

1 **Synchronizing the Greenland ice core and radiocarbon  
2 timescales over the Holocene - Bayesian wiggle-matching  
3 of cosmogenic radionuclide records**

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8

9 **Abstract**

10 Investigations of past climate dynamics rely on accurate and precise chronologies of the  
11 employed climate reconstructions. The radiocarbon dating calibration curve (IntCal13) and  
12 the Greenland ice core chronology (GICC05) represent two of the most widely used  
13 chronological frameworks in paleoclimatology of the past ~50,000 years. However,  
14 comparisons of climate records anchored on these chronologies are hampered by the  
15 precision and accuracy of both timescales. Here we use common variations in the production  
16 rates of  $^{14}\text{C}$  and  $^{10}\text{Be}$  recorded in tree-rings and ice cores, respectively, to assess the  
17 differences between both timescales during the Holocene. Compared to earlier work, we  
18 employ a novel statistical approach which leads to strongly reduced and yet, more robust,  
19 uncertainty estimates. Furthermore, we demonstrate that the inferred timescale differences are  
20 robust independent of (i) the applied ice core  $^{10}\text{Be}$  records, (ii) assumptions of the mode of  
21  $^{10}\text{Be}$  deposition, as well as (iii) carbon cycle effects on  $^{14}\text{C}$ , and (iv) in agreement with  
22 independent estimates of the timescale differences. Our results imply that the GICC05  
23 counting error is likely underestimated during the most recent 2,000 years leading to a dating  
24 bias that propagates throughout large parts of the Holocene. Nevertheless, our analysis  
25 indicates that the GICC05 counting error is generally a robust uncertainty measurement but  
26 care has to be taken when treating it as a nearly Gaussian error distribution. The proposed  
27 IntCal13-GICC05 transfer function facilitates the comparison of ice core and radiocarbon  
28 dated paleoclimate records at high chronological precision.

29

1    **1    Introduction**

2    Paleoclimatology can provide significant insights into natural climate changes and thus,  
3    improve our understanding of the climate system. Besides the reconstruction of past climate  
4    itself, a precise chronology of each paleoclimate record is crucial to reliably assess the  
5    dynamics of the inferred changes. Furthermore, consistent chronologies across multiple  
6    paleoclimate records are required to assess the spatiotemporal evolution of climatic events  
7    and thus, to test for potential leads and lags within the climate system and ultimately improve  
8    the understanding of the underlying processes of past climate change. Two independent key  
9    timescales in paleoclimatology of the past 50,000 years are the radiocarbon- (IntCal13,  
10   Reimer et al., 2013) and the Greenland ice core timescale (GICC05, Andersen et al.,  
11   2006;Rasmussen et al., 2006;Seierstad et al., 2014;Svensson et al., 2008;Vinther et al., 2006).  
12   To be able to infer leads and lags between paleoclimatic changes anchored on these  
13   chronologies at high precision, it is crucial to test the consistency between the timescales and  
14   establish climate-independent isochrones and thus, reduce the influence of their absolute  
15   dating uncertainties (e.g., Lane et al., 2013). One method to compare and synchronize  
16   different timescales is the use of cosmogenic radionuclide records, such as  $^{10}\text{Be}$  and  $^{14}\text{C}$   
17   (Muscheler et al., 2014a;Muscheler et al., 2014b;Muscheler et al., 2008;Southon, 2002).

18   Cosmogenic radionuclides such as  $^{10}\text{Be}$  and  $^{14}\text{C}$  are produced in the atmosphere through a  
19   nuclear cascade mainly triggered by incoming galactic cosmic rays (GCR, Lal and Peters,  
20   1967). The flux of GCR reaching the atmosphere is in turn modulated by the strength of the  
21   helio- and geo- magnetic fields resulting in varying production rates of  $^{10}\text{Be}$  and  $^{14}\text{C}$  (Masarik  
22   and Beer, 2009, 1999;Kovaltsov et al., 2012;Kovaltsov and Usoskin, 2010). Thus, increased  
23   (decreased) intensity of the solar- and/or geomagnetic field will result in decreased  
24   (increased) cosmogenic radionuclide production rates. Therefore,  $^{14}\text{C}$  and  $^{10}\text{Be}$  production  
25   rates co-vary globally due to external processes, making them a powerful synchronization  
26   tool.

27   After production,  $^{14}\text{C}$  oxidizes to  $^{14}\text{CO}_2$  that enters the global carbon cycle and gets stored in  
28   various environmental archives such as tree rings, sediments, and speleothems.  $^{10}\text{Be}$  attaches  
29   to aerosols which are deposited within 1-2 years (Raisbeck et al., 1981) by wet and dry  
30   deposition processes and is stored in sediments including polar ice sheets. These ‘system  
31   effects’ (i.e., non-production influences on  $^{10}\text{Be}$  and  $^{14}\text{C}$  records such as the mixing, transport,  
32   and deposition of  $^{14}\text{C}$  and  $^{10}\text{Be}$ ) can challenge an unequivocal reconstruction of cosmogenic

1 radionuclide production rates from paleoarchives and thus, synchronization efforts based on  
2 cosmogenic radionuclides.

3 Due to the large actively exchanging carbon reservoirs, changes in the atmospheric  $^{14}\text{C}/^{12}\text{C}$   
4 ratio are attenuated and delayed compared to the corresponding  $^{14}\text{C}$  production rate variations  
5 (Oeschger et al., 1975). In comparison,  $^{10}\text{Be}$  is a more direct recorder of production rate  
6 changes. Thus, when comparing  $^{14}\text{C}$  and  $^{10}\text{Be}$  records directly, this difference in  
7 geochemistry has to be taken into account by using carbon cycle models (Muscheler et al.,  
8 2004b). However, to be fully realistic, these corrections would require prior knowledge on  
9 the variable state of the carbon cycle, which is often difficult to quantify (Köhler et al., 2006).  
10  $^{10}\text{Be}$  records (for example from ice cores) can be affected by non-production related  
11 processes as well. Firstly, it depends on the assumed mode of deposition (wet vs. dry)  
12 whether the  $^{10}\text{Be}$  concentration (all wet deposition) or the  $^{10}\text{Be}$  flux (all dry deposition) is the  
13 better measure of atmospheric  $^{10}\text{Be}$  concentration changes (Alley et al., 1995; Delaygue and  
14 Bard, 2010). In reality, both modes of deposition contribute to the accumulation of  $^{10}\text{Be}$  on  
15 the ice sheet. Today, wet deposition processes dominate over dry deposition which accounts  
16 for about one third or less of the deposited  $^{10}\text{Be}$  in Greenland (Heikkilä et al., 2011; Elsässer  
17 et al., 2015). However, this dry/wet deposition ratio has likely been variable over time (Alley  
18 et al., 1995). Secondly, a variety of climatic influences can leave an imprint in ice core  $^{10}\text{Be}$   
19 records. Atmospheric circulation changes and air mass precipitation history (i.e.,  $^{10}\text{Be}$   
20 scavenging by precipitation prior to the arrival of the air mass at the ice core site) may, for  
21 example, modulate the transport path and efficiency of  $^{10}\text{Be}$  delivery to the ice core site  
22 (Heikkilä and Smith, 2013; Pedro et al., 2012; Pedro et al., 2011b). Furthermore, changes in  
23 the exchange rates between stratospheric (high  $^{10}\text{Be}$  concentrations) and the tropospheric  
24 (low  $^{10}\text{Be}$  concentrations) air masses can affect the tropospheric  $^{10}\text{Be}$  budget (Pedro et al.,  
25 2011a). Thirdly, contrary to  $^{14}\text{C}$ ,  $^{10}\text{Be}$  might not be hemispherically well mixed owing to its  
26 short atmospheric residence time. This has led to the proposition of a so-called “polar bias” in  
27 ice core  $^{10}\text{Be}$  records, stating that if polar  $^{10}\text{Be}$  records were dominated by  $^{10}\text{Be}$  produced at  
28 high latitudes, the anisotropy of the geomagnetic shielding would lead to an enhanced solar-  
29 and an attenuated geomagnetic modulation signal in polar  $^{10}\text{Be}$  records. There is  
30 contradicting evidence from data and modelling studies to whether this is the case (Field et  
31 al., 2006; Bard et al., 1997; Pedro et al., 2012; Muscheler and Heikkilä, 2011; Heikkilä et al.,  
32 2009; Elsässer et al., 2015).

1 In summary, to be able to use  $^{10}\text{Be}$  and  $^{14}\text{C}$  as synchronization tools, ‘system effects’ on each  
2 radionuclide have to be assessed and corrected for. If successful, this method has the  
3 advantage that it can provide near-continuous estimates of time scale differences as opposed  
4 to discrete tie-points obtained from tephrochronology (Abbott and Davies, 2012;Lane et al.,  
5 2013) or changes in atmospheric trace gases during Dansgaard-Oeschger events (Blunier et  
6 al., 1998;Buizert et al., 2015).

7

## 8 **1.1 Aim of this study**

9 Recently, Muscheler et al. (2014a) assessed the differences of the radiocarbon and ice core  
10 time scales for the past 14,000 years by comparing GRIP  $^{10}\text{Be}$  (Yiou et al., 1997;Muscheler et  
11 al., 2004b;Vonmoos et al., 2006) and IntCal13  $^{14}\text{C}$  data (Reimer et al., 2013). Here, we revisit  
12 this approach using a different statistical framework (Bronk Ramsey et al., 2001) that is  
13 computationally less expensive and provides improved error estimates for the inferred  
14 timescale differences as compared to the method used in Muscheler et al. (2014a).  
15 Furthermore, we test the robustness of the obtained results with respect to the use of different  
16 ice core  $^{10}\text{Be}$  records as well as potential ‘system effects’ on the radionuclide records. We  
17 focus our analysis on the period where dendrochronologically dated high quality  $^{14}\text{C}$   
18 measurements on tree rings are available. While this is theoretically the case back to 12,560  
19 calBP (calibrated before present, AD1950, Friedrich et al., 2004), the accuracy of the oldest  
20 part of tree-ring chronology has recently been questioned (Hogg et al., 2013) causing a gap in  
21 the  $^{14}\text{C}$  records underlying IntCal13 around 12,000 calBP (Reimer et al., 2013). Hence, we  
22 limit our analysis to the Holocene where dendrochronological and  $^{14}\text{C}$ -data replication is high  
23 and most robust (Reimer et al., 2013;Friedrich et al., 2004).

24

## 25 **2 Methods**

### 26 **2.1 Data**

27 The key data used in this paper is shown in figure 1. The GRIP  $^{10}\text{Be}$  record (Vonmoos et al.,  
28 2006;Muscheler et al., 2004b;Yiou et al., 1997) covers almost the entire Holocene with a gap  
29 between 9,400 and 10,800 years BP (Before Present 1950 AD) and no data for sections  
30 younger than 300 years BP. We use the data as presented in Vonmoos et al. (2006) that

1 includes a 61-point binomial filter (roughly corresponding to a 20 year low-pass filter or a  
2 decadal sampling resolution) minimizing weather related noise in the  $^{10}\text{Be}$  data. The GISP2  
3  $^{10}\text{Be}$  record (Finkel and Nishiizumi, 1997) has a gap between 7980 and 9400 years BP and no  
4 data for sections younger than 3270 years BP. We used the GISP2  $^{10}\text{Be}$  record on the  
5 GICC05 timescale (Seierstad et al., 2014). Its temporal resolution varies between 20 to 60  
6 years with an average of one sample every 35 years. Hence, no smoothing filter was applied.  
7 The GISP2  $^{10}\text{Be}$  concentrations have been normalized to the same standard used for the GRIP  
8  $^{10}\text{Be}$  measurements (NIST SRM 4325, see Yiou et al., 1997; Muscheler et al., 2004b). The  
9 resulting GRIP and GISP2  $^{10}\text{Be}$  records differ by on average 0.12  $10^4$  atoms/g of ice. To avoid  
10 inhomogeneities when splicing the records together, we adjusted the GISP2  $^{10}\text{Be}$  data  
11 accordingly by adding 0.12  $10^4$  atoms/g to the GISP2  $^{10}\text{Be}$  record (see figure 1). We note that  
12 reconciling the  $^{10}\text{Be}$  records through normalization instead of addition does not affect the  
13 results shown here. The lower panel in figure 1 shows atmospheric  $\Delta^{14}\text{C}$  (that is  $^{14}\text{C}/^{12}\text{C}$  after  
14 correction for fractionation and decay relative to a standard) as reconstructed from  
15 dendrochronologically dated tree rings (Friedrich et al., 2004) and presented in IntCal13 in 5-  
16 year resolution while the underlying data has typically a resolution of 10 years for most of the  
17 Holocene (Reimer et al., 2013).

18

## 19 **2.2 Statistical method**

20 In the following section we will describe the statistics used for the  $^{14}\text{C}/^{10}\text{Be}$  comparison. To  
21 be able to compare both radionuclides quantitatively, we converted the ice core  $^{10}\text{Be}$  records  
22 into  $\Delta^{14}\text{C}$  variations using a box-diffusion carbon cycle model (Siegenthaler et al.,  
23 1980; Muscheler et al., 2004b). The details of this conversion and its uncertainties are  
24 addressed in more detail in section 2.4. In the following we will refer to these modelled  $\Delta^{14}\text{C}$   
25 variations as „ $^{10}\text{Be}$ -based  $\Delta^{14}\text{C}$  anomalies“.

26 We employ a statistical approach that is commonly used in the ‘wiggle-match dating’ of  $^{14}\text{C}$   
27 records that have an initial relative chronology, i.e. the age differences between neighbouring  
28 samples are known, such as tree-rings (Bronk Ramsey et al., 2001). Contrary to classical  $^{14}\text{C}$ -  
29 age calibration we use  $\Delta^{14}\text{C}$  anomalies, since  $^{10}\text{Be}$  cannot provide information on absolute  
30  $\Delta^{14}\text{C}$  (and hence,  $^{14}\text{C}$ -ages) which depends on  $^{14}\text{C}$  production rates and the state of the carbon  
31 cycle long before the investigated period. Given the results shown in section 3.1 we employ  
32 centennial (<500 year FFT high-pass filter)  $\Delta^{14}\text{C}$  anomalies of the tree-ring and the  $^{10}\text{Be}$ -

1 based  $\Delta^{14}\text{C}$  records for this comparison as shown in figure 3. The mathematical formulation  
 2 remains however, unchanged. The calibration record, IntCal13 (Reimer et al., 2013),  
 3 describes  $\Delta^{14}\text{C}$  anomalies for each point in time,  $R(t)$ , with an associated uncertainty,  $\delta R(t)$ .  
 4 This can be compared to  $^{10}\text{Be}$ -based  $\Delta^{14}\text{C}$  anomalies ( $R_{i:n}$ ) for which we know the absolute  
 5 age differences ( $\Delta t_i$ ) between each sample from ice core layer counting. We can estimate the  
 6 probability ( $P_i$ ) for different assumed time scale differences between the records ( $t_s$ ) for each  
 7 sample by using equation 8 in Bronk Ramsey et al. (2001):

$$8 P_i(t_s + \Delta t_i) \propto \frac{\exp\left(-\frac{(R_i - R(t_s + \Delta t_i))^2}{2(\delta R_i^2 + \delta R^2(t_s + \Delta t_i))}\right)}{\sqrt{\delta R_i^2 + \delta R^2(t_s + \Delta t_i)}} \quad (1)$$

9 Using Bayes' theorem to combine the probabilities for each individual measurement we can  
 10 obtain an overall probability ( $P_s$ ) for each time scale difference between GICC05 and  
 11 IntCal13 (equation 9 in Bronk Ramsey et al., 2001):

$$12 P_s(t_s) \propto \prod_{i=1}^n P_i(t_s + \Delta t_i) \quad (2)$$

13 To allow a continuous comparison, all records have been interpolated to annual resolution.  
 14 However, since the ice core sampling resolution is in reality lower we do not obtain truly  
 15 independent probability distributions for each sample. Consequently, we correct for the  
 16 reduced degrees of freedom by scaling  $P_s$  as:

$$17 P_{s_{scaled}}(t_s) = P_s(t_s)^{1/r} \quad (3)$$

18 where  $r$  is the original sample spacing (years/sample) of the ice core  $^{10}\text{Be}$  records. This  
 19 scaling effectively widens the obtained probability distribution and thus, increases the derived  
 20 uncertainties. For the filtered GRIP  $^{10}\text{Be}$  record, we assume a decadal resolution.

21 This ‘wiggle-matching’ is done for predefined windows of IntCal13 and GRIP and hence,  
 22 yields a probability distribution ( $P_{s_{scaled}}(t_s)$ ) for their time scale difference for each window.  
 23 We apply this method to 1,000 year windows of  $^{14}\text{C}/^{10}\text{Be}$  data and investigate one window  
 24 every 50 years back in time. For each window we test for time scale differences (shifts) of  $\pm$   
 25 150 years without stretching or compression of the timescale within this window. Hence, in  
 26 analogy to  $^{14}\text{C}$ -wiggle-match dating, each window could be seen as a single 1,000 year long  
 27 “tree” that is being calibrated. We tested different window sizes between 500 and 2,000 year  
 28 length and the corresponding results are consistent within error. The choice of a 1,000 year  
 29 window represents a trade-off between (i) an increasing statistical robustness and hence,

1 smaller uncertainties, and (ii) a loss of detail (variability) in the final transfer function (see  
2 also section 2.5) with increasing window length.

3 It can be seen from equation 1, that contrary to the correlation analysis employed by  
4 Muscheler et al. (2014a) this method favours  $^{10}\text{Be}/^{14}\text{C}$  linkages with a direct 1:1 relationship  
5 between IntCal13 and  $^{10}\text{Be}$ -based  $\Delta^{14}\text{C}$  records. Hence, the  $^{14}\text{C} : ^{10}\text{Be}$  production rate ratio has  
6 to be assessed. Furthermore, the uncertainty for the  $^{10}\text{Be}$ -based records and the  $^{10}\text{Be} : ^{14}\text{C}$   
7 conversion is quantitatively included in the calculation and hence, needs to be estimated. In  
8 the following sections we will outline how these factors can be initially assessed.

9

## 10 **2.3 Assessment of uncertainties due to climatic influences on $^{10}\text{Be}$**

11 As outlined in the introduction, ice core  $^{10}\text{Be}$  records can be affected by various climatic  
12 influences that can ‘contaminate’ the production signal. To account for these effects, we use  
13 four different versions of the GRIP and GISP2  $^{10}\text{Be}$  records throughout the manuscript. We  
14 use  $^{10}\text{Be}$  concentrations and fluxes ( $^{10}\text{Be}$  concentration multiplied by snow accumulation and  
15 ice density) as endmembers of the assumed mode of  $^{10}\text{Be}$  deposition (wet vs. dry,  
16 respectively) on the ice sheet. To address the role of climate influences on  $^{10}\text{Be}$  mixing and  
17 transport to the ice sheet, we additionally generated “climate corrected” versions of the  
18 concentrations and fluxes. For this purpose, we performed multiple linear regression analysis  
19 between  $^{10}\text{Be}$  and climate proxy time series from the GRIP and GISP2 ice cores. Using ice  
20 accumulation rates (Seierstad et al., 2014),  $\delta^{18}\text{O}$  (Johnsen et al., 1995; Stuiver et al., 1997),  
21 and ion data (Mayewski et al., 1997) as predictors, we linearly detrended the  $^{10}\text{Be}$   
22 concentrations and fluxes. This procedure removes covariance between  $^{10}\text{Be}$  and climate  
23 proxy data and may thus, diminish the climate influences in the  $^{10}\text{Be}$  record. It should be  
24 noted, that this is a ‘blind’ empirical approach that does not aim for a process based  
25 understanding of the climate influences on  $^{10}\text{Be}$ . This method would, for example, confound  
26 solar ( $^{10}\text{Be}$ ) variations that had an influence on climate as climate influences on  $^{10}\text{Be}$   
27 (Adolphi et al., 2014). Hence, these ‘climate corrected’ versions should rather be seen as  
28 sensitivity tests for our analysis than as improved estimates of past  $^{10}\text{Be}$  production rates per  
29 se. In summary, we use four (concentrations, fluxes, and “climate corrected” versions  
30 thereof) different versions of the GRIP and GISP2  $^{10}\text{Be}$  data. Each version represents a  
31 plausible endmember of the  $^{10}\text{Be}$  production rate history, depending on the assumed mode of

1 deposition and climatic impacts on  $^{10}\text{Be}$  and can thus, be used to assess the sensitivity of our  
2 analysis to these processes.

3

4 **2.4 Assessment of uncertainties due to  $^{10}\text{Be} - ^{14}\text{C}$  conversion**

5 **2.4.1 Carbon cycle modelling**

6 To be able to compare  $^{10}\text{Be}$  to  $^{14}\text{C}$  records, we converted the  $^{10}\text{Be}$  records into  $\Delta^{14}\text{C}$   
7 anomalies using a box-diffusion carbon cycle model (Oeschger et al., 1975; Siegenthaler et  
8 al., 1980). The model was run under pre-industrial conditions and has been shown to yield  
9 consistent results with more complex carbon cycle models for our purposes (Muscheler et al.,  
10 2007). As outlined in the introduction, the unknown state and dynamics of the carbon cycle  
11 introduce uncertainty to the comparison of  $^{10}\text{Be}$  and  $^{14}\text{C}$ . To test for the sensitivity to these  
12 effects, we conducted four experiments (table 1). Each experiment was forced with an  
13 idealized 200 year  $^{14}\text{C}$  production rate cycle of  $\pm 20\%$  approximately corresponding to a  
14 solar de Vries cycle. For two of the experiments we perturbed the state of the carbon cycle by  
15 increasing (S1) or decreasing (S2) the air-sea gas exchange constant by 50 % mimicking  
16 changes in wind speed and/or sea ice extent. In the scenarios S3 and S4 the ocean diffusivity  
17 parameter (ocean ventilation) was increased and decreased by 50 %, respectively. Each  
18 experiment was spun up for 50,000 years under preindustrial conditions until all  $^{14}\text{C}$   
19 reservoirs were in steady state. Subsequently the investigated parameter was changed linearly  
20 from its preindustrial to its perturbed value within 50 years (transition 1). The perturbed state  
21 was then maintained for 25,000 years to reach equilibrium again (steady state) before linearly  
22 changing the perturbed parameter back to preindustrial values within 50 years (transition 2).  
23 We use these different sensitivity experiments to obtain an uncertainty estimate of the  
24 modelled ( $^{10}\text{Be}$ -based)  $\Delta^{14}\text{C}$  records due to carbon cycle effects.

25

26 **2.4.2  $^{10}\text{Be}/^{14}\text{C}$  production rate ratio**

27 To compare tree ring and ice core radionuclide records we used the normalized  $^{10}\text{Be}$  records  
28 as  $^{14}\text{C}$  production rate input for the carbon cycle model. This yields a  $^{10}\text{Be}$ -based  $\Delta^{14}\text{C}$   
29 anomaly record that can be directly compared to the tree-ring data. Hence, we have to assume  
30 a ratio between the production rates of  $^{14}\text{C}$  and  $^{10}\text{Be}$ . This ratio depends on the radionuclide

1 production cross sections and the energy spectrum of the incoming GCR. Model estimates of  
2 relative  $^{14}\text{C}$ : $^{10}\text{Be}$  production rate increases for a change in the solar modulation parameter  
3 from 700 to 0 MeV at modern geomagnetic field strength differ between 1.34 (Masarik and  
4 Beer, 2009) and 1.04 (Kovaltsov et al., 2012; Kovaltsov and Usoskin, 2010). Similarly, the  
5 predicted  $^{14}\text{C}$ : $^{10}\text{Be}$  production rate ratios for changes in the geomagnetic field strength are  
6 model dependent for unresolved reasons (Cauquoin, 2014).

7 Furthermore, the  $^{14}\text{C}$ : $^{10}\text{Be}$  production rate ratio depends on the presence of a potential ‘polar  
8 bias’ (see introduction). If a ‘polar bias’ was present (Bard et al., 1997; Field et al., 2006) the  
9 ratio between  $^{14}\text{C}$  and ice core  $^{10}\text{Be}$  variations could be biased towards lower values. (Bard et  
10 al. (1997) report a value of 0.65 for the South Pole  $^{10}\text{Be}$  record). For Greenland, however,  
11 high resolution  $^{10}\text{Be}$  records do not support such a strong polar bias but would instead be  
12 consistent with a well mixed atmosphere (Pedro et al., 2012; Muscheler and Heikkilä, 2011).  
13 Simply comparing the standard deviations of centennial variations of IntCal13 and  $^{10}\text{Be}$ -  
14 based  $\Delta^{14}\text{C}$  anomalies leads to ratios between 0.95 and 1.05 ( $\sigma^{14}\text{C}_{\text{IntCal}}/\sigma^{14}\text{C}_{^{10}\text{Be}}$ ) depending on  
15 which ice core (GRIP/GISP2) and which version of the  $^{10}\text{Be}$  records (concentration, flux,  
16 climate corrections) is used. Thus, we start with a  $^{14}\text{C}$ : $^{10}\text{Be}$  production rate ratio of 1:1 and  
17 test the sensitivity of our results to this assumption by repeating the calculations outlined in  
18 section 2.2 using  $^{14}\text{C}$ : $^{10}\text{Be}$  ratios of 1.5:1 and 0.5:1.

19

## 20 **2.5 Timescale transfer function**

21 The methodology outlined in section 2.2 yields a probability estimate of the IntCal13-  
22 GICC05 timescale difference every 50 years. These probability distributions are however not  
23 fully independent since neighbouring 1,000 year windows overlap and are, hence, largely  
24 based on the same data. To create a timescale transfer function we employed a Monte-Carlo  
25 procedure that creates 20,000 possible transfer functions based on independent, i.e. non-  
26 overlapping, windows. Each iteration, i) randomly selects one of the youngest (most recent)  
27 20 windows and ii) randomly samples from the probability distribution  $P_{s_{\text{scaled}}}(t_s)$  of this  
28 window as well as the older non-overlapping windows (i.e. one window every 1,000 years so  
29 that the selected windows are fully independent with respect to the data points they contain).  
30 The resulting transfer functions are then interpolated to annual resolution and converted into  
31 probability distributions for the timescale difference at each point in time. For each transfer  
32 function we assume that both timescales are correct at 0 BP (i.e. AD 1950).

1

## 2 **2.6 Iterative structure of the synchronization method**

3 The separate aspects of our synchronization method outlined above are applied in an iterative  
4 manner to obtain robust and self consistent error estimates for our results. The different steps  
5 involved are carried out in the following order:

- 6 i. We create four versions of both ice core  $^{10}\text{Be}$  records as endmembers of plausible  
7  $^{10}\text{Be}$  production rate histories (see section 2.3).
- 8 ii. We convert these  $^{10}\text{Be}$  records into  $\Delta^{14}\text{C}$  using a box-diffusion carbon cycle model  
9 (section 2.4.1) assuming a  $^{14}\text{C} : {}^{10}\text{Be}$  production rate ratio of 1 (see section 2.4.2).
- 10 iii. The difference between the different  $^{10}\text{Be}$ -based  $\Delta^{14}\text{C}$  records, and results from the  
11 carbon cycle sensitivity experiments (see section 2.4.1) serve as initial uncertainty  
12 estimates for the  $^{10}\text{Be}$ -based  $\Delta^{14}\text{C}$  records.
- 13 iv. We then compare the tree ring and  $^{10}\text{Be}$ -based  $\Delta^{14}\text{C}$  records with respect to their  
14 timescale differences using the statistics outlined in section 2.2. We test for the  
15 robustness of these results by using all four different  $^{10}\text{Be}$  versions of GRIP and  
16 GISP2 separately as well as  $^{10}\text{Be}$ - $^{14}\text{C}$  conversion factors of 0.5 and 1.5 (see section  
17 2.4.2).
- 18 v. Calculating an initial timescale transfer function (see section 2.5) we then  
19 synchronize IntCal13 and GICC05. This enables us to directly compare tree ring and  
20  $^{10}\text{Be}$ -based  $\Delta^{14}\text{C}$  records and estimate the optimal  $^{14}\text{C} : {}^{10}\text{Be}$  production rate ratio, as  
21 well as uncertainties for the  $^{10}\text{Be}$ -based  $\Delta^{14}\text{C}$  record.
- 22 vi. Based on these posterior estimates of the  $^{14}\text{C} : {}^{10}\text{Be}$  ratio and the uncertainty of the  
23  $^{10}\text{Be}$  records, we repeat the calculations outlined in sections 2.2 and 2.5 yielding our  
24 final estimates of the IntCal13-GICC05 timescale differences over the Holocene.

25

## 26 **3 Results**

### 27 **3.1 Climate and Carbon cycle related uncertainties in the GRIP and GISP2** 28 **$^{10}\text{Be}$ records**

29 Figure 2 displays the different  $^{10}\text{Be}$  production rate scenarios from GRIP (top two panels) and  
30 GISP2 (lower two panels)  $^{10}\text{Be}$  concentrations (Conc), fluxes (Flux) and their climate

1 corrected versions ( $\text{Conc}_{\text{clim}}$  and  $\text{Flux}_{\text{clim}}$ , respectively). Dividing the  $^{10}\text{Be}$  records into a  
2 centennial ( $<500$  years) and millennial ( $>500$  years) variations indicates that the different  
3  $^{10}\text{Be}$  versions mainly differ in the low frequency range. These millennial differences can  
4 systematically affect the modelling of  $\Delta^{14}\text{C}$  since the carbon cycle acts as an integrator over  
5  $^{14}\text{C}$  production rate variations. The centennial changes in the GRIP  $^{10}\text{Be}$  versions, however,  
6 are highly coherent and indicate a limited climate influence on  $^{10}\text{Be}$  on these timescales and  
7 the same holds true for the GISP2  $^{10}\text{Be}$  versions. This is in agreement with Adolphi et al.  
8 (2014) who showed that centennial GRIP  $^{10}\text{Be}$  variations are dominated by solar activity  
9 changes and indicate only little sensitivity to the assumed mode of  $^{10}\text{Be}$  deposition even over  
10 large deglacial climatic transitions. It should be noted that this statement solely refers to the  
11 filtered centennial  $^{10}\text{Be}$  variations investigated here. Other potential climatic influences on  
12  $^{10}\text{Be}$  such as changes in the stratosphere-troposphere exchange rates are, however, difficult to  
13 assess from climate proxy data and will thus, not be removed by our detrending technique.  
14 Thus, in the following we will focus on centennial ( $<500$  years) changes in  $^{10}\text{Be}$  and  $^{14}\text{C}$   
15 production rates to avoid systematic errors originating from uncertainties in the millennial  
16  $^{10}\text{Be}$  production rate history.

17 The left hand panels in figure 3 show the corresponding modelled  $\Delta^{14}\text{C}$  anomalies from the  
18 centennial  $^{10}\text{Be}$  variations indicated in figure 2 assuming a  $^{14}\text{C} : {}^{10}\text{Be}$  production rate ratio of  
19 1:1. As expected, similar to the  $^{10}\text{Be}$  records these variations are highly coherent. The right  
20 panels in figure 3 display histograms of the maximal  $\Delta^{14}\text{C}$  difference between the different  
21 production rate histories (i.e. the absolute  $\Delta^{14}\text{C}$  difference between the highest and the lowest  
22 modelled  $\Delta^{14}\text{C}$  version at each point in time). It can be seen that the different  $^{10}\text{Be}$  versions  
23 translate into a modelled  $\Delta^{14}\text{C}$  uncertainty of about  $\pm 3\text{ \%}$  ( $1\sigma$ ) for GRIP (figure 3 a, d) and  
24 GISP2 (figure 3 b, e). Similarly, the  $\Delta^{14}\text{C}$  anomalies modelled from GRIP and GISP2  $^{10}\text{Be}$   
25 agree within  $\pm 2.5\text{ \%}$  ( $1\sigma$ , figure 3 c, f).

26 As outlined in the introduction, the state and the dynamics of the carbon cycle impose an  
27 uncertainty on the  $^{10}\text{Be}$ - $^{14}\text{C}$  comparison that is difficult to quantify from the data itself  
28 (Köhler et al., 2006; Muscheler et al., 2004b). Figure 4 shows the results from the performed  
29 carbon cycle sensitivity experiments (see section 2.4.1, table 1). It can be seen that the  
30 millennial  $\Delta^{14}\text{C}$  variations are substantially altered by carbon cycle perturbations (figure 4 b).  
31 Changes in ocean ventilation (experiments S3 and S4) and well as air-sea gas exchange  
32 (experiments S1 and S2) can cause  $\Delta^{14}\text{C}$  anomalies larger than the amplitude of  $\Delta^{14}\text{C}$   
33 anomalies induced by  $^{14}\text{C}$  production rate changes only (control). However, as before, the

1 centennial  $\Delta^{14}\text{C}$  variations are considerably less affected by these perturbations (figure 4 c).  
2 The increase (decrease) of air-sea gas exchange or ocean ventilation does lead to a decrease  
3 (increase) in the amplitude of the modelled centennial  $\Delta^{14}\text{C}$  variations. However, these  
4 changes in amplitude are largely limited to about  $\pm 3\text{ \%}$  (figure 4, panel d) except for about  
5 200-300 years around the timing of the carbon cycle perturbation itself (figure 4, transitions 1  
6 and 2). Importantly, the phase of the centennial  $\Delta^{14}\text{C}$  variations is not affected by the imposed  
7 carbon cycle changes. Since the applied carbon cycle changes in our sensitivity experiments  
8 are likely unrealistically large for Holocene conditions (Köhler et al., 2006; Roth and Joos,  
9 2013), we conservatively assume a  $1\sigma$  uncertainty of  $\pm 3\text{ \%}$  (see figure 4, panel d, ‘steady  
10 state’) for the modelled  $\Delta^{14}\text{C}$  records due to carbon cycle effects.

11 Adding the uncertainties due to climate impacts on  $^{10}\text{Be}$  ( $\pm 3\text{ \%}$ ) and the carbon cycle ( $\pm 3\text{ \%}$ )  
12 in quadrature we thus, obtain an initial uncertainty estimate of about  $\pm 4.5\text{ \%}$  for the  
13 modelled  $\Delta^{14}\text{C}$  records.

14

### 15 **3.2 Sensitivity of the synchronization method to uncertainties in the $^{10}\text{Be}$ - $^{14}\text{C}$ 16 conversion**

17 In the following we will compare the centennial  $\Delta^{14}\text{C}$  (i.e., <500 years, separated by an FFT-  
18 based high-pass filter) anomalies reconstructed from tree rings (IntCal13) and ice cores  
19 (GRIP/GISP2  $^{10}\text{Be}$ -based) with respect to their timescale differences. The choice of a 500  
20 year high-pass filter results from the climate and carbon cycle related uncertainties shown in  
21 section 3.1 which increase on longer timescales. We use the statistical framework outlined in  
22 section 2.2 and assign an initial uncertainty of  $\pm 4.5\text{ \%}$  to the  $^{10}\text{Be}$ -based  $\Delta^{14}\text{C}$  records. The  
23 uncertainties for the tree-ring based  $\Delta^{14}\text{C}$  anomalies are taken from IntCal13 (Reimer et al.,  
24 2013). For this purpose we spliced the GISP2  $^{10}\text{Be}$  versions into the corresponding GRIP  
25  $^{10}\text{Be}$  versions to fill the gap in the GRIP record between 9,400 and 10,800 years BP and  
26 create a continuous record for the entire Holocene. Hence, in the following “GRIP” refers to  
27 this combination of GRIP and GISP2 data, while results for the GISP2 data are only shown  
28 for periods where they have not been used to fill the gap in the GRIP record.

29 Figure 5 displays the obtained probability distributions  $P_{s_{scaled}}(t_s)$  for each sliding window,  
30 centred on its mean age. The results are shown for all four GRIP  $^{10}\text{Be}$  versions (panel a), in  
31 comparison to results based on GISP2 data only (panel b), as well as for different assumed

1  $^{14}\text{C}$ : $^{10}\text{Be}$  production rate ratios (panel c). The different GRIP  $^{10}\text{Be}$  versions yield consistent  
2 estimates of the IntCal13-GICC05 timescale differences throughout the Holocene. The only  
3 marked difference occurs around the 8.2 ka BP event (Blockley et al., 2012). During this  
4 period the  $^{10}\text{Be}$  flux indicates a more rapid increase in the IntCal13-GICC05 timescale  
5 difference as compared to all other  $^{10}\text{Be}$  versions. As noted by Muscheler et al. (2004a) the  
6 accumulation rate anomaly associated to the climate oscillation around 8,200 years ago  
7 appears to lead to an ‘over correction’ of the  $^{10}\text{Be}$  deposition during flux calculation. This  
8 leads to a worse agreement between  $^{14}\text{C}$  and  $^{10}\text{Be}$  fluxes as compared to  $^{14}\text{C}$  and  $^{10}\text{Be}$   
9 concentrations (see figure 3 in Muscheler et al., 2004a). This is corroborated by the fact that  
10 results based on the “climate corrected”  $^{10}\text{Be}$  flux follow the probability estimates of  $^{10}\text{Be}$   
11 concentrations (figure 5a).

12 Comparing GRIP based results to GISP2 based estimates indicates consistent estimates of the  
13 timescale differences. The larger uncertainties of the GISP2 based results are due to the lower  
14 sampling resolution of the GISP2  $^{10}\text{Be}$  record (see equation 3).

15 Figure 5c shows the sensitivity of our results to the assumed  $^{14}\text{C}$ : $^{10}\text{Be}$  production rate ratio. It  
16 can be seen that the inferred timescale differences are relatively insensitive to the assumed  
17  $^{14}\text{C}$ : $^{10}\text{Be}$  ratio. However, the derived uncertainty of  $P_{s_{scaled}}(t_s)$  does increase with lower  
18  $^{14}\text{C}$ : $^{10}\text{Be}$  ratios. This can easily be understood by imagining a scaling of zero for the  $^{10}\text{Be}$ -  
19 based record which would result in an infinitely wide probability distribution.

20 In summary, our method of estimating the IntCal13-GICC05 timescale difference is i) largely  
21 robust for all versions of the GRIP  $^{10}\text{Be}$  record, ii) consistent for GRIP and GISP2  $^{10}\text{Be}$  data,  
22 and iii) independent of the assumed  $^{14}\text{C}$ : $^{10}\text{Be}$  production rate ratio. However, this analysis  
23 also shows that it is important to compare  $^{10}\text{Be}$  concentrations and fluxes to identify potential  
24 caveats as seen around the 8.2 ka BP event. Furthermore, while the estimate of the most  
25 likely timescale difference (i.e. the location of the maximum of  $P_{s_{scaled}}(t_s)$ ) may not be  
26 affected by the assumed  $^{14}\text{C}$ : $^{10}\text{Be}$  ratio, the uncertainty of this estimate is. Hence, in the  
27 following section we will derive a posterior estimate of the  $^{14}\text{C}$ : $^{10}\text{Be}$  ratio, as well as a refined  
28 uncertainty estimate of the  $^{10}\text{Be}$ -based  $\Delta^{14}\text{C}$  records.

29

### 3.3 Posterior estimate of the $^{14}\text{C} : {}^{10}\text{Be}$ production rate ratios and uncertainties

As shown in the previous section, our estimates of the most likely timescale difference between IntCal13 and GICC05 are largely independent of which  ${}^{10}\text{Be}$  record (GRIP/GISP2) and which version thereof (concentration, flux, climate corrections) is used, as well as which  $^{14}\text{C} : {}^{10}\text{Be}$  ratio is assumed. Hence, we calculated an initial GICC05-IntCal13 transfer function (section 2.5) and synchronized the tree ring based and  ${}^{10}\text{Be}$ -based  $\Delta^{14}\text{C}$  record. This enables us to compare the records with respect to the most likely  $^{14}\text{C} : {}^{10}\text{Be}$  ratio. In addition, we can derive a posterior estimate of the modelled  ${}^{10}\text{Be}$ -based  $\Delta^{14}\text{C}$  uncertainty.

After synchronization we can compare tree ring and  ${}^{10}\text{Be}$ -based  $\Delta^{14}\text{C}$  sample pairs assuming different  ${}^{10}\text{Be}$  scaling factors (i.e.  $^{14}\text{C} : {}^{10}\text{Be}$  ratios) between zero and two. The difference between tree ring and  ${}^{10}\text{Be}$ -based  $\Delta^{14}\text{C}$  sample pairs ( $\delta(t)$ ) is a function of the uncertainty of IntCal13 ( $\delta_{IC}(t)$ ) and the uncertainty of the  ${}^{10}\text{Be}$ -based records ( $\delta_{Be}(t)$ ) in the form that:

$$\delta(t) = \sqrt{\delta(t)_{IC}^2 + \delta(t)_{Be}^2} \quad (4)$$

Hence, we can rearrange equation 4 and use the quoted uncertainties of IntCal13 to derive  $\delta(t)_{Be}$ :

$$\delta(t)_{Be} = \sqrt{\delta(t)^2 - \delta(t)_{IC}^2}; \quad \delta(t) > \delta(t)_{IC} \quad (5)$$

$$\delta(t)_{Be} = 0; \quad \delta(t) \leq \delta(t)_{IC}$$

These uncertainties can be summarized to the rooted mean square error ( $\text{RMSE}_{10\text{Be}}$ ). This way we can obtain the optimal  ${}^{10}\text{Be}$  scaling factor (where the  $\text{RMSE}_{10\text{Be}}$  minimizes) and the associated uncertainty of the  ${}^{10}\text{Be}$ -based  $\Delta^{14}\text{C}$  records (the minimum of the  $\text{RMSE}_{10\text{Be}}$ ).

Figure 6 displays the results of this analysis indicating an optimal  ${}^{10}\text{Be}$  scaling factor of around 0.7. Assuming that the centennial  ${}^{10}\text{Be}$  and  $^{14}\text{C}$  production rate changes are mainly modulated through solar activity this low scaling factor would point to a strong polar bias of the GRIP GISP2  ${}^{10}\text{Be}$  records (see sections 1 and 2.4.2). However, when investigating the  $\Delta^{14}\text{C}$  time series it becomes apparent, that this low scaling leads to an underestimation of the amplitude of virtually all grand solar maxima and minima (i.e. large  $\Delta^{14}\text{C}$  anomalies) in the  ${}^{10}\text{Be}$ -based  $\Delta^{14}\text{C}$  record (figure 7, top). This bias is induced by the fact, that the  $\Delta^{14}\text{C}$  anomalies are normally distributed around 0 % leading to a majority of the  $\Delta^{14}\text{C}$  values lying close to zero dominating the  $\text{RMSE}_{10\text{Be}}$ . Hence, for these values a low scaling of the  ${}^{10}\text{Be}$ -based  $\Delta^{14}\text{C}$  records will simply act to reduce noise from the record and thus, reduce the  $\text{RMSE}_{10\text{Be}}$ .

1 To avoid this bias, we performed a binned regression analysis. We divided the tree ring and  
2  $^{10}\text{Be}$ -based  $\Delta^{14}\text{C}$  sample pairs into bins of 2.5 ‰ (defined based on the tree ring  $\Delta^{14}\text{C}$   
3 anomalies) and calculated the  $\text{RMSE}_{10\text{Be}}$  for each bin ( $\text{RMSE}_{10\text{Be\_bin}}$ ). These uncertainties for  
4 each bin can then be summarized to an overall  $\text{RMSE}_{10\text{Be}}$  as:

$$5 \quad \text{RMSE}_{10\text{Be}} = \sqrt{\text{RMSE}_{10\text{Be\_bin}}^2} \quad (6)$$

6 This binning leads to an equal weighting of small and large  $\Delta^{14}\text{C}$  anomalies in the  
7 comparison of the  $\Delta^{14}\text{C}$  records. It can be seen that this method indicates a larger  $^{14}\text{C}:\text{Be}^{10}$   
8 ratio of about 1.1 (figure 8) and avoids the systematic underestimation of large amplitude  
9  $\Delta^{14}\text{C}$  anomalies (figure 7, bottom). Depending on the production rate model used, this scaling  
10 indicates a weak (Masarik and Beer, 2009, 1999) or no (Kovaltsov et al., 2012; Kovaltsov and  
11 Usoskin, 2010) polar bias in the Greenland  $^{10}\text{Be}$  records. In addition, it can be seen that the  
12 minimum of the  $\text{RMSE}_{10\text{Be}}$  becomes larger than without binning, indicating an uncertainty of  
13 about 4 ‰ for the  $^{10}\text{Be}$ -based  $\Delta^{14}\text{C}$  records. This is due to the above described effect, that the  
14 noise is not artificially suppressed and can be seen by comparing the decadal scale peaks in the  
15 top and bottom panels of figure 7. The larger  $^{10}\text{Be}$  scaling factor makes the  $^{10}\text{Be}$  record  
16 appear noisier. However, firstly, this noise may represent remaining influences of ‘system  
17 effects’ on ice core  $^{10}\text{Be}$  records and hence, represent an uncertainty that has to be taken into  
18 account. Secondly, it should be kept in mind that IntCal13 is a stack of multiple  $^{14}\text{C}$  datasets  
19 which will inevitably result in smoothing. This smoothing may also reduce the amplitude of  
20 ‘real’  $\Delta^{14}\text{C}$  variations instead of merely reducing noise, since the differences between the  
21 underlying raw data sets of IntCal13 are potentially in part systematic (Stuiver et al.,  
22 1998; Adolphi et al., 2013).

23 In conclusion we use a  $^{14}\text{C}:\text{Be}^{10}$  ratio of 1.1:1 and an uncertainty of 4 ‰ for the modelled  
24  $\Delta^{14}\text{C}$  record to derive a final IntCal13-GICC05 transfer function in the next section. It should  
25 be noted that this uncertainty estimate is only valid for the centennial (<500 year) variations  
26 studied here.

27

### 28 **3.4 IntCal13-GICC05 transfer function**

29 Using the estimated  $^{14}\text{C}:\text{Be}^{10}$  ratio of 1.1 and a  $^{10}\text{Be}$ -based  $\Delta^{14}\text{C}$  error of  $\pm 4$  ‰ ( $\pm 1\sigma$ ) (see  
30 previous section) we recalculated the ‘wiggle-match’ probability distributions ( $P_{s_{\text{scaled}}}(t_s)$ ,

1 equation 3) for the IntCal13-GICC05 timescale difference (figure 9, grey shading). For these  
2 calculations we used the mean of all GRIP<sup>10</sup>Be-based  $\Delta^{14}\text{C}$  versions (concentration, flux,  
3 climate corrections) and filled the gap between 9,400 and 10,800 yrBP using the GISP2 data.  
4 Based on these probability distributions we modelled the IntCal13-GICC05 transfer function  
5 as described in section 2.5. The resulting transfer function (figure 9 solid lines) averages out  
6 some short-term fluctuations in the timescale difference compared to the initial ‘wiggle-  
7 match’ probability distributions. As described in section 2.5 this is due to the used window  
8 length of 1,000 years to determine  $P_{s_{scaled}}(t_s)$  at each point in time, preventing an  
9 independent assessment of faster changes in the timescale difference. Nevertheless, the  
10 estimated uncertainties of the timescale transfer function (thin black lines in figure 9)  
11 encompass the uncertainties of the ‘wiggle-match’ probability distribution at each point in  
12 time.

13 Figure 10 shows three examples of GRIP <sup>10</sup>Be based  $\Delta^{14}\text{C}$  anomalies before (grey) and after  
14 (black) synchronization to IntCal13 (red). The examples encompass (i) a period of relatively  
15 low  $\Delta^{14}\text{C}$  variability ( $\pm 5\text{--}7\%$ ) but good agreement between GRIP and IntCal13 (figure 10, a),  
16 (ii) a period of large  $\Delta^{14}\text{C}$  variability ( $\pm 10\%$ ) but less good agreement between GRIP and  
17 IntCal13 (figure 10, b), and (iii) a section of large  $\Delta^{14}\text{C}$  ( $\pm 10\%$ ) variability and excellent  
18 agreement between GRIP and IntCal13 (figure 10, c). It can be seen, that in all cases the fit  
19 between GRIP and IntCal13 is improved when applying the proposed GICC05-IntCal13  
20 transfer function. However, figure 10 (b) also shows, that short periods of disagreement (i.e.,  
21 around 7,250 – 7,500 years BP) may remain, as they cannot be reliably resolved by our  
22 method which matches 1,000 year-long sections. It should, however, be noted that matching  
23 these short sections would (i) represent a serious violation of the GICC05 counting error  
24 which is minimal over these short periods of time ( $\pm 6$  years at  $2\sigma$  between 7,250 – 7,500  
25 years BP), and (ii) not account for the possibility that <sup>10</sup>Be and <sup>14</sup>C may simply not agree due  
26 to the caveats outlined in the introduction. Furthermore, the applied shift of GICC05 in figure  
27 10 (b) leads to an improved agreement between <sup>14</sup>C and <sup>10</sup>Be after and prior to 7,250 and  
28 7,500, respectively. Hence, we consider it unlikely that for this short period of time the  
29 timescale difference deviates significantly from the estimate for the entire window.

30

1    **4 Discussion**

2    Figure 11 shows the obtained estimate of the IntCal13-GICC05 timescale difference in  
3    comparison to the results obtained by using the method of Muscheler et al. (2014a, re-run  
4    with a 1,000 year window length) and age markers that have been independently anchored on  
5    both timescales.

6    Our results are fully consistent with the results obtained by Muscheler et al. (2014a). While  
7    this is expected to some extent, as our study and the work by Muscheler et al. (2014a) are  
8    based on the same data, it shows that the statistical approach used here leads to similar results  
9    as the Monte-Carlo lag-correlation analysis but is computationally much less expensive.  
10   Furthermore, as shown in figure 5, we obtain similar results when using the GISP2  $^{10}\text{Be}$   
11   instead of the GRIP  $^{10}\text{Be}$  record lending additional support to the robustness of our results.  
12   The additional modelling of the transfer function employed here (sections 2.5 and 3.4) leads  
13   to a smoother development of the timescale difference which is more realistically reflecting  
14   limitations of the method imposed by the window size of the  $^{14}\text{C}$ - $^{10}\text{Be}$  comparison. The  
15   difference between the timescale transfer functions around 8,200 years BP is induced by the  
16   fact that Muscheler et al. (2014a) based their calculations on  $^{10}\text{Be}$  fluxes which are influenced  
17   by accumulation rate changes around this time as discussed in section 3.2 and in Muscheler et  
18   al. (2004a).

19   The largest difference between the results presented here and by those of Muscheler et al.  
20   (2014a) is seen in the derived error estimates. We obtain strongly reduced uncertainties for  
21   the estimated timescale differences. This is likely due to the fact, that Muscheler et al.  
22   (2014a) used a comparably ad-hoc and highly conservative method to derive their  
23   uncertainties. By taking the distribution of the mean  $r^2$ -values of all iterations Muscheler et al.  
24   (2014a) do not include the results of the Monte-Carlo analysis of the “Best Fits” in their error  
25   estimate. Thus,  $^{14}\text{C}$ - $^{10}\text{Be}$  matches that may not be the most likely solution in any of the  
26   iterations become included in the uncertainty envelope. In comparison, the statistics  
27   employed here allow a direct analytical assessment of the synchronization uncertainties.  
28   Hence, while our uncertainty estimates are significantly smaller, we consider them more  
29   robust. Theoretically, systematic errors from undetected biases in the  $^{10}\text{Be}$  record could lead  
30   to erroneous results. However, the results shown in section 3.2 demonstrate the consistency  
31   of GRIP and GISP2  $^{10}\text{Be}$ -based calculations as well as for different climate corrections and  
32   do, thus, not indicate such biases (see figure 5). In conclusion, while largely consistent, we

1 regard the method employed here a significant improvement to the approach by Muscheler et  
2 al. (2014a).

3 Comparing our results to independent estimates of IntCal13-GICC05 timescale differences  
4 further supports our analyses (figure 11, symbols). Two major solar proton events (“775 and  
5 994 AD events”) leaving well defined spikes in the  $^{14}\text{C}$  content of dendrochronologically  
6 dated trees (Miyake et al., 2013; Miyake et al., 2012; Gütter et al., 2015) as well as in  
7 Greenland ice core  $^{10}\text{Be}$  records (Mekhaldi et al., 2015; Sigl et al., 2015) indicate an IntCal13-  
8 GICC05 timescale difference of  $-7 \pm 2$  ( $2\sigma$ ) years for both events (Sigl et al., 2015).  
9 Consistent with these findings, we obtain IntCal13-GICC05 differences of  $-4 \pm 4$  and  $-6 \pm 5$   
10 years ( $2\sigma$ ) for the 994 and 775 AD event, respectively. It should be noted that these annual  
11 radionuclide excursions are not present in the data used here, which is of lower resolution,  
12 and are hence, independent estimates of the timescale difference.

13 Based on tephra findings in the GRIP ice core (Barbante et al., 2013) the historically dated  
14 AD 79 eruption of Vesuvius has been used as a reference point in the GICC05 chronology  
15 (Vinther et al., 2006). However, our results indicate a timescale offset of  $-11 \pm 6$  ( $2\sigma$ ) years at  
16 AD 79 (1871 years BP, see figure 11). Assuming that the tree-ring chronologies are correct at  
17 this time, this would imply an age of AD  $90 \pm 6$  for the GRIP tephra layer – incompatible  
18 with an attribution to the age of the Vesuvius eruption within  $2\sigma$ . This result is in agreement  
19 with the analysis by Sigl et al. (2015) who recently counted annual layers in the NEEM and  
20 NEEM-2011-S1 ice cores and dated this marker horizon to AD 87 and 89, respectively.

21 The age of the Minoan eruption of Santorini has long been debated and the presence of an  
22 unequivocally attributable signal in the ice core records has been questioned (Pearce et al.,  
23 2004; Hammer et al., 1987; Hammer et al., 2003; Friedrich et al., 2006). The GICC05 age of  
24  $3591 \pm 5$  BP of an identified tephra horizon is incompatible with the radiocarbon based age of  
25  $3563 \pm 14$  calBP of the Santorini eruption ( $\Delta = -28 \pm 15$  yrs). Our results indicate a  
26 chronology difference of  $-20 \pm 5$  years around this time, reconciling the two aforementioned  
27 ages (see figure 11, open diamond). Hence, at least from a chronological point of view, it  
28 cannot be ruled out that the ice core tephra may be ascribable to the Santorini eruption  
29 (Muscheler, 2009).

30 Volcanic glass shards from the Saksunarvatn ash have been found in the GRIP ice core  
31 (Grönvold et al., 1995), as well as in multiple marine, lacustrine and terrestrial sites, of which  
32 the Lake Kråkenes record provides the highest resolution radiocarbon based age for the

1 deposit (Lohne et al., 2013). The dating difference of  $-86 \pm 35$  years between the radiocarbon  
2 based age by Lohne et al. ( $10,210 \pm 35$  calBP,  $\pm 1\sigma$ ) and the GICC05 age (10,296 BP, Abbott  
3 and Davies, 2012) of the Saksunarvatn ash is consistent with our estimated timescale  
4 difference of  $-66 \pm 10$  years during this time interval.

5 In summary, our results are consistent within uncertainties with all independent age markers  
6 that link the GICC05 and IntCal13 timescales over the Holocene.

7 Figure 12 displays the inferred IntCal13-GICC05 timescale differences in comparison to the  
8 GICC05 maximum counting error (Rasmussen et al., 2006; Vinther et al., 2006). Assuming  
9 that the tree-ring chronologies underlying IntCal13 are accurate throughout the Holocene our  
10 results imply an underestimation of the absolute dating uncertainty of GICC05 for large parts  
11 of the Holocene. Furthermore, it can be seen that the counting error appears to be systematic,  
12 in that most uncertain years (counted as  $0.5 \pm 0.5$  years, Rasmussen et al., 2006) have indeed  
13 not been true calendar years during the Holocene (i.e., a systematic over-counting of years).  
14 Nevertheless, when comparing the rate of change of the inferred IntCal13-GICC05 timescale  
15 difference to the rate of change of the maximum counting error (i.e. the relative maximum  
16 counting error) it can be seen that – even though systematic – the identification of uncertain  
17 years in the ice core records is accurate. Except for the most recent 2,000 years where  
18 (potentially erroneous) fix-points like the Vesuvius eruption are used to constrain GICC05  
19 the relative layer counting uncertainty appears to be an accurate uncertainty estimate. This  
20 can be seen in figure 12 (lower panel) which indicates that the rate of change of the GICC05  
21 maximum counting error is consistent within error with the rate of change of the IntCal13-  
22 GICC05 timescale difference prior to 2,000 years BP. This is important to note as it generally  
23 supports the GICC05 layer counting methodology and uncertainty which forms the basis of  
24 GICC05 back to 60,000 years BP (Svensson et al., 2008), even though the systematic nature  
25 of the derived timescale differences challenges the use of the maximum counting error as a  
26 nearly Gaussian distributed  $2\sigma$  uncertainty during the Holocene (Andersen et al., 2006). It  
27 can, however, not be assumed that the counting error continues to be systematic beyond this  
28 period, since the parameters used for layer identification as well as the sources of uncertainty  
29 (e.g. melt layers) differ back in time under changed climatic conditions (Rasmussen et al.,  
30 2006).

31 Alternatively, uncertainties in the dendrochronologies underlying IntCal13 could contribute  
32 to the growing discrepancy between IntCal13 and GICC05 over the Holocene. This appears,  
33 however, unlikely since the tree-ring chronologies have been cross-dated back to 7,272 calBP

1 to the Irish Oak Chronology (Pilcher et al., 1984) and back to 9,741 calBP using  
2 independently constructed German Oak Chronologies (Friedrich et al., 2004; Spurk et al.,  
3 2002). Furthermore, the gradual development of the timescale difference appears consistent  
4 with a counting uncertainty, while a dendrochronological mismatch could be expected to  
5 cause sudden ‘jumps’ in the timescale difference. However, consistently missing tree rings in  
6 both German oak chronologies for the period older than 7,272 calBP could theoretically  
7 contribute to the growing timescale difference.

8

## 9 **5 Conclusions**

10 We employed a novel approach to infer timescale differences between two of the most widely  
11 used chronologies in Holocene paleoclimatology, the radiocarbon (IntCal13, Reimer et al.,  
12 2013) and Greenland ice core (GICC05, Svensson et al., 2008) timescales. Our results are  
13 largely consistent with the results of Muscheler et al. (2014a) but yield significantly smaller  
14 and more robust uncertainty estimates. The inferred timescale differences are consistent with  
15 independent tie-points obtained from volcanic tephras and solar proton events. However, in  
16 agreement with Sigl et al. (2015) our analyses indicate that the attribution of an ice core  
17 tephra to the AD 79 eruption of Vesuvius (Barbante et al., 2013) may be erroneous which  
18 leads to a propagating ice core dating bias that affects large parts of the Holocene.  
19 Nevertheless, the identification of uncertain years in the ice core during the Holocene is  
20 otherwise generally accurate as expressed in the relative counting error (figure 12 lower  
21 panel). This is important to note as it, in principle, supports the layer counting method and  
22 uncertainty estimates also beyond the period investigated here. Furthermore, it should be  
23 noted that these conclusions are based on the assumption that the tree-ring time scale is  
24 accurate.

25 Independent of the accuracy of either of the two chronologies we provided a high precision  
26 transfer function between the radiocarbon and Greenland ice core timescales. This allows  
27 radiocarbon dated and ice core paleoclimate records to be compared at high chronological  
28 precision which will improve studies of leads and lags within the climate system throughout  
29 the Holocene (Bronk Ramsey et al., 2014). Furthermore, the methodology outlined here can  
30 be applied to link high resolution  $^{14}\text{C}$  records such as floating tree-ring chronologies to ice  
31 core time scales and thus, aid in testing and improving the glacial radiocarbon dating  
32 calibration curve.

1 The proposed GICC05-IntCal13 transfer function shown in figure 9, 11 and 12 is available as  
2 a supplementary file to this paper and on NOAA.

3

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8 accumulation rates.

9

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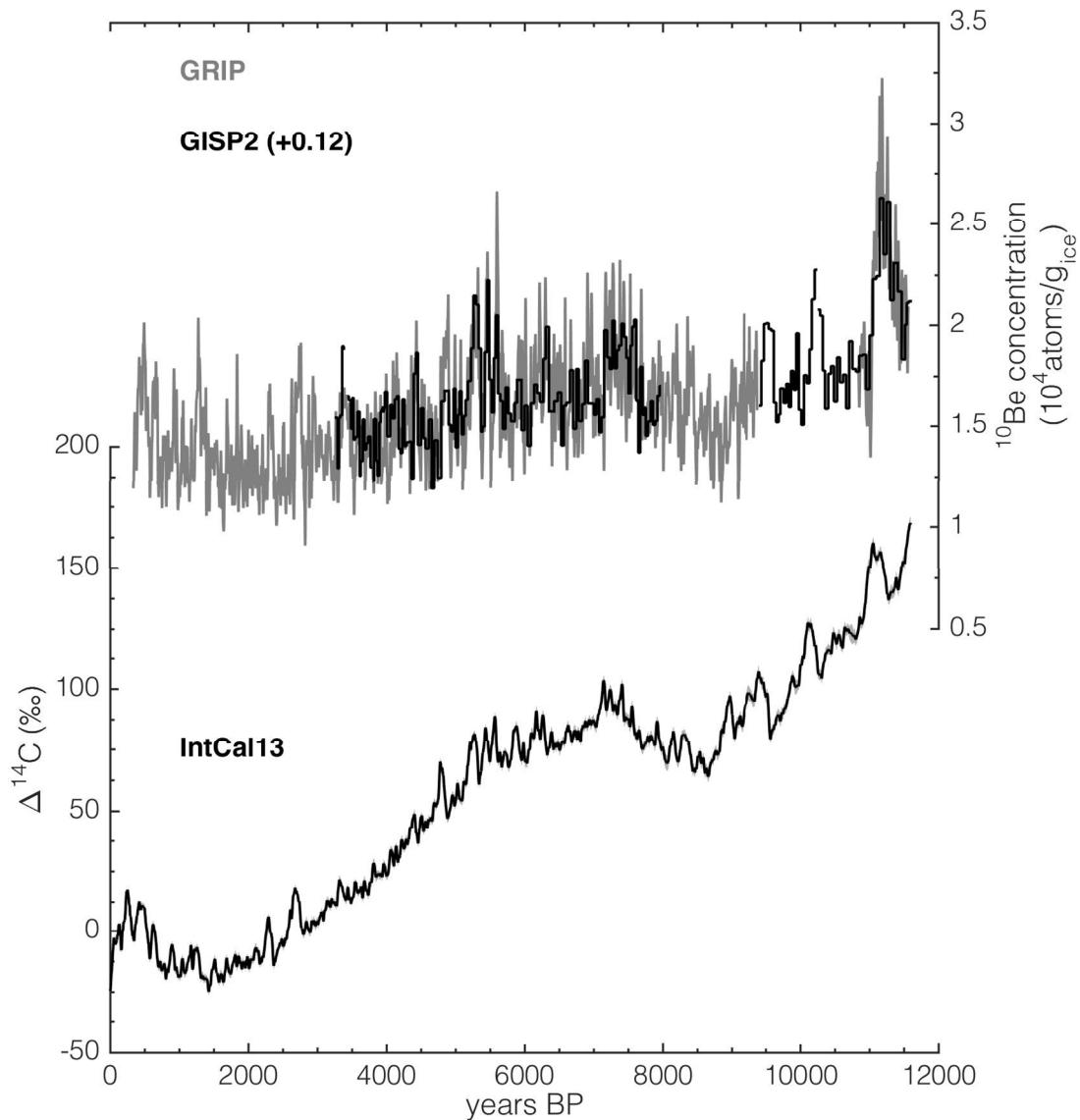
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1 Table 1. Performed carbon cycle sensitivity experiments. All percentage values refer to the  
2 control simulation under pre-industrial conditions.

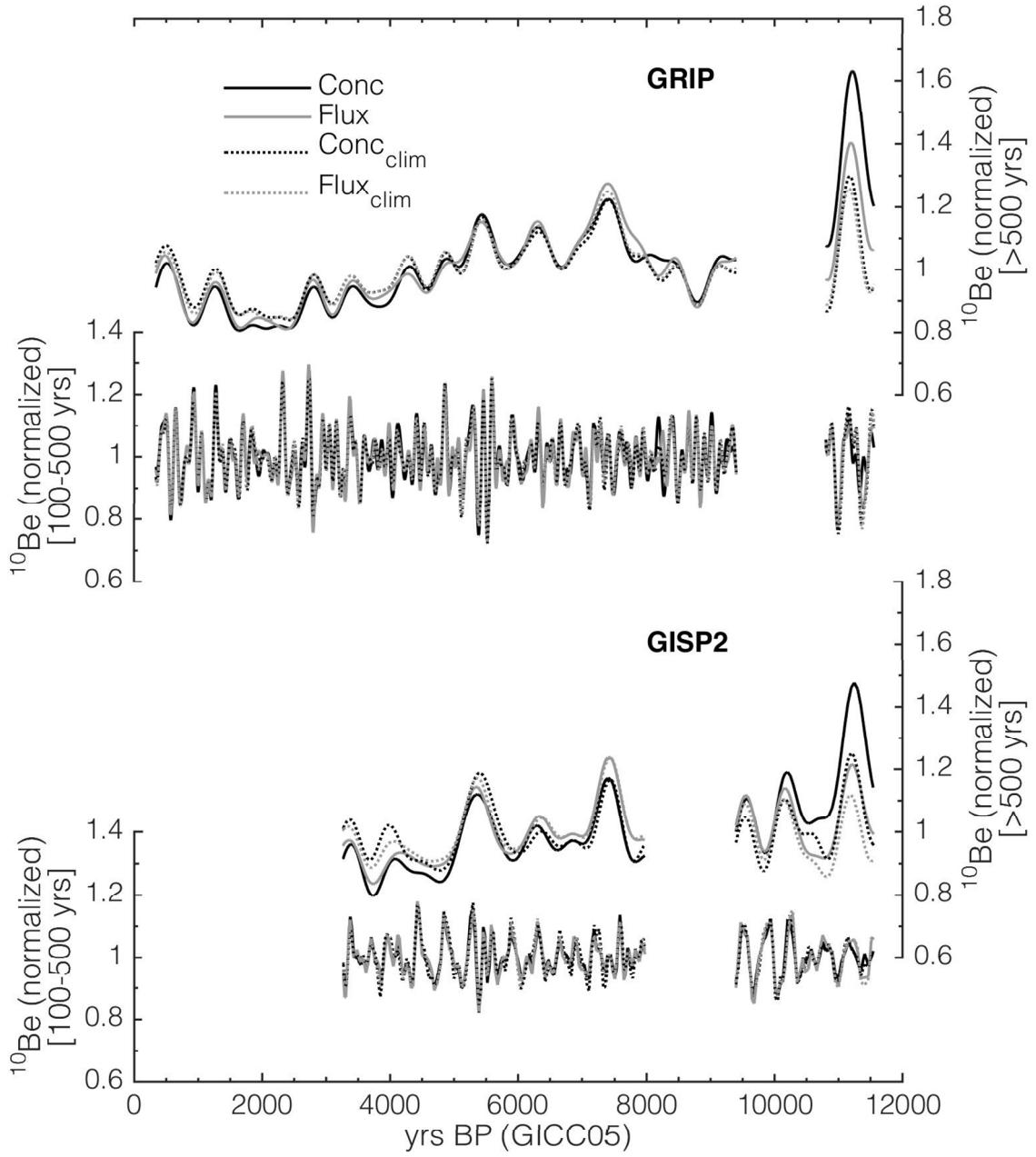
	Control	S1	S2	S3	S4
Air/Sea Exchange	100 %	150 %	50 %	100 %	100 %
Ocean ventilation	100 %	100 %	100 %	150 %	50 %

3  
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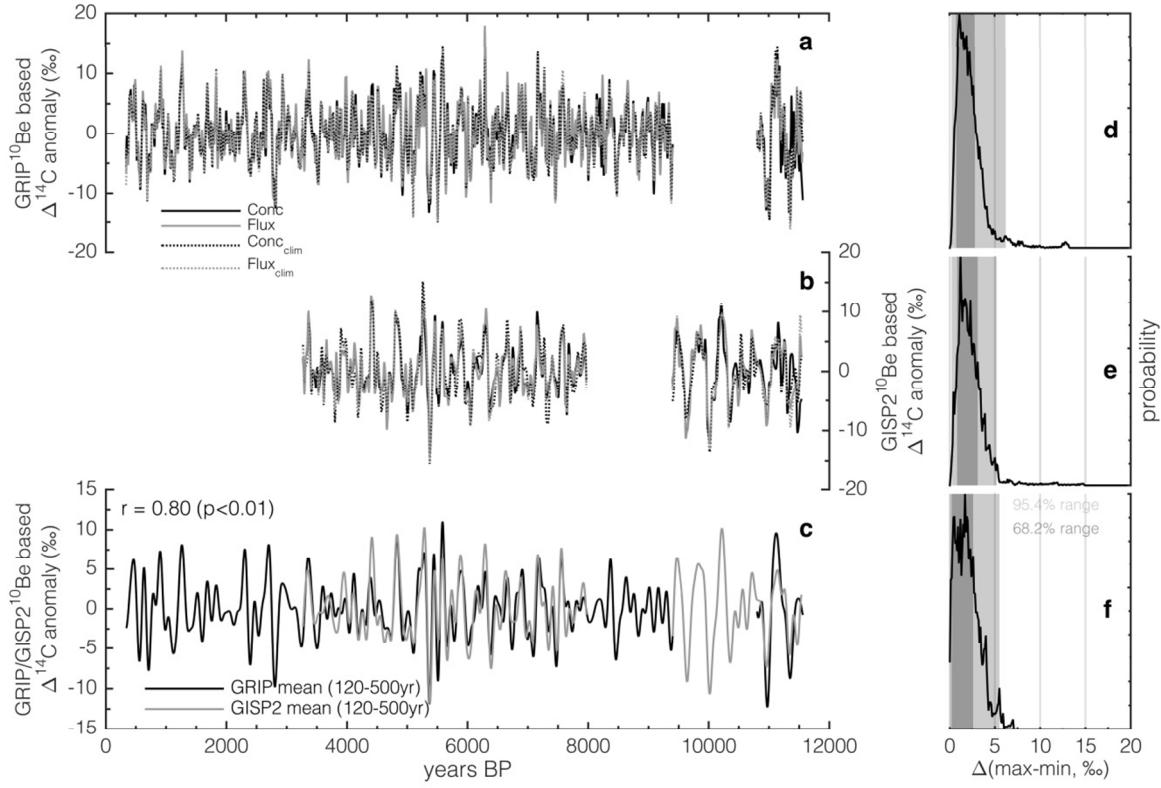


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2 Figure 1: *Top*: GRIP (grey, Vonmoos et al., 2006) and GISP2 (black, Finkel and Nishiizumi,  
3 1997) Holocene  $^{10}\text{Be}$  concentrations. The GRIP  $^{10}\text{Be}$  record is smoothed by a 61-pt binomial  
4 filter (see Vonmoos et al., 2006). The GISP2  $^{10}\text{Be}$  record has been shifted by  
5  $+0.12 \times 10^4 \text{ atoms/g}$  to correct for a difference in the mean of the GRIP and GISP2  $^{10}\text{Be}$   
6 records. *Bottom*: Atmospheric  $\Delta^{14}\text{C}$  as reconstructed from tree rings (Reimer et al., 2013 and  
7 references therein).  
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9  
10

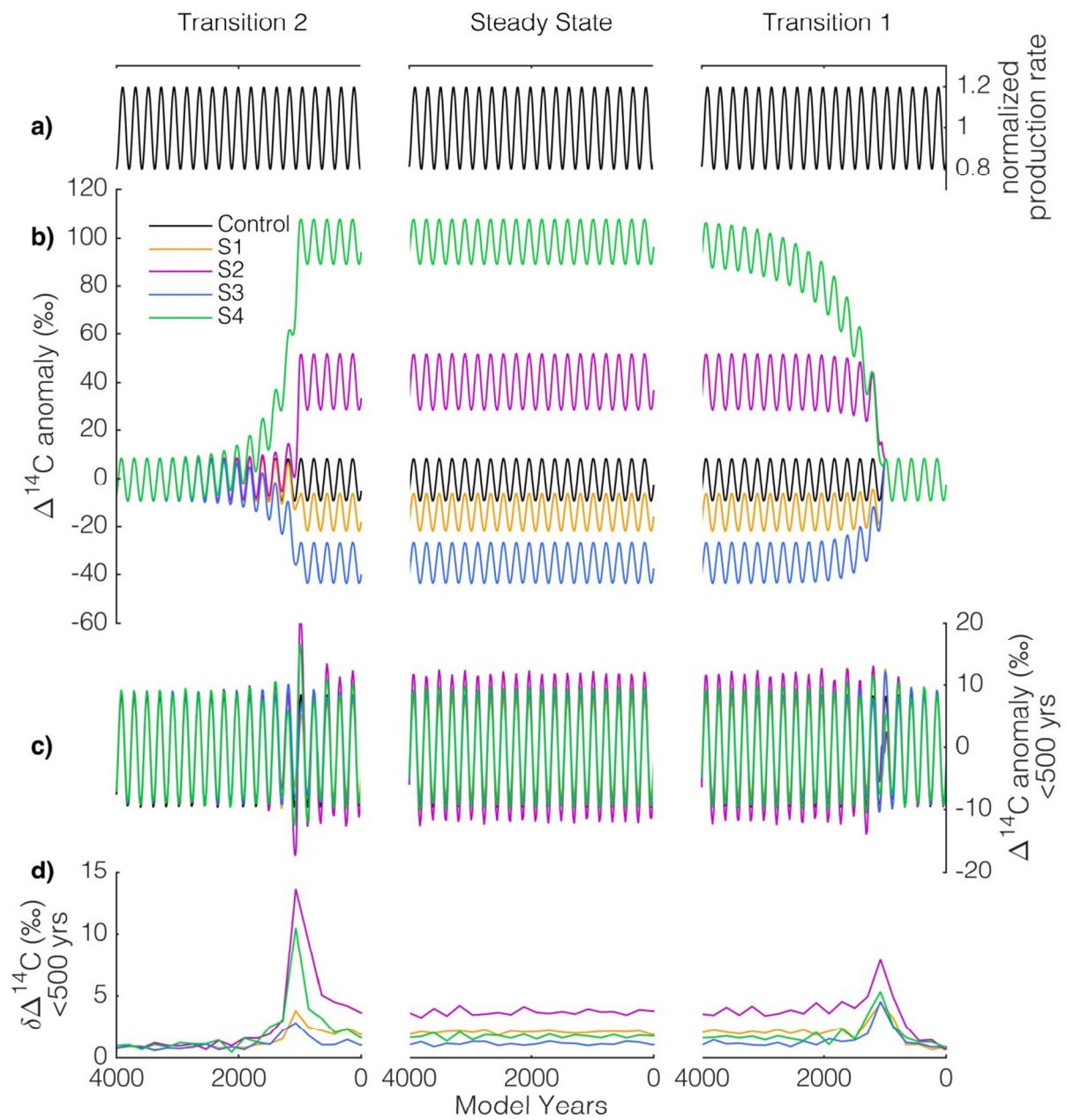




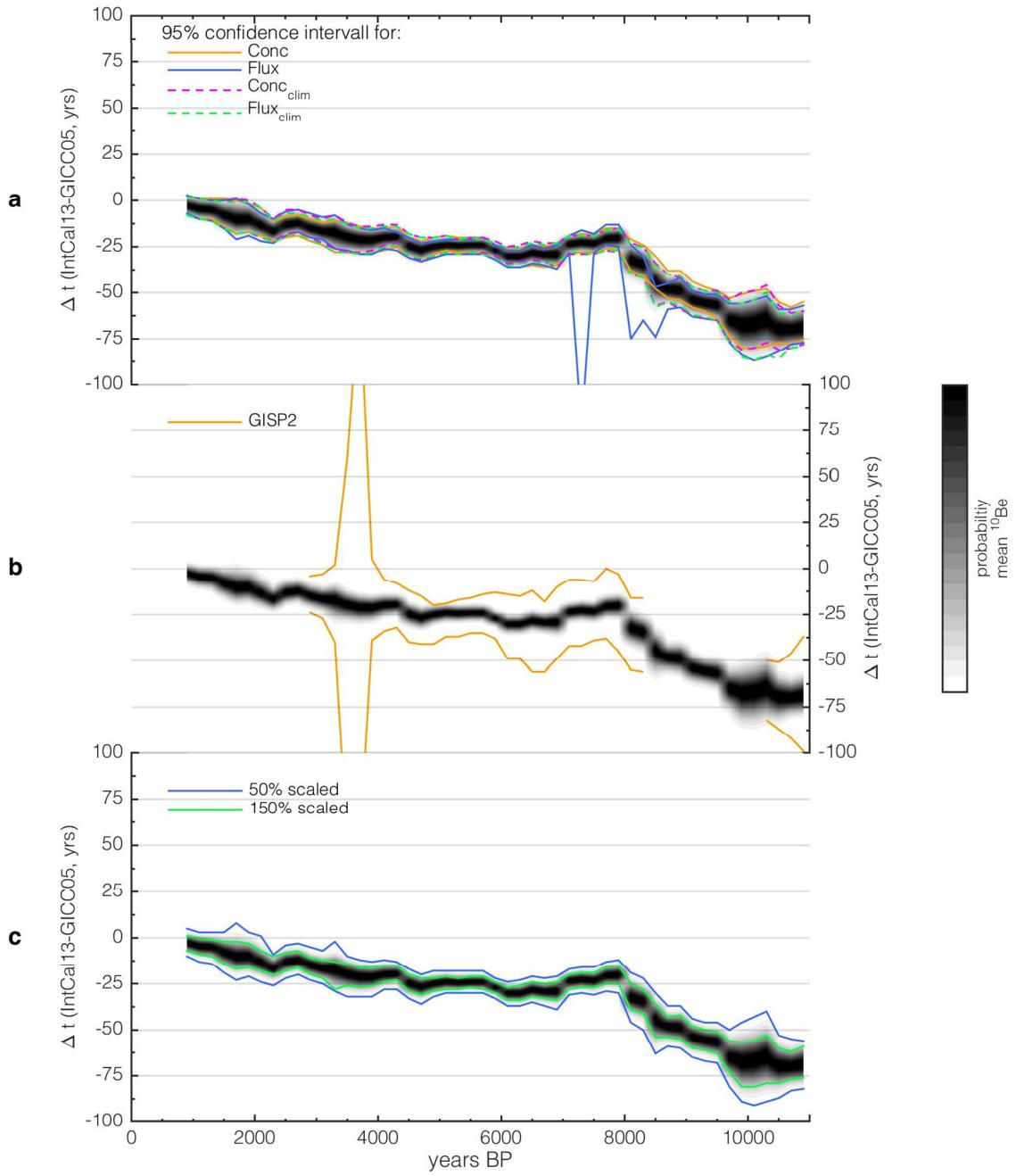
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 2 Figure 2: Comparison of  $^{10}\text{Be}$  fluxes and concentrations over the Holocene. Solid black and  
 3 grey curves denote  $^{10}\text{Be}$  concentrations and fluxes, respectively. Dotted lines refer to the  
 4 “climate corrected” (see text) versions of concentrations and fluxes with similar colour  
 5 coding as solid lines. The top two panels show GRIP  $^{10}\text{Be}$  for variations on time scales longer  
 6 (top) than 500 years, and for wavelengths between 100-500 years (below). The 100 year cut-  
 7 off has been applied for clarity of the figure. The bottom two panels show GISP2  $^{10}\text{Be}$  for the  
 8 same wavelengths as for GRIP.



1  
 2 Figure 3. Centennial (<500 years)  $\Delta^{14}\text{C}$  variations modelled from GRIP and GISP2  $^{10}\text{Be}$  data.  
 3 Panels a and b show the modelled  $\Delta^{14}\text{C}$  variations from  $^{10}\text{Be}$  concentrations (solid black),  
 4 fluxes (solid grey), “climate corrected” concentrations (dotted black), and “climate corrected”  
 5 fluxes (dotted grey) for the GRIP (a) and GISP2 (b)  $^{10}\text{Be}$  records. Panels d and e on the right  
 6 side depict the probability density functions for the maximum  $\Delta^{14}\text{C}$  difference between  
 7 curves shown in panels a and b, respectively. Panel c shows the mean of all GRIP (black) and  
 8 GISP2 (grey)  $^{10}\text{Be}$  based  $\Delta^{14}\text{C}$  anomalies shown in panels a and b, respectively. Panel f  
 9 shows the corresponding probability density function of their maximum  $\Delta^{14}\text{C}$  differences. For  
 10 this comparison both ice core records have been band-pass filtered [120 – 500 years] to  
 11 minimize inconsistencies arising from their different sampling resolution. The correlation  
 12 between the GRIP and GISP2 records is given in panel c together with its p-value.  
 13



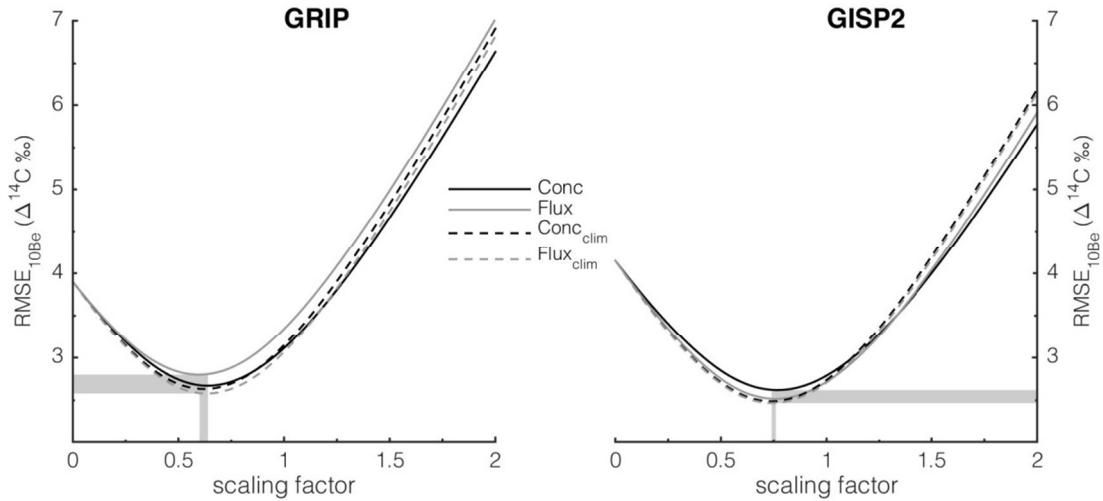
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 2 Figure 4. Carbon cycle sensitivity experiments. a) Normalized  $^{14}\text{C}$  production rate input to  
 3 the model. b) Modelled  $\Delta^{14}\text{C}$  anomaly. c) Centennial ( $<500$  year) anomalies of modelled  
 4  $\Delta^{14}\text{C}$  shown in panel b. d) differences in the centennial  $\Delta^{14}\text{C}$  variations (panel c) from the  
 5 control run. All model runs and panels are shown for the transition from preindustrial to  
 6 perturbed conditions (transition 1, right), steady state of the perturbed conditions (steady  
 7 state, middle), and the transition back to preindustrial carbon cycle conditions (transition 2,  
 8 left). See also section 2.4.1.



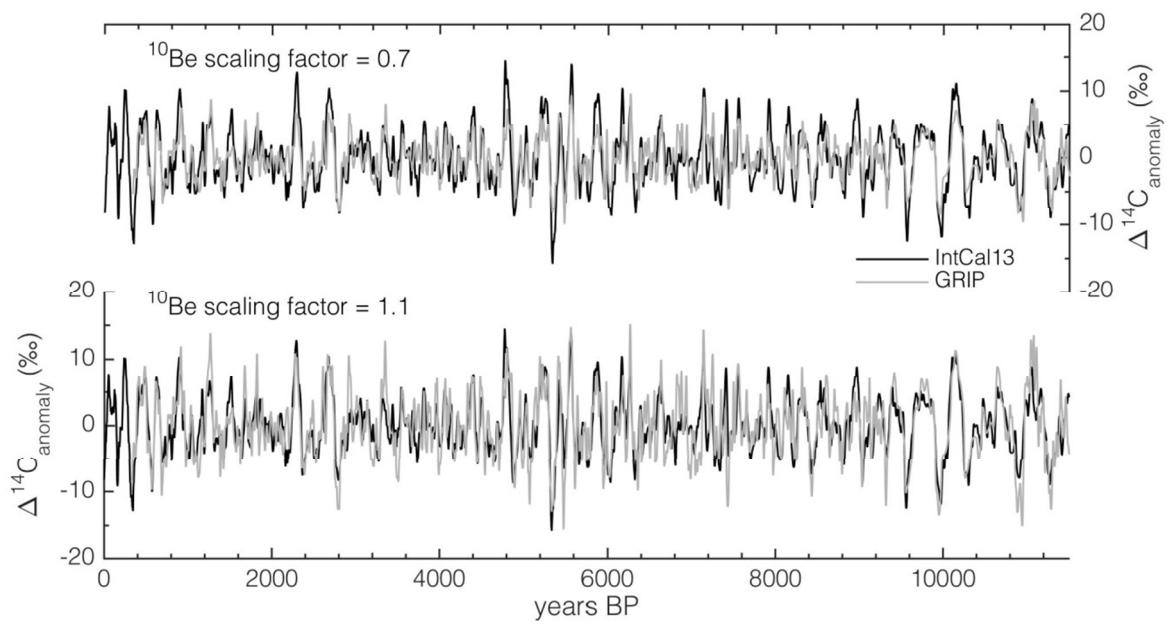
1 Figure 5. Probability distributions for IntCal13-GICC05 timescale differences ( $P_{s_{scaled}}(t_s)$ ,  
2 see section 2.1) for each 1,000-year window based on the mean of GRIP  $^{10}\text{Be}$  concentrations,  
3 fluxes, and their climate corrected versions (grey-scale patches in all panels). The gap in the  
4 GRIP  $^{10}\text{Be}$  record between 9,400 and 10,800 BP has been filled with data from the GISP2 ice  
5 core. Each probability distribution is centred on the mean age of the investigated window. a)  
6 Comparison to 95% probability intervals based on GRIP  $^{10}\text{Be}$  concentrations (solid orange),  
7 fluxes (solid blue) and their “climate corrected versions (dashed pink and green lines). b)  
8 Comparison to 95% confidence intervals based on the mean of GISP2  $^{10}\text{Be}$  concentrations,  
9

1 fluxes, and their climate corrected versions. Results for GISP2 are only shown for periods  
2 where it has not been used to fill the gap in the GRIP record. c) Comparison to results based  
3 on a different scaling (factors of 0.5 and 1.5 shown as blue and green lines, respectively) of  
4 the GRIP  $^{10}\text{Be}$  record.

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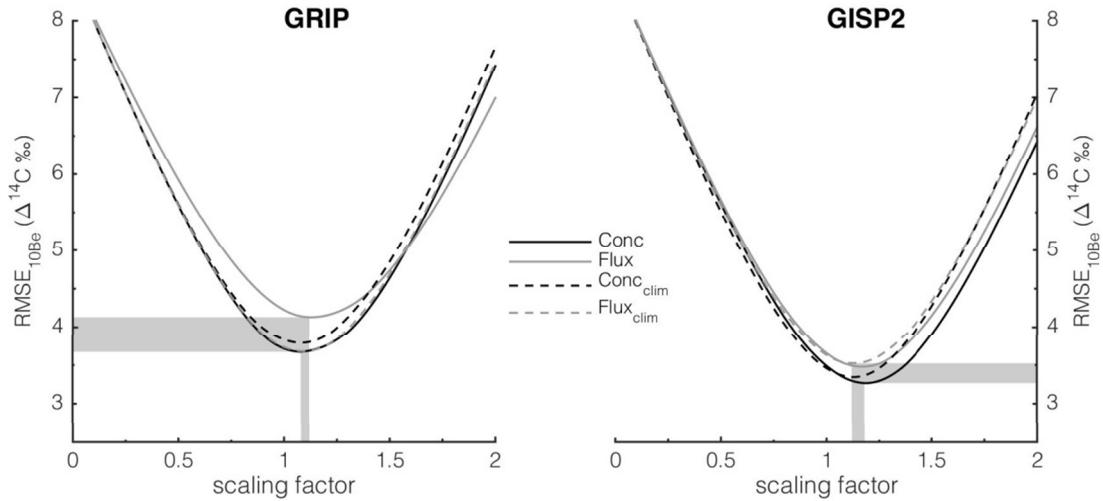
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2 Figure 6. Rooted mean square error ( $\text{RMSE}_{10\text{Be}}$ , see text) of synchronized centennial IntCal13  
3 and  $^{10}\text{Be}$ -based  $\Delta^{14}\text{C}$  variations as a function of different  $^{10}\text{Be}$ -scaling factors ( $^{14}\text{C} : {}^{10}\text{Be}$   
4 ratios). Results for the different versions of the GRIP  $^{10}\text{Be}$  record are shown on the left, while  
5 GISP2  $^{10}\text{Be}$ -based results are shown on the right.  
6



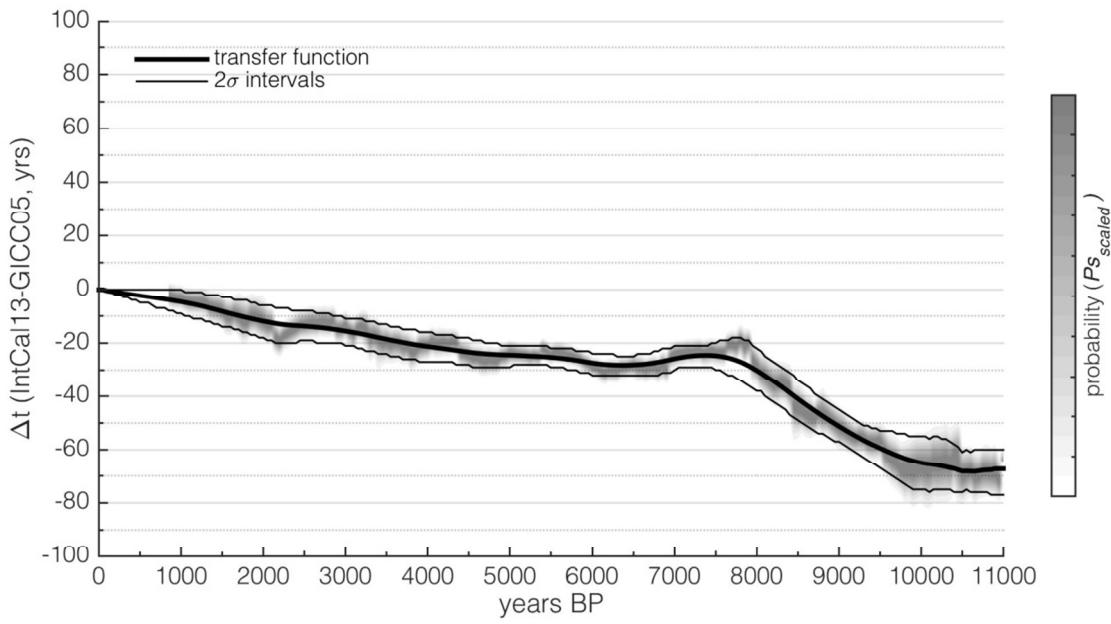
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2 Figure 7. Comparison of synchronized tree-ring (black) and ice core (grey) based  $\Delta^{14}\text{C}$   
 3 anomalies for  $^{14}\text{C}:^{10}\text{Be}$  ratios of 0.7 (top) and 1.1 (bottom).

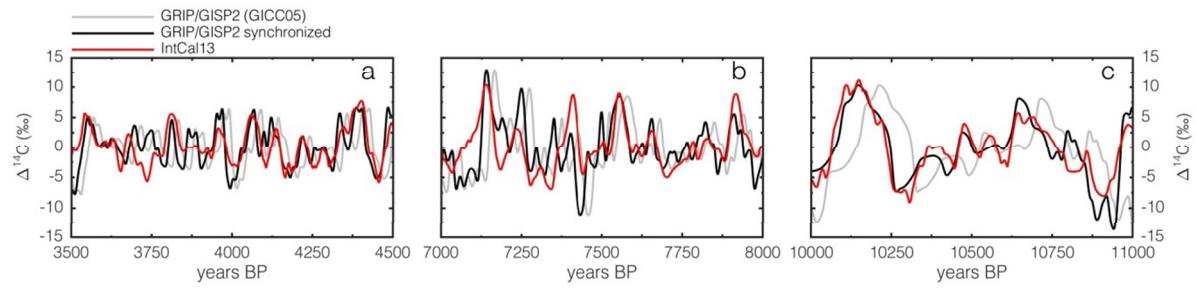
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2 Figure 8. Rooted mean square error (RMSE<sub>10Be</sub>) of IntCal13  $\Delta^{14}\text{C}$  and  $^{10}\text{Be}$  based  $\Delta^{14}\text{C}$   
3 records from GRIP (left) and GISP2 (right) for different scalings of the  $^{10}\text{Be}$  based data after  
4 synchronization. The RMSE<sub>10Be</sub> has been calculated for binned data (bin size = 2.5 ‰, see  
5 text) taking IntCal  $\Delta^{14}\text{C}$  errors into account.  
6

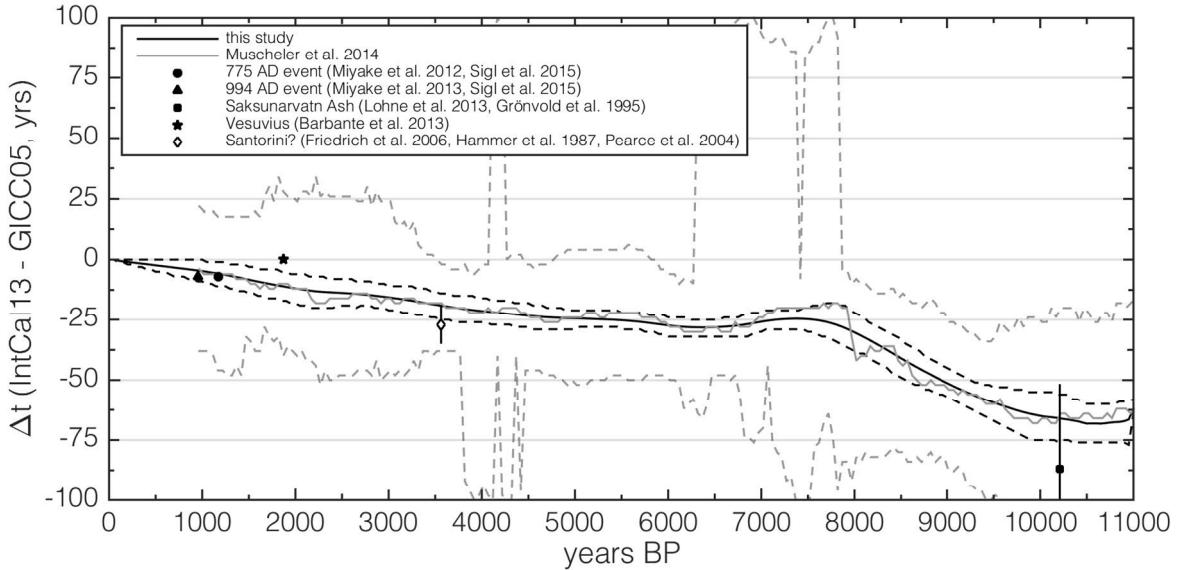


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2 Figure 9. IntCal13-GICC05 age transfer function (thick black line) and its 2 $\sigma$  confidence  
3 intervals (thin black lines) based on the probability distributions ( $P_{sscaled}(t_s)$ , grey shading)  
4 obtained from comparing the GRIP  $^{10}\text{Be}$ -based  $\Delta^{14}\text{C}$  (mean of concentration, flux and climate  
5 corrections) and IntCal13  $\Delta^{14}\text{C}$  records.  
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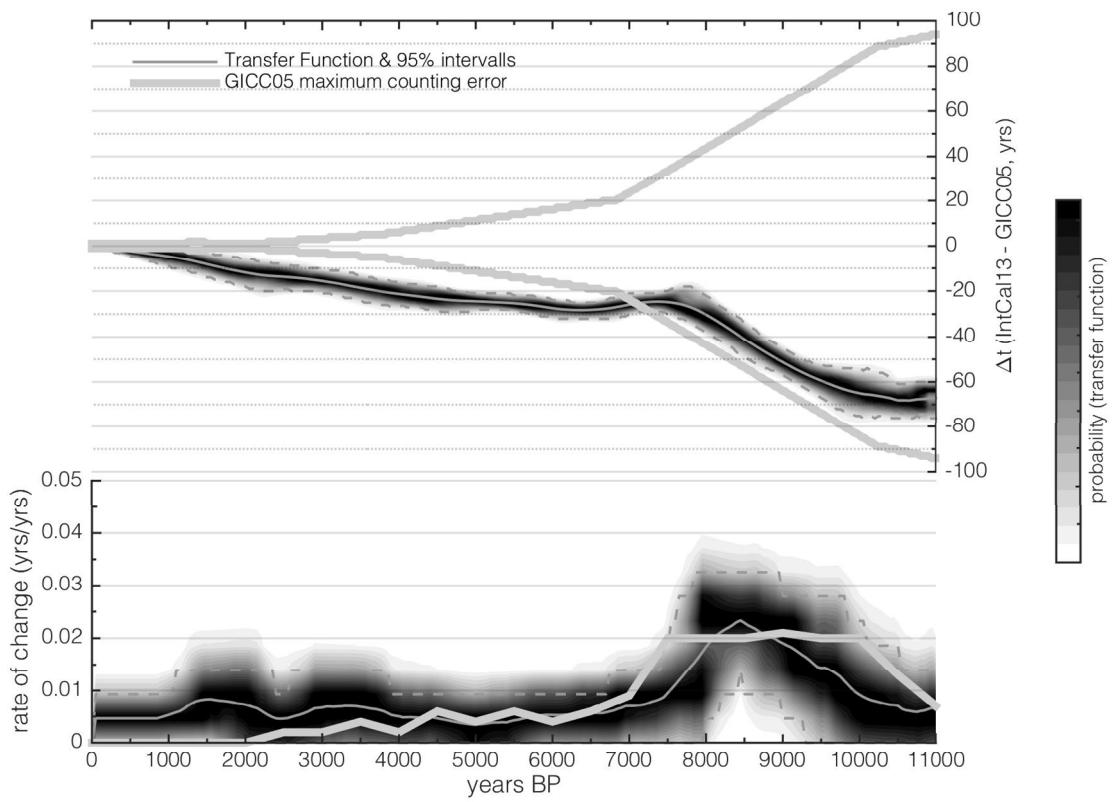
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2 Figure 10. GRIP/GISP2  $^{10}\text{Be}$  based  $\Delta^{14}\text{C}$  before (grey) and after (black) synchronization to  
3 IntCal13 (red) for the sections a) 3,500-4,500 years BP, b) 7,000-8,000 years BP, c) 10,000-  
4 11,000 years BP.

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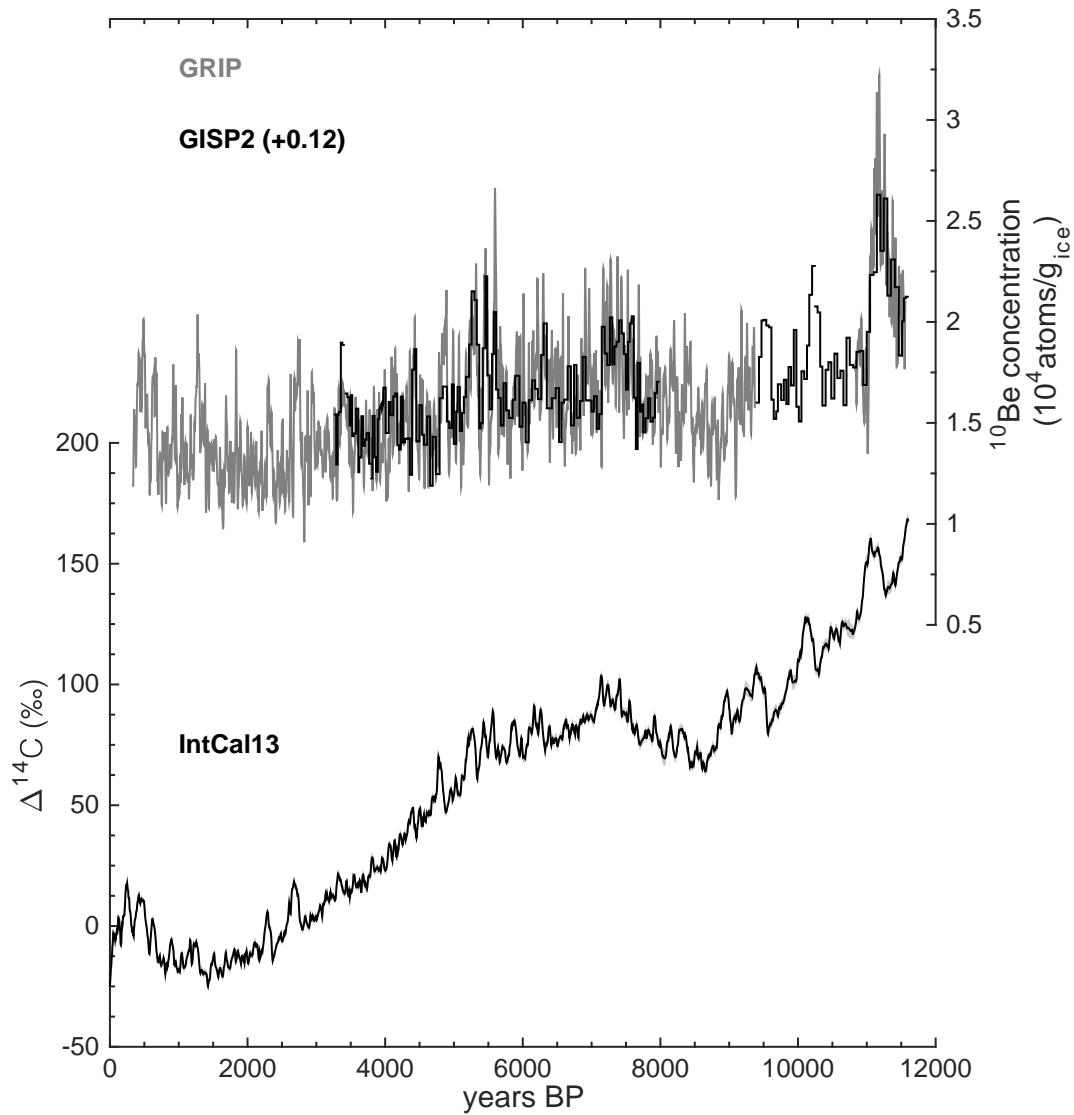


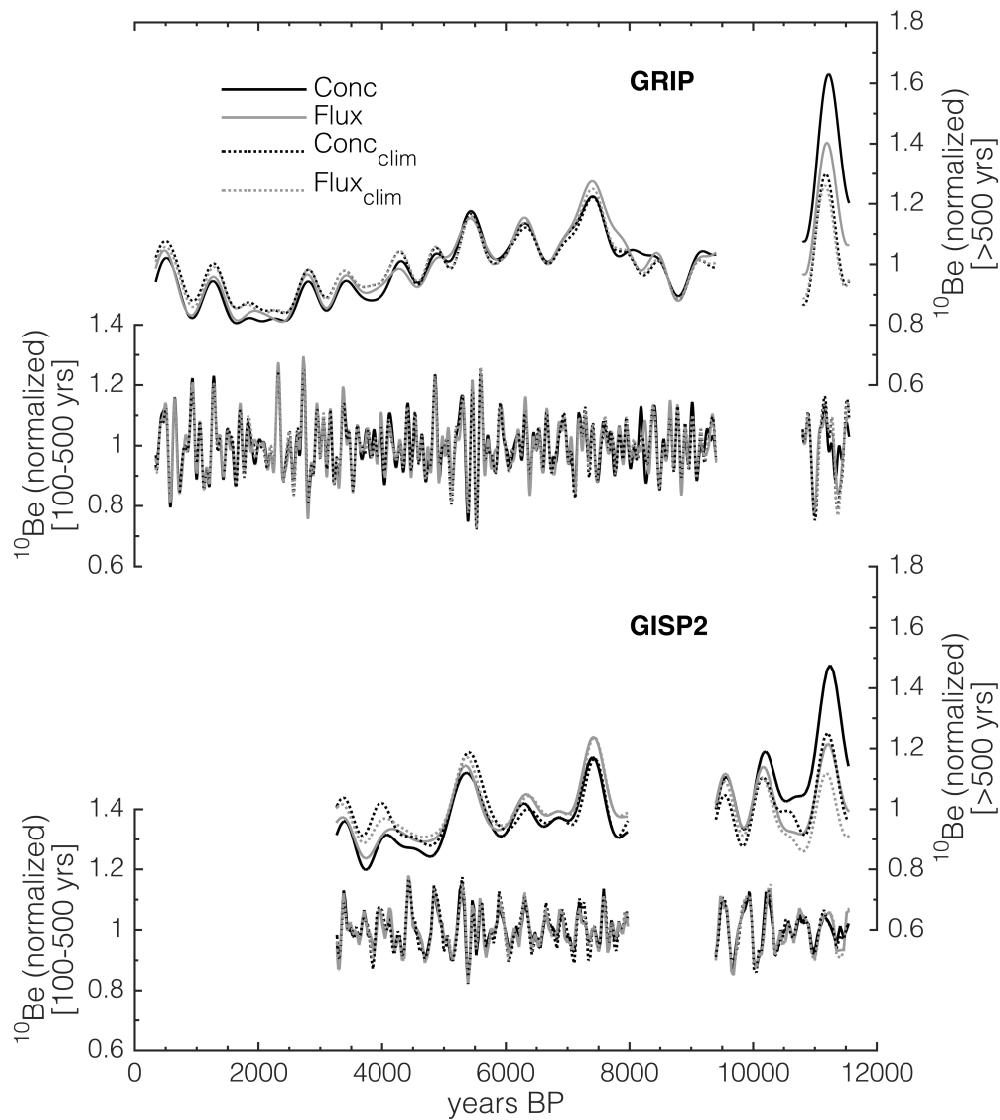
1 Figure 11. Comparison of the derived IntCal13-GICC05 timescale transfer function (black  
2 lines, this study) to the results by Muscheler et al. (2014, grey lines), and independent age  
3 markers that have been linked independently to the IntCal13 and GICC05 timescales at high  
4 precision (symbols). The results of this study and Muscheler et al. are shown with their  
5 respective 95 % confidence intervals (dashed lines). The independent age markers are plotted  
6 as the difference between their estimated ages based on radiocarbon dating (Saksunarvatn  
7 Ash, Santorini), historical documents (Vesuvius) and dendrochronology (775 and 994 AD  
8 events), and their respective GICC05-ages. The plotted  $1\sigma$  error bars largely reflect  
9 uncertainties in the radiocarbon-dating and calibration of the Saksunarvatn Ash (Lohne et al.,  
10 2013) and the Santorini eruption (Friedrich et al., 2006). Note that the identification of the  
11 Santorini tephra in ice cores has been challenged based on its geochemistry (Pearce et al.,  
12 2004).  
13

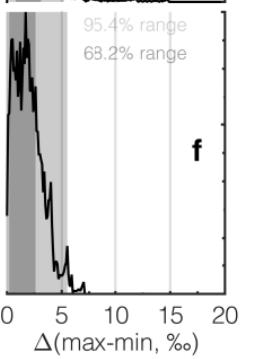
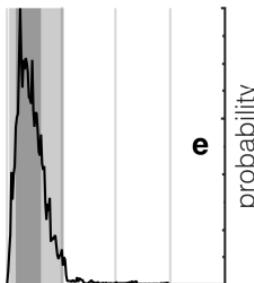
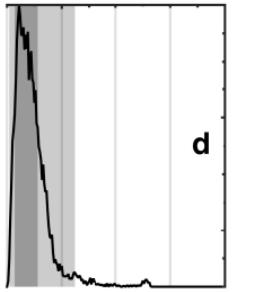
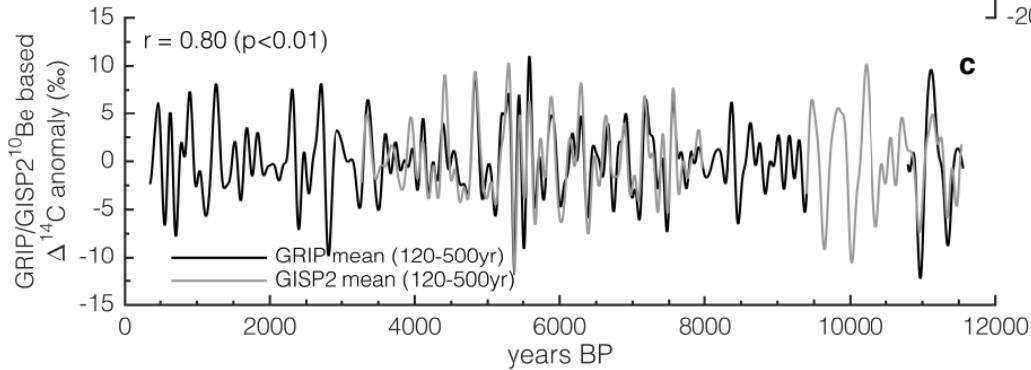
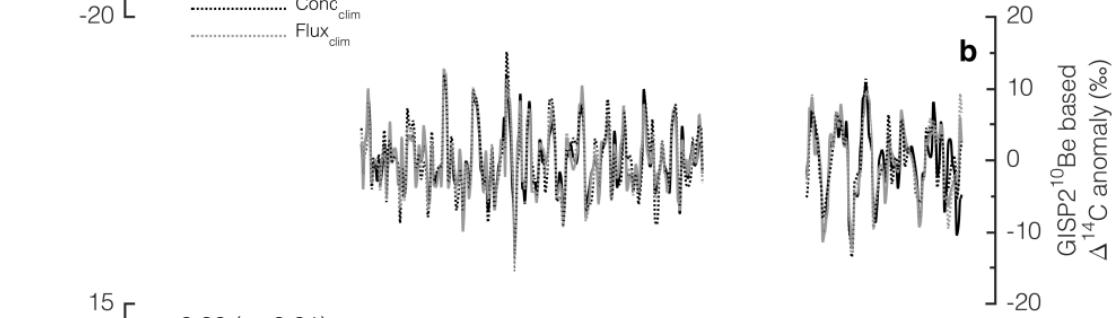
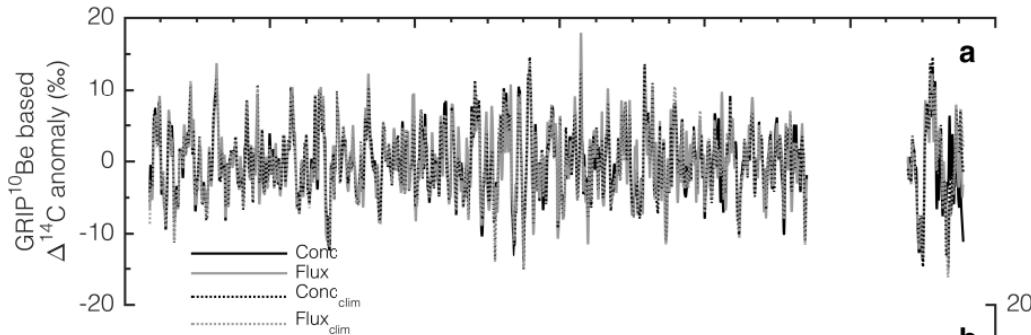
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2 Figure 12. Top: Comparison of the derived IntCal13-GICC05 transfer function (thin grey  
3 lines and shading, dashed lines denote the 95% confidence interval) to the GICC05 maximum  
4 counting error (bold grey lines). Bottom: Same as above but expressed as the rate of change  
5 (yrs/yr) of the GICC05 maximum counting error and the derived timescale transfer function.  
6







Transition 2

Steady State

Transition 1

