12 scanned and synchronized using CoRecorder software (Larsson 2012) (Cybis
- 13 Dendrochronology 2014) to identify the position and exact dating of each ring. The tree-ring
- 14 width was measured, at 0.01 mm precision, using LINTAB table (Rinn 2005). Prior to
- 15 detrending, COFECHA (Holmes 1983) was used to assess the cross-dating of all
- 16 measurement series.
- 17 An additional set of 95 samples from three sites was provided by the project CLI96-1862
- 18 (Creus et al. 1992, Saz 2003) i.e., the outermost rings range from 1992 to 1993. Finally, a set
- 19 of 125 samples from five sites was downloaded from the International Tree Ring Data Bank
- $20 \hspace{0.5cm} (ITRDB, \hspace{0.5cm} http://www.ncdc.noaa.gov/data-access/paleoclimatology-data/datasets/tree-ring). \\$
- 21 These data were developed in the 1980s by K. Richter and collaborators, i.e. the outermost
- 22 rings range from 1977 to 1985.
- 23 In order to attempt a climate reconstruction for the western IR from this tree-ring network, we
- 24 perform an exploratory analysis of the 11 tree-ring chronologiessites by creating a correlation
- 25 matrix of the raw ehronologies TRW series for each site and the correlation with a composite
- 26 regional chronology. Calculations are computed for the common period (1842-1977) and for
- 27 the full period (1465-2012).

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2.2.1 Standardization methods

- 29 The key concept in dendroclimatology is referred to as the standardization process (Fritts,
- 30 1976; Cook et al., 1990) where the aim is to preserve as much of the climate-related

1 information as possible while removing the non-climatic information from the raw TRW 2 measurements. However, with most of the standardization methods a varying proportion of 3 the low-frequency climatic information is also lost in the process (Grudd, 2008). When the 4 aim is to use tree-ring chronologies as a proxy for climatic reconstructions, an adequate 5 standardization is critical and the best method should preserve high to low frequency 6 variations (Büntgen et al., 2004). It is common practice to calculate a mean value function as 7 the best estimate of the trees' signal at a site (Frank et al., 2006). 8 We here applied four standardization methods to the 316 TRW measurement series to develop 9 a single tree-ring index chronology. (i) To emphasize inter-decadal and higher frequency 10 variations, each ring width series was fitted with a cubic spline with a 50% frequency 11 response cut off at 67% of the series length (Cook et al., 1990). A bi-weight robust mean was 12 calculated to assemble the ArstanSTD regional chronology. (ii) A residual chronology 13 (ArstanRES) is produced after removing first-order autoregression to emphasize high-14 frequency variability. (iii) To preserve common inter-decadal and lower frequency variations, 15 Regional Curve Standardization (RCS) was applied (Mitchell, 1967; Briffa et al., 1992, 1996; 16 Esper et al., 2003). RCS is an age-dependent composite method and involves dividing the size 17 of each tree-ring by the value expected from its cambial age. To assemble the chronology, all 18 the series are aligned by cambial age. A single growth function (regional curve, RC) 19 smoothed using a spline function of 10% of the series length is fit to the mean of all agealigned series. A biweigth biweight robust mean was applied to develop the RCS chronology 20 21 (RCS). (iv)To preserve high to low frequency variance, we additionally applied a novel 22 standardization method based on the principles of RCS. However, instead of using the 23 cambial age of the trees as the independent variable, we used their sizes, calculated as the 24 square of the basal area of the tree in the year prior to ring formation. Then, a Poisson 25 regression model was used to fit the individual tree-ring widths. Standardized indices were 26 calculated as the ratio between the observed and predicted values, and a biweigthbiweight 27 robust mean was used to develop the Basal Area Poisson chronology (BasPois). To evaluate uncertainty of the mean chronologies running interseries correlations (Rbar) and 28 29 the express population signal (EPS) were calculated (Wigley et al., 1984). Rbar is a measure 30 of the strength of the common growth 'signal' within the chronology (Wigley et al. 1984; 31 Briffa and Jones, 1990), here calculated in a 50-year window sliding along the chronology.

- 1 EPS is an estimate of the chronology's ability to represent the signal strength of a chronology
- 2 on a theoretical infinite population (Wigley et al., 1984).

2.3 Climatic data, calibration and climate reconstruction

- 4 Monthly temperature (mean, maximum, and minimum) and precipitation values from the
- 5 gridded CRU TS v.3.22 dataset (0.5° resolution) dataset for the period 1945-2012 were used
- 6 (Harris et al. 2014). The three grid points closest to the tree-ring network were averaged to
- 7 develop a regional time series (Fig. 1). In addition, we calculate a cumulative monthly mean
- 8 for each of the four parameters (max., min., mean temperature, and monthly precipitation).
- 9 The cumulative mean is calculated by adding the months gradually. First the previous month
- 10 is added, and then further months are included up to 36 previous months. For the calculations
- we take into account the current and the previous year.
- 12 For calibration, we correlated the four chronologies (ArstanSTD, ArstanRES, RCS, and
- 13 | BasPois) with monthly climate data and the cumulative monthly mean derived. However, to
- be consistent statistically, the two chronologies which highlight high frequency variations,
- 15 ArstanRES and ArstanSTD, were correlated with the detrended climatic data. To assess the
- stability of the correlation, we calculated a 30-year moving correlation shifted along 1945-
- 17 2012 with the cumulative monthly mean from the current and the previous year. In addition,
- 18 the maximum and minimum differences between the moving correlations were calculated. As
- 19 a result, the climatic variable chosen for the reconstruction is supported by having the highest
- 20 moving correlation with the least difference between the maximum and the minimum over the
- 21 moving correlation period.

- A split calibration/verification approach was perform over the periods 1945-1978 and 1979-
- 23 2012 to evaluate the accuracy of the transfer model considering the following metrics;
- 24 Pearson's correlation (r), coefficient of determination (r²), reduction of error (RE), mean
- 25 | square error (MSE), and the sign test (Cook et al., 1994) and the Durbin-Watson test (Durbin
- 26 and Watson, 1951). R is a measure of the linear correlation between the chronology and
- 27 climatic variable. R² indicates how well the data fit a statistical model. An r² of 1 indicates
- 28 that the regression line perfectly fits the data; an r² of 0 indicates that there is not fit at all. RE
- 29 is a measure of shared variance between actual and estimated series and provides sensitive
- 30 measure of the reliability of a reconstruction (Cook et al., 1994; Akkemik et al., 2005;
- Büntgen et al., 2008); it ranges from +1 indicating perfect agreement, to minus infinity. MSE

- 1 estimates the difference between the modelled and measured while sign test compares the
- 2 number of agreeing and disagreeing interval trends, from year-to-year, between the observed
- and reconstructed series (Fritts et al., 1990; Cufar et al., 2008). To verify that there is no
- 4 autocorrelation in the residuals we perform the Durbin-Watson test. Additionally, a
- 5 Superposed Epoch Analysis (SEA; Panofsky and Brier, 1958) was performed using dplR
- 6 (Bunn, 2008) to assess post-volcanic cooling signals in our reconstruction. The approach has
- been used in studies of volcanic effect on climate (Fischer et al., 2007; D'Arrigo et al., 2009;
- 8 Esper et al. 2013a, 2013b). The major volcanic events chosen for the analysis were those
- 9 identified by Crowley (2000).
- 10 To transfer the TRW chronology into a temperature reconstruction a linear regression model
- was used. The magnitude and the spatial extent of the climate signal are evaluated considering
- 12 the CRU TS v. 3.22 gridded dataset for Europe.

14 3 Results

- 15 The correlation matrix (Fig. 3) shows <u>not only</u> the high inter-correlation between sampling
- sites and tree species but also the high correlation between each chronology and the regional
- 17 chronology. The highest correlation is found between *Pinus uncinata* (VIN and CAV) located
- 18 at the highest altitude. On the other hand, the weakest correlation is found between one of the
- 19 lowest sites (s006) and the highest (VIN). The mean correlation among all sampling sites is r
- 20 = 0.51 over the common period (1842-1977) is 0.51, and r = 0.46 over the full period of
- 21 overlap, revealing a regionally common, external forcing controlling tree growth and
- 22 justifying the development of a single chronology integrating the data from this IP tree-ring
- 23 network.
- 24 The model (regional curve) of the RCS standardization method and the model of the BasPois
- 25 method are presented in Fig.4. BasPois model (Fig.4a) indicates a growth of 130 mm when
- 26 the size of the basal area is near 0 and a growth of 8mm when it reaches the maximum basal
- area. RCS model (Fig.4b) presents values of 250 mm of growth when the cambial age is 0
- 28 with a gradual decline of the growth until the cambial of 450. Cambial age from 500 to 550
- 29 has a slight increase in growth most likely derived by low replication regarding trees with this
- 30 age.

Calibration of the four differently detrended mean chronologies reveals a highly negative correlation with maximum temperatures (Fig. 5). The ArstanRES chronology shows moderate correlations with previous-year September (r = -0.2539), and the ArstanSTD chronology correlates at r = -0.3856 with June and September and October temperature of the previous year- with a cumulative monthly mean of 21 months. Considering the RCS chronology, the previous-year September signal increases to r = -0.49 with a cumulative monthly mean of 21 months. Finally, the best correlations or revealed for the BasPois chronology reaching r = -0.78 with maximum September temperature of the previous year with a cumulative mean of 21 months, which is, in fact a two year cumulative monthly mean. Even though the signals show the same seasonal patterns among the chronologies, the BasPois record always shows the highest correlations. Accordingly, we used the BasPois chronology for the calibration and reconstruction process. The final BasPois network chronology (Fig.6) is based on 316 TRW series of *Pinus uncinata* and Pinus sylvestris spanning the 1465-2012 period. Since this chronology is derived from only living trees, mean chronology age increases from 47 years in 1966 to 528 in 1465. The mean sensitivity is 0.21, and first-order autocorrelation is 0.83. The inter-series correlation (Rbar) reaches 0.26, and the first principal component explains about 35% of the variance. The network chronology's signal to noise ratio is 48.52, and EPS exceeds 0.85 after 1602, constraining the reconstruction period to 410 years until 2012. The selection of the best climate parameter to develop the reconstruction is presented in the Figure 7. Correlations where correlations between -0.54 and -0.86 representing only the most significant values are shown. Four parameters reveal the highest correlations over the full calibration period: October of the current year with a cumulative monthly mean of 22 months; September of the previous year with a cumulative monthly mean of 20-months; September of the previous year with a cumulative monthly mean of 21 months; and October of the previous

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September of the previous year with a cumulative monthly mean of 20-months; September of the previous year with a cumulative monthly mean of 21 months; and October of the previous year with a cumulative monthly mean of 21 months. The stability of the correlation and therefore the consistency of the signal are tested considering the minimum difference between the maximum and minimum correlation (Fig. 7b) over the full running correlation period. The smallest difference (0.24) is reached for September of the previous year with a cumulative monthly mean of 21months. Therefore, this parameter is chosen for the climate reconstruction. According to the 30-year moving correlations, maximum values are reached from 1973-2003 (r = -0.80), whereas the lowest 30-year correlation (r = -0.60) is reached from

- 1 1956-1986. In addition, the relationship between September of the previous year with a
- 2 cumulative monthly mean of 21months is spatially consistent throughout the Iberian
- 3 Peninsula, reaching into southern France and northern Africa (Fig.11).

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- 5 The transfer model is validated by the high correlation (r = -0.78) and correlation
- 6 $\frac{\text{coefficientes}}{\text{coefficient coefficient of determination}}$ ($r^2 = 0.61$) over the full period 1945-
 - 2012. Through the split calibration/verification process, considering 1945-1978 and 1979-
- 8 2012, the temporal robustness was tested revealing highly significant correlations for both
- 9 periods (r^2 =0.41 and r^2 =0.55 respectively) and verifying the final reconstruction (Table 2 and
- 10 Fig. 8). The Durbin-Watson test for the full period (1.45 p<0.0001) indicates no substantial
- 11 <u>autocorrelation in the residuals.</u> To develop the final reconstruction spanning 1602-2012, we
- used a lineal regression model over the full period 1945-2012 with maximum temperature of
- 13 September of the previous year with a cumulative monthly mean of 21months (Eq.1),
- 14 denominated IR2T_{max}:
- 15 $R2T_{max} = -3.9759 * BasPoisChron + 15.769(r^2adj = 0.61; p < 0.0001).$
- 16 $IR2T_{\text{max}} = -3.9759 * BasPoisChron + 15.769(r^2 = 0.61; p < 0.0001)_{\underline{.}}$ (1)

Field Code Changed

3.1 IR2T_{max} reconstruction

- 18 IR2T_{max} describes 410 years of maximum temperature of September with a cumulative
- 19 monthly mean of 21-months meaning it has memory of the last two years. Temperature
- 20 ranges from 13.52°C (-2.13°C with respect to the mean) in 1603 to 17.64°C n (+1.94°C with
- 21 respect to the mean) in 2005 (Fig. 9). It is remarkable that the 12 years of the XXI century
- 22 happen to be within the 25 warmest years. IR2T_{max} covers a part of the Little Ice Age (Grove,
- 23 1988) from 1602 to the end of the XIX century. The year-to-year temperature variability is
- 24 3.92°C in the seventeenth century, 2.89°C in the eighteen century, 3.17°C in the nineteenth
- 25 century and 3.07°C in the twentieth century. The seventeenth and eighteen centuries were the
- 26 coldest of the reconstruction with 73% and 80% of the years with temperatures below the
- 27 long-term mean, respectively. On the other hand, the nineteenth and the twentieth centuries
- were the warmest with 66% and 78% of the years exceeding the mean.
- 29 The main driver of the large-scale character of the warm and cold episodes may be changes in
- 30 the solar activity (Fig.9). The beginning of the reconstruction starts with the end of the Spörer

- 1 Minimum. The Maunder minimum, from 1645 to 1715 (Luterbach et al., 2001) seems to
- 2 cohere with a cold period from 1645 to 1706. In addition, the Dalton minimum from 1796 to
- 3 1830, is detected for the period 1810 to 1838. However, a considerably cold period from 1778
- 4 to 1798 is not in consonance with a decrease in the solar activity. Four warm periods, 1626-
- 5 1637, 1800-1809, 1845-1859 and 1986-2012, have been identified to cohere with increased
- 6 solar activity. Overall, the correlation between the reconstruction and the solar activity is 0.34
- 7 (p < 0.0001), and increases to r = 0.49 after 11-year low pass filtering the series, thought the
- 8 degrees of freedom are substantially reduced due to the increase autocorrelation.
- 9 The SEA (Fig.10) indicates some impact of volcanic eruptions on the short-term temperature
- variability within the reconstruction. It shows significance (p < 0.05) decrease in September's
- 11 temperature with a lag of three years.
- 12 Figure 11 shows the spatial correlation between the reconstruction and the CRU TS v.3.22 for
- Europe and northern Africa. High adjusted correlations coefficient of determination ($r^2 > 0.4$, p
- $| < 0.0001 \rangle$ indicate indicates a robust agreement and spatial extend of the reconstruction over
- 15 the Iberian Peninsula (IP), especially for the central and Mediterranean Spain. The spatial
- 16 correlation, however, decreases towards the southwest of the IP and the north of Europe.

4 Discussion and conclusion

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- 19 Based on a coherent network of 11 tree-ring sites in the IR including 316 TRW series we
- 20 developed a 410-year maximum September temperature reconstruction. This record is the first
- 21 climate reconstruction for the IR filling the gap between the temperature reconstructions
- developed for the north IP (Büntgen et al., 2008; Dorado-Liñán et al., 2012a, Esper et al.
- 23 2015a) and for the southern IP (Dorado-Liñán et al, 2014). The IR2T_{max} has been achieved
- 24 using TRW as well as for the southern IP (Dorado-Liñán et al, 2014). However, for the
- 25 Pyrenees, MXD (Büntgen et al., 2008, Dorado-Liñán et al., 2012a) or stable isotopes (Esper et
- al. 2015a) are needed to get skilful records for a temperature reconstruction.
- 27 The main statistics used to verify the accuracy of the reconstruction present similar values to
- 28 those developed for the IP. For instance, the best RE coefficient is 0.99 for the split
- 29 calibration/verification modelledperiod 1945-2012 is 0.56 meaning that the reconstruction has
- 30 almost the perfect skillindeed useful skills to develop a reconstruction. A relatively high
- 31 signal to noise ratio indicates there is meaningful climatic information in the chronology. The

- 1 mean correlation between sites for the common period (r = 0.51, Fig. 3) reveals substantial
- 2 agreement between the sites and species. Correlation is strongest among high elevation sites
- 3 including the sites VIN and CAV which are both derived from Pinus uncinata. The mean
- 4 chronology, with 35.40% of the first component variance and 48.52 of signal to noise ratio,
- 5 captures the regional climate signal accurately, which highlights the beauty of regional
- 6 averages (Briffa et al., 1998).
- 7 The original, raw chronology extended over the 1465-2012 period, some 150 years longer
- 8 than the final reconstruction. However, due to low EPS values prior to 1602, which is related
- 9 to the low number of samples the final reconstruction was developed for the period 1602-
- 10 2012.
- 11 A novel detrending approach, considering a Basal Area-Poisson model instead of the
- 12 traditional regional curve (Esper et al. 2003) has certainly improved the skill of the
- 13 reconstruction and enabled retaining high-to-low frequency climate variance. The traditional
- 14 approach of using RCS with the mean TRW curve of the age-aligned data only reached
- 15 correlations with the maximum temperature of September with a cumulative monthly mean of
- 16 21 months up to r = -0.557, while with the new approach reached r = -0.78.
- 17 It is usually difficult to determine the extent to which the effects of environmental factors on
- 18 tree growth depend on age (genetic control) and/or on size (physiological control), but recent
- 19 investigations suggest that it is often the size, and not the age, that is important (Mencuccini et
- 20 al. 2005; Peñuelas 2005). In fact, climate variability is more size-dependent than age or
- 21 species (De Luis et al., 2009). Hence, the size-based standardization considered here
- 22 maximizes the common signal. In addition, when combining TRW series from different sites
- 23 and species, as done here, the heterogeneity in responses might be large. Therefore, size
- 24 standardization may be a commendable solution to develop unbiased chronologies. Finally,
- 25 the new method should be tested in other locations since it may help to maximizes responses
- 26 especially in heterogeneous areas.
- 27 Taking The development of climate parameters retaining temperature information of the past 2
- 28 years is certainly unusual and distinctive. However, memory effects in TRW data can arise
- 29 <u>from physiological processes already suggested by Schulman (1956) and Matalas (1962).</u>
- 30 Moreover, taking into account that TRW growth is conditioned by the storage of starch and
- 31 sugar in parenchyma ray tissue, the remobilization of carbohydrates from root structures, and
- 32 the development of needle enduring several growing seasons, influencing the radial increment

beyond the instant impact of temperature variability (Pallardy, 2010), we added led us to add the cumulative monthly mean to the climate parameters. In fact, we demonstrated that the signal in the study area is magnified with a memory of 21 months from the previous September... Memory effects in TRW data have been also studied regarding the delayed response in TRW (1~5 years) to post volcanic eruptions associated with a decrease in current's year temperature (D'Arrigo et al., 2013, Esper et al., 2014). Thus, developing the two year memory IR2T_{max} allowed us to maintain not only the low frequency signal, highlighting the warm and cold phases, which may be explained by the high correlation with solar activity during 410 years (0.34, p<0.001), but also the high frequency signal, emphasizing the memory effects of the volcanic eruptions in TRW, already studied by Briffa et al. (1998) and recently by Esper et al. (2015b). According to the SEA (Fig.9), the volcanic eruptions have a significance reduction (95% confidence) of September's temperature (-1.98°C) with a three years lag. However, the IR2T_{max} is already considering the two previous year's temperature, which means the temperature decrease occurred the year after the extreme volcanic event in consistency with (Frank et al., 2007a). The stability of the signal was assessed by a 30-y moving correlation from 1945 to 2012, which shows a better correlation for the period 1979-2012 in agreement with the raise of temperatures observed for last decades which may be limiting TRW growth and therefore magnifying the climate signal. However, the relationship between the chronology and the climate parameter chosen never drops from -0.54 within the calibration period 1945-2012. The negative correlation with maximum temperature of previous September is in concordance with the values detected in Cazorla by Dorado-Liñán et al. 2014. Presumably, a continuous rise in temperatures, as suggested by the IPCC (2013), will trigger an incessant decrease in the tree-ring growth. Even though the CRU dataset extents the 1901-2013 period, the general distribution of meteorological observatories in Spain did not begin until the mid-twentieth century (Gonzalez-Hidalgo et al. 2011). In fact, the closest instrumental weather station, located in Vinuesa (Fig.1), began in 1945. However, due to the large amount of gaps in the time series, the CRU dataset was used instead for the split calibration/verification approach for the period 1945-2012. The advantages of regional climatic averages were already addressed by Blasing et al. (1981) stating that the average climatic record of the gridded dataset over the study area is representative of the regional climatic conditions, and does not reflect microclimate conditions which may be characteristic of the climatic record at a single station. Tree-ring data might therefore have more variance in common with the regionally averaged climatic

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- 1 record than with the climatic record of the nearest weather station. Generally, studies have
- 2 shown that the measurements of MXD produce chronologies with an improved climatic signal
- 3 (Briffa et al., 2002) as it was revealed for summer temperature reconstructions (Hughes et al.,
- 4 1984; Büntgen et al. 2008; Matskosvsky and Helama, 2014). However, based on a TRW
- 5 chronology, it is remarkable the high correlation coefficient for the full calibration period and
- 6 the CRU dataset (r = -0.78).
- 7 Throughout the IR2T_{max} reconstruction we identified the main warm and cold phases
- 8 (Maunder minimum, Dalton minimum) related with long-term temperature variability
- 9 generally attributed to changes in cycles of activity (Lean et al., 1995; Lassen et al. 1995;
- Haigh et al. 2015). In addition, similar cold and warm phases are observed comparing with
- 11 the Pyrenees (Büntgen et al. 2008) and Cazorla (Dorado-Liñán et al. 2014) reconstructions.
- 12 However, previously to the Dalton minimum, a warm phase is detected in IR2T_{max} and the
- 13 Cazorla reconstruction although it is not present in the Pyrenees or in the Alps (Büntgen et al.,
- 14 2011).

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- 15 Through the spatial extent and magnitude of the $IR2T_{max}$ reconstruction over Europe it can be
- 16 acknowledged that the reconstruction is effective and usable for most of the Spanish Iberian
- 17 Peninsula. Working especially for the central and Mediterranean IP with very high
- 18 | $\frac{\text{correlations}}{\text{coefficient of determination (r}^2>0.4)}$.

Acknowledgements

- 21 This study was supported by the Spanish government (through the projects 'CGL2011-
- 22 28255)28255', 'CGL2015-69985' and the government of Aragon throughout the Program of
- 23 research groups (group Clima, Cambio Global y Sistemas Naturales, BOA 147 of 18-12-
- 24 2002) and FEDER funds. Ernesto Tejedor is supported by the government of Aragon with a
- 25 Ph.D. grant. Fieldwork was carried out in the province of Soria; we are most grateful to its
- authorities, for supporting the sampling campaign. We are thankful to Klemen Novak, Edurne
- 27 Martinez, Luis Alberto Longares, and Roberto Serrano for help during fieldwork.

5 References

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Table 1. Tree ring sites characteristics

Code	Site	Source	Lat	Long	Elevation	Species	Tree	Sample	Tree-	Period
							no	no	rings	
s047	Urbión	ITRDB	41.98	-2.87	1750	PISY	15	31	6549	1567-
	Covaleda									1983
s048	Urbión	ITRDB	42.02	-2.90	1840	PISY	8	17	3590	1671-
	Duruelo									1983
s049	Urbión	ITRDB	42.03	-3.03	1840	PISY	12	27	4713	1593-
	Quintenar									1985
s050	Urbión	ITRDB	42.00	-2.85	1750	PISY	4	8	1942	1681-
	Vinuesa									1983
s006	Urbión	ITRDB	42.03	-2.7	1634	PISY	11	22	2397	1842-
										1977
CAV	Castillo de	UNIZAR	42.01	-2.75	1900	PIUN	18	36	9236	1593-
	Vinuesa									2012
COV	Covaleda	IPE-	41.93	-2.83	1500	PISY	16	48	14696	1568-
		CSIC-								1993
		UNIZAR								
HER	Barranco	IPE-	41.94	-2.84	1500	PISY	25	32	9347	1562-
	de las	CSIC-								1993
	heridas	UNIZAR								
NEI	Neila	IPE-	42.05	-3.08	1850	PISY	9	15	4822	1587-
		CSIC-								1992
		UNIZAR								

URB	Picos Urbión	de	UNIZAR	41.96	-2.82	1750	PISY	28	60	11328	1733- 2012
VIN	Castillo Vinuesa		IPE- CSIC- UNIZAR	42.03	-2.73	1900	PIUN	13	20	7653	1465- 1992
							Total	159	316	76273	

UNIZAR University of Zaragoza, IPE-CSIC Spanish National Research Council, ITRDB International Tree-Ring
Databank

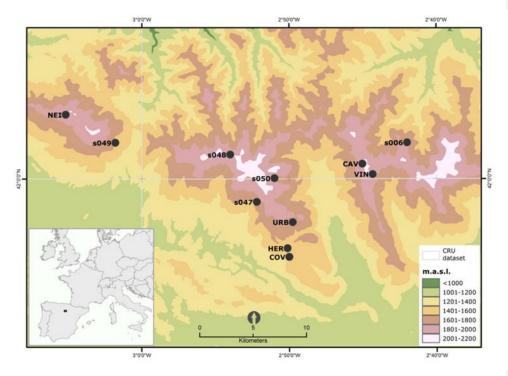


Figure 1. Map showing the tree ring study sites and the climate data (CRU TS v.3.22) grid points in the Western Iberian Range (Soria).

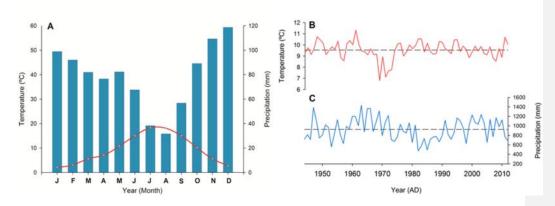


Figure 2. Climate diagram (A), mean temperature (B), mean precipitation (C) calculated using
data from CRU TS v.3.22 over the period 1944-2012 (Harris et al 2014).

	CHRONOLOGY	VIN(1900)	CAW(1900)	NEI(1850)	s048(1840)	s049(1840)	s047(1750)	s050(1750)	URB(1750)	s006(1630)	COV(1500)	HER(1500)		_ %
CHRONOLOGY	1	0.47	0.65	0.69	0.69	0.74	0.69	0.63	0.83	0.7	0.83	0.71		1
VIN(1900)	0.6	1	0.83	0.33	0.18	0.17	0.26	0.33	0.42	0.04	0.27	0.34	ı	- 0.8
CAV(1900)	0.73	0.83	1	0.37	0.29	0.28	0.42	0.41	0.55	0.3	0.46	0.4		- 0.6
NEI(1850)	0.74	0.51	0.56	1	0.59	0.7	0.39	0.42	0.53	0.24	0.61	0.57		- 0.4
s048(1840)	0.74	0.32	0.49	0.49	1	0.46	0.37	0.34	0.56	0.4	0.72	0.67		- 0.2
s049(1840)	0.57	0.13	0.23	0.58	0.39	1	0.48	0.54	0.7	0.4	0.53	0.57		- 0
s047(1750)	0.64	0.08	0.06	0.36	0.11	0.31	1	0.6	0.67	0.37	0.61	0.55		
s050(1750)	0.65	0.4	0.41	0.45	0.44	0.3	0.39	1	0.55	0.23	0.47	0.44		0.2
URB(1750)	0.7	0.37	0.34	0.45	0.42	0.34	0.55	0.51	1	0.41	0.69	0.66		0.4
s006(1630)	0.7	0.04	0.3	0.24	0.4	0.4	0.37	0.23	0.41	1	0.54	0.25		0.6
COV(1500)	0.75	0.12	0.2	0.5	0.53	0.48	0.8	0.56	0.68	0.54	ī	0.62		0.8
HER(1500)	0.69	0.41	0.39	0.47	0.54	0.45	0.39	0.53	0.6	0.25	0.56	1		1

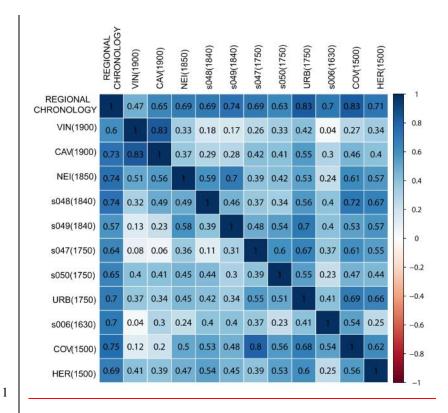
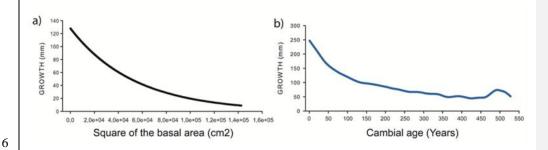


Figure 3. CorrelationInter correlation of the raw chronologies between sites and the regional chronology, sorted by elevation. Top right shows the correlations calculated over the common period 1842-1977. Bottom left shows the correlation over the full period of overlap between pairs of chronologies



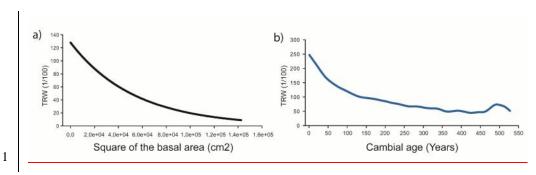
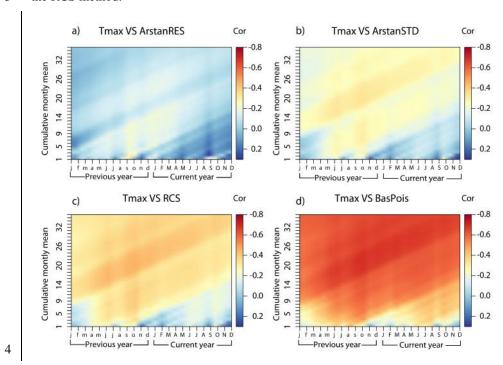


Figure 4. a) Represents the model of the BasPois method, b) represents the regional curve of the RCS method.



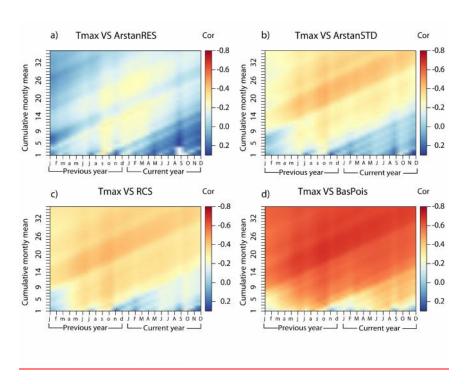


Figure 5. Correlation between the maximum temperature (from January of the previous year to December of the current year with a cumulative monthly mean from 1 to 36 months) and the residual Arstan chronology (a), the standard Arstan chronology (b), the RCS standard chronology (c) and the Basal Area-Poisson standard chronology (d).

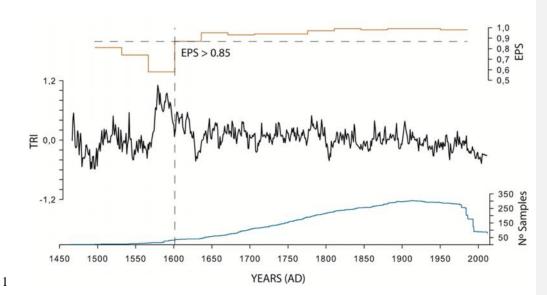


Figure 6. BasPois chronology (in black), number of samples (blue) and EPS statistic (computed over 30-y window lagged by 15 years) back to 1465. Vertical dashed line highlights the EPS=0.85 threshold in 1602.

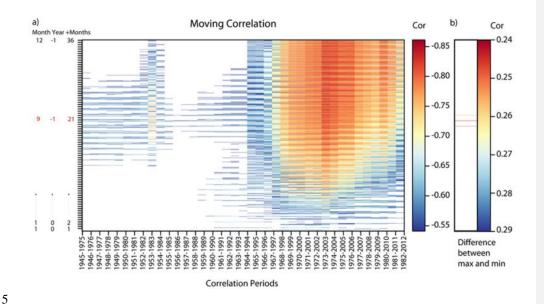


Figure 7.a) 30-year moving correlation from 1945 to 2012 between the maximum temperature, from January of the current year (1,0,1) to December of the previous year (12, -1, 36) with a cumulative monthly mean from 1 to 36 months and the BasPois chronology. Red numbers indicates the chosen climatological parameter; 9, September, -1, previous year, 21,

months used for the cumulative monthly mean. b) The four best parameters are represented. Reddish line indicates the least difference between the maximum and minimum correlation in the correlation periods.

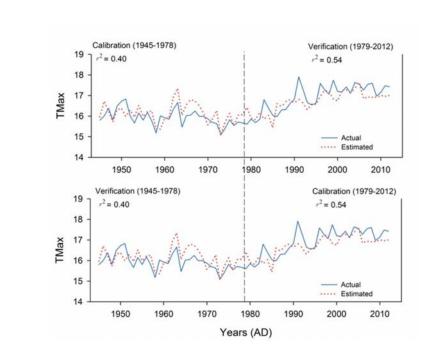


Figure 8. Calibration and verification results of the CRU data based Tmax_{Sep-1} reconstruction

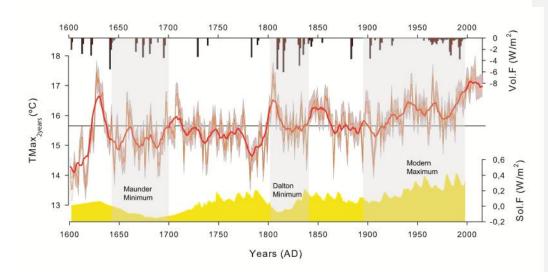


Figure 9. IR2T_{max} reconstruction since AD 1602 for the Iberian Range. Bold red curve is a 11year running mean, purple shading indicates the mean square error based on the calibration period correlation. Yellow shading at the bottom show solar forcing and bars on top indicate volcanic forcings (Crowley 2000).

	Calibration	Verification	Calibration	Verification	Period	→ ¬¬¬¬	Formatted: Centered
	1945-1978	1978-2012	1979-2012	1945-1978	1945-2012		Formatted Table
Years	34	34	34	34	68	4	Formatted: Centered
Correlation	-0.64	0.73	-0.74	0.64	-0.78	4	Formatted: Centered
\mathbb{R}^2	0.41	0.55	0.55	0.41	0.61	4	Formatted: Centered
MSE	0. 09 43	0. 66 42	0. 18 <u>42</u>	0. 29 43	0. 37<u>43</u>	4	Formatted: Centered
Reduction of	0. 99 40	0. 99 <u>65</u>	0. 99 65	0. 99<u>40</u>	0. 99 <u>56</u>	4	Formatted: Centered
error							
Sing test	28+/6-	24+/10-	28+/6-	24+/10-	52+/16-		
Durbin-	1.31 p<0.01	1.53 p<0.05	1.53 p<0.05	1.31 p<0.01	1.45 p<0.001	15	Formatted: Centered
	1.61 p <0.01	1.00 p x0.05	2.00 p x0.00	2.6.1 p x0.01	2e p x0.001	11/	Formatted: Centered
Watson						_ '','	Formatted: Centered

Table 2. Calibration/verification statistics of the Tmax_{Sep-1} reconstruction

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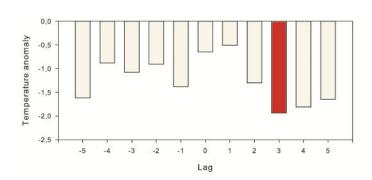
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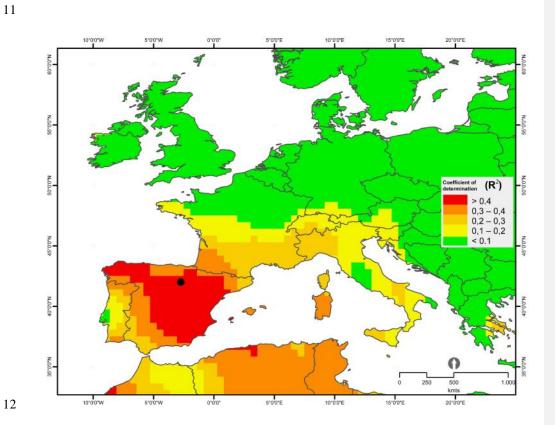
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9 Figure 10. Superposed epoch analysis with a back and forward lag of 5 years. Significance (p<0.05) at 3 years after the extreme volcanic event.



- 1 Figure 11. Map showing the spatial correlation patterns of the BasPois chronology with the
- 2 gridded September of the previous year with a cumulative monthly mean of 21months data.
- 3 Correlation values are significant at p<0.0001.