

**ENSO flavors in  
a tree-ring  $\delta^{18}\text{O}$   
record**

K. Schollaen et al.

# ENSO flavors in a tree-ring $\delta^{18}\text{O}$ record of *Tectona grandis* from Indonesia

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Title Page

Abstract

Introduction

Conclusions

References

Tables

Figures



Back

Close

Full Screen / Esc

Printer-friendly Version

Interactive Discussion





## ENSO flavors in a tree-ring $\delta^{18}\text{O}$ record

K. Schollaen et al.

Title Page

Abstract

Introduction

Conclusions

References

Tables

Figures



Back

Close

Full Screen / Esc

Printer-friendly Version

Interactive Discussion



(Kao and Yu, 2009) or Cold Tongue El Niño (Kug et al., 2009; Ren and Jin, 2011), exhibits SST anomalies localized in the eastern equatorial Pacific. The El Niño variant with maximum SST anomalies located in the central equatorial Pacific is referred to as the central Pacific (CP) El Niño (Kao and Yu, 2009), Warm Pool (WP) El Niño (Kug et al., 2009; Ren and Jin, 2011), date line El Niño (Larkin and Harrison, 2005) or El Niño Modoki (Ashok et al., 2007; Takahashi et al., 2011). In this study we use the terms Cold Tongue (CT), and Warm Pool (WP) El Niño (Ren and Jin, 2011) to describe these two ENSO flavors.

Identifying the mechanisms responsible for the CT and WP ENSO flavors is an active field of research. At present, there is no consensus on whether the increased frequency of WP ENSO events in recent decades (Ashok et al., 2007; Kao and Yu, 2009; Kug et al., 2009; Lee and McPhaden, 2010) are a result of anthropogenic greenhouse gas (GHG) forcing (Yeh et al., 2009), or natural variability (McPhaden et al., 2011; Newman et al., 2011). In addition, the simulation of ENSO flavors in Global Climate Models (GCMs) is still subject to limitations in our understanding of the phenomenon. Consequently, there is much uncertainty in whether ENSO activity will be enhanced or damped in the future, or if the relative frequency of ENSO flavors will change (Collins et al., 2010). Long records of ENSO activity are essential for identifying trends and multidecadal changes in the patterns of sea surface temperature associated with ENSO, making palaeoclimate reconstructions particularly attractive for shedding light onto the past and future of ENSO flavors.

Recent research on ENSO-proxy teleconnections recommends, that when interpreting proxy data, details in the differences between ENSO flavors, with regards to SST's, precipitation and salinity, should be taken into account (Karamperidou et al., 2014). Certain regions like Java lie in key locations where interannual precipitation variability is significantly correlated to one ENSO flavor but not the other (see Fig. 1). Thus, long-term rainfall proxies from Java can be useful for distinguishing between ENSO flavors, and for studying their relation to monsoon variability.

## ENSO flavors in a tree-ring $\delta^{18}\text{O}$ record

K. Schollaen et al.

Title Page

Abstract

Introduction

Conclusions

References

Tables

Figures



Back

Close

Full Screen / Esc

Printer-friendly Version

Interactive Discussion



Over the last decade there have been several attempts to reconstruct continuous time series of ENSO variability using different proxy archives such as corals (e.g. Abram et al., 2008; Charles et al., 2003; Cobb et al., 2013; Evans et al., 2002; Linsley et al., 2004; Pfeiffer et al., 2009; Quinn et al., 2006; Wilson et al., 2006), tree-ring widths (e.g. D'Arrigo et al., 2005; Fowler et al., 2012; Stahle, 1998) or tree-ring stable isotopes (Sano et al., 2012). Furthermore, several multi-proxy reconstructions of ENSO variability are available (e.g. Braganza et al., 2009; D'Arrigo et al., 2006; Emile-Geay et al., 2013; Mann et al., 2000; Wilson et al., 2010). However, many of these reconstructions are based on extratropical proxy records, particularly from tree-ring widths, and thus do not represent ENSO activity directly.

Tree-ring stable isotopes often provide additional climate information where the more commonly used tree-ring proxies (e.g., ring width and maximum latewood density) do not, or where the teleconnection signal is weak. In tropical regions, oxygen isotope data from tree rings ( $\delta^{18}\text{O}_{\text{TR}}$ ) are often more sensitive to precipitation than ring width (e.g. Brienen et al., 2012; Schollaen et al., 2013).  $\delta^{18}\text{O}_{\text{TR}}$  data are primarily controlled by the isotopic composition of precipitation, i.e. the source water, and relative humidity (e.g. Barbour, 2007; McCarroll and Loader, 2004). The isotopic composition of precipitation ( $\delta^{18}\text{O}_{\text{Pre}}$ ) depends on a number of factors, the so-called “kinetic isotope effects” (Araguás-Araguás et al., 2000). One of these effects, “the amount effect”, is the inverse correlation between rainfall amount and  $\delta^{18}\text{O}_{\text{Pre}}$  values, and a crucial driver in determining  $\delta^{18}\text{O}_{\text{Pre}}$  values in the tropics (e.g. Brienen et al., 2012; Zhu et al., 2012). Thus  $\delta^{18}\text{O}_{\text{TR}}$  records offer a promising approach to examine monsoon activity, and large-scale climate variations such as ENSO.

In previous studies we investigated relationships between seasonal rainfall variability and tree-ring stable isotope records from Javanese teak trees on inter- to intra-annual time scales (Schollaen et al., 2014, 2013). In this study we explore the signal strength of ENSO flavors in our annually resolved  $\delta^{18}\text{O}_{\text{TR}}$  record from Java, the only well replicated, centennial  $\delta^{18}\text{O}$  record from Javanese teak in existence. We place particular emphasis on the time stability of the teleconnected  $\delta^{18}\text{O}$ /ENSO relationship. To the



## ENSO flavors in a tree-ring $\delta^{18}\text{O}$ record

K. Schollaen et al.

Title Page

Abstract

Introduction

Conclusions

References

Tables

Figures



Back

Close

Full Screen / Esc

Printer-friendly Version

Interactive Discussion



phase (La Niña events) brings excess rain to the region (Sarachik and Cane, 2010). In this study, we further show that precipitation anomalies in Java are sensitive to ENSO flavors. Figure 1 shows the relationship between precipitation data and the WP and CT El Niño indices (see Sect. 2.2 for definition of the indices) for the IMC and Pacific region. WP El Niños are associated with drought over Java (Fig. 1, upper panel), and have a strong influence on the Australian–Indonesian monsoon system (e.g. Kumar et al., 2006; Taschetto and England, 2009). On the other hand, Java lies on the nodal line of influence of CT El Niños (Fig. 1, lower panel), which makes it a key location for obtaining records able to distinguish between the two ENSO flavors.

The growing season for teak in Central and Eastern Java occurs mostly during the wet season, from October to May (Coster, 1928, 1927; Geiger, 1915; Schollaen et al., 2013). In all subsequent analysis, we use the Southern Hemisphere convention, which assigns to each tree ring the year in which radial growth begins (Schulman, 1956). Thus lag-0 refers to the year  $n$  where tree growth starts:  $\text{Oct}_n - \text{Sep}_{n+1}$ . Lag-1 refers to  $\text{Oct}_{n-1} - \text{Sep}_n$ . ENSO-flavor indices are averaged over  $\text{Jan}_{n+1}\text{Feb}_{n+1}$ .

## 2.2 Definition of ENSO flavors

We use the global SST dataset of Kaplan et al. (1998) to calculate ENSO indices, and the coordinate transform of the NINO3-NINO4 phase space by Ren and Jin (2011) to define the two ENSO flavors (Eq. 1). No significant differences were found when using alternative indices for calculating ENSO flavors (not shown here) since the NINO3-NINO4 SST anomalies are so closely associated with rainfall anomalies in the Java region.

$$\begin{aligned} N_{\text{CT}} &= N_3 - \alpha N_4 \\ N_{\text{WP}} &= N_4 - \alpha N_3 \end{aligned}, \quad \alpha = \begin{cases} 2/5 & \text{if } N_3 N_4 > 0 \\ 0 & \text{otherwise} \end{cases} \quad (1)$$

For subsequent analyses we use the January–February ( $\text{Jan}_{n+1}\text{Feb}_{n+1}$ ) time-averaged indices for years 1900–2007. We focus on the JF period which represents the maximum

rainy season in Java, when our  $\delta^{18}\text{O}_{\text{TR}}$  record correlates the best with regional rainfall data (Schollaen et al., 2013). We classify each year as CT, or WP when  $N_{\text{CT}}$ , or  $N_{\text{WP}}$  are greater than one standard deviation of the respective monthly index. We classify a year as La Niña (LN) when NINO4 is negative by less than one standard deviation of the monthly NINO4 index. Table 1 shows the list of years classified as CT, WP, and LN according to the above criteria.

Distinguishing between the two corresponding types of La Niña events, as advocated by Kao and Yu (2009) and Ashok and Yamagata (2009), may not be necessary because the SST and precipitation patterns of the two La Niña types are not very distinctive (Kug and Ham, 2011).

### 2.3 ENSO signal assessment

To assess the long-term temporal stability of the ENSO signal, running 31 year correlations were calculated between the  $\delta^{18}\text{O}_{\text{TR}}$  record and the varying ENSO flavors. A Kalman filter analysis was also used as a time-dependent regression-modelling tool to test the temporal stability of the relationship between the  $\delta^{18}\text{O}_{\text{TR}}$  record and the two ENSO flavors. In contrast to the running correlation procedure, the Kalman filter method uses maximum likelihood estimation to objectively test for the identification of time-dependence between predictor and predicted variables (see Visser and Molenaar (1988) for details, and Cook et al. (2002, 2013) or Wilson et al. (2013) for examples).

Furthermore, probability density functions of the correlation between  $\delta^{18}\text{O}_{\text{TR}}$  variability and the different ENSO phases (WP, CT and LN), as well as during neutral conditions, were calculated. Finally, the spectral properties of the  $\delta^{18}\text{O}_{\text{TR}}$  proxy time series were analyzed (Schulz and Mudelsee, 2002) and wavelet coherency analysis performed (Grinsted et al., 2004; Torrence and Compo, 1998).

CPD

10, 3965–3987, 2014

## ENSO flavors in a tree-ring $\delta^{18}\text{O}$ record

K. Schollaen et al.

Title Page

Abstract

Introduction

Conclusions

References

Tables

Figures



Back

Close

Full Screen / Esc

Printer-friendly Version

Interactive Discussion



### 3 Results

Monthly and seasonal correlations between the Javanese  $\delta^{18}\text{O}_{\text{TR}}$  record (Fig. 2, green line in all plots) and ENSO flavors (see Sect. 2.2) were computed for both the concurrent year (lag-0) and the year prior to tree growth (lag-1) (Table 2). Statistically significant (95% level or higher) positive correlations were found between WP El Niño and the concurrent rainy season (Oct<sub>*n*</sub> – May<sub>*n+1*</sub>,  $r = 0.26$ ). Correlations are strongest when averaged over Jan<sub>*n+1*</sub>Feb<sub>*n+1*</sub> ( $r = 0.33$ ), the period of maximum rainy season precipitation. Furthermore, there is a significant correlation with lag-1 January precipitation (Jan<sub>*n*</sub>,  $r = 0.22$ ), indicating a WP El Niño influence on tree growth in the following year. Statistically significant negative correlations were found for La Niña events in January and February (Jan<sub>*n+1*</sub>Feb<sub>*n+1*</sub>,  $r = -0.24$ ) (Table 2). No positive correlation was found between the tree-ring proxy and the CT El Niño flavor (Table 2). As noted, this El Niño flavor has a weaker influence over Java (Fig. 1), therefore we expected the lag-0 correlation to be insignificant.

Although the  $\delta^{18}\text{O}_{\text{TR}}$  record correlates significantly ( $p < 0.05$ ) with ENSO flavors, the response is not stationary. Figure 2 presents the running 31 year correlation and Kalman filter analysis between the varying ENSO flavors and the tree-ring proxy for the period of highest correlation (see Table 2). The teleconnection with JF WP El Niño is strong and significantly positive from the 1950s till present, with running correlations reaching 0.6, and an overall  $r$  of 0.43 ( $p < 0.001$ ) (Fig. 2a). However, before 1950 the correlation falls to zero, and even becomes negative. The Kalman filter time-varying regression coefficients (beta weights) follow the same trend as the correlation values and reinforce the time dependency of the teleconnection. From 1950 onwards, the lower limits do not cross zero, which means that the beta weights are considered statistically significant. However, the correlation weakens slightly again in the beginning of the 21st century. The teleconnection with JF La Niña index (Fig. 2c) is also time dependent with weak correlations before 1950 and after 2000, but a significant negative relationship in the second half of the century with  $r = -0.38$  ( $p < 0.01$ ).

## ENSO flavors in a tree-ring $\delta^{18}\text{O}$ record

K. Schollaen et al.

[Title Page](#)

[Abstract](#)

[Introduction](#)

[Conclusions](#)

[References](#)

[Tables](#)

[Figures](#)



[Back](#)

[Close](#)

[Full Screen / Esc](#)

[Printer-friendly Version](#)

[Interactive Discussion](#)





## ENSO flavors in a tree-ring $\delta^{18}\text{O}$ record

K. Schollaen et al.

Title Page

Abstract

Introduction

Conclusions

References

Tables

Figures



Back

Close

Full Screen / Esc

Printer-friendly Version

Interactive Discussion



The fingerprints of the ENSO flavors in the  $\delta^{18}\text{O}_{\text{TR}}$  record can be seen in the probability density function (PDF) of  $\delta^{18}\text{O}_{\text{TR}}$  anomalies (Fig. 3). The WP probability mass is skewed towards positive anomalies associated with dry conditions. By contrast, the PDF for CT El Niño events exhibits bimodality with peaks in both positive and negative  $\delta^{18}\text{O}_{\text{TR}}$  anomalies, suggesting this record is not a good proxy for CT El Niño variability. The PDF for the previous (lag-1) rainy season CT El Niño (dashed line) is also skewed towards negative  $\delta^{18}\text{O}_{\text{TR}}$  anomalies, similar to that of lag-0 La Niña events (blue line), supporting the idea that the lag-1 correlation reflects subsequent La Niña events.

To further investigate expressions of ENSO variability in the  $\delta^{18}\text{O}_{\text{TR}}$  record we performed spectral analysis (Fig. 4a). Spectral analysis of the  $\delta^{18}\text{O}_{\text{TR}}$  record reveals a broad peak at 2–4 years, falling within the classic ENSO bandwidth (Sarachik and Cane, 2010) as well as significant, decadal-to-multidecadal variability (12.5 years). Wavelet coherence analysis between the proxy record and WP El Niño (Fig. 4b) and La Niña (Fig. 4c) indicates that their coherence varies in time across most spectral bands. The periods of greatest coherence in time occur on inter-annual timescales (2–4 years), again spanning the classic ENSO bandwidth.

## 4 Discussion

The positive correlation pattern between the  $\delta^{18}\text{O}_{\text{TR}}$  record and the WP El Niño flavor, as well as the negative correlation with La Niña events, supports the conclusion in Schollaen et al. (2013) that the formation of annual  $\delta^{18}\text{O}$  in Javanese teak trees is dominated by precipitation patterns. El Niño events are linked to drought conditions over the IMC coinciding with increased  $\delta^{18}\text{O}$  values in the tree-ring proxy (Figs. 2 and 3). The opposite occurs during La Niña events. The PDFs illustrate a clear WP El Niño and a less strong La Niña signal, with really dry years linked to WP El Niños. In contrast, no clear CT El Niño signal is preserved in the  $\delta^{18}\text{O}_{\text{TR}}$  record. The bimodality in the PDF illustrates the uncertainty in cases with no strong signal. The different seasonal

rainfall signals (wet and dry season rainfall) in the  $\delta^{18}\text{O}_{\text{TR}}$  record are damped in the annually resolved proxy due to seasonally alternating isotope signatures in  $\delta^{18}\text{O}$  of precipitation (Schollaen et al., 2013). Thus, CT El Niño signals seem to be obscured when followed by a La Niña event. This is the case for the strong CT El Niño event in 1982/83 that was followed by a La Niña, resulting in a low  $\delta^{18}\text{O}_{\text{TR}}$  value (Fig. 2b and c). High-resolution intra-annual  $\delta^{18}\text{O}_{\text{TR}}$  analyses help to disentangle the contrasting isotope effects of dry and rainy season rainfall patterns, as demonstrated in Schollaen et al. (2014). We conclude that the annually resolved tree-ring proxy is suitable for distinguishing between WP El Niño and La Niña, but not for CT El Niños. Overall, the strongest and most significant ENSO signal in the tree-ring proxy data is that of WP El Niño.

Our experiments show that the teleconnections described above are not stationary (Figs. 2 and 4). There is a drop in correlation in the first half of the 20th century. One can speculate this weakening teleconnection is related to the pattern of relatively weak and irregular ENSO activity in the middle of the 20th century (Tudhope et al., 2001). Arguably, there may be other factors (e.g. Indian Ocean Dipole Mode) determining wetter or drier conditions in this period and the ENSO phenomenon may play a secondary role. In recent decades, a climate regime transition has preceded periods of strong and sustained ENSO events (e.g. O’Kane et al., 2014), leading to a stronger ENSO fingerprint in the  $\delta^{18}\text{O}_{\text{TR}}$  record. Furthermore Chang et al. (2004) reveal an interdecadal trend of increasing correlations between Indonesian monsoon rainfall and ENSO beginning in the late 1970s.

Several analyses of Indonesian rain gauge data show that Indonesian rainfall is poorly correlated with ENSO events during the wet monsoon season, but reveal highest coherence during the dry season and transition months prior to the wet season (June to November) (Haylock and McBride, 2001; Hendon, 2003). This is especially true for January, which has consistently insignificant correlations over Indonesia (1979–2002) (Chang et al., 2004). However, taking the IMC and surrounding oceanic rainfall into account, rainfall during the wet season is related to ENSO (see Fig. 8a in Jourdain

**ENSO flavors in a tree-ring  $\delta^{18}\text{O}$  record**

K. Schollaen et al.

[Title Page](#)[Abstract](#)[Introduction](#)[Conclusions](#)[References](#)[Tables](#)[Figures](#)[Back](#)[Close](#)[Full Screen / Esc](#)[Printer-friendly Version](#)[Interactive Discussion](#)

## ENSO flavors in a tree-ring $\delta^{18}\text{O}$ record

K. Schollaen et al.

et al. (2013) and Fig. 1 of this study). The  $\delta^{18}\text{O}_{\text{TR}}$  record is a rainfall indicator for wet and dry season rainfall, albeit largely dominated by the wet season signal (Schollaen et al., 2013). Note, that the “amount effect” leads to different isotopic signatures in  $\delta^{18}\text{O}_{\text{TR}}$  values during wet and dry season. Thus, the dry season rainfall signal, which tends to have the highest coherence with ENSO, is damped in the annually resolved  $\delta^{18}\text{O}_{\text{TR}}$  record by the following wet season signal. This may explain the low correlation between the tree-ring proxy and June to November ENSO indices. To distinguish the causes of inter-annual rainfall variability across Java future work needs to focus on high-resolution  $\delta^{18}\text{O}_{\text{TR}}$  records.

### 5 Conclusions

In this study we used a  $\delta^{18}\text{O}_{\text{TR}}$  chronology from teak (*Tectona grandis*) that correlates significantly with regional precipitation over Java (Schollaen et al., 2013) to examine various manifestations of ENSO. This is the first time a high-resolution  $\delta^{18}\text{O}_{\text{TR}}$  record is used to detect signals of ENSO flavors in palaeoclimatic data as argued by Karamperidou et al. (2014). These results indicate the significant potential for generating reconstructions of different ENSO flavors from the  $\delta^{18}\text{O}_{\text{TR}}$  records in Indonesian teak. Such palaeoclimatic records may help answering the many remaining questions surrounding the diversity of ENSO activity. In addition, the conclusions of our study call for caution when doing model-proxy comparisons using ENSO indices that are not able to distinguish between the two flavors (e.g. single standard indices such as NINO3.4). Performing such comparisons may confound attempts to reconcile models with proxies.

More emphasis is needed on sampling long-term terrestrial  $\delta^{18}\text{O}_{\text{TR}}$  records at seasonal resolution from eastern Indonesia to reveal a robust reconstruction of wet and dry season rainfall with its teleconnection to the different ENSO flavors.

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[Title Page](#)[Abstract](#)[Introduction](#)[Conclusions](#)[References](#)[Tables](#)[Figures](#)[Back](#)[Close](#)[Full Screen / Esc](#)[Printer-friendly Version](#)[Interactive Discussion](#)



## ENSO flavors in a tree-ring $\delta^{18}\text{O}$ record

K. Schollaen et al.

[Title Page](#)

[Abstract](#)

[Introduction](#)

[Conclusions](#)

[References](#)

[Tables](#)

[Figures](#)



[Back](#)

[Close](#)

[Full Screen / Esc](#)

[Printer-friendly Version](#)

[Interactive Discussion](#)



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## ENSO flavors in a tree-ring $\delta^{18}\text{O}$ record

K. Schollaen et al.

[Title Page](#)

[Abstract](#)

[Introduction](#)

[Conclusions](#)

[References](#)

[Tables](#)

[Figures](#)



[Back](#)

[Close](#)

[Full Screen / Esc](#)

[Printer-friendly Version](#)

[Interactive Discussion](#)



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## ENSO flavors in a tree-ring $\delta^{18}\text{O}$ record

K. Schollaen et al.

[Title Page](#)

[Abstract](#)

[Introduction](#)

[Conclusions](#)

[References](#)

[Tables](#)

[Figures](#)



[Back](#)

[Close](#)

[Full Screen / Esc](#)

[Printer-friendly Version](#)

[Interactive Discussion](#)



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- 30

## ENSO flavors in a tree-ring $\delta^{18}\text{O}$ record

K. Schollaen et al.

[Title Page](#)

[Abstract](#)

[Introduction](#)

[Conclusions](#)

[References](#)

[Tables](#)

[Figures](#)



[Back](#)

[Close](#)

[Full Screen / Esc](#)

[Printer-friendly Version](#)

[Interactive Discussion](#)



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## ENSO flavors in a tree-ring $\delta^{18}\text{O}$ record

K. Schollaen et al.

Title Page

Abstract

Introduction

Conclusions

References

Tables

Figures



Back

Close

Full Screen / Esc

Printer-friendly Version

Interactive Discussion



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**ENSO flavors in a tree-ring  $\delta^{18}\text{O}$  record**

K. Schollaen et al.

**Table 2.** Correlation values between the annually resolved  $\delta^{18}\text{O}_{\text{TR}}$  record and climate months of different ENSO flavors for the period from the year prior to growth (lag-1) to the current year (lag-0) and seasonal means (calculated over the 1900–2007 period). (\*\*:  $p < 0.001$ , \*:  $p < 0.01$ , bold:  $p < 0.05$ ).

Climate months lag-1  lag-0	WP El Niño	CT El Niño	La Niña
Oct <sub>n-1</sub>   Oct <sub>n</sub>	0.12  <b>0.2</b>	-0.18  -0.0	0.00  -0.1
Nov <sub>n-1</sub>   Nov <sub>n</sub>	0.17  0.1	<b>-0.23</b>   -0.0	0.00  -0.1
Dec <sub>n-1</sub>   Dec <sub>n</sub>	0.18  0.1	<b>-0.20</b>   0.0	-0.01  -0.1
Jan <sub>n</sub>   Jan <sub>n+1</sub>	<b>0.22</b>   <b>0.35</b> **	-0.19  -0.0	-0.06  <b>-0.25</b> *
Feb <sub>n</sub>   Feb <sub>n+1</sub>	0.15  <b>0.29</b> *	-0.15  -0.06	-0.05  <b>-0.2</b>
Mar <sub>n</sub>   Mar <sub>n+1</sub>	0.05  <b>0.2</b>	-0.17  -0.12	0.03  -0.1
Apr <sub>n</sub>   Apr <sub>n+1</sub>	0.14  <b>0.2</b>	<b>-0.21</b>   -0.0	-0.02  -0.1
May <sub>n</sub>   May <sub>n+1</sub>	0.12  <b>0.2</b>	-0.14  -0.1	-0.03  -0.1
Jun <sub>n</sub>   Jun <sub>n+1</sub>	0.14  0.0	-0.07  -0.0	-0.08  -0.0
Jul <sub>n</sub>   Jul <sub>n+1</sub>	0.12  0.1	-0.02  -0.0	-0.09  -0.1
Aug <sub>n</sub>   Aug <sub>n+1</sub>	0.11  0.1	-0.02  -0.0	-0.10  -0.0
Sep <sub>n</sub>   Sep <sub>n+1</sub>	0.17  0.0	-0.05  -0.0	-0.11  0.0
peak wet season (Jan <sub>n+1</sub> Feb <sub>n+1</sub> )	<b>0.33</b> **		<b>-0.24</b>
wet season (Oct <sub>n</sub> – May <sub>n+1</sub> )	<b>0.26</b> *		

Title Page

Abstract

Introduction

Conclusions

References

Tables

Figures

◀

▶

◀

▶

Back

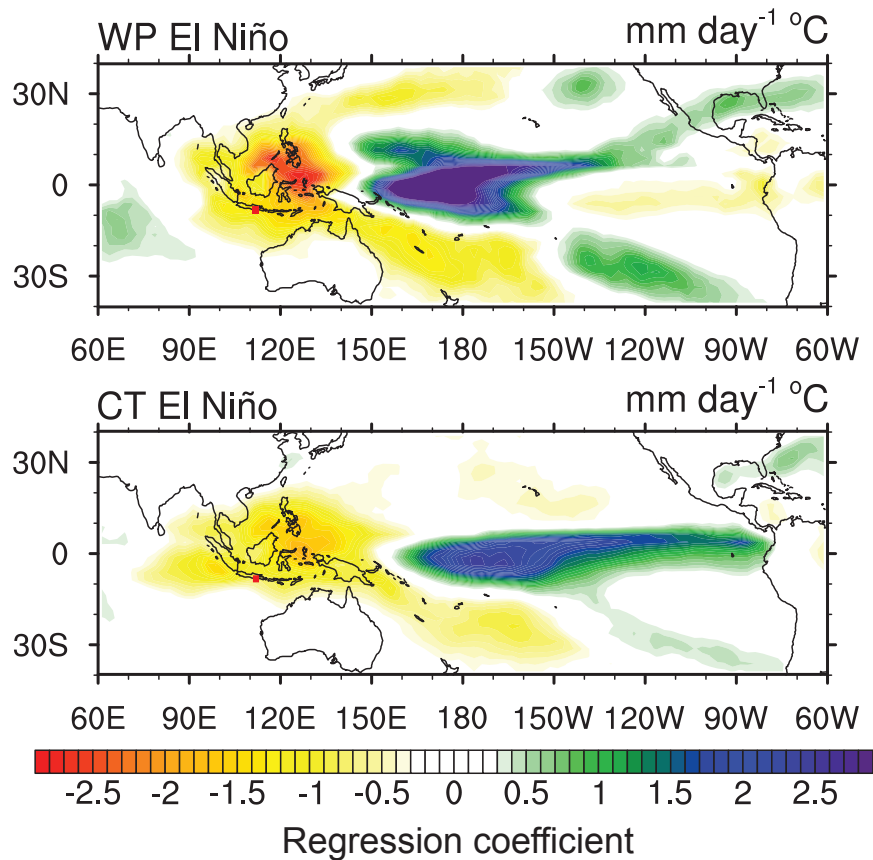
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Full Screen / Esc

Printer-friendly Version

Interactive Discussion

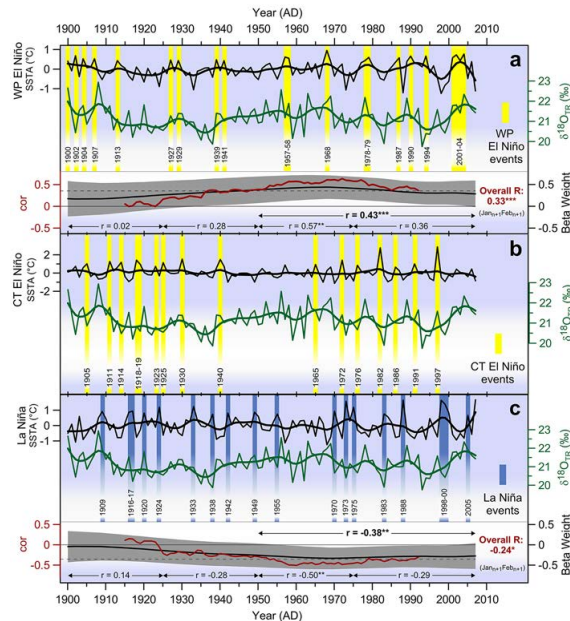




**Figure 1.** Regression coefficients (mm day<sup>-1</sup> °C) of precipitation on the Warm Pool (WP) and the Cold Tongue (CT) El Niño index. The two indices are computed as per Ren and Jin (2011) (1). Precipitation data are from the GPCP (Adler et al., 2003), for the period 1987–2010. The tree-ring site is marked with a red squared.

## ENSO flavors in a tree-ring $\delta^{18}\text{O}$ record

K. Schollaen et al.



**Figure 2.** Time series of the  $\delta^{18}\text{O}_{\text{TR}}$  chronology (green) and the January–February ( $\text{Jan}_{n+1}\text{Feb}_{n+1}$ ) time-averaged indices of **(a)** Warm Pool (WP) El Niño (black), **(b)** Cold Tongue (CT) El Niño (black), and **(c)** La Niña (black). The WP and CT El Niño indices are computed as per Ren and Jin (2011) (1). Thick lines denote 10 year cubic smoothing spline. In the lower part of each figure the running 31 year correlation (red) is shown. Dashed horizontal line indicates the 75 % confidence level. Also shown are the results from a Kalman filter analysis (black line) used as a dynamic regression modeling tool. Grey shading denotes  $\pm 2$  standard error limits of the beta weights. Where the limits do not cross zero, the regression relationship are considered statistically significant ( $p = 95\%$ ). ENSO events based on classification of Table 1 are highlighted in yellow (El Niño) and blue (La Niña), respectively. (\*\*  $p < 0.01$ , \*\*\*  $p < 0.001$ , \*  $p < 0.05$ ).

Title Page

Abstract

Introduction

Conclusions

References

Tables

Figures



Back

Close

Full Screen / Esc

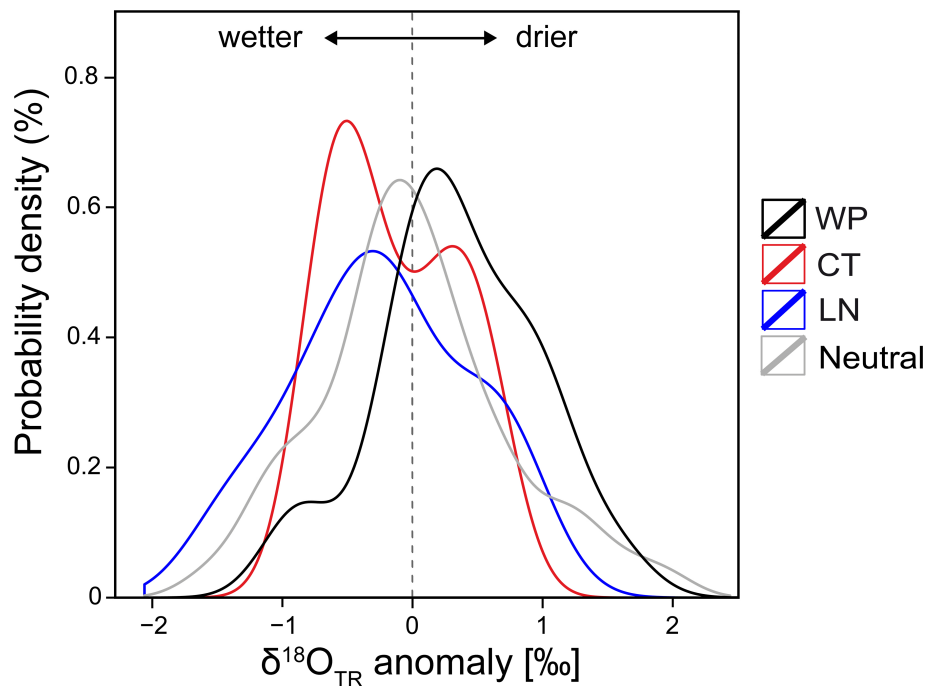
Printer-friendly Version

Interactive Discussion



## ENSO flavors in a tree-ring $\delta^{18}\text{O}$ record

K. Schollaen et al.



**Figure 3.** Probability density function of tree-ring  $\delta^{18}\text{O}$  variability from different ENSO types: Warm Pool El Niño (WP, black line), Cold Tongue El Niño (CT, red line), La Niña (LN, blue line) and neutral conditions (grey line). The January to February ( $\text{Jan}_{n+1}\text{Feb}_{n+1}$ ) time-averaged indices are shown. Plotted are  $\pm 1$  standard deviation values.

Title Page

Abstract

Introduction

Conclusions

References

Tables

Figures

◀

▶

◀

▶

Back

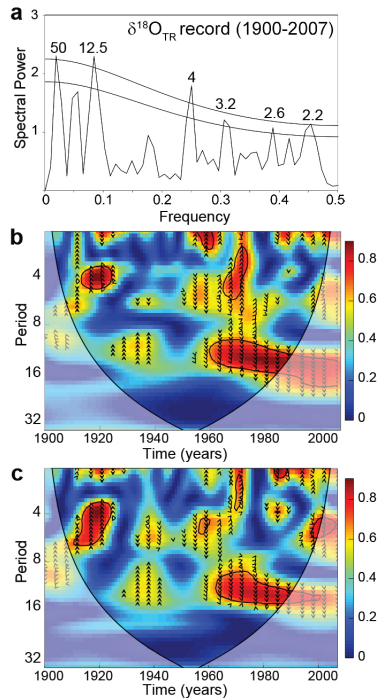
Close

Full Screen / Esc

Printer-friendly Version

Interactive Discussion





**Figure 4.** (a) Spectral analysis (Schulz and Mudelsee, 2002) of the  $\delta^{18}\text{O}_{\text{TR}}$  chronology from 1900 to 2007. 90 and 95% confidence levels are indicated. (b) Wavelet coherence transform comparing shared frequency between  $\delta^{18}\text{O}_{\text{TR}}$  record and Warm Pool (WP) El Niño index ( $\text{Jan}_{n+1}\text{Feb}_{n+1}$ ), and (c) La Niña index ( $\text{Jan}_{n+1}\text{Feb}_{n+1}$ ) for 1900 to 2007. The wavelet coherence illustrating temporal frequency coherence between the time series at given periods. The thick black contour designates where time series share significant coherence ( $p = 95\%$ ) and the cone of influence where edge effects might distort the picture is shown as a lighter shade. Arrows indicate the phase relationship between series with in-phase pointing right and antiphase pointing left.

ENSO flavors in a tree-ring  $\delta^{18}\text{O}$  record

K. Schollaen et al.

Title Page

Abstract Introduction

Conclusions References

Tables Figures

◀ ▶

◀ ▶

Back Close

Full Screen / Esc

Printer-friendly Version

Interactive Discussion

